

A STATISTICAL RANK TEST FOR ANALYSING BIOMEDICAL DATA

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DATA

By

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ABSTRACT

In the analysis of biomedical data, a question commonly asked by researchers involves the determination of the "best" or "worst" member of a group of results and an associated measure of the probability that this member is the "best" or "worst". Commonly, analysis of variance is suggested as the test of choice. Unfortunately, this test does not exactly answer the original question and further testing must be done to satisfy the question completely. This thesis presents a non-parametric rank test which directly answers the question of "best" or "worst".

Before applying this test to biomedical problems, the probability tables associated with this test are expanded and the methods used are presented and discussed. An analogous parametric test is then described and compared in performance with the non-parametric test throughout the remainder of the thesis. Power curves for both the non-parametric and parametric test are developed for several population distributions and the results compared. The three areas of application are; chromosome frequencies in the culture of human melanoma tissue; scoring patterns among evaluators of letters of applications to medical school; and the determination of outliers when relating

vital capacity to ventilatory response.

It was found that except for cases where the number of objects was less than 10, the parametric test has equal or greater power than the non-parametric test when analysing continuous data, regardless of the population distribution. For less than 10 objects, the non-parametric test had greater power regardless of population distribution. Subsequent to analysis in the three areas cited, it was concluded that the two tests agreed very highly in selecting extreme deviates although the non-parametric test was consistently more conservative in its probability measure. The problem of ties was found to weaken the power of the non-parametric test as did the ranking procedure itself but its ease of application and superior power with small sample sizes is a distinct advantage. The robustness of the parametric test is obvious throughout the examples. A method of selecting data values which are second or third most extreme was tested and it became obvious that the data must be displayed to show its distributional characteristics before this type of analysis could be carried out or interpreted.

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CHAPTER I

INTRODUCTION

1.1 Aims

The aims of this thesis are: -

a) to introduce the idea that many of the problems in the biomedical sciences which are evaluated by analysis of variance can be more successfully analysed by a technique which measures the deviation of the most extreme member of the group from the others;

b) to expand the capabilities of an existing non-parametric technique which measures the deviation of the most extreme member of the group from the others and to compare it with its parametric analog;

c) to show how these tests may be used to analyse various types of biomedical data.

1.2 Relevance

Biomedical researchers are often faced with the problem of determining which of several treatments is "best". Another problem which appears to be unrelated to the selection of "best", is the determination of outliers. These two problems require the researcher to determine the probability that the "best" or most extreme result is dif-

ferent from all the other results in the experiment. Often, a two way analysis of variance is performed with the assumption that the data are normally distributed and that the results will reflect, at least in part, how significantly the best or most extreme member of the data deviates from the other data values. The assumption of normality with respect to biomedical data can at times be very risky and, while statistical methods based on normally distributed data are quite robust, an appropriate non-parametric test should be used in those situations where the data are known to be non-normal. The use of analysis of variance to answer the question of which value is most extreme is not particularly useful. If the extreme data value is not significantly extreme, and the data are approximately normally distributed and all from the same population, then analysis of variance will answer the question regarding extremeness. If, on the other hand, there exists a significantly extreme data value, then traditional analysis of variance would be unable to determine a significance level for this deviant result.

This thesis presents two statistical tests, one parametric and the other non-parametric, which have been developed specifically to solve problems of selecting the most extreme data value in the context of a two way classification of data.

1.3 Thesis Format

The first part of this thesis discusses the need for, and the general approach to extreme value analysis in the biomedical sciences. Also, a general description of the historical development of extreme value analysis is presented and two particular tests are described in detail.

The second part is concerned with the expansion of the non-parametric test for the extreme deviate and a comparison with the parametric analog. The third part presents a detailed analysis of three problems in the biomedical sciences using these methods. The thesis ends with a discussion of the research questions and several general conclusions.

1.4 Research Questions

The following research questions were posed.

1. What types of data are most suitable for analysis by the non-parametric test?
2. Under what conditions is the parametric test superior to the non-parametric test?
3. Is it beneficial to use the non-parametric test to determine second or even third most extreme deviates?

CHAPTER II

REVIEW OF LITERATURE RELATING TO EXTREME VALUE ANALYSIS

2.1 The Development of Extreme Value Analysis

The distant beginning of extreme value analysis is Student's t-test. Here the objective is to determine whether the difference between, say, the number of post-operative infections in two comparable groups of patients is indicative of a true difference in treatments or if the difference is the result of pure chance. A test statistic "t" determined such that, in the absence of a true treatment effect, the Studentized difference will exceed t with a probability of α . Student's t test may not be valid when the data are not normally distributed. This has led to the development of distribution-free methods of assessing the significance of differences between two data sets. The Wilcoxon two sample and paired tests being strong competitors of the t-test for this purpose.

The next level of complexity leads to the comparison of means of several samples. The F-test (one way analysis of variance) provides a test of the equality of several treatment means. The analogous non-parametric method for a one way classification has been developed by

Kruskal and Wallis (1952). The F-test is also useful in a two way classification and the analogous non-parametric methods are Friedman's χ^2 . (Friedman, 1937) and Kendall's concordance coefficient W (Kendall, 1948).

At times it is not sufficient to determine whether the null hypothesis of equal treatment means has been rejected, but it is also necessary to identify which treatment mean leads to the rejection of the null hypothesis. An approach to answering this question is as follows: in scanning the results of an experiment, one's attention will naturally be drawn to the largest contrasts and the smaller ones will not be examined. This may be stated in the following way: what value will be exceeded by a sample contrast with a prescribed conditional probability, the condition being that the contrast was large enough to attract attention in the first place from among a defined set of contrast effects? Techniques which attempt to formulate and answer this question are known as multiple comparison methods. Nemenyi (1963), in his thesis, surveyed the entire subject of distribution free multiple comparisons, providing a comprehensive list of techniques and sample computations for each. He partially expanded several tables of probability values for some of the tests, but individual methods were not examined in detail.

Extreme value analysis is one type of multiple comparison analysis, and is synonymous with "outlier stati-

stics" and "slippage measures".

The Studentized maximum absolute deviate introduced by Halperin et al (1955) is a multiple comparison test used in extreme value analysis. It is a normal theory statistic for one and two way classifications, which may be used to augment the F-test in determining which treatment has a mean different from the rest. A one tailed version of this test has also been developed by Nair (1948). The non-parametric extreme value rank sum test proposed by Doornbos-Prins (1958) and Youden (1960) and developed by Thompson-Willke (1963) and Willke (1964) can be used to augment Friedman's χ^2 or Kendall's concordance coefficient in the same way, but only for a two way classification. A non-parametric extreme value rank sum test which deals with a one way classification has been proposed by Odeh (1967) and developed by McDonald and Thompson (1967).

In a logical and chronological sense, statistical methods for the analysis of variance have developed from the simple t-test to the F-test and finally to a technique for selecting extreme values. Tests based on normal distribution theory and comparable distribution free methods have developed in parallel. The following two sub-sections will describe Nair's test and Youden's test and will serve as background to the remainder of the thesis.

2.2 Nair's Test

Let x_1, x_2, \dots, x_k be independent normally distributed variates, each with mean μ and variance σ^2 . Denote the ordered values of the above variates as Y_1, Y_2, \dots, Y_k ; then the Nair statistics may be defined as

$$d = \frac{Y_k - \bar{Y}}{s_v} \text{ or } \frac{\bar{Y} - Y_1}{s_v} ; \bar{Y} = \sum_{i=1}^k (Y_i) / k$$

where s_v is the sample estimate of σ based on v degrees of freedom and the two expressions provide for the choice of either upper or lower tail for a one tailed test.

The following two examples have been taken from Nair (1948) but the setting has been altered in keeping with the biomedical nature of this thesis.

Example 1.

A randomized block experiment has been designed to select the best of four bacteriocidal methods for purifying drinking water. Five replications gave the following mean bacteria counts per 100 ml. for the methods A; B, C and D

A	B	C	D
34.4	34.8	33.7	28.4

and an error variance of $\sigma^2 = 2.19$ based on 12 d.f. Despite the similarity of the first three means, an analysis of variance (Table 2.1) showed significant differences among the four means.

The variance ratio F for treatments against error is $44.82/2.19 = 20.5$ which is much larger than $F_{0.001} = 10.8$. Nair points out that while the large F value is probably attributable largely to the small mean associated with method D, a test of significance of the difference between 28.4 and the mean of 34.4, 34.8 and 33.7 by the usual t -test

$$t = \frac{34.3-28.4}{\sqrt{(2.19(1/15+1/5))}} = 7.7 ,$$

with 12 degrees of freedom is probably not valid as 28.4 has been selected as the smallest mean instead of being selected at random. The appropriate criterion in this case is the Studentized extreme deviate used by Nair as follows

$$\frac{\bar{x}-x_1}{s_v} = \frac{32.8-28.4}{\sqrt{1/5 \times 2.19}} = 6.7$$

with $n = 4$ and $v = 12$. Referring to Hartley and Pearson

(1969) Vol. 1, Table 26 we find that this value far exceeds the critical value at $\alpha = 0.001$ which is 4.1.

Source of Variation	Degrees of Freedom	Sum of Squares	Variance
Replications	4	21.46	5.36
Methods	3	134.45	44.82
Error	12	26.26	2.19
Total	19	182.17	-

Table 2.1

Analysis of Variance for Example 1

Having concluded, by the above procedure, that the smallest mean 28.4 is significantly smaller than the other three means, we are justified in saying that method D is definitely superior to A, B and C.

Example 2.

This is artificially created from example 1 by changing the error variance from 2.19 to 13.00. An error variance of 13.00 gives a standard error per replication of 11% which is high but not unrealistic for this type of procedure. The variance ratio for methods against error is $F = 44.82/13.00 = 3.45$ which is not significant at the 5% level. One obviously concludes, therefore, that there are

no significant differences among the means of the four procedures A, B, C and D.

But, if we compare the smallest mean against the general mean and calculate the studentized extreme deviate

$$\frac{\bar{x} - x_1}{s_v} = \frac{32.8 - 28.4}{\sqrt{1/5 \times 13.00}} = 2.7$$

we find that the probability of getting this or a larger value when $n = 4$, $v = 12$ lies between 2.5% and 1%. On the 2.5% level, therefore, 28.4 is significantly smaller than the general mean, indicating that D is superior to A, B, and C. Although this situation was contrived, the possibility of its actual occurrence is real, and care should therefore be taken in using the F test as a screening procedure for the detection of extreme deviation.

2.3 Youden's Test

Youden's extreme rank sum test is the non-parametric analogue of Nair's test and is described by Thompson and Willke (1963) as follows. Let I objects be ranked independently by each of J judges and let r_{ij} denote the rank of object i assigned by judge j . Place r_{ij} in the i th row of the j th column of a table of ranks. Each column of this table will contain a permutation of the first I integers.

Let $r_i = \sum_j r_{ij}$ denote the sum of the j ranks for the i th object. These rank sums will be our test statistics. Note that $\bar{r} = \sum r_i / I = N(I+1)/2$. The null hypothesis to be tested, is that the ranks are assigned at random by each of the judges. More precisely: H_0 : for each judge, every one of the $I!$ permutations of the ranks is equally likely. If H_0 is true, then r_1, \dots, r_I are identically, but not independently, distributed with expectations \bar{r} , and the marginal distribution of each rank sum is symmetric about \bar{r} . If, however, one of the objects tends to rank higher (or lower) than the others, then all permutations are not equally likely, and the distribution of the rank sum for that object will be skewed accordingly. Hence the rejection region for the test is taken to be the event that at least one rank sum occurs which is extreme enough to be unlikely under H_0 .

Example 1.

Four different hospital laboratories have been collaborating on a quality control program. The results shown below indicate the values obtained by each hospital analysing for a particular constituent in each of six unassayed controls labelled, A, B, C, D, E and F. It is required to know whether any of the labs has produced results extremely different from the others. The hospitals are represented as I, II, III and IV, the values in brackets are the within column rankings and the rank sums of each

hospital are shown in the far right column.

Control Lab	A	B	C	D	E	F	R _S
I	14.8(1)	16.0(1)	12.2(1)	21.3(2)	18.5(2)	22.3(1)	8
II	15.3(3)	16.3(2)	12.7(3)	22.0(3)	18.8(3)	23.0(3)	17
III	15.4(4)	16.7(3)	12.8(4)	21.1(1)	18.9(4)	23.1(4)	20
IV	15.1(2)	17.0(4)	12.3(2)	22.9(4)	18.0(1)	22.5(2)	15

Table 2.2

Quality Control results of four Laboratories
on six unknowns

Using the rank sums, Friedman's χ^2 shows

$$\begin{aligned}\chi^2 &= \frac{12}{Nk(k+1)} \sum_{j=1}^k (R_j)^2 - 3N(k+1) \\ &= \frac{12}{4 \times 6 \times (6+1)} \times 978 - 3 \times 4 \times (6+1) = 7.8\end{aligned}$$

which, with 3 degrees of freedom is significant at the 5% level. Kendall's concordance coefficient also shows significance at the same level.

$$W = \frac{S}{\frac{1}{12} k^2 (N^3 - N)}$$

$$s = \sum_{j=1}^k (R_j - \bar{R})^2$$

$$W = \frac{78}{\frac{1}{12} \times 6^2 (4^3 - 4)} = .43$$

This degree of association (0.43) is significant at the 5% level.

By invoking Youden's test and using the same rank sums, we observe (Appendix A, $I = 4$, $J = 6$) that a rank sum of 8 or less occurs about 3% of the time under the null hypothesis and therefore it can be said that the minimum rank sum of 8 is significant at the 3% level. This suggests that the results from lab I are significantly different from the other labs. We may also check the other end of the distribution by examining the largest rank sum and here we find that a rank sum of 20 is not significant at the 10% level.

Example 2.

As in section 2.2, this example is a modification of example 1, designed to show using non-parametric statistics that even though overall variance analysis is not significant, there can still exist a significantly extreme deviate.

Suppose the values in column six were modified as

in Table 2.3

Control Lab	A	B	C	D	E	F	R _S
I	14.8(1)	16.0(1)	12.2(1)	21.3(2)	18.5(2)	22.5(2)	9
II	15.3(3)	16.3(2)	12.7(3)	22.0(3)	18.8(3)	23.0(3)	17
III	15.4(4)	16.7(3)	12.8(4)	21.1(1)	18.9(4)	22.3(1)	17
IV	15.1(2)	17.0(4)	12.3(2)	22.9(4)	18.0(1)	23.1(4)	17

Table 2.3

Quality Control results of four Laboratories
on six unknowns (modified)

The rank sums appear to show that lab I varies considerably from the others and yet both Friedman's χ^2 (4.8) and Kendall's W ($s = 48$) fail to recognize an overall variation significant at the 10% level. On the other hand, Youden's test recognizes a rank sum of 9 as being significantly extreme at the 8% level.

2.4 Current Approaches to Extreme Value Analysis in the Biomedical Sciences

On reviewing current biomedical literature, it is apparent that tests of extremeness are almost never used in data analysis in this area. In the simplest situations, where a test group is compared with a control group and the

t-test applied in determining the probability of a difference, it could also be said that the degree of extremeness of the test group from the control is being assessed. But beyond this, tests of extremeness are rarely employed. Moreover, a large amount of biomedical research today asks such questions as; what is the best ..., what is the largest ..., the smallest ..., the strongest ..., the cheapest The search for these superlatives argues for the use of tests of extremeness.

Many of the problems in the biomedical sciences are problems of selection but are not formulated as such. They are formulated instead as hypothesis testing questions, with the result that the real questions are not addressed directly.

Both Youden's test and Nair's test are able to measure extremeness in a two way classification but are not strictly selection techniques as there is an underlying null hypothesis of no difference. They are, though, more specific than the popular analysis of variance methods in that they do specify a significance level for the extreme deviate.

The use of these techniques in the biomedical sciences, with particular attention to Youden's test for one tailed probabilities is developed in the remaining chapters.

CHAPTER III

THE EXPANSION OF TABLES FOR YODEN'S TEST FOR THE EXTREME DEVIATE

3.1 Introduction

As far as can be determined, there are only two tables published for Youden's Test for the Extreme Deviate. Thompson and Willke produced a table of two-sided percentage points for nominal one, three and five percent values over the range three to fifteen for both I and J. Subsequently Willke published a table of one-sided percentage points for the same nominal values and ranges of I and J. This chapter will describe the various techniques used to produce a table of one-sided percentage points for one, three, five and ten percent values over the range two to twenty-five for I and three to twenty-five for J. In addition, an attempt has been made to supply four digits of precision as opposed to three, or an upper and lower bound where this degree of precision is unattainable.

The second part of this chapter deals with a comparison of Youden's Test and its parametric analog, a test developed by Nair. A simulation technique is used to compare the two tests in terms of their operating characteristics curves for three different types of parent populations.

3.2 Techniques For Generating Tables Of One-Sided Percentage Points

3.2.1 Introduction

This section deals with the various methods used to generate the table of one-sided percentage points. Also discussed are two approximation techniques for obtaining percentage points which are beyond the bounds of the table.

Table 3.1 outlines which techniques were used and in what areas of the overall table they were used. The numbers in each of the areas refer to the sub-heading numbers for a description of the techniques found in the text. For example; 3.2.2, which refers to exhaustive enumeration, was used to calculate the values in the table for I from 3 to 11 while J varied from 3 to 8 for each I.

3.2.2 Exhaustive Enumeration

When the problem of developing a table for Youden's test was first approached, this technique was thought of first and while it reflects a basic understanding of the use of Youden's rank sum test, exhaustive enumeration is soon found to be of little use generally.

Consider I objects and J judges. Each of the judges can independently assign I! different rank arrangements to the I objects. Therefore, in total there are $(I!)^J$ different sets of rankings of the I objects by the J judges. For each

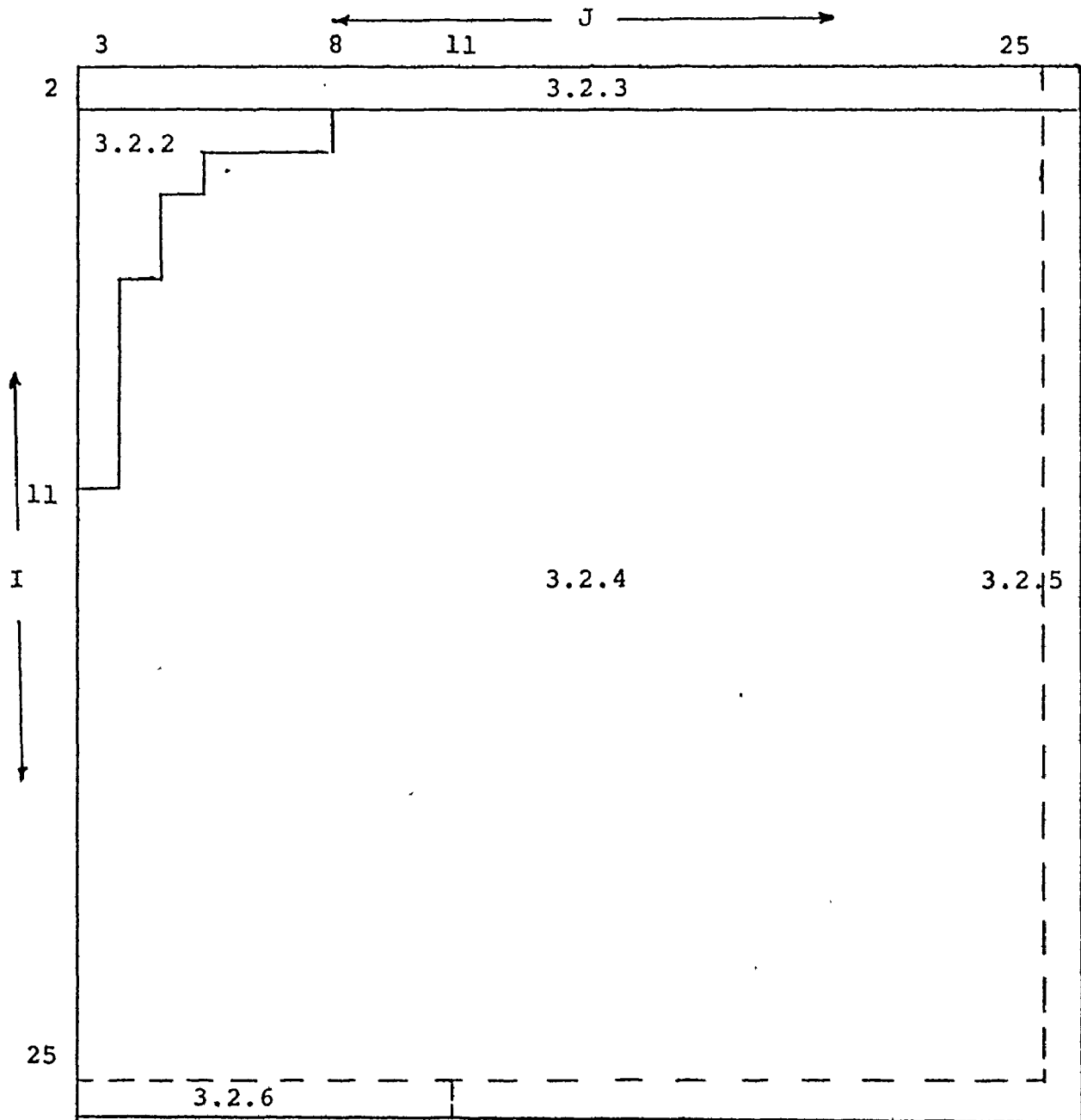


Table 3.1

A Mapping of Techniques to Youden's Table of One-Sided Percentage Points

of these sets of rankings it is required to produce the I rank sums, determine the size of the smallest or largest rank sum (for the one-sided tables) and increment the appropriate rank sum counter by one. The number of sets can be reduced to $(I!)^{J-1}$ by setting the first judge's ranking as constant and ultimately dividing the rank sum counts by $(I!)^{J-1}$ rather than $(I!)^J$, to produce the individual rank sum probabilities. The limitation of this technique is the sheer amount of calculation required. For example, to calculate the probability distribution for ten objects and eight judges would require the evaluation of more than $(10!)^7 = 8.2 \times 10^{45}$ sets of rankings of eighty rank values each. The number of arrangements produced by each judge is also greater than 3.6×10^6 . If a computer could evaluate one million sets per second, it would take about 2.6×10^{32} years to complete the job for this one distribution alone.

The main benefit of exhaustive enumeration lies in the calculation of distributions for very small values of I and J where limiting type approximations are notoriously poor and also to aid the statistician in "getting a feel" for what occurs as these rank sets are produced. The latter benefit is helpful in the development of techniques for approximating the actual distributions.

Example:

The case of $I = 3, J = 2$.

Keeping the first judge constant we have $(3!)^{2-1} = 6$ different sets, from which we extract and tabulate minimum rank sums.

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Min. Rank Sum	Frequency	Probability	Cum. Prob.
2	2	2/6 (.3)	2/6 (.3)
3	3	3/6 (.5)	5/6 (.83)
4	1	1/6 (.16)	6/6 (1.0)

As I becomes large it becomes increasingly complicated to produce all possible arrangements of the I ranks. To simplify this, the program described in Appendix B (1) which carries out exhaustive enumeration can use either the transposition method of M.B. Wells to initialize an array

with all arrangements of I ranks, or the lexicographical method of D.H. Lehmer which produces the i th permutation of I integers directly ($1 \leq i \leq I!$).

3.2.3 A Probability Generating Function For The Case $I = 2$

For the case of two objects ($I = 2$) and J judges it was possible to develop a probability generating function. In this situation each of the J judges can only assign one of two possible ranks to each of the two objects and therefore we have a binomial or two-class population. The binomial distribution is the sampling distribution of the proportions we might observe in random samples drawn from a two-class population. Since the proportion of cases expected in one of the categories is $1/2$, the probability of obtaining a given rank sum R , for either object, is

$$p(R) = \binom{J}{R-J} \cdot \left(\frac{1}{2}\right)^{R-J} \cdot \left(\frac{1}{2}\right)^{J-R+J} = \binom{J}{R-J} \cdot \left(\frac{1}{2}\right)^J$$

for

$$J < R < 2J$$

and for the one-tailed situation

$$p(R) = 2 \binom{J}{R-J} \cdot \left(\frac{1}{2}\right)^J .$$

The cumulative sampling distribution is therefore

$$2 \cdot \sum_{i=0}^x \binom{J}{i} \cdot \left(\frac{1}{2}\right)^J$$

where

$$x = R - J$$

$$J \leq R \leq 2J$$

The binomial coefficients can then be used to produce the one-tailed percentage points for any J and $I = 2$ in the following way:

1. Generate the J th order coefficients using Pascals algorithm.
2. Double the first $\lceil \frac{J+1}{2} \rceil$ coefficients.
3. Divide the first $\lfloor \frac{J}{2} \rfloor + 1$ coefficients by 2^J .
4. Accumulate these results such that the i th cumulant is the sum of the first i terms.

Example:

Suppose we wish to generate the cumulative probability distribution for the case $I = 2$ and $J = 6$.

1. Using Pascal's algorithm

	ORDER
1 1	1
1 2 1	2
1 3 3 1	3
1 4 6 4 1	4
1 5 10 10 5 1	5
1 6 15 20 15 6 1	6

2. Double the first $\lfloor \frac{6+1}{2} \rfloor = 3$ coefficients

2 12 30 20 15 6 1

3. Divide the first $\lfloor \frac{6}{2} \rfloor + 1 = 4$ coefficients by $2^6 (64)$

2/64 12/64 30/64 20/64

4. Accumulate these results

2/64 14/64 44/64 64/64

and the resulting distribution for $I = 2, J = 6$ is:

R_{\min}	R_{\max}	Probability
6	12	2/64 (.03125)
7	11	14/64 (.21875)
8	10	44/64 (.68750)
9	9	64/64 (1.0000)

Appendix B (2) contains a listing of a computer program

which will generate probabilities by the above method for specified values of J .

3.2.4 Approximation Using An Iterative Procedure Based On Bonferroni's Inequalities

This technique has been used by Thompson and Willke (1963) to approximate tables of two-sided percentage points and later by Willke (1964) to produce corresponding one-sided tables.

Let $r_i = \sum_j r_{ij}$ denote the sum of the ranks for the i th object, where r_{ij} is the rank in the i th row and j th column of a table of ranks. These rank sums become the test statistics for testing the hypothesis H_0 : The I objects are indistinguishable. For each i we have $J \leq r_i \leq IJ$. Let R be defined, so that under the null hypothesis

$$P[A(R)] = 1 - \alpha$$

where $A = A(R) = \{r_i : J+R < r_i < IJ-R, i = 1, \dots, I\}$ and α is an appropriate significance level. \bar{A} , the complement of A provides a two-sided symmetric region for rejecting H_0 .

$$\bar{A} = \{r_i : r_{\min} < J+R \text{ or } r_{\max} > IJ-R\},$$

and a one-sided region for rejecting H_0 is provided by

using r_{\min} or r_{\max} in the above equation, where r_{\min} and r_{\max} are respectively the minimum and maximum rank sums. Upper and lower bounds for $P[A(R)]$ can be obtained by the following procedure:

Define the event $A_i\{r_i: J+R < r_i < IJ-R\}$ and denote the complement by \bar{A}_i ; then $\bar{A} = \bigcup_i \bar{A}_i$.

By Bonferroni's inequalities we have under H_0

$$IP(\bar{A}_1) \geq P(\bar{A}) \geq IP(\bar{A}_1) - \binom{I}{2} P(\bar{A}_1, \bar{A}_2) \quad (1)$$

To compute the upper bound let $p_J(k) = P(r_1=k)$ for J judges. If $p(i)$ is the probability that the first object receives a rank of i from the J th judge, then we have the recursion relationship.

$$p_J(k) = \sum_i p(i) p_{J-1}(k-i) \quad (2)$$

Under the null hypothesis, $p(i) = \frac{1}{I}$. When $k < \frac{J}{2}$ (for J odd) or $k < \frac{J}{2}-1$ (for J even), then the upper bound $IP(\bar{A}_1)$ as calculated by $p_J(k)$ is the exact probability of a given rank sum occurring. This fact is particularly useful for small I , since the 10% level is usually found before either of these inequalities is violated. When $k > \frac{J}{2}$ (for J odd) or $k > \frac{J}{2}-1$ (for J even) then $IP(\bar{A}_1)$ as calculated by $p_J(k)$ is no longer exact but becomes a conservative upper bound. That is, it will always be higher than the exact value. In

this case a lower bound is required. In light of the fact that $IP(\bar{A}_1)$ has been shown to be exact under certain conditions, a better lower bound has been developed. If we let q represent the rank sum beyond which $p_J(k)$ is no longer exact, then the lower bound is defined as

$$p_J^{(I)}(k) - p_J^{(I-1)}(k-q-1+j)$$

the superscripts I and $I-1$ indicating the number of objects. Appendix B (3) contains a listing of a computer program which will calculate one-sided probabilities by this technique.

3.2.5 Approximation By A Normal Distribution For Large J

Using the moment values

$$Er_i = \frac{1}{2}J(I+1) = \bar{r}'$$

$$\text{var } r_i = J(I+1)(I-1)/12$$

$$\text{cov}(r_i, r_{i'}) = -J(I+1)/12$$

and asymptotic normality we have, according to Thompson and Willke (1963), the following:

For large J the probability is approximately $1 - \alpha$ that all row sums r_i simultaneously lie in the interval whose endpoints are $\frac{1}{2}J(I+1) \pm J^{1/2}a(I, \alpha)$. Here

I	1 PERCENT		3 PERCENT	
	YOU DEN	NORMAL	YOU DEN	NORMAL
5	54	55	57	56
10	93	93	98	98
15	131	131	138	138
20	169	168	177	177
25	206	204	216	216

I	5 PERCENT		10 PERCENT	
	YOU DEN	NORMAL	YOU DEN	NORMAL
5	58	59	60	60
10	100	101	104	104
15	141	141	146	147
20	182	182	188	189
25	222	221	230	230

Table 3.2

A Comparison of Critical Values Between Tabled Values and The Normal Approximation For $J = 25$

$$a(I, \alpha) = \left[\frac{I(I+1)}{12} \right]^{1/2} h(I, \alpha)$$

and $h(I, \alpha)$ is as defined by Halperin et al (1955) as the significance point of the maximum absolute deviate in normal samples. Values for $h(I, \alpha)$ may be found in Biometrika Tables for Statisticians, Hartley and Pearson (1968). Table 3.2 gives an idea of how good the normal approximation is to the tabled values for $J = 25$ and various values of I . The minimum rank sums for nominal 1, 3, 5 and 10 percent levels are compared.

3.2.6 Approximation By A Uniform Distribution For Large I

Thompson and Willke (1963) have developed the justification for a uniform approximation in the following way. Under H_0 , r_{ij} has a discrete uniform distribution over the integers 1, ..., I . Thus for a large I a continuous uniform distribution ought to provide a good approximation. Also, r_{ij} and $r_{i,j}$ are dependent random variables, but one suspects that this dependency should vanish as $I \rightarrow \infty$. Subsequently, Thompson and Willke derive the following:

$$\alpha_u = IP(\bar{A}_i) \doteq \frac{I}{J!} \left(\frac{J+2R+1}{2I} \right)^J ; \frac{J+2R+1}{2I} \leq 1$$

and solving for R ,

$$R \doteq I \left(\frac{\alpha_u^{J!} I^{1/J}}{I} \right) - \frac{J+1}{2}; \frac{\alpha_u^{J!}}{2I} < 1$$

Table 3.3 compares the uniform approximation to tabled values for $I = 25$ and selected values of J as large as ten. Where $J > 10$ and I is large, the normal approximation seems superior to the uniform approximation, but this has not been thoroughly investigated.

3.3 The Comparison Of Youden's Test and Nair's Test

3.3.1 Introduction

As a basis for comparison, it was felt that the power curves for the two tests would provide a simple yet useful way of attaining this goal. A simulation technique was used to produce the curves under various conditions, as the derivation of the theoretical curves was beyond the scope of this thesis.

Separate pairs of curves were generated for values of $I = 5, 15$ and 25 and for values of $J = 5, 15$ and 25 . Alpha levels of both one and five percent were specified. The above specifications were carried out under three separate assumptions regarding the variance of the parent populations. In the first case, both parent populations were normal, with equal variances. In the second case, they were both uniform with equal variances and in the third case the distributions were also uniform, but the variance of the

J	1 PERCENT		3 PERCENT	
	YOU DEN	UNIFORM	YOU DEN	UNIFORM
3	5	4	6	6
4	9	9	12	12
5	16	16	19	19
7	31	31	35	35
10	56	56	63	62

J	5 PERCENT		10 PERCENT	
	YOU DEN	UNIFORM	YOU DEN	UNIFORM
3	7	7	8	8
4	13	13	15	15
5	21	21	24	24
7	38	38	41	41
10	66	65	70	70

Table 3.3

A Comparison of Critical Values Between Tabled Values and The Uniform Approximation For $I = 25$

extreme deviate population increased as the difference between the two means increased.

The simulation was performed in the following manner. Values for I and J were specified prior to execution along with the type of parent population, and the number of separate evaluations to be done at each delta value. All data presented here represent 200 evaluations at each delta level. The delta value is defined as the true difference between the population mean and the mean of the extreme deviate. This value is incremented from zero to three in steps of 0.2, proportionate to the population standard deviation such that a delta of 1.4, for example, indicates that the mean of the extreme deviate is 1.4 standard deviations away from the population mean. The simulation generates and evaluates in turn, each of the 200 separate data sets for each of the specified delta values within the constraints of I, J and alpha. The data sets are selected in a pseudo-random fashion from the specified distribution. Each member of the Ith row is then incremented by an amount equal to delta. This arbitrarily defines the Ith row as the extreme deviate with known mean and variance. The data set is then analysed by Nair's and Youden's tests and scores are kept of how often each of these tests, working at a specified alpha level and delta value, correctly identify the known extreme deviate. A percentile version of these scores, plotted against the delta values (0.0 to 3.0 by 0.2) produces

the power curves for Nair and Youden under identical conditions.

A source code listing of this simulation program can be found in Appendix B (4).

The actual curves were approximated by fitting a logistic function to the data generated by the simulation. The function was of the form

$$y = \frac{1}{1 + e^{-f(x)}}$$

where $f(x)$ represents a first, second or third degree polynomial. The transformation

$$f(x) = \ln\left(\frac{y}{1-y}\right)$$

is then subjected to a least squares fit for each of the first three orders of the polynomial. That function with the minimum sum of squared differences based on the original data was selected. Generally, the mean deviation of the data points from the function never exceeded three percent, and this was felt to be acceptable for a comparative demonstration based on a simulation. This also agreed with the maximum binomial standard error for $n = 200$. In every case, the order of the polynomial was also found to be the same for Nair and Youden, given that the same data had been evaluated.

Only the data for alpha equal to five percent are shown in the following tables, as the one percent values were found to be entirely comparable, but with the usual overall decreases in power.

3.3.2 A Comparison of Power Curves Derived From A Simulation of Normally Distributed Data

Figure 3.4 presents the power curves based on normally distributed data. Each graph presents the curves for a different value of I and for J equal to 5, 15 and 25. Where there is only one curve for a particular value of J, the power curves are identical and are both represented by the single line.

Several interesting observations can be made from the curves aside from using them to determine power. For the case of $I = 5$ and $J = 5$ we see that the non-parametric test is slightly superior and this is probably due to the low degrees of freedom (16) used in the parametric test. This superiority however is soon overcome as either I or J increases, and the tests show roughly equal power for any value of I or J greater than ten. The observation that power is generally independent of I except for very small values, is consistent with the findings of Willke. Siegel, also suggests that when N is small, the non-parametric test is usually preferred as the calculated probability values are usually exact and in this instance ($I = 5, J = 5$) such is the case.

Figure 3.4

Power Curves for Nair's and Youden's Tests Based on Normally Distributed Data

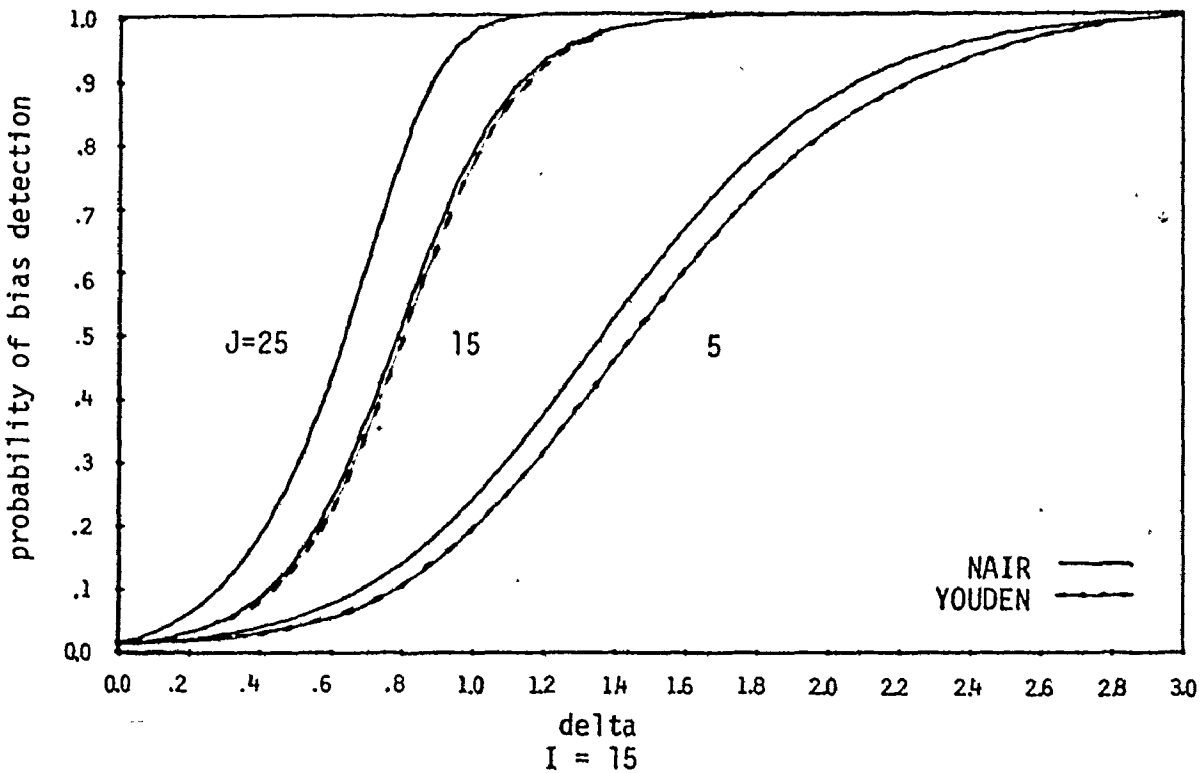
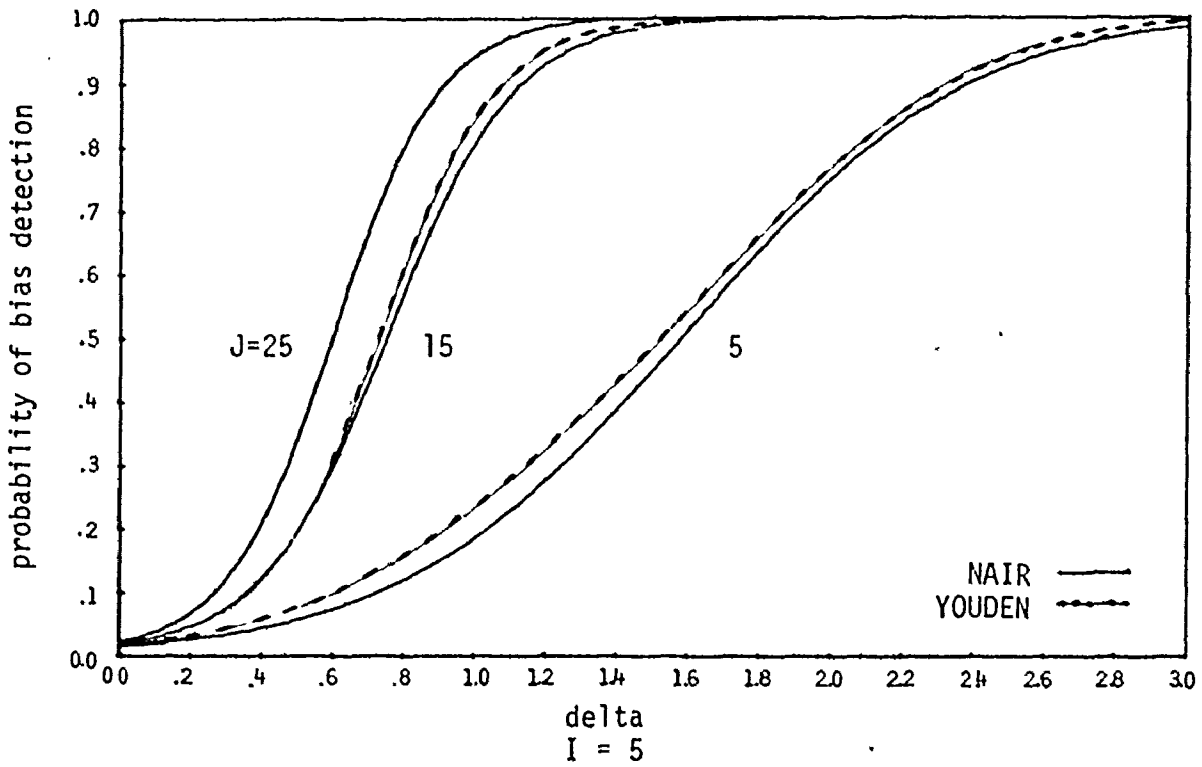
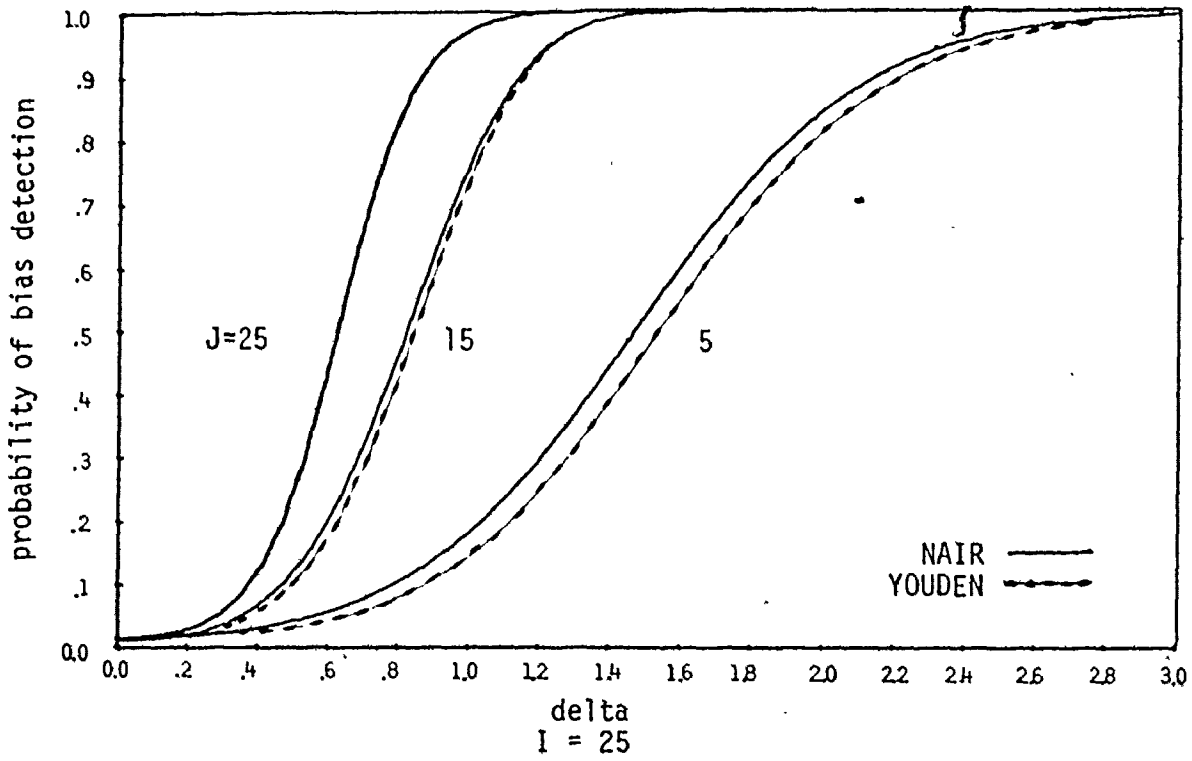


Figure 3.4
(continued)



3.3.3 A Comparison of Bias Detection in Non-Normally Distributed Data

Figure 3.5 presents the power curves based on uniformly distributed data. The uniform distribution was arbitrarily selected as a simple yet representative non-normal distribution.

The similarity of these power curves with those previously described is somewhat unexpected. Generally, it was felt that once the parent distribution strayed significantly from normality, that the non-parametric test would show superior power. However, this is obviously not the case. In fact, not only are the comparisons between parametric and non-parametric identical for both distributions, but the actual values of corresponding power curves are essentially the same. This may be accounted for in part by the fact that the uniform distribution and the normal distribution are both symmetrical and that the uniform is a rough approximation to a normal distribution with a large variance. Nevertheless, the apparent robustness of the parametric test is probably the single most impressive aspect of this comparison.

Figure 3.5

Power Curves for Nair's and Youden's Tests Based on Uniformly Distributed Data

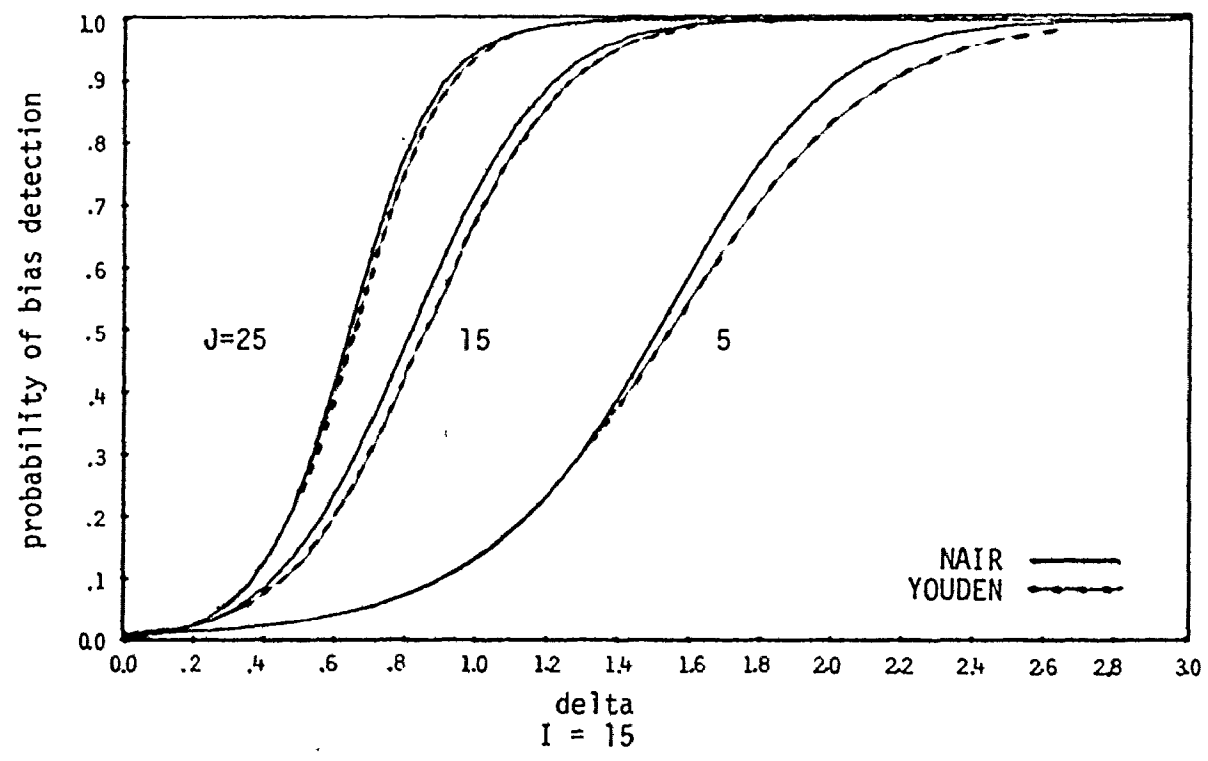
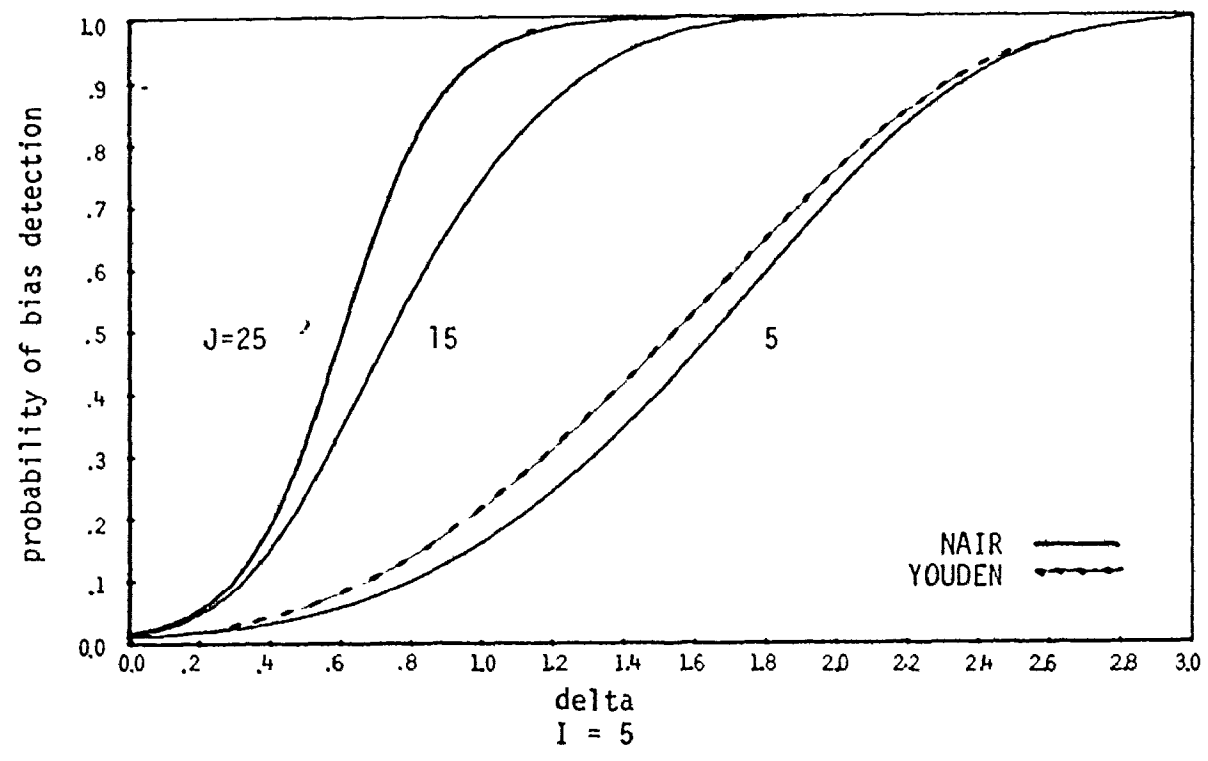
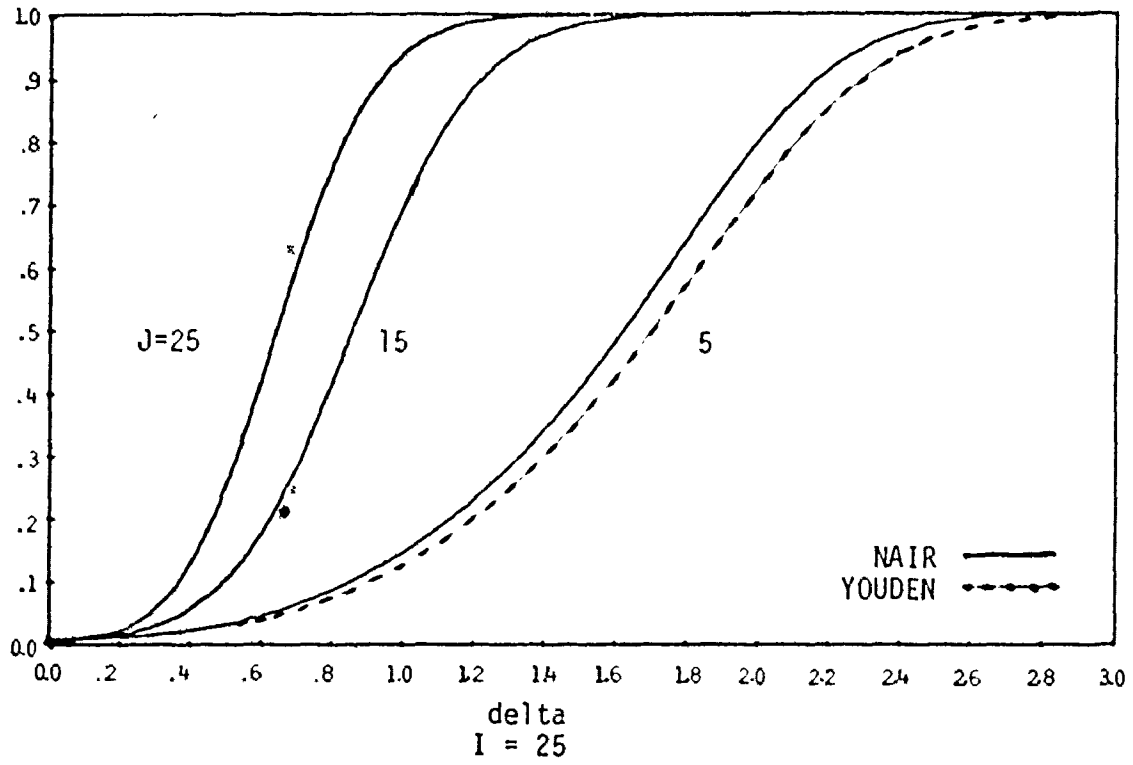


Figure 3.5
(continued)

3.3.4 A Comparison of Bias Detection in Non-Normally Distributed Data With Differing Variances

Figure 3.6 shows the power curves based on uniformly distributed data in which the extreme deviate has a variance which varies directly as the difference between the means. The previous two cases demonstrated a situation known as "slippage" where the extreme deviate has a parent population identical to the other samples, except that the mean is different. By relating the variance of the extreme deviate to the difference between the means (δ) one is able to look at a more complex and probably more realistic situation found in extreme value analysis.

Looking at the power curves, two things are immediately apparent. First, the power curves have a lower "slope" than the previous two cases, indicating that the power-efficiency is decreased. This is predictable, since the overlap between the two distributions decreases more slowly as δ increases, due to increasing variance of the deviant population and consequently it is more difficult to distinguish between distributions.

Second, we see that the relationship between members within a given pair has not changed from the previous two cases. For $I = 5$ and $J = 5$, the non-parametric test is still slightly better, and in all other cases the two tests are either identical or the parametric test shows a slight improvement. Again, the robustness of the parametric test is evident.

Power Curves for Nair's and Youden's Tests Based on Uniformly Distributed Data with Variable Variance

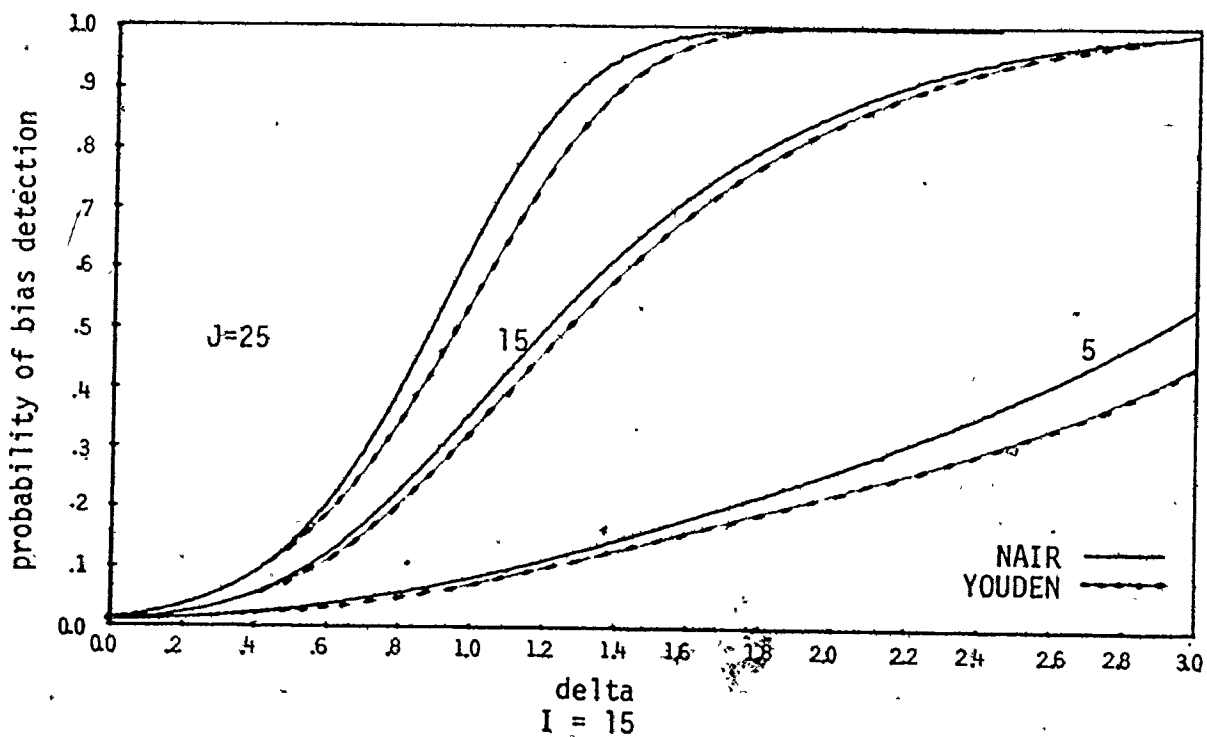
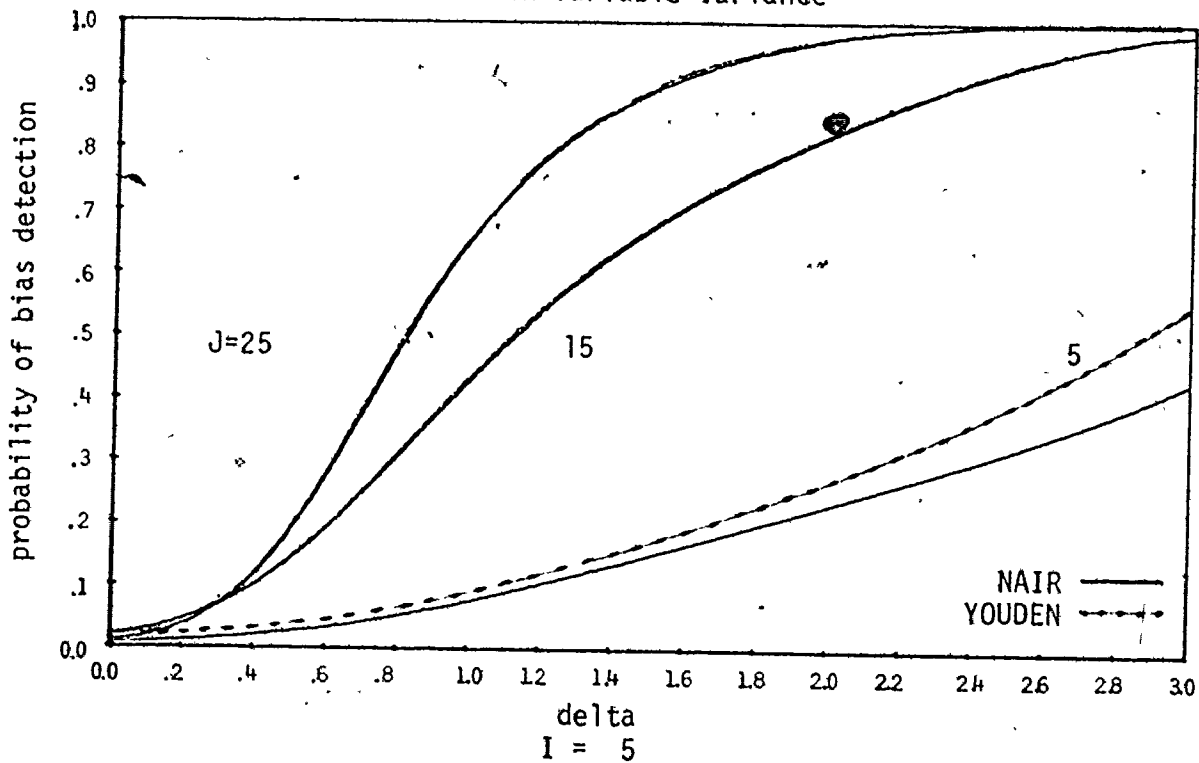
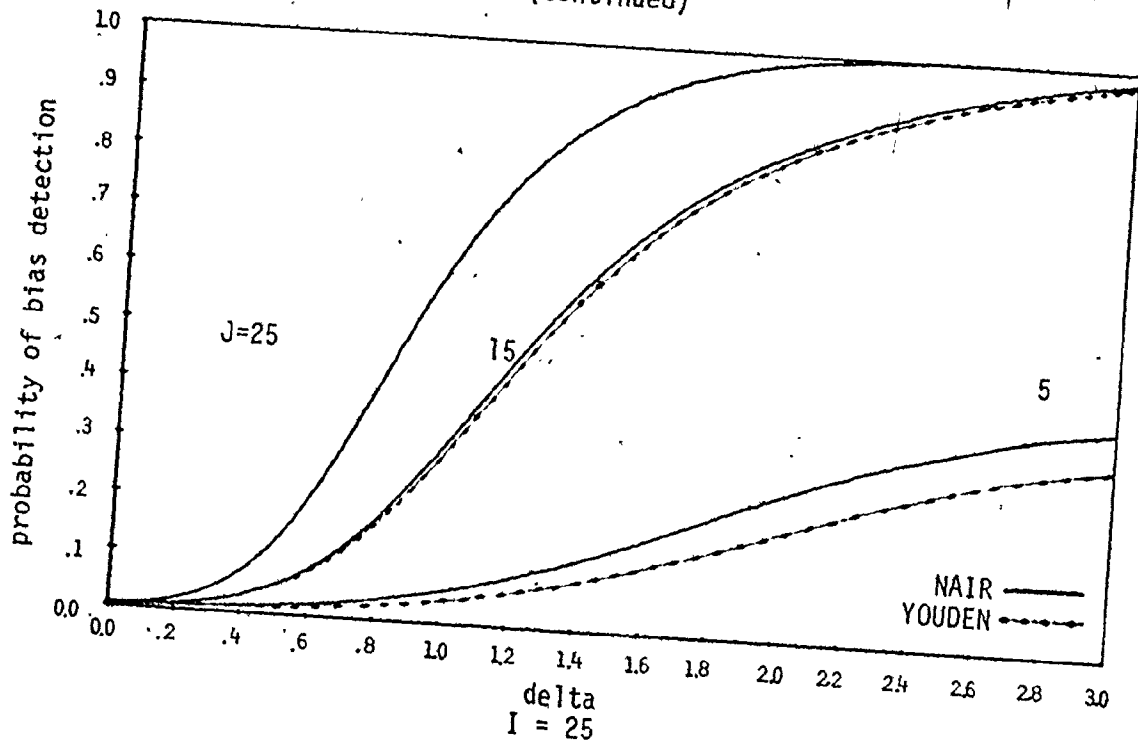


Figure 3.6 ,
(continued)



CHAPTER IV

AN ANALYSIS OF CHROMOSOMAL CHARACTERISTICS OF HUMAN MELANOMA

4.1 Introduction

Many experiments in cancer research involve the artificial growth of different types of cancerous tissue and the examination of the cells during this growth. The following discusses such an experiment after a brief description of some of the less common terms and concepts.

With the exception of cells used in sexual reproduction (sperm and ova) all normal human cells contain twenty-three pairs of chromosomes which are confined within the nucleus. These twenty-three pairs of chromosomes are classified as autosomes (22 pairs, in which the two members of each pair are visually very similar) and sex chromosomes, (1 pair, the female having two X chromosomes and the male having one X and one Y chromosome). Within the human body, these cells are able to reproduce by a process known as mitosis which results in two daughter cells for each parent cell. During one of the stages of mitosis, called the metaphase, the chromosomes, which are normally a tightly tangled mass within the nucleus become separated and clearly visible under magnification. In the laboratory, it

is possible to cause certain types of human cells to reproduce artificially in a controlled environment (in vitro). It is also possible to stop reproduction at the metaphase, and after certain preparation, the chromosomes are visible for examination. One technique often used is to photograph a cell's chromosomes with suitable magnification, produce a black and white print, then cut out the chromosomes and arrange them into the twenty-three pairs for further examination. This orderly arrangement of pairings is known as a karyotype.

McCulloch et al (1976) have recently completed an experiment designed to select specific chromosomal characteristics of cultured human melanoma. This study was undertaken to characterize several cell lines so that "in-house" immunological investigations could be performed on pure cultures from a defined origin.

Basically, the experiment consisted of culturing eight different strains of human malignant melanoma and then karyotyping ten of the best of about fifty metaphases from each of the eight strains. Analysis was carried out in two areas. First, within strains, chromosomes were examined for any unusual, strain specific marker conditions. In this case, a marker condition refers to the consistent occurrence of a chromosome which could not be normally classified. Such a unique occurrence, if valid, could be used to identify this strain in other situations. Second, the frequency

of specific chromosomes was examined across strains. The objective in this case was to determine if there was a specific chromosome which occurred much more or less frequently than the others for melanoma, as represented by the eight selected strains. It is this second objective which can benefit from extreme value analysis and was, in fact, the exact problem which suggested this thesis.

4.2 Description of the Data

Table 4.1 shows the total counts of normal chromosomes found in ten cells at metaphase for each of the eight selected strains of melanoma. Each row represents a particular autosome as described in the first column. The sex chromosomes were not used in the analysis, as the sex of the original donors was not known for some of the strains. The strains are represented by the columns and are identified in the first row. The column at the right shows the rank sum for each chromosome. The chromosome counts were ranked from 1 to 22 within each strain and the ranks were summed across strains for each of the chromosomes. To maintain clarity, the individual ranks are not shown.

4.3 Analysis and Results

The objective, as stated by McCulloch et al (1976), was to determine if one or more of the chromosomes occurred significantly more or less than the others. If the findings

CELL LINES CHROMOSOME	M-1	M-2	M-3	M-4	M-5	M-6	M-7	73-61	RANK SUM
1	9	29	21	26	26	48	9	17	93
2	29	22	19	30	18	60	16	20	105
3	5	28	20	32	22	64	27	20	118
4	18	18	12	32	8	30	22	15	61
5	24	22	19	22	16	45	29	10	90
6	22	20	16	28	22	39	29	15	91
7	10	42	32	40	30	71	42	25	157
8	26	16	16	29	17	31	26	10	68
9	19	24	19	11	20	44	30	13	92
10	8	17	15	20	14	36	20	10	40
11	13	29	14	11	18	35	18	25	74
12	10	23	24	32	27	42	28	20	115
13	26	31	20	24	19	25	9	10	76
14	20	28	24	39	22	42	23	20	124
15	11	12	26	20	22	30	20	17	70
16	13	25	18	18	15	35	15	17	60
17	17	13	11	6	11	38	23	22	57
18	19	13	18	24	20	50	22	20	87
19	16	15	17	32	16	44	11	17	71
20	22	27	17	28	23	60	26	10	105
21	7	28	26	30	21	34	39	30	112
22	32	31	31	40	31	66	28	17	158

Table 4.1

Total Number of normal chromosomes found in ten cells at metaphase for each of the eight selected strains of melanoma

were positive, then the relative frequency of occurrence of these particular chromosomes could be used to help classify unknown melanoma tissue.

In the case of non-parametric analysis, the rank sums as shown in Table 4.1 have been chosen as the test statistics. A visual representation of the rank sums is shown in Figure 4.2. A real number line is drawn and the integer values from 20 to 160 by 10's are marked. Each of the twenty-two rank sums is located on this line with a vertical mark. In the case of identical rank sums, the marks are placed very close together about the actual rank sum.

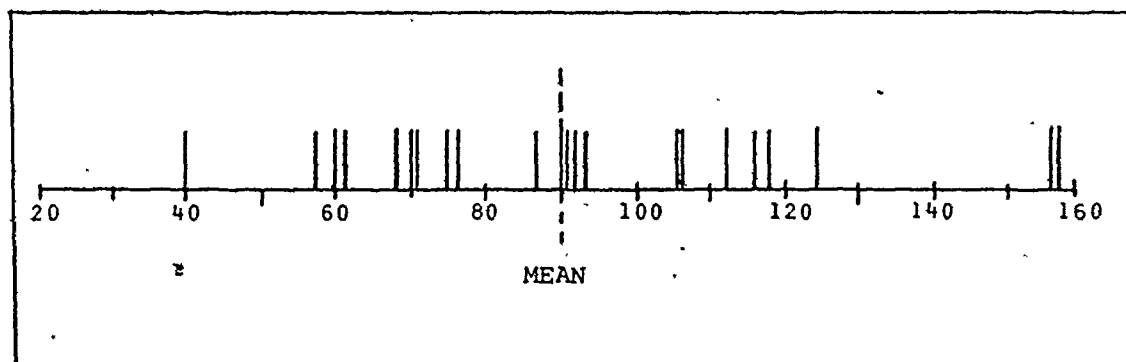


Figure 4.2

The distribution of rank sums for all chromosome counts in the melanoma data

This representation of the rank sums allows the investigator to get an intuitive feel for the rank sum distribution and an expectation with regard to those values which will

be selected as extreme.

Both Youden's test and Nair's test were used to analyse Table 4.1 for the occurrence of extremely low or high chromosome counts and the results are shown in Table 4.3.

CHROMOSOME	DIRECTION OF EXTREMENESS	YOU DEN		NAIR					
		SIGNIF	R _s	ROWS	COLS	SIGNIF	d	n	df
10	MIN	0.03	10	22	8	0.08	2.59	22	147
7	MAX(1)	<<0.01	158	22	8	<0.001	5.21	22	147
22	MAX(2)	<<0.01	177	21	8	<0.001	4.73	21	140

Table 4.3

Analysis of melanoma data for extreme deviate chromosome counts, using Youden's and Nair's tests

Chromosomes which are felt to be extreme ($p < 0.05$) are listed in column one and column two indicates which tail of the distribution was examined. With regard to Youden's test, the significance value of the selected row and its rank sum (R_s) are shown along with the number of rows and columns in the data set. For Nair's test, the significance value is also shown as is the test statistic (d), the number of objects under consideration (n) and the degrees of freedom (df).

A technique for determining second, third, ..., etc. most extreme deviates was also employed and this is the

reason why two maxima are shown in Table 4.3. After determining the most extreme deviate in the maximum direction, this row of data was then deleted from the data set and the identical calculation was again performed on the modified data set. The two maxima shown are ordered (1) and (2) with (1) being the most extreme. It also accounts for the decrease in the number of rows, n , and degrees of freedom in the bottom line of Table 4.3.

A two way analysis of variance was performed on the complete data set (Table 4.4) for comparison with the results of Nair's test, and Friedmans χ^2 was calculated for comparison with Youden's test.

SOURCE OF VARIATION	DEGREES OF FREEDOM	SUM OF SQUARES	VARIANCE	F
CHROMOSOMES	21	4031.41	191.97	3.99
STRAINS	7	11753.	1679.	34.91
ERROR	147	7069.78	48.09	
TOTAL	175	22854.19	-	-

Table 4.4

Analysis of variance for melanoma data

Looking at figure 4.2 it appears as though there are two maximum extreme deviates and possibly one minimum extreme deviate according to the distribution of the rank sums. The F-test indicates very strongly (Table 4.4) that the chromosome counts are not all from the same distribution

($p < 0.001$), and Friedman's $\chi^2 = 58$ ($p < 0.001$) is in agreement. Table 4.3 indicates that Youden's test and Nair's test agree in selecting the extreme deviates but differ slightly on the level of significance for the extreme minimum, as shown in the first row of the table.

4.4 Conclusions

Interpreting figure 4.2 visually, three different chromosomes are of immediate interest, one occurring less frequently than the others and two which both occur considerably more frequently than the rest. Both the F test and Friedman's χ^2 indicate very strongly that the mean counts for the different chromosomes come from more than one distribution. Nair's test and Youden's test are in agreement with the visual inspection and with the analysis of variance. They also agree with one another as to order and direction of extremeness for the three outlying chromosomes and do not differ appreciably as to the significance level of each.

Given the high degree of consistency among the techniques used and the associated high probability values, the conclusion must be that chromosome numbers seven and twenty-two occur much more frequently than all others for these eight strains of melanoma. Also, chromosome number ten occurs less frequently than all others, although the confidence in this statement is not as great as for the

maximum extreme deviates.

Chromosomes seven, twenty-two, and to a lesser extent, ten could then be used, by virtue of their relative frequencies, as marker chromosomes indicative of melanoma as defined by this selection of cell-lines.

CHAPTER V

AN ANALYSIS OF SCORING PATTERNS AMONG EVALUATORS OF LETTERS OF APPLICATION TO MEDICAL SCHOOL

5.1 Introduction

Several medical schools in Canada have recently modified their admission procedures by using some non-academic data, as well as the traditional marks of academic achievement to choose their new students. This change stems from the concept that, given a basic level of academic ability, other characteristics of the applicant may be equally important. These other characteristics are often referred to as personal qualities.

The primary tool used to collect information about these personal qualities at McMaster University is a letter written by the applicant about himself. The letter is up to eight hundred words long and should attempt to answer the questions posed by the medical school. These questions are purposefully left vague and simply ask, "describe who you are, where you are going and how you will get there".

The letters are assessed by three readers working independently. There are commonly fifty teams, each of three readers, to process a total of about twenty-two hundred letters. Each team consists of a faculty person, a

medical student and a member of the community; each member having been randomly selected. This three member concept is repeated at each step of the admission process in the belief that admission decisions are not the sole prerogative of the faculty.

The scoring of the letter is centered around a form which first asks the reader to assess the presence or absence of a set of personal qualities as shown by the letter. The tabulation of these personal qualities assists the reader in selecting a global score which ranges from 1 (poor) to 4 (excellent). No attempt is made to formalize the assignment of a global score from the personal qualities, nor to imply that other factors should not be considered. The sum of the global scores from the three readers of each letter is used in selecting candidates for interview, which is the second step in the admissions procedure.

In an effort to compare readers, five "control" letters are sent to every reader in such a way that they cannot be distinguished from the other letters they are asked to read. The "control" letters are selected from all applicants' letters to provide a representative cross-section. One letter is selected which should be rated very high, another should be scored very low, a third letter is selected because it is highly controversial (readers will tend to score this letter very high or very low) and two other letters are selected which are average.

The readers are non-expert at this method of evaluation and change from year to year. Also, they donate their time and come from all parts of the community near the medical school. With this in mind, evaluation of the readers is vital to ensure that applicants are unlikely to be incorrectly scored by a poor reader or team of readers.

Before an analysis of the readers and teams can be carried out, several assumptions must be made.

1. The best estimate of an applicant's score is the mean score of all readers who read the letter.

2. The overall mean score (3 readers reading 45 randomly assigned letters) should be approximately the same for all teams.

3. The individuals and teams giving the most extreme results with the control letters are most likely to give the most extreme results generally.

With the above assumptions, extreme value analysis can be used to select individual readers and teams which deviate most from the mean (best estimate of the true value) and are therefore most likely to score applicants incorrectly.

5.2 Description of the Data

Table 5.1 lists the forty-eight teams and the total scores for the five control letters by team member. Also shown are the team totals and the totals for each type of

TEAM	FACULTY	STUDENT	COMMUNITY	TOTAL
1	11	14	14	39
2	14	11	12	37
3	11	12	10	33
4	13	11	10	34
5	11	14	11	36
6	9	10	LO 7	LO 26
7	13	11	11	35
8	15	11	9	35
9	13	13	11	37
10	9	10	15	34
11	HI 17	13	16	HI 46
12	11	12	8	31
13	13	10	11	34
14	12	13	11	36
15	13	10	13	36
16	10	9	9	28
17	11	11	10	32
18	11	9	HI 17	37
19	13	8	HI 14	35
20	10	13	11	34
21	10	14	11	35
22	13	8	10	31
23	12	14	10	36
24	9	8	13	30
25	13	9	12	34
26	13	15	14	42
27	12	9	12	33
28	11	13	11	35
29	11	11	12	34
30	10	11	14	35
31	13	10	13	36
32	14	10	13	37
33	13	9	13	35
34	13	13	11	37
35	13	12	9	34
36	12	11	13	36
37	12	14	13	39
38	LO 8	11	9	28
39	12	14	9	35
40	11	10	11	32
41	12	13	11	36
42	11	10	15	36
43	12	9	15	36
44	10	12	9	31
45	10	12	13	35
46	10	HI 16	11	37
47	9	LO 6	14	29
48	11	9	11	31
TOTAL	560	538	562	1660

Table 5.1

Total scores on five control letters given by teams and reader types

TEAM	MEAN SCORE	TEAM	MEAN SCORE	TEAM	MEAN SCORE
1	7.02	17	6.28	33	7.48
2	6.98	18	HI 8.10	34	6.88
3	6.39	19	7.41	35	6.93
4	7.80	20	6.72	36	6.27
5	6.26	21	7.40	37	7.28
6	LO 5.27	22	6.95	38	6.02
7	7.30	23	6.41	39	6.45
8	6.82	24	7.24	40	6.50
9	6.64	25	7.02	41	5.79
10	6.65	26	7.42	42	7.20
11	7.93	27	6.55	43	7.63
12	6.90	28	7.07	44	6.75
13	6.19	29	7.43	45	7.17
14	7.00	30	6.56	46	6.41
15	7.28	31	7.91	47	6.49
16	7.07	32	5.75	48	6.15

Table 5.2

Mean scores for each team based on all letters read

team member, i.e. student, faculty, community. Although the individual scores of each letter by reader were available, they are not shown because of their great number and minimal use in this discussion. Table 5.2 lists the mean scores of the actual data for the same forty-eight teams. Each team read between forty-one and forty-six letters and the mean is shown in this table to standardize for the number of letters read.

5.3 Analysis and Results

The evaluation of the readers' performance was considered at two levels; the individual reader, and the team. It is most important to know that each of the teams is evaluating correctly, for it is the total, or team score, which is used to determine whether or not the applicant should be considered further. The measurement of the individual reader's performance is of secondary importance but should be useful in determining weaknesses in the training of the readers or in pinpointing the reason for a poor team performance. Using extreme value analysis, precision and accuracy of both individuals and teams will be measured to determine if any significantly extreme deviates exist for these particular data. Table 5.3 shows Table 5.1 in the form of three histograms. The histograms represent the total scores given to the five control letters by each of the readers. The readers are separated by type

into faculty, student and community. The results of analysis by Youden's and Nair's tests are shown in Table 5.4. The faculty member of team 11, and the community member of team 18 both appear to be scoring significantly higher than all other members of their type. These particular individuals are indicated in Figure 5.3 by the hatched areas. It is interesting to note that there are no significantly extreme ($p < 0.05$) deviates in the minimum direction. The student from team 47 definitely appears to be extreme according to Figure 5.3 and Table 5.1. Nair's test even gives a significance level of 0.02 but this could be caused by the more or less consistent skew to the right shown in these data. Youden's test should not be affected by the nature of the distribution and indicates a significance level of > 0.10 .

These results show that two of the individual readers are significantly different from others of their group with respect to accuracy in determining the "true" scores of the five control letters. It is also useful to note that both of these individuals tend to err on the high side of the "true" value.

The other important measure of performance in individual readers is that of precision. In terms of correctly evaluating a letter, it is not so important to measure the spread about an individual reader's mean, but to measure the individuals' spread about the group mean or best

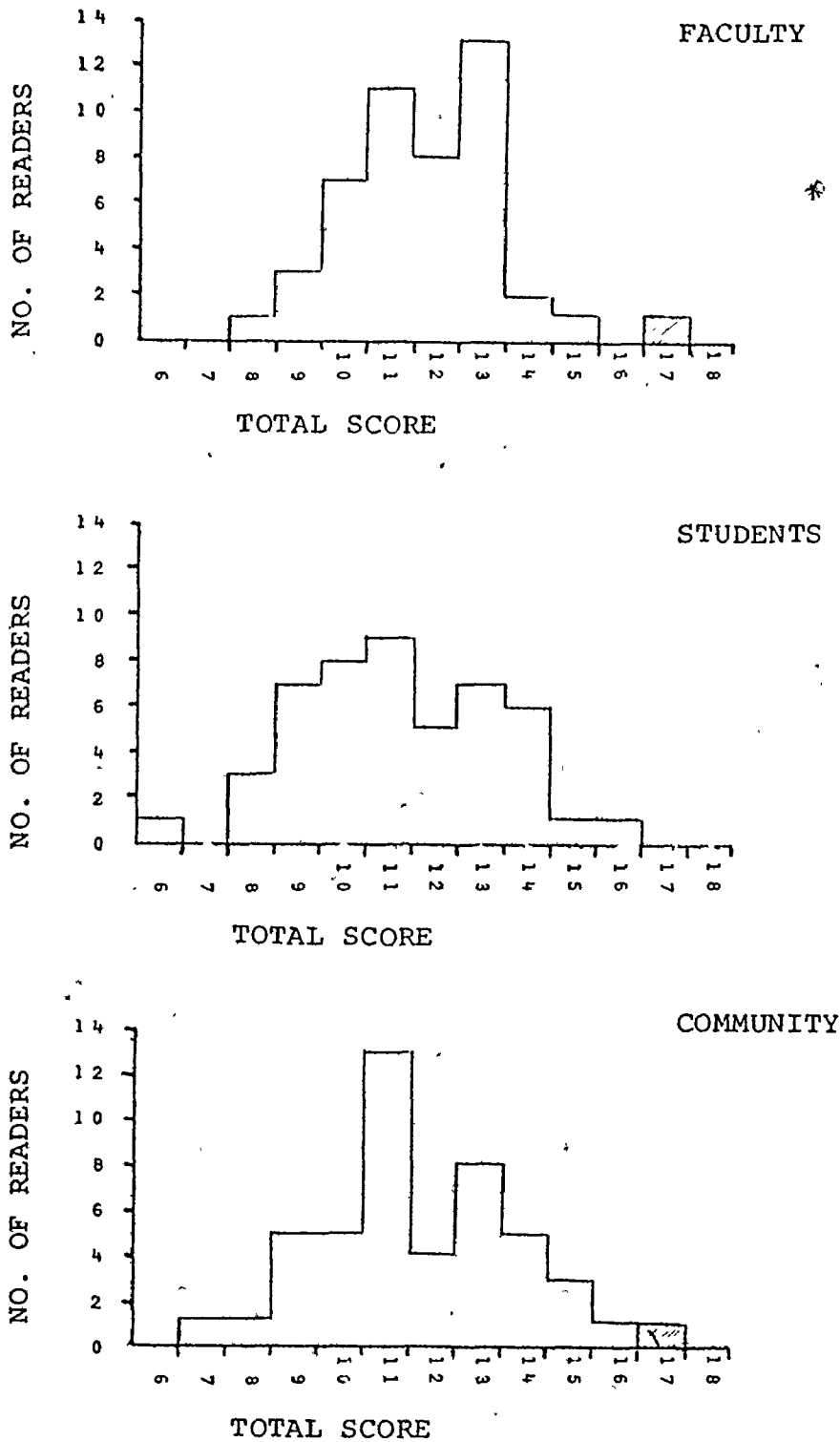


Figure 5.3

Frequency of total scores for 5 control letters by type of reader

TYPE OF READER	MAXIMUM DEVIATE			MINIMUM DEVIATE		
	TEAM	*Y	*N	TEAM	*Y	*N
FACULTY	11	0.03	0.02	38	>>0.10	>>0.10
STUDENT	46	0.10	0.03	47	0.10	0.02
COMMUNITY	18	0.05	0.03	6	>0.10	0.07

(*Y = YOUDEN, N = NAIR)

Table 5.4

Maximum and Minimum extreme deviates in the three types of readers (Mean Scores)

estimate of the true value. This approach is able to point out not only those who are consistently high or low scorers, but also those who for some reason always seem to score in the opposite direction to everyone else. This type of reader would score good applicants as poor and poor applicants as good and is therefore unlike those who score everyone as good or poor. Table 5.5 shows the absolute differences between the group total and the individual total for each reader. Frequency histograms of these data are found in Figure 5.6. These data are then analysed by Nair's test and Youden's test and the results are displayed in Table 5.7.

The histograms displayed in Figure 5.6 indicate a definite skewing to the right in all three cases which probably invalidates Nair's test as a method of analysing these data. Visually, there appears to be at least one suspect extreme deviate in each case but the results shown

TEAM	FACULTY	STUDENT	COMMUNITY	TOTAL
1	1.84	2.72	3.17	7.73
2	2.72	1.96	3.17	7.85
3	1.84	LO 1.42	3.30	6.56
4	2.42	3.12	3.75	9.29
5	LO 1.58	2.72	1.58	LO 5.88
6	2.54	2.96	4.54	10.04
7	4.88	3.58	3.59	12.05
8	4.17	3.58	4.75	12.50
9	2.42	1.72	2.75	6.89
10	3.84	4.12	4.17	12.13
11	5.47	2.42	4.72	12.61
12	HI 6.05	1.42	HI 6.84	HI 14.31
13	3.59	4.12	3.84	11.55
14	2.33	3.96	1.58	7.87
15	3.33	4.12	1.72	9.17
16	2.84	4.30	3.12	10.26
17	2.42	3.59	2.38	8.39
18	3.14	3.38	5.47	11.99
19	1.26	3.54	3.17	9.97
20	2.00	2.42	3.12	7.54
21	3.30	3.17	1.84	8.31
22	2.17	4.84	2.12	9.13
23	2.33	2.47	2.12	6.92
24	3.12	3.54	3.14	9.80
25	2.89	4.30	2.96	10.15
26	3.96	3.72	3.17	10.85
27	1.88	3.12	2.14	7.14
28	3.14	2.63	2.42	8.19
29	1.58	1.84	2.71	6.13
30	2.00	1.58	3.17	6.75
31	2.17	2.96	3.71	8.84
32	2.47	2.58	3.72	8.77
33	2.42	2.54	2.42	7.38
34	3.33	3.71	3.59	10.63
35	2.42	2.96	2.54	7.92
36	2.14	1.58	2.42	6.14
37	1.88	3.17	2.89	7.94
38	3.54	2.75	3.00	9.29
39	2.14	2.72	3.12	7.98
40	3.14	2.00	3.13	8.27
41	3.87	2.17	LO 1.58	7.62
42	1.58	2.58	3.72	7.88
43	2.33	3.12	3.47	8.92
44	2.84	2.58	3.00	8.42
45	3.30	1.42	1.72	6.44
46	3.68	4.47	1.96	10.11
47	3.38	HI 5.54	2.47	11.39
48	2.75	3.12	4.05	9.92

Table 5.5

Absolute differences between group total and individual total
by reader

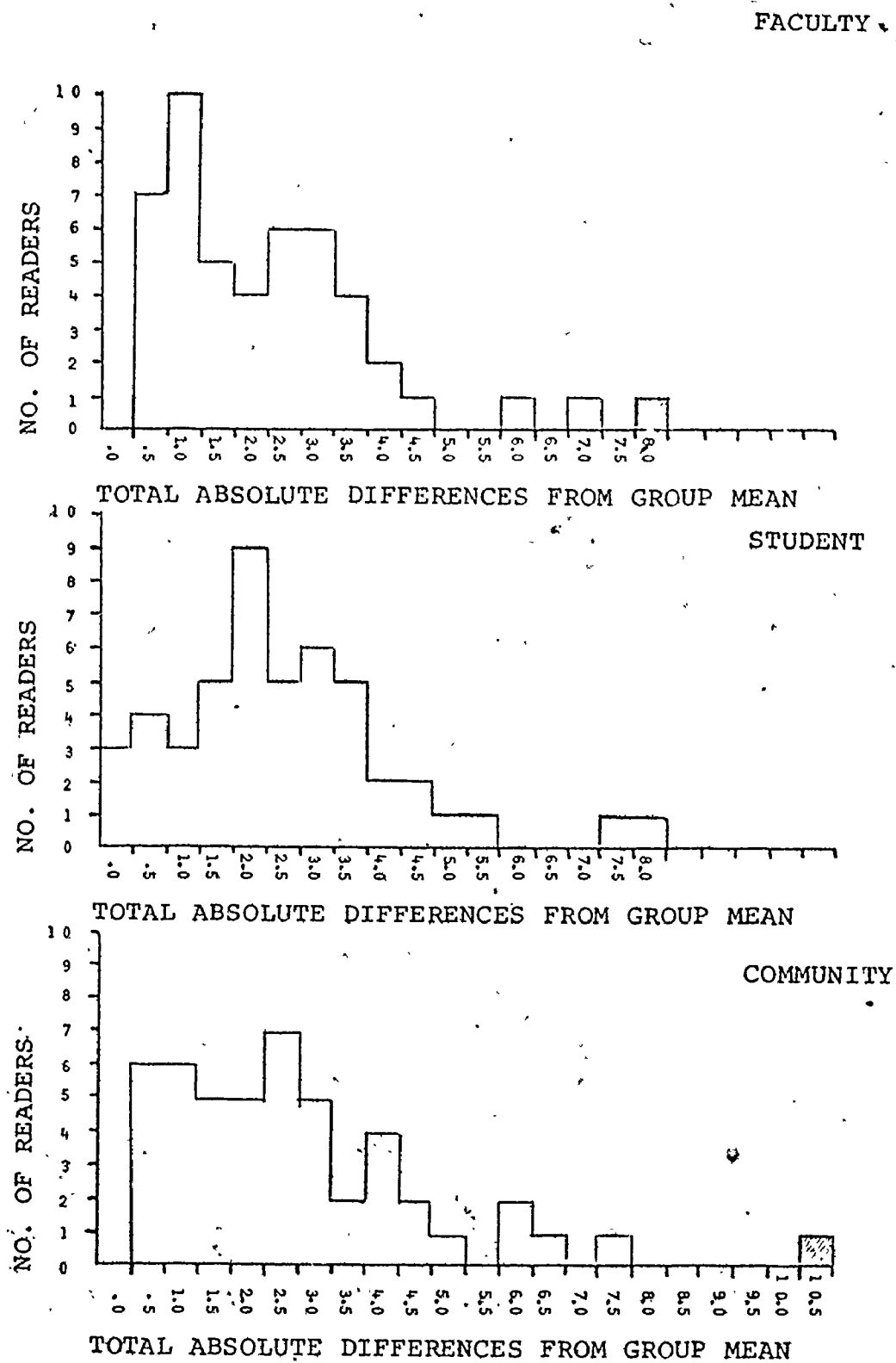


Figure 5.6

Frequency of total absolute differences from group mean for 5 control letters by type of reader

TYPE OF READER	MAXIMUM DEVIATE			MINIMUM DEVIATE		
	TEAM	*Y	*N	TEAM	*Y	*N
FACULTY	12	0.01	0.002	5	>>0.10	>>0.10
STUDENT	47	>0.10	>0.10	3	>>0.10	>>0.10
COMMUNITY	12	0.03	0.003	41	>>0.10	>>0.10

(*Y = YOUDEN, N = NAIR)

Table 5.7

Minimum and Maximum extreme deviates in the three types of readers
(absolute differences from the group mean)

in Table 5.7 indicate that this is not exactly true, keeping in mind that Nair's test is weakened by the skewness of the data. Youden's test indicates that both the faculty member of team 12 and the community member of team 12 are significantly extreme from the other members of their groups in respect of the degree of dispersion of their scores about the group mean. Again, there are no extreme members of the student group which appears to contradict a visual interpretation, but this is probably due to the mean student variation being somewhat higher than that in the other two groups of readers. This would tend to keep values of 7.5 - 8.0 more within the body of the distribution.

The selection of the faculty member from team 12 as an extreme variance deviate is of particular interest because it points out that readers can be found who score high when all others score low and vice versa. This reader's

total score for all five control letters is 11, which is very close to the group mean of 11.7 indicating that there is no significant difference between this reader and the rest. However, an analysis of the total absolute difference between the group means and the individual reader's scores shows this reader to be significantly different from all other readers. On the other hand, the community member from team 12 whose dispersion measure about the group mean was also significant, can be recognized in the first analysis as having had the second lowest total score.

The analysis of the reading teams is carried out in much the same way as the analysis of individual readers. Figure 5.8 shows the distribution of the total scores given by each team for each of the control letters and Table 5.9 summarizes the analysis of these data by Nair's and Youden's tests. Both tests select team 11 as the maximum extreme deviate and team 6 as the minimum extreme deviate and their relative locations in Figure 5.8 are indicated by shading. Team 26, while appearing suspect as a second most extreme deviate has a significance value in excess of 0.10 according to both tests and is therefore not considered as extreme.

A tabulation of the absolute differences between group mean total and team total was also carried out, and the results are displayed as a histogram in Figure 5.10. The results of analysis for extreme difference show team 11

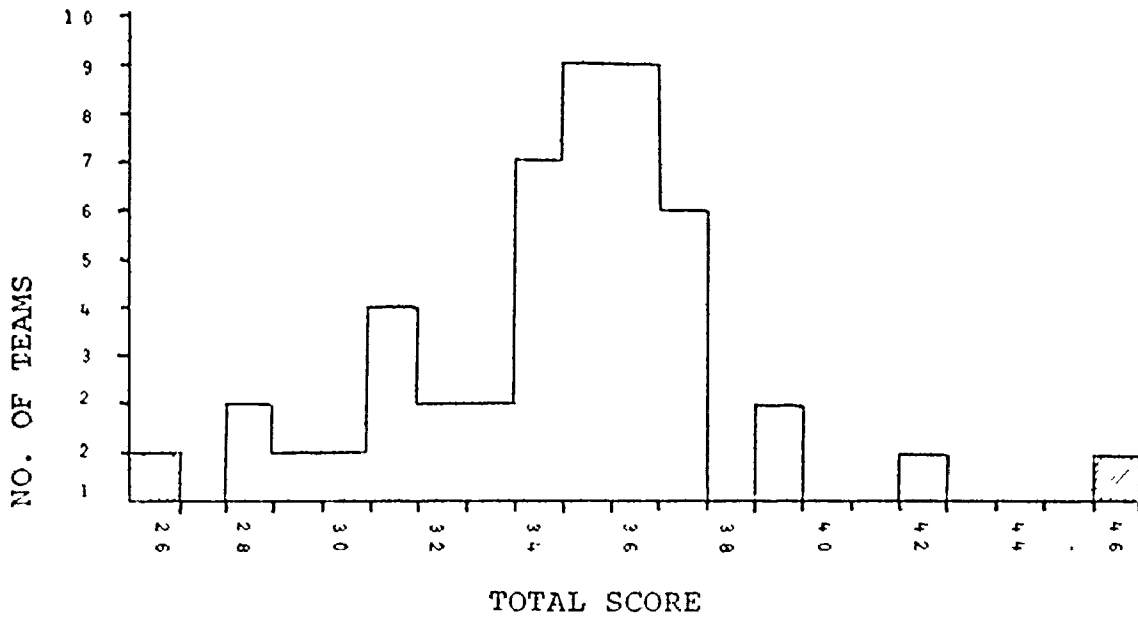


Figure 5.8

Frequency of total scores for 5 control letters (all teams)

MAXIMUM DEVIATE			MINIMUM DEVIATE		
TEAM	YOUDEN	NAIR	TEAM	YOUDEN	NAIR
11	<0.01	<0.001	6	0.01	0.01

Table 5.9

Maximum and minimum extreme deviates in reading teams
(Mean Scores)

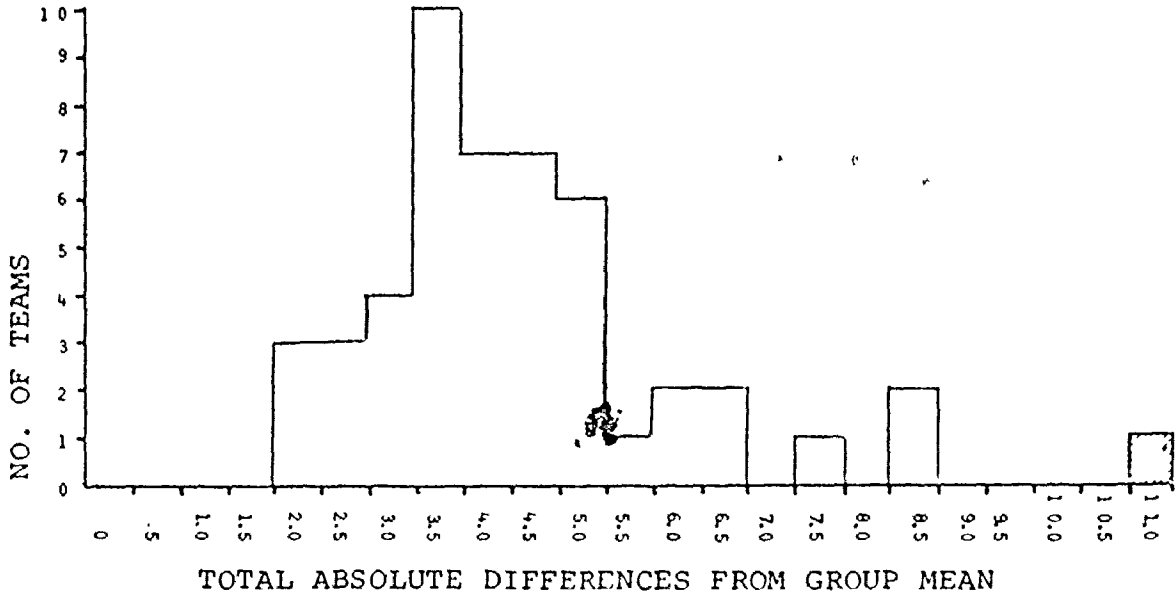


Figure 5.10

Frequency of total absolute differences from group mean for 5 control letters (all teams)

MAXIMUM DEVIATE			MINIMUM DEVIATE		
TEAM	YODEN	NAIR	TEAM	YODEN	NAIR
11	0.05	<0.01	36	>>0.10	>>0.10

Table 5.11

Maximum and minimum extreme deviates in reading teams (dispersion about the group mean)

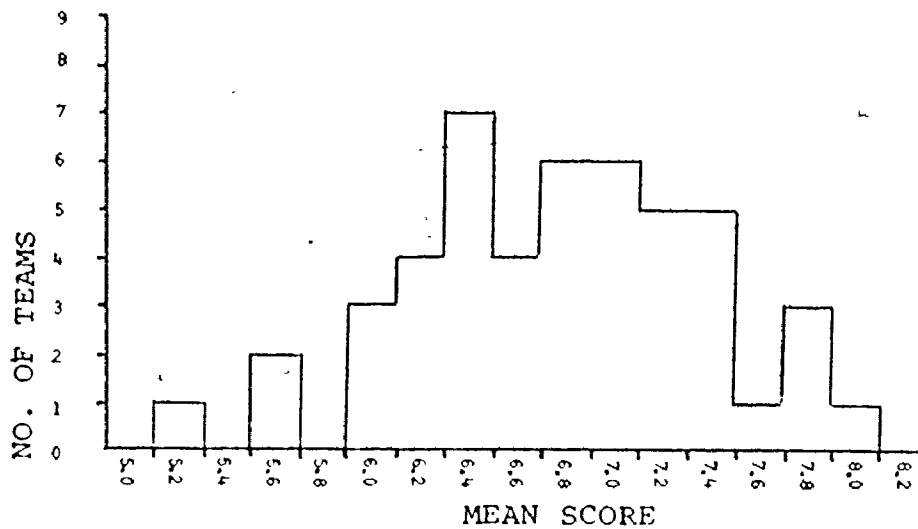


Figure 5.12

Frequency of mean team scores for all letters

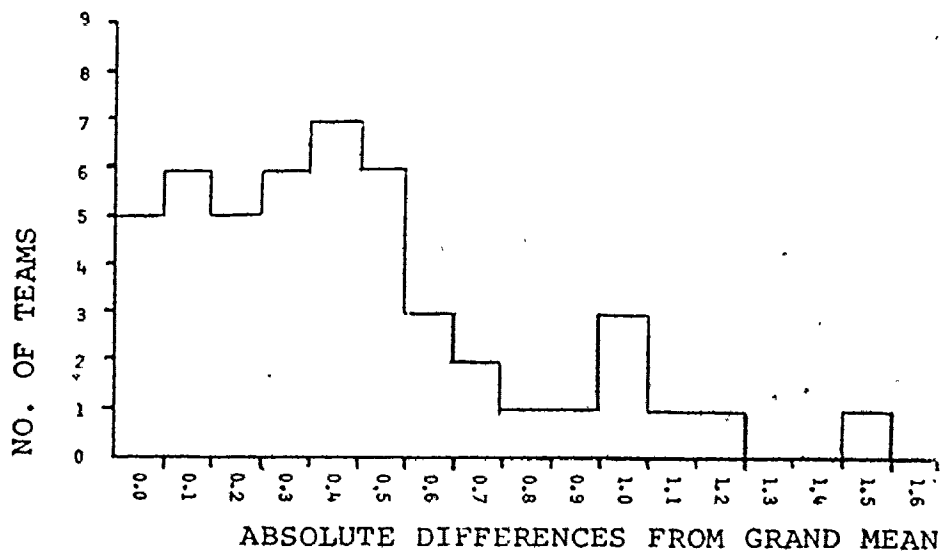


Figure 5.13

Frequency of absolute differences from the grand mean for all letters, by team

to be the significant maximum deviate with no significant minimum deviate. Analysis of the data for second most extreme deviate in the maximum direction failed to show significance.

Now that the analysis of the control letter data is complete it is worthwhile examining the complete set of letter scores to see how the deviant teams performed overall. Figure 5.12 displays the frequency of mean team scores for all letters and Figure 5.13 shows the frequency of absolute differences from the grand mean for the same data. Table 5.14 summarizes the teams which were extreme in evaluation of both control letters and total letters.

ANALYSIS	CONTROL LETTERS	ALL LETTERS
MAXIMUM MEAN	team 11	team 18
MINIMUM MEAN	team 6	team 6
MAXIMUM * $ \bar{X}_g - X $	team 11	team 6
MINIMUM * $ \bar{X}_g - X $	none	none

* absolute difference of the team score from the group team

Table 5.14

Summary of teams showing extreme deviance in control letter data and comparison with all letter data

Team 18 is shown to have the maximum mean score for all letters and considering that this team was very close to the mean for both measures of the control letter data, it

is reasonable to assume that the letters allocated to this team were unusually good. Team 11, which had the highest mean with the control letters also had the third highest mean for all letters which suggests that team 11 is probably the highest scoring team. On the other hand, team 6, consistent in both control letters and all letters, is definitely the lowest scoring team. The problem now is to decide which team will take honours for being most deviant from the group mean; high scoring team 11 or low scoring team 6. Team 6 was most deviant in the all letters category and third most deviant in the control letter evaluation while team 11 was most deviant in the control letter study and about average for all letters. Undoubtedly team 6 is the most deviant team from the group mean.

5.4 Conclusion

From the results obtained in this chapter, it is obvious that extreme value analysis can be a useful tool in evaluating letter readers and teams in the context of a medical school admissions procedure. The use of control letters allows for comparison of readers and the results can be used to reasonably predict team performance in the sense of extreme variation. ~~The analysis of absolute differences from the group mean is superior to the analysis of mean scores in the determination of extreme deviates,~~ because extreme value analysis of these measurements will

not only pick out extreme means (consistently low or high) but will also select those who are extremely different from the group (consistently different in the extreme from the group mean). Finally, Youden's test is probably more useful in analysing these data than Nair's test, because of the skewed nature of the distribution of the absolute difference from the group mean.

CHAPTER VI

THE DETERMINATION OF OUTLIERS IN A STUDY OF VENTILATORY RESPONSE TO CHANGES IN OXYGEN AND CARBON DIOXIDE CONCENTRATION

6.1 Introduction

Of particular interest to respirologists is the fact that increasing levels of carbon dioxide and decreasing levels of oxygen in the blood, will stimulate a person to increase their ventilation. This may occur either by a change in the depth (an increase in the tidal volume), or by an increase in the frequency at which they are breathing. In the actual situation, most people will exhibit a response which is mixed with respect to an increase in tidal volume and also an increase in frequency.

The rate at which ventilation increases with either increasing levels of carbon dioxide or decreasing levels of oxygen, exhibits a wide range of variation among subjects. It is possible to regress the change in ventilation against either changing partial pressures of carbon dioxide in the blood or changing levels of oxygen saturation. Within the biologic range of either carbon dioxide tension or oxygen saturation in this context, a linear model appears to be a fairly valid way of describing the data. It is not biologi-

cally feasible to interpret the intercept of these regression lines because this would mean extrapolating back to almost zero levels of either carbon dioxide or oxygen in the blood, which clearly do not occur under normal circumstances. However, it is possible to interpret the slope of the regression line and compare it among subjects or even within the same subject under different circumstances. Experimentally this may involve the interposing of some obstruction to breathing, simulating airway obstruction or by restricting the subject's chest movement by applying a chest binder (simulating various "restrictive" disorders of lung and chest wall elasticity).

The factors affecting the slope of the regression line relating change in ventilation to change in either carbon dioxide or oxygen (under conditions where the other variable, be it carbon dioxide or oxygen, is held constant while the test variable is changing) include the person's sex, genetic constitution, vital capacity and athleticism. These biological variations among subjects may be interpreted by looking at the components of the ventilatory change which, as mentioned previously, include tidal volume and frequency responses.

As both increasing levels of carbon dioxide and decreasing levels of oxygen provoke an increase in ventilation, it would be reasonable to assume that there would be a further relationship between the slope of the regression

associating ventilation change with change in both of these variables. This does not necessarily appear to be the case, although there is a loose association. However, some people who have a brisk ventilatory response to increasing levels of carbon dioxide, may have a relatively shallow and slow response to decreasing levels of oxygen. The reasons for this variability of response are not clearly understood, but can be analysed in terms of the components of ventilation, namely the response of tidal volume and the response of frequency to the stimulus. A further way in which they could be examined is by relating the slopes of the regression lines associating ventilation with change in a particular variable, to vital capacity.

The problem investigated in this chapter relates to just such an examination. The slope of the ventilatory response, regressed against changing levels of oxygen saturation was examined in relationship to vital capacity in a group of eight subjects (figure 6.1). Most of the subjects appear to fall within the confidence interval of a regression of ventilatory response against vital capacity. However, two subjects were outside the confidence interval, and the question arose as to whether these subjects constituted a separate subset of the population or whether they were simply indicating that the normal distribution about the fitted line was much wider than had first been assumed and that the error associated with the linear model was

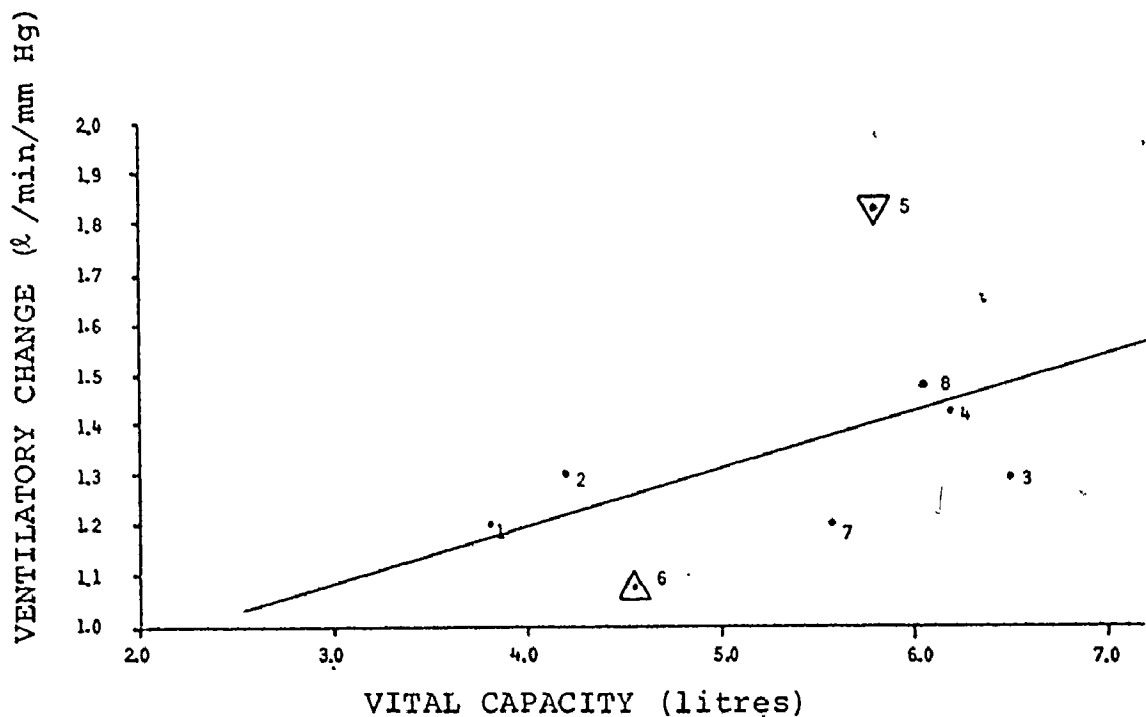


Figure 6.1

Rate of ventilatory change due to decreasing O_2 and fixed CO_2 concentrations plotted against vital capacity

artificially narrow.

The analysis carried out in section 6.3 will attempt to look at the way in which the different components of ventilation, namely tidal volume and frequency response relate to overall ventilatory response in the eight subjects. This will be considered separately for each of the following three sets of conditions:

- 1) concurrent decrease of oxygen and carbon dioxide concentrations;

2) fixed carbon dioxide concentration and decreasing oxygen concentration;

3) fixed oxygen concentration and increasing carbon dioxide concentration.

At each step, a maximum and minimum extreme deviate is selected and the associated probability values determined. Finally, all the results are evaluated for consistency, or lack thereof, among the subjects to determine if the sample represents more than one population.

6.2 Description of the Data

Table 6.2 is a tabulation of all the data used in this chapter. For each of the eight subjects the following is displayed:

Column 1: subject identification number;

Column 2: vital capacity in litres. Vital capacity is the greatest volume of air which can be exhaled in one breath subsequent to a maximal inhalation;

Column 3: the change in ventilatory response under conditions of progressively decreasing oxygen and carbon dioxide concentrations. Measured in litres per min per percentage change in oxygen saturation;

Column 4: the change in ventilatory response under conditions of progressively decreasing oxygen concen-

- trations and fixed carbon dioxide levels.
Measured in litres per minute per percentage change in oxygen saturation;
- Column 5: the change in ventilatory response under conditions of increasing carbon dioxide concentration but fixed oxygen saturation. Measured in litres per minute per millimetre of mercury pressure of carbon dioxide;
- Column 6: the change in ventilation frequency under the same conditions as column three. Measured in respirations per minute per percentage change in oxygen saturation;
- Column 7: the change in tidal volume under the same conditions as column three. Measured in litres per percentage change of oxygen saturation;
- Column 8: the change in ventilation frequency under the same conditions as column four. Measured in respirations per minute per percentage change of oxygen saturation;
- Column 9: the change in tidal volume under the same conditions as column four. Measured in litres per percentage change of oxygen saturation;
- Column 10: the change in ventilation frequency under the same conditions as column five. Measured in respirations per minute per millimetre of mercury pressure of carbon dioxide;

Column No.	1	2	3	4	5	6	7	8	9	10	11
Subject No.	VC	S_{aO_2} (ET)	S_{aO_2} (PCO ₂)	SCO ₂	$\frac{\Delta f}{\Delta S_{aO_2}}$ (ET)	$\frac{\Delta V_t}{\Delta S_{aO_2}}$ (ET)	$\frac{\Delta f}{\Delta S_{aO_2}}$ (PCO ₂)	$\frac{\Delta V_t}{\Delta S_{aO_2}}$ (PCO ₂)	$\frac{\Delta f}{\Delta CO_2}$	$\frac{\Delta V_t}{\Delta CO_2}$	
1	3.8	0.46	1.20	1.88	0.12	0.018	0.36	0.032	0.57	0.042	
2	4.2	0.56	1.30	2.14	0.22	0.025	0.50	0.030	0.72	0.036	
3	6.5	0.22	1.30	2.08	0.45	0.016	0.12	0.068	0.32	0.081	
4	6.2	0.53	1.44	2.91	0.14	0.030	0.22	0.055	0.50	0.084	
5	5.8	0.60	1.84	3.30	0.19	0.022	0.50	0.043	0.73	0.063	
6	4.6	0.28	1.06	1.06	0.04	0.031	0.35	0.038	0.41	0.025	
7	5.6	0.44	1.20	2.39	0.12	0.018	0.20	0.038	0.67	0.050	
8	6.1	0.22	1.48	2.93	0.10	0.070	0.26	0.042	0.63	0.066	

Table 6.2

Data collected to show change in ventilatory response with changing O₂ and CO₂ concentrations

Column 11: the change in tidal volume under the same conditions as column five. Measured in litres per millimetre of mercury pressure of carbon dioxide.

6.3 Analysis and Results

The problem described in the introduction and displayed in figure 6.1 is to determine whether the two subjects farthest from the regression line, namely members five and six, are from the same population as the other subjects or from some other population. If it could be shown that these two subjects still appear extreme when another measure of ventilatory change or response is plotted against vital capacity, then the probability of these two subjects belonging to a population different from the other subjects is improved. In fact, two other measures of ventilatory response have been used; the response to concurrent decrease in oxygen and carbon dioxide concentrations and the response to increasing carbon dioxide concentration with fixed oxygen saturation. The information regarding these two tests is displayed in figures 6.3 and 6.4. The data for figure 6.3 are taken from columns two and three of table 6.2 and the data for figure 6.4 are taken from columns two and five of the same table.

To organize the information from these three measures of ventilatory response in such a way that Youden's test

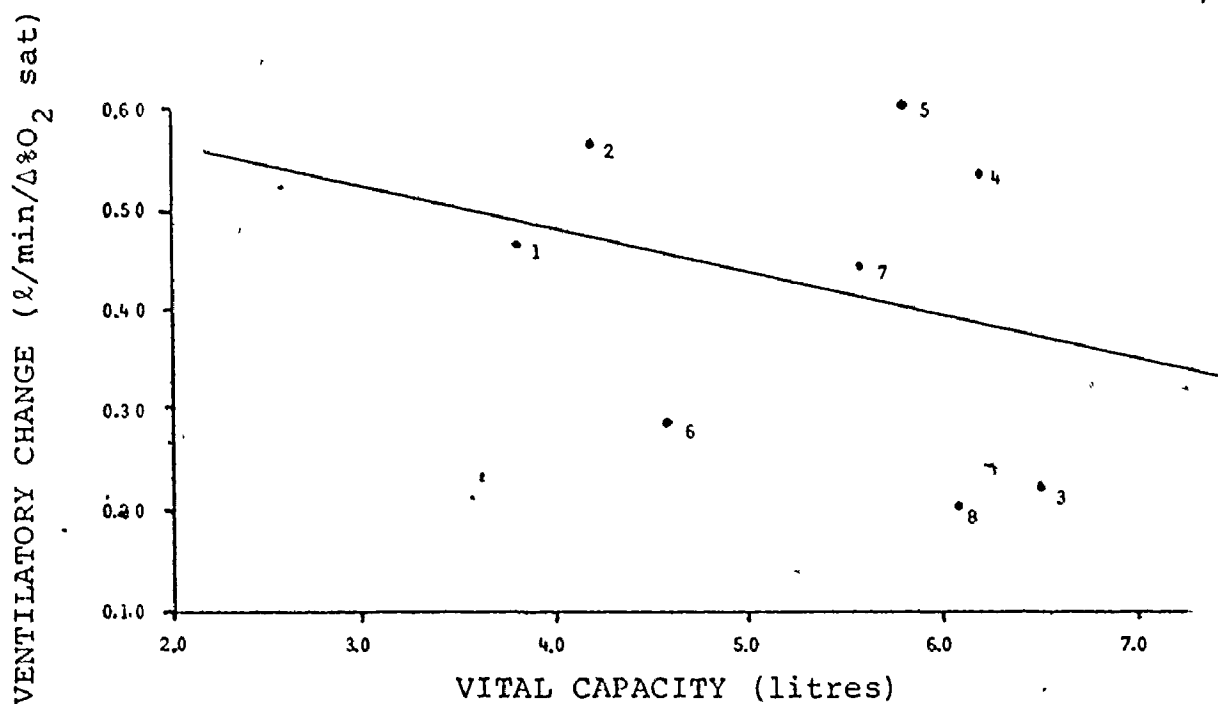


Figure.6.3

Rate of ventilatory change due to decreasing O₂ (with "freefall" CO₂) concentrations plotted against vital capacity

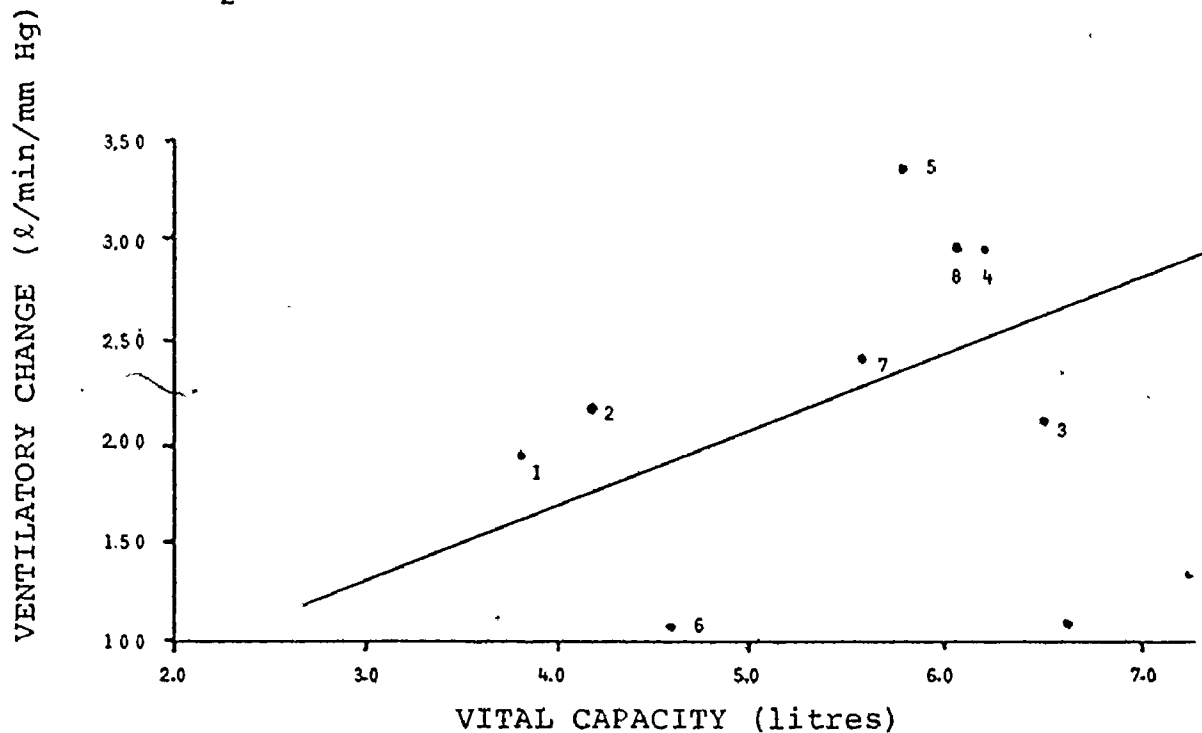


Figure 6.4

Rate of ventilatory change due to increasing CO₂ (fixed O₂) concentrations plotted against vital capacity

could be applied, it was necessary to develop a two way classification of the eight subjects by the three measures of ventilatory response. This was done by performing a linear regression on each of the three sets of data and using, as a measure of extremeness, the signed difference between the actual ventilatory response and the calculated ventilatory response as a percentage of the calculated ventilatory response.

$$\text{EXTREMENESS} = \frac{(\text{VR}_C - \text{VR}_A) \times 100}{\text{VR}_C}$$

where: VR_C is the calculated ventilatory response
 VR_A is the actual ventilatory response.

The measures of extremeness determined by the above formula for all subjects and for all three measures of ventilatory response are shown in table 6.5. Youden's test selected subject number five as the maximum extreme value ($p = .016$) as did Nair's test ($p = .01$). The minimum extreme deviate was subject number six according to both tests but the associated probabilities were not as convincing (Youden, $p > 0.10$; Nair, $p = .05$).

Another way of determining those subjects who are extremely different from the rest with regard to ventilatory response is to look at the relationship between change in

SUBJECT	MEASURES OF VENTILATORY RESPONSE			RANK SUMS
	O ₂ ↑ CO ₂ ↓	O ₂ ↓ CO ₂ ↔	O ₂ ↔ CO ₂ ↑	
1	- 4.8 (4)*	1.8 (5)	16.9 (5)	14
2	20.4 (6)	6.3 (7)	21.7 (7)	20
3	-38.4 (2)	-12.2 (3)	-20.5 (3)	8
4	42.6 (7)	- 0.5 (4)	16.1 (4)	15
5	53.7 (8)	31.1 (8)	40.0 (8)	24**
6	-37.2 (3)	-16.4 (1)	-44.4 (2)	6
7	10.0 (5)	-13.0 (2)	-47.3 (1)	8
8	-46.8 (1)	2.9 (6)	18.7 (6)	13

* Numbers in brackets show the rank of each value within column

Table 6.5

Extremeness values and associated rankings for the eight subjects by the three measures of ventilatory response when regressed against vital capacity

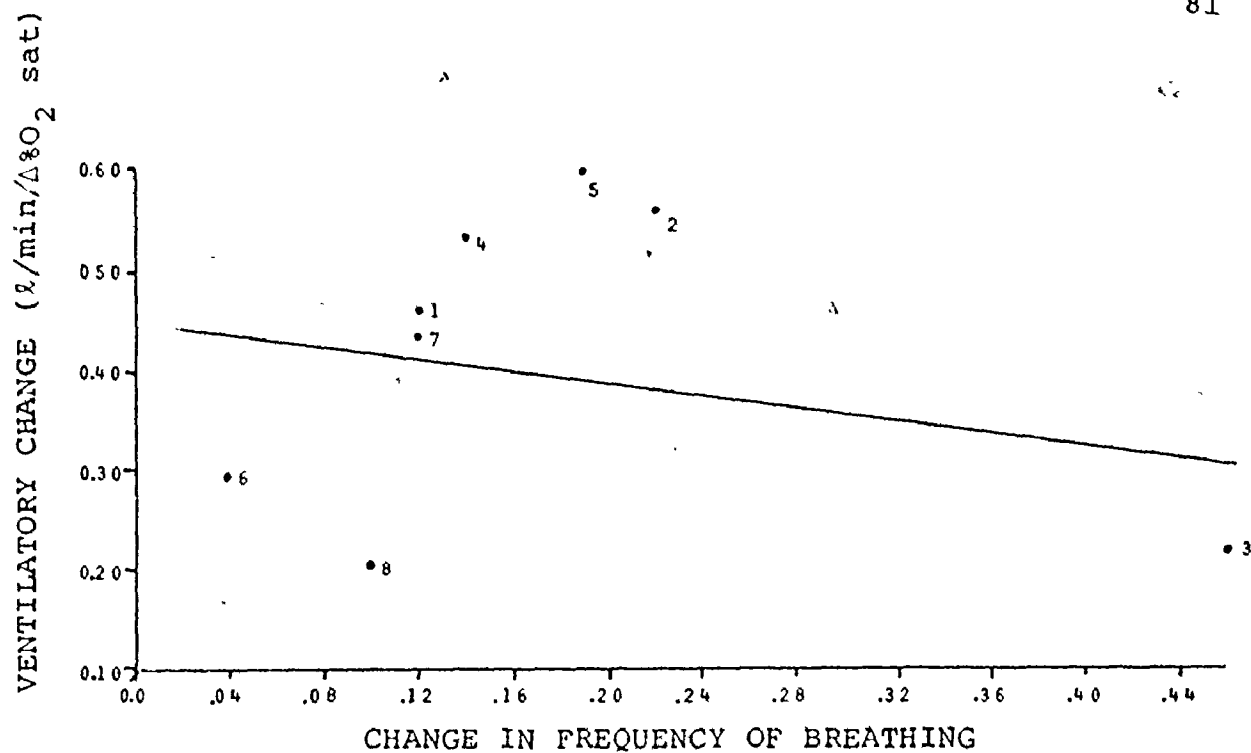


Figure 6.6

Rate of ventilatory change due to decreasing O_2 and CO_2 concentration plotted against change in frequency of respirations

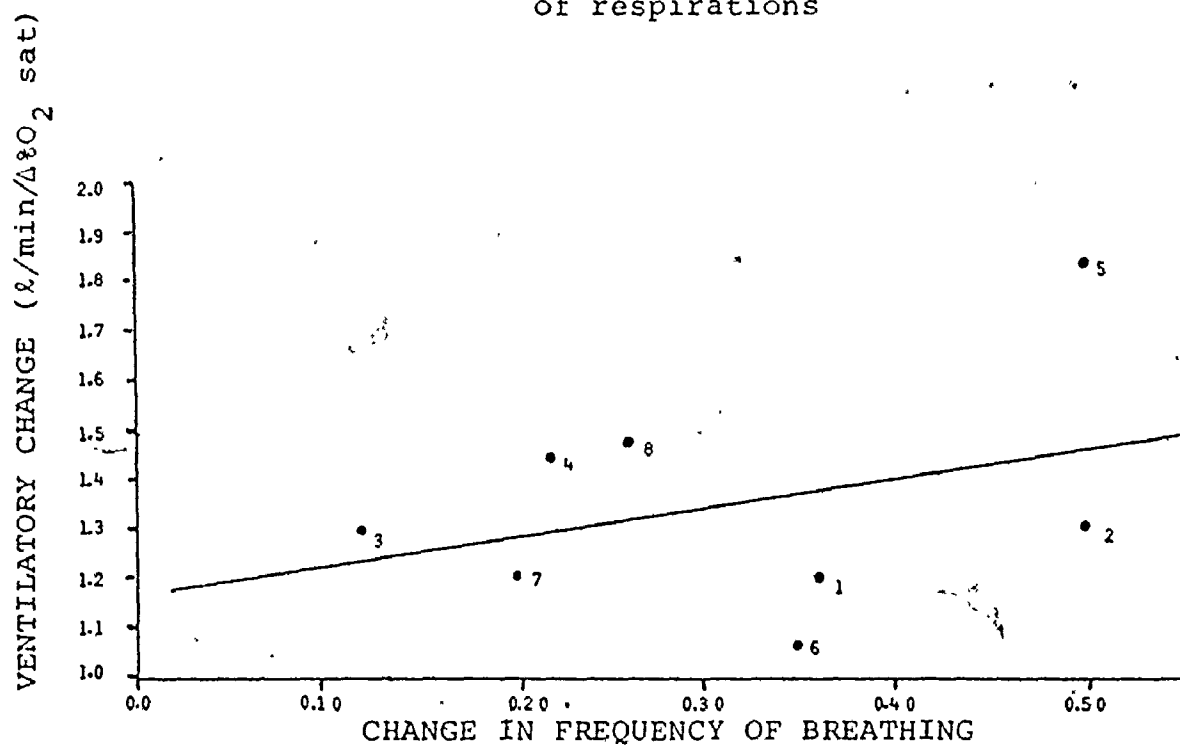


Figure 6.7

Rate of ventilatory change due to decreasing O_2 and fixed CO_2 concentrations plotted against change in frequency of respirations

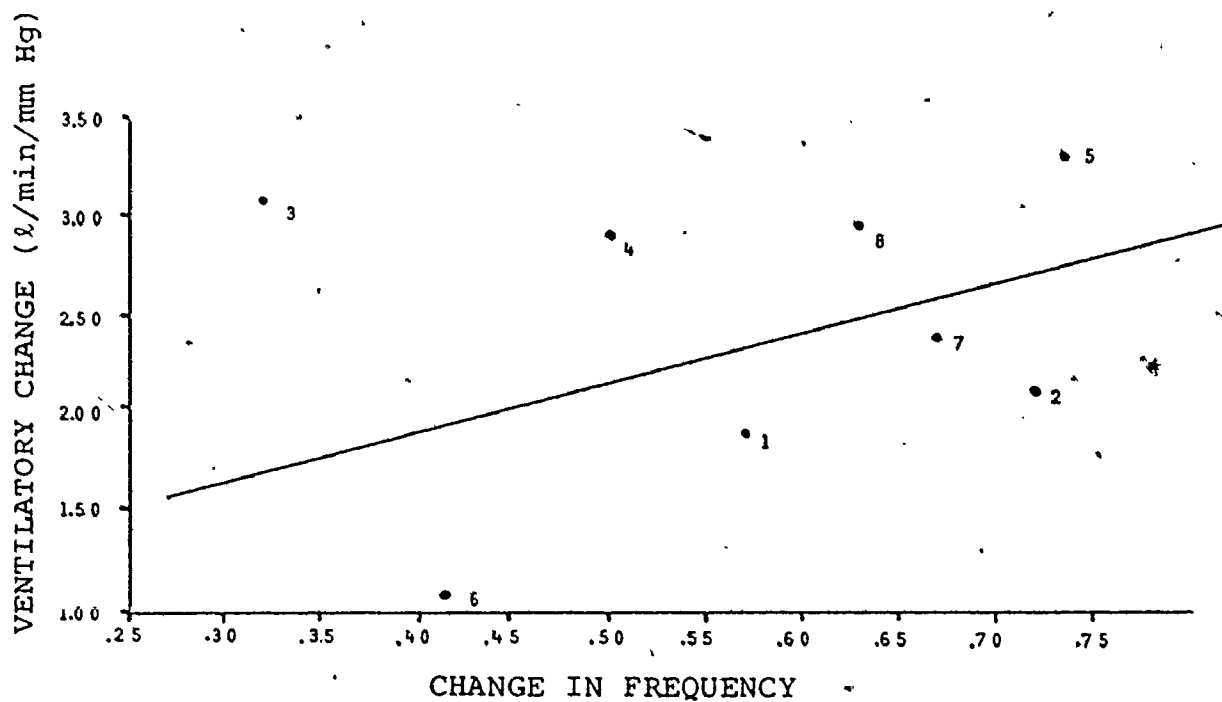


Figure 6.8

Rate of ventilatory change due to fixed O_2 and increasing CO_2 concentrations plotted against change in frequency of respirations

SUBJECT	MEASURES OF VENTILATORY RESPONSE			RANK SUMS
	O ₂ + CO ₂ +	O ₂ + CO ₂ ↔	O ₂ ↔ CO ₂ +	
1	9.5 (5)*	-12.9 (2)	-19.6 (3)	10
2	38.8 (7)	-10.7 (3)	-21.3 (2)	12
3	-39.8 (2)	4.5 (5)	22.1 (7)	14
4	4.7 (6)	10.7 (6)	34.6 (8)	20
5	46.9 (8)	26.2 (8)	20.1 (6)	22
6	-35.3 (3)	-22.7 (1)	-45.1 (1)	5
7	4.7 (4)	- 6.8 (4)	- 7.8 (4)	12
8	-52.7 (1)	11.9 (7)	17.5 (5)	13

*Numbers in brackets show the rank of each value within column

Table 6.9

Extremeness values and associated rankings for the eight subjects by the three measures of ventilatory response when regressed against change in frequency of breathing

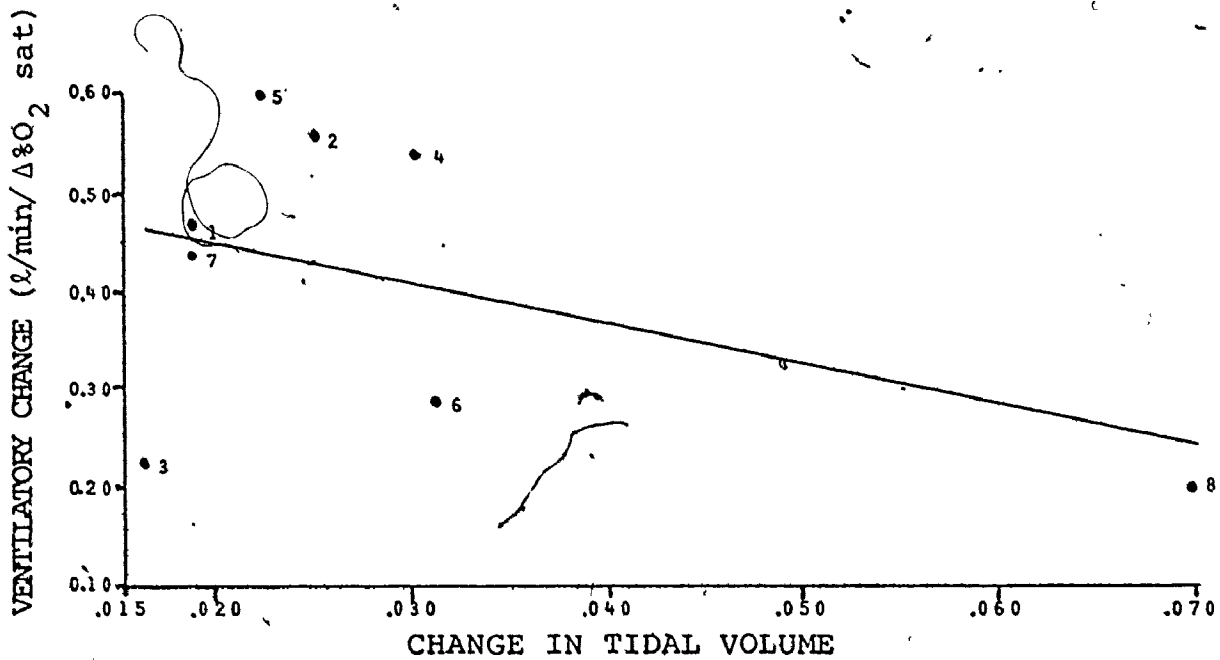


Figure 6.10

Rate of ventilatory change due to decreasing O_2 and CO_2 concentrations plotted against change in tidal volume

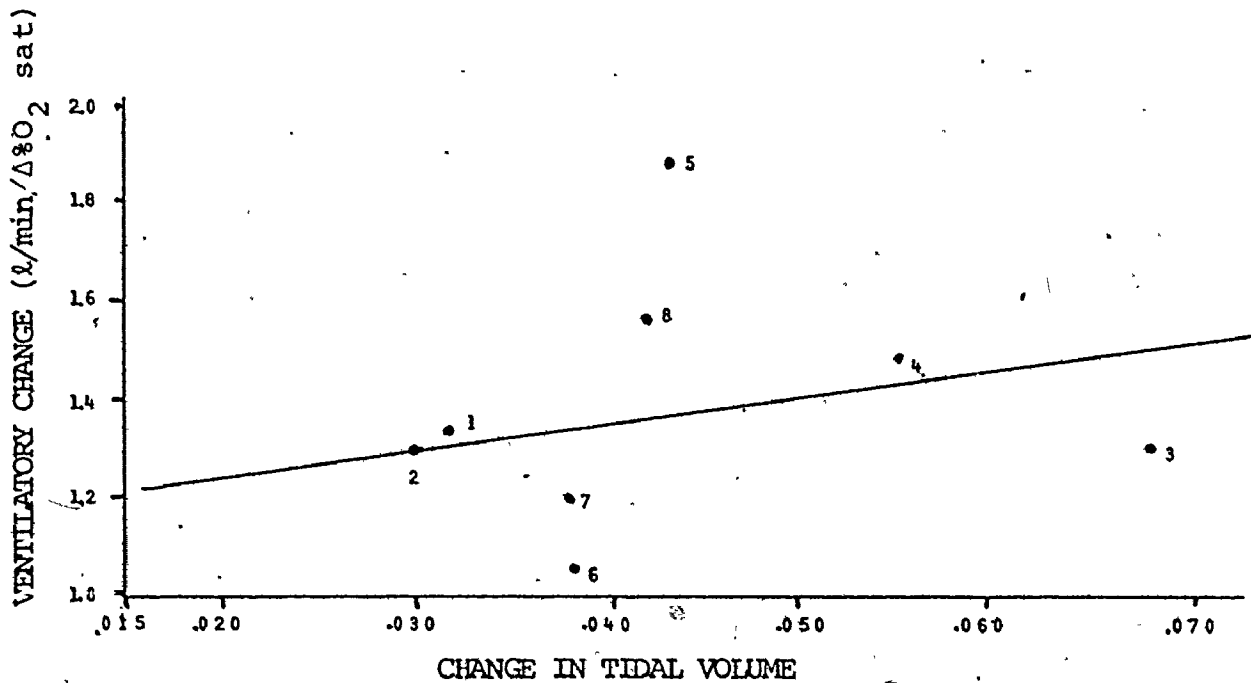


Figure 6.11

Rate of ventilatory change due to decreasing O_2 and fixed CO_2 concentrations plotted against change in tidal volume

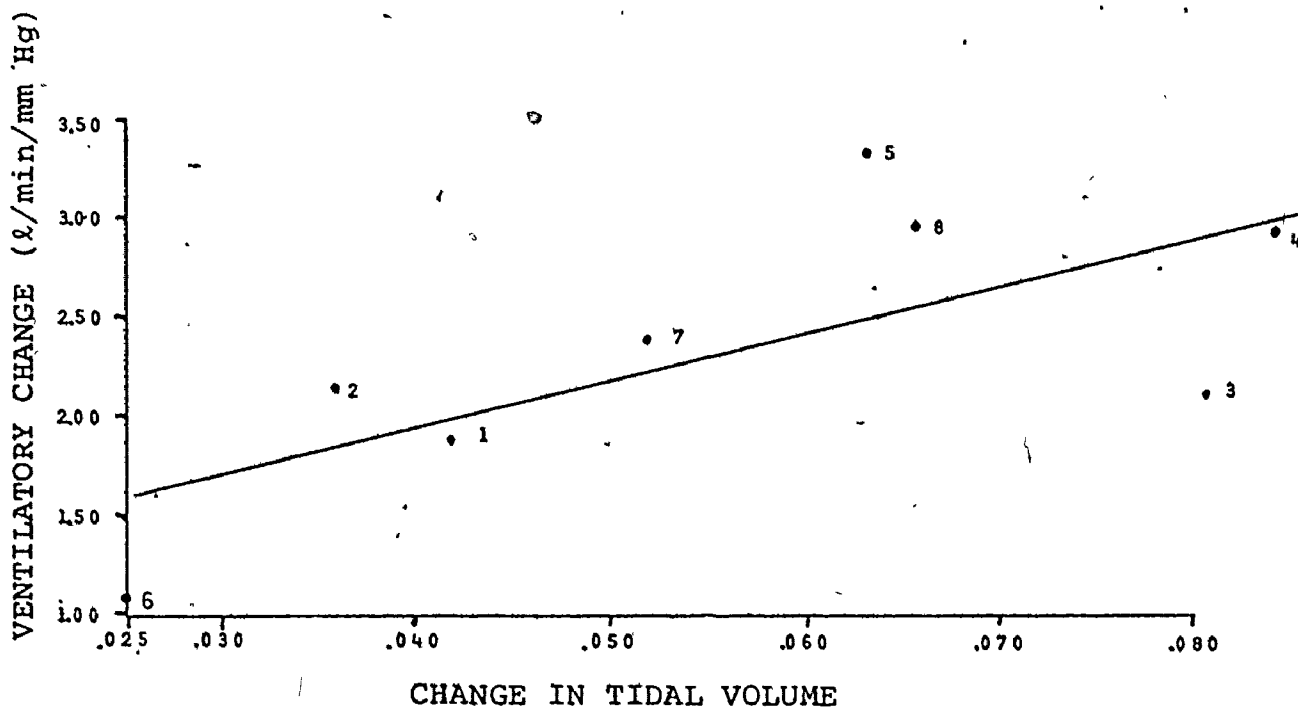


Figure 6.12

Rate of ventilatory change due to fixed O_2 and increasing CO_2 concentrations plotted against change in tidal volume.

SUBJECT	MEASURES OF VENTILATORY RESPONSE			RANK SUMS
	O ₂ † CO ₂ †	O ₂ † CO ₂ ↔	O ₂ ↔ CO ₂ †	
1	0.7 (5)	- 8.3 (4)	- 6.5 (3)	12
2	31.1 (7)	- 0.1 (5)	14.2 (6)	18
3	-52.6 (1)	-10.1 (2)	-28.4 (2)	5
4	30.5 (6)	3.0 (6)	- 2.2 (4)	16
5	36.4 (8)	36.1 (8)	32.3 (8)	24**
6	-30.3 (2)	-20.4 (1)	-34.5 (1)	4**
7	- 3.6 (4)	- 9.9 (3)	6.6 (5)	12
8	-15.9 (3)	9.8 (7)	14.3 (7)	17

*Numbers in brackets show the rank of each value within the column.

Table 6.13

Extremeness values and associated rankings for the eight subjects by the three measures of ventilatory response when regressed against change in tidal volume

frequency of respiration and overall ventilatory response under the same three sets of conditions mentioned previously. Similarly, tidal volume may also be related to overall ventilatory response. Figure 6.6, 6.7 and 6.8 display the data which relate change in frequency of respiration to overall ventilatory response for the three sets of conditions. Table 6.9 shows the extremeness values for the eight subjects under the three sets of conditions. These results are determined by the same methods as those in Table 6.5. Here again subject five is the maximum extreme deviate and subject six the minimum extreme deviate but the probabilities for both subjects is $>.10$ as determined by both Nair's test and Youden's test.

Figures 6.10, 6.11 and 6.12 display information much the same as that just analysed except that the change in frequency of respirations has been replaced with change in tidal volume. The extremeness values for each subject are similarly summarized in table 6.13.

Here, the evidence points strongly to subject five being the maximum extreme deviate (Nair, $p = .0025$; Youden, $p = .016$) and subject six being the minimum extreme deviate (Nair, $p = .01$; Youden, $p = .06$).

6.4 Conclusions

The initial problem of determining extreme deviation in subjects when relating ventilatory response to vital

capacity has been approached by developing a two way classification of the eight subjects against three different measures of ventilatory response. The measure of extremeness in each of the three cases was defined as the distance from the measured value of ventilatory response to the calculated value as determined by a linear model showing the relationship between vital capacity and the particular measure of ventilatory response. The data were then analysed using both Youden's and Nair's tests. Subjects number five and six were the original extremes as shown in figure 6.1. Confirmation of this extremeness was supplied when data from the other two measures of ventilatory response were incorporated in the analysis as shown in table 6.5. In particular, subject five was selected as a maximum extreme deviate by both tests ($p \approx .01$). Subject six was selected as the minimum extreme deviate but with a probability which makes its significance questionable ($.10 > p > .05$). Of the three measures of ventilatory response, the first one (decreasing oxygen and carbon dioxide concentrations), is probably the weakest because both gas concentrations are decreasing at the same time. A rank of three for subject six under these conditions is probably higher than the true value and if a value of two were used instead, then both Youden and Nair would agree that subject six was the minimum extreme deviate with $p = .05$.

Further analysis has shown how these two subjects

differ from the others in their association of ventilatory response with either change in frequency of breathing or tidal volume. Table 6.9 shows the results of regressing change in frequency of breathing with ventilatory response and measuring for extremeness from the fitted line. Again, subjects five and six are selected as extreme but both Nair's and Youden's tests suggest a probability value $>.10$ for both subjects.

In the case of regressing change in tidal volume with ventilatory response and measuring for extremeness about the fitted line, table 6.13 shows that the rank sums for subjects five and six are definitely extreme. The probability values determined by Youden's test are; subject five: maximum ($p = .016$), subject six: minimum ($p = .06$). Nair's test supplies more extreme values; subject five: maximum ($p = .0025$) subject six: minimum ($p = .01$).

In conclusion, subject six appears extreme in the minimum direction with a probability value of 0.05 and subject five extreme in the maximum direction with a probability value of 0.02.

It should be pointed out, however, that this study has been used to demonstrate the application of extreme value analysis and does not lead automatically to more general biological conclusions about the distribution of ventilatory response in the population. This technique and a larger random sample would be needed to determine if a real population subset does, in fact, exist.

CHAPTER VII

DISCUSSION AND CONCLUSIONS

7.1 Introduction

In the previous chapters, Youden's and Nair's tests have been introduced and used in the analysis of three different biomedical problems. Youden's test in particular has been emphasized and chapter two dealt exclusively with the various methods used to increase the range of the associated probability tables and their accuracy. In this chapter, each of the three research questions posed in chapter one will be discussed based on the experience gained in analysing the three problems presented in chapters four, five and six. The final section will draw general conclusions regarding the use of Youden's test in analysing biomedical data.

7.2 Research Questions

7.2.1 Research Question 1

What types of data are most suitable for analysis by the non-parametric test?

In discussing the type of data most suitable for analysis by Youden's test, it is important that we look at

both the nature of the distribution of the data and the nature of the values themselves, i.e. whether they are continuous or discrete measurements.

With regard to the nature of the underlying distribution of the data, chapter three presents power curves for comparing Youden's test and Nair's test assuming three different types of population distributions. It is observed that when the number of objects (identified by the letter I) is small, and the number of judges (identified by the letter J) is small, then Youden's test has the greater power, independent of the data distribution used. As J increases and I remains small, the two tests appear to approach the same power independent of the data distribution. As I increases in value, the parametric test of Nair is either of greater power, or of power equal to, the non-parametric test of Youden for the three distributions examined.

The simulation program which generated these power curves used continuous data as opposed to discrete data putting Youden's test to a slight disadvantage. In this case it is probably correct to conclude that Youden's test has superior power in situations where I is small, otherwise Nair's test has power at least equal to that of Youden's test.

Youden's test, like all other rank tests, suffers from the ranking process which is usually required to transform the data prior to analysis. This process causes a

considerable decrease in power compared to parametric tests when analysing data whose parent population is approximately normally distributed. For example, if two subjects have values of 4.001 and 4.002 they could be ranked 1 and 2, but under different conditions these same two subjects could have values of 4.001 and 86.2 and the ranking would still be 1 and 2. Parametric tests do not have this problem and are able to take into account the magnitude of the difference between subjects. Another problem encountered when using non-parametric tests is the method of solving tied values. Depending on the technique used, either loss of power or additional calculation of a correction factor, is involved.

Although power curves were not calculated using rank data, Youden's test would obviously perform best in this situation as the probability tables are based on approximations of all possible rankings. On the other hand, Nair's test has greater power when analysing continuous data.

7.2.2 Research Question 2

Under what conditions is the parametric test superior to the non-parametric test?

From the simulations performed in chapter three it is obvious that Nair's test is superior in analysing normal-

ly distributed data when the sample size is not small (say greater than 10). The considerable robustness shown when the data were not normally distributed indicates that this test can also be used when the data distribution is strictly non-normal (but similar to a normal distribution), providing the sample size exceeds 10. If the data contain a large number of tied values then Nair's test is probably superior, as the technique used for resolving ties in Youden's test is to randomly assign the N tied values the next N ranks. If the number of judges is small, then there is a high probability of one of the objects attaining an unusually high or low rank sum which could erroneously classify it as a significantly extreme deviate. With Nair's test there is no requirement that the scores be unique and so the problem of ties does not arise.

7.2.3 Research Question 3

Is it beneficial to use the non-parametric test to determine second and third most extreme deviates?

When this question was originally formulated, it was thought that it would be useful to develop a technique which would select those objects which were second and third most extreme from the other objects. The technique involves deleting the most extreme object from the two way classification of data and re-ranking to select the most extreme

deviate. This object would be the second most extreme deviate. By a similar technique the third most extreme deviate can also be found. The most important thing learned regarding this procedure is that one must display the data in a way which shows its distributional characteristics prior to interpreting the meaning of second and third most extreme deviates. In chapter four, Figure 4.2 displays the melanoma data in such a way that two extreme maximum deviates and one extreme minimum deviate appear obvious. Analysis shows this to be the case (see Table 4.3). The technique worked for this particular set of data because the majority of the values were clustered around the mean with very few far from the mean. However, if the data are more widely spread with no clustering, it is possible to show each value to be significantly extreme from the others as either a first, second, third or more, extreme deviate. Results like this are generally meaningless and may be avoided by visually examining the sample distribution.

Another problem regarding sample distribution is bi-modality or possibly multi-modality. A histogram or graph of some form will usually show several tight clusters of values with relatively vacant areas in between. In this case it is not only questionable as to whether second and third most extreme deviates should be determined, but questionable as to whether extreme value analysis would be meaningful.

7.3 Conclusions

An existing non-parametric test has been expanded by increasing the size and accuracy of the associated probability table. The methods used to expand the table have been described and the updated table presented. Power curves for the non-parametric test and its parametric analog were estimated by a simulation technique, and presented graphically for visual comparison. The ability of this test to solve problems of a biomedical nature has been demonstrated in three separate instances and comparison with its parametric analog has been made throughout. Answers to the research questions were presented and discussed.

As a final comment on the application of Youden's test to biomedical problems, it has been found that reformulation of the problem or reorganization of the data is a very important step in applying this test to situations which at first glance seem inappropriate.

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APPENDIX A

TABLES OF ONE-SIDED PERCENTAGE POINTS AT NOMINAL
LEVELS OF 1, 3, 5 AND 10%

I = 2

J	1% MIN	1% MAX	UPPER EXACT	LOWER EXACT	3% MIN	3% MAX	UPPER EXACT	LOWER EXACT	5% MIN	5% MAX	UPPER EXACT	LOWER EXACT	10% MIN	10% MAX	UPPER EXACT	LOWER EXACT
3																
4	4	8	.1250	.1250	4	8	.1250	.1250	4	8	.1250	.1250	4	8	.1250	.1250
5	5	10	.0625	.0625	5	10	.0625	.0625	5	10	.0625	.0625	5	10	.0625	.0625
6	6	12	.0313	.0313	6	12	.0313	.0313	6	12	.0313	.0313	6	12	.0313	.0313
7	7	14	.0156	.0155	7	14	.0155	.0155	7	14	.0156	.0156	7	14	.0156	.0156
8	8	16	.0078	.0078	8	16	.0078	.0078	8	16	.0156	.0156	8	16	.0156	.0156
9	10	17	.0391	.0391	10	17	.0391	.0391	9	15	.0703	.0703	9	15	.0703	.0703
10	11	19	.0215	.0215	11	19	.0215	.0215	10	17	.0391	.0391	10	17	.0391	.0391
11	12	21	.0117	.0117	12	21	.0117	.0117	11	19	.0215	.0215	11	19	.0215	.0215
12	13	23	.0063	.0063	13	22	.0385	.0385	13	20	.0654	.0654	12	18	.1094	.1094
13	15	24	.0225	.0225	14	22	.0385	.0385	14	22	.0386	.0386	13	20	.0654	.0654
14	16	26	.0129	.0129	15	24	.0225	.0225	15	24	.0225	.0225	15	21	.1460	.1460
15	17	28	.0074	.0074	16	26	.0129	.0129	17	25	.0574	.0574	16	23	.0923	.0923
16	19	29	.0213	.0213	18	27	.0352	.0352	18	27	.0352	.0352	17	25	.0574	.0574
17	20	31	.0127	.0127	19	29	.0213	.0213	20	28	.0768	.0768	19	26	.1185	.1185
18	21	33	.0075	.0075	20	31	.0127	.0127	21	30	.0490	.0490	20	28	.0768	.0768
19	23	34	.0192	.0192	22	32	.0309	.0309	22	32	.0309	.0309	22	29	.1435	.1435
20	24	36	.0118	.0118	23	34	.0192	.0192	24	33	.0636	.0636	23	31	.0963	.0963
21	25	38	.0072	.0072	25	35	.0414	.0414	25	35	.0414	.0414	24	33	.0636	.0636
22	27	39	.0169	.0169	26	37	.0265	.0265	27	37	.0266	.0266	26	34	.1153	.1153
23	28	41	.0106	.0106	27	39	.0169	.0169	28	38	.0525	.0525	27	36	.0784	.0784
24	29	43	.0066	.0066	29	40	.0347	.0347	29	40	.0525	.0525	29	37	.1338	.1338
25	31	44	.0146	.0146	30	42	.0227	.0227	31	41	.0347	.0347	30	39	.0931	.0931
					32	43	.0433	.0433	32	43	.0639	.0639	31	41	.0639	.0639
													33	42	.1078	.1078

I = 3

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	9	3	9	3	9	3	9
4	4	12	4	12	4	12	4	12
5	5	15	5	14	5	14	6	14
6	6	18	7	17	7	17	8	16
7	8	20	9	19	9	19	9	19
8	9	23	11	21	11	21	11	21
9	11	25	12	24	12	24	13	23
10	13	27	14	26	14	26	15	25
11	14	30	15	29	15	29	16	28
12	16	32	17	31	17	31	18	30
13	18	34	19	33	19	33	20	32
14	19	37	21	35	21	35	22	34
15	21	39	23	37	23	37	24	36
16	23	41	25	39	25	39	25	39
17	24	44	26	42	26	42	27	41
18	26	45	27	44	27	44	29	43
19	28	48	29	47	30	45	31	45
20	30	50	31	49	32	48	33	47
21	31	53	33	51	34	50	35	49
22	33	55	35	53	35	53	36	52
23	35	57	37	55	37	55	38	54
24	37	59	39	57	39	57	40	56
25	38	62	41	59	41	59	42	58

J	1%		3%		5%		10%	
	UPPER EXACT	LOWER EXACT	UPPER EXACT	LOWER EXACT	UPPER EXACT	LOWER EXACT	UPPER EXACT	LOWER EXACT
3	.1111	.0370	.1111	.0370	.1111	.0370	.1111	.0370
4	.0370	.0123	.0370	.0123	.0370	.0123	.0370	.0123
5	.0123	.0041	.0288	.0041	.0741	.0288	.0741	.0288
6	.0041	.0110	.0110	.0110	.0288	.0494	.0494	.0494
7	.0110	.0041	.0205	.0041	.0494	.0718	.0718	.0718
8	.0041	.0064	.0322	.0064	.0718	.0322	.0718	.0322
9	.0064	.0140	.0453	.0140	.0322	.0453	.0453	.0453
10	.0140	.0060	.0209	.0060	.0453	.0594	.0594	.0594
11	.0060	.0094	.0288	.0094	.0594	.0288	.0594	.0288
12	.0094	.0136	.0375	.0136	.0288	.0375	.0375	.0375
13	.0136	.0062	.0184	.0062	.0375	.0469	.0469	.0469
14	.0062	.0088	.0239	.0088	.0469	.0569	.0569	.0569
15	.0088	.0118	.0298	.0118	.0569	.0672	.0672	.0672
16	.0118	.0057	.0363	.0057	.0672	.0363	.0672	.0363
17	.0057	.0076	.0191	.0076	.0363	.0432	.0432	.0432
18	.0076	.0098	.0233	.0098	.0432	.0504	.0504	.0504
19	.0098	.0123	.0278	.0123	.0504	.0578	.0578	.0578
20	.0123	.0063	.0327	.0063	.0578	.0655	.0655	.0655
21	.0063	.0079	.0377	.0079	.0655	.0377	.0655	.0377
22	.0079	.0097	.0212	.0097	.0377	.0431	.0431	.0431
23	.0097	.0117	.0247	.0117	.0431	.0486	.0486	.0486
24	.0117	.0063	.0283	.0063	.0486	.0543	.0543	.0543
25	.0063				.0543			

I = 4

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	12	3	12	3	12	3	12
4	4	16	5	15	5	15	5	15
5	6	19	5	19	5	19	7	18
6	7	23	8	22	8	22	9	21
7	9	26	10	25	10	25	11	24
8	11	29	12	28	12	28	13	27
9	13	32	14	31	15	30	15	30
10	15	35	16	34	17	33	18	32
11	17	38	18	37	19	36	20	35
12	19	41	20	40	21	39	22	38
13	21	44	22	43	23	42	24	41
14	23	47	24	46	25	45	26	44
15	25	50	27	48	27	48	29	46
16	27	53	29	51	30	50	31	49
17	29	56	31	54	32	53	33	52
18	31	59	33	57	34	56	35	55
19	33	62	35	60	35	59	37	58
20	36	64	37	63	38	62	40	60
21	38	67	40	65	41	64	42	63
22	40	70	42	68	43	67	44	66
23	42	73	44	71	45	70	46	69
24	44	76	46	74	47	73	49	71
25	46	79	48	77	49	76	51	74

1%		3%		5%		10%	
UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER
.0625	.0625	.0625	.0625	.0525	.0633	.0625	.1289
.0156	.0156	.0156	.0280	.0781	.0578	.0781	.1150
.0234	.0234	.0234	.0265	.0234	.0526	.0820	.1025
.0068	.0068	.0273	.0248	.0273	.0477	.0789	.0913
.0088	.0088	.0293	.0231	.0293	.0431	.0742	.0813
.0101	.0101	.0297	.0213	.0297	.0389	.0688	.1273
.0108	.0108	.0292	.0389	.0297	.0645	.0688	.1125
.0110	.0110	.0280	.0350	.0688	.0644	.0688	.0996
.0110	.0110	.0265	.0315	.0633	.0575	.0688	.0882
.0107	.0107	.0265	.0283	.0633	.0512	.0688	.0781
.0103	.0103	.0248	.0255	.0633	.0457	.0688	.1131
.0098	.0098	.0231	.0229	.0633	.0407	.0688	.1001
.0098	.0098	.0213	.0229	.0633	.0615	.0688	.0887
.0092	.0092	.0213	.0229	.0633	.0545	.0688	.0786
.0086	.0086	.0389	.0363	.0633	.0485	.0688	.1090
.0080	.0080	.0350	.0324	.0633	.0431	.0688	.0967
.0074	.0074	.0315	.0249	.0633	.0384	.0688	.0964
.0068	.0068	.0283	.0288	.0633	.0384	.0688	.0964
.0123	.0123	.0255	.0288	.0633	.0384	.0688	.0964
.0111	.0111	.0229	.0288	.0633	.0384	.0688	.0964
.0100	.0100	.0363	.0288	.0633	.0384	.0688	.0964
.0091	.0091	.0324	.0288	.0633	.0384	.0688	.0964
.0082	.0082	.0249	.0288	.0633	.0384	.0688	.0964
.0074	.0074	.0257	.0288	.0633	.0384	.0688	.0964
.0074	.0074	.0230	.0288	.0633	.0384	.0688	.0964

I = 5

J	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT
3	3	15	.0400	.0400	3	15	.0400	.0400	3	15	.0400	.0400	3	15	.0400	.0400
4	4	20	.0080	.0400	5	19	.0400	.0400	6	18	.1200	.1188	6	18	.1200	.1188
5	6	24	.0096	.0336	7	23	.0336	.0336	8	22	.0896	.0894	8	22	.0896	.0894
6	8	29	.0090	.0269	9	27	.0269	.0672	10	26	.0672	.0669	10	26	.0672	.0669
7	10	32	.0077	.0211	11	31	.0211	.0502	12	30	.0502	.1057	12	30	.0502	.1057
8	12	36	.0063	.0375	14	34	.0375	.0375	14	34	.0375	.0774	14	34	.0375	.0774
9	15	39	.0126	.0280	16	38	.0280	.0572	17	37	.0572	.1074	17	36	.1080	.1074
10	17	43	.0096	.0209	19	42	.0209	.0422	19	41	.0422	.0788	19	40	.0792	.0788
11	19	47	.0073	.0312	21	45	.0312	.0584	22	44	.0584	.1026	22	43	.1030	.1026
12	22	50	.0117	.0231	23	49	.0231	.0431	24	48	.0431	.0758	24	47	.0762	.0758
13	24	54	.0088	.0319	25	52	.0319	.0564	27	51	.0564	.0948	27	50	.0952	.0948
14	27	57	.0128	.0237	28	56	.0237	.0418	29	55	.0418	.1142	29	54	.1151	.1142
15	29	61	.0096	.0312	31	59	.0312	.0529	32	58	.0529	.0860	31	57	.0863	.0860
16	31	65	.0071	.0232	33	63	.0232	.0395	34	62	.0395	.1017	33	61	.1024	.1017
17	34	68	.0098	.0296	35	66	.0296	.0485	37	65	.0485	.1184	36	64	.1191	.1184
18	36	72	.0073	.0365	39	69	.0365	.0584	40	68	.0584	.0900	39	67	.0905	.0900
19	39	75	.0097	.0275	41	73	.0275	.0441	42	72	.0441	.1037	41	71	.1042	.1037
20	42	78	.0125	.0333	44	76	.0333	.0522	45	75	.0522	.1174	44	74	.1183	.1174
21	44	82	.0094	.0252	46	80	.0252	.0397	47	79	.0397	.0906	47	78	.0909	.0906
22	47	85	.0117	.0301	49	83	.0301	.0464	50	82	.0464	.1020	49	81	.1026	.1020
23	49	89	.0088	.0354	52	86	.0354	.0536	53	85	.0536	.1140	52	84	.1146	.1140
24	52	92	.0109	.0270	54	90	.0270	.0411	55	89	.0411	.0886	55	88	.0890	.0886
25	54	95	.0082	.0315	57	93	.0315	.0471	58	92	.0471	.0986	57	91	.0990	.0986

I = 6

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	18	3	18	3	19	4	17
4	5	23	5	23	5	22	6	22
5	7	28	8	27	8	27	9	26
6	9	33	10	32	11	31	12	30
7	11	38	13	36	13	35	14	35
8	14	42	15	41	15	40	17	39
9	16	47	18	45	19	44	20	43
10	19	51	21	49	22	48	23	47
11	22	55	24	53	25	52	26	51
12	25	59	26	58	27	57	29	55
13	27	64	29	62	30	61	32	59
14	30	68	32	66	33	65	35	63
15	33	72	35	70	35	69	38	67
16	36	76	39	74	39	73	41	71
17	39	80	41	78	42	77	44	75
18	42	84	44	82	45	81	47	79
19	44	89	47	86	48	85	50	83
20	47	93	50	90	51	89	53	87
21	50	97	53	94	54	93	56	91
22	53	101	56	98	57	97	60	94
23	56	105	59	102	60	101	63	98
24	59	109	62	106	64	104	66	102
25	62	113	65	110	67	108	69	106

J	1%		3%		5%		10%	
	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER
3	.0276	.0278	.0279	.0277	.0278	.0278	.0694	.0691
4	.0231	.0231	.0231	.0231	.0694	.0691	.0972	.0967
5	.0162	.0432	.0432	.0432	.0432	.0432	.1181	.1165
6	.0108	.0270	.0270	.0270	.0594	.0591	.0724	.0721
7	.0071	.0365	.0365	.0365	.0366	.0366	.0821	.0815
8	.0107	.0227	.0227	.0227	.0447	.0446	.0893	.0888
9	.0066	.0278	.0278	.0277	.0512	.0510	.0943	.0935
10	.0089	.0320	.0320	.0320	.0562	.0560	.0976	.0970
11	.0109	.0355	.0355	.0355	.0600	.0599	.0994	.0986
12	.0127	.0225	.0225	.0224	.0383	.0382	.1002	.0996
13	.0080	.0244	.0244	.0244	.0404	.0403	.1002	.0994
14	.0091	.0260	.0260	.0260	.0420	.0419	.0994	.0994
15	.0100	.0273	.0273	.0272	.0431	.0430	.0994	.0994
16	.0108	.0282	.0282	.0282	.0439	.0437	.0981	.0974
17	.0115	.0289	.0289	.0288	.0442	.0441	.0964	.0959
18	.0120	.0293	.0293	.0293	.0442	.0441	.0943	.0937
19	.0078	.0295	.0295	.0295	.0440	.0439	.0920	.0916
20	.0082	.0295	.0295	.0295	.0435	.0435	.0896	.0890
21	.0085	.0295	.0295	.0295	.0431	.0430	.0870	.0866
22	.0087	.0293	.0293	.0292	.0424	.0423	.1160	.1151
23	.0088	.0290	.0290	.0289	.0416	.0415	.1116	.1110
24	.0090	.0285	.0285	.0285	.0570	.0569	.1073	.1065
25	.0090	.0281	.0281	.0280	.0554	.0552	.1031	.1025

I = 7

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	21	3	21	3	21	4	20
4	5	27	6	26	5	26	7	25
5	7	33	8	32	9	31	10	30
6	10	38	11	37	12	36	13	35
7	12	44	14	42	15	41	16	40
8	15	49	17	47	18	46	19	45
9	18	54	20	52	21	51	22	50
10	21	59	23	57	24	56	26	54
11	24	64	26	62	27	61	29	59
12	27	69	30	66	31	65	32	64
13	30	74	33	71	34	70	36	68
14	34	78	36	76	37	75	39	73
15	37	83	39	81	41	79	43	77
16	40	88	43	85	44	84	46	82
17	43	93	46	90	47	89	50	86
18	47	97	49	95	51	93	53	91
19	50	102	53	99	54	98	56	96
20	53	107	56	104	58	102	60	100
21	57	111	50	108	61	107	64	104
22	60	115	63	113	65	111	67	109
23	63	121	66	118	68	116	71	113
24	67	125	70	122	72	120	74	118
25	70	130	73	127	75	125	78	122

UPPER EXACT		LOWER EXACT		UPPER EXACT		LOWER EXACT	
.0204	.0204	.0204	.0204	.0204	.0204	.0204	.0204
.0140	.0437	.0436	.0437	.0437	.0436	.0437	.0436
.0987	.0233	.0233	.0233	.0525	.0523	.0525	.0523
.0125	.0275	.0274	.0275	.0550	.0547	.0550	.0547
.0067	.0291	.0291	.0291	.0542	.0540	.0542	.0540
.0078	.0291	.0290	.0291	.0515	.0513	.0515	.0513
.0083	.0280	.0280	.0280	.0479	.0477	.0479	.0477
.0085	.0254	.0263	.0254	.0438	.0435	.0438	.0435
.0084	.0244	.0244	.0244	.0395	.0395	.0395	.0395
.0080	.0356	.0355	.0356	.0552	.0549	.0552	.0549
.0076	.0318	.0317	.0318	.0485	.0485	.0485	.0485
.0115	.0293	.0282	.0293	.0428	.0427	.0428	.0427
.0105	.0251	.0250	.0251	.0552	.0550	.0552	.0550
.0094	.0329	.0329	.0329	.0480	.0479	.0480	.0479
.0085	.0283	.0283	.0283	.0418	.0417	.0418	.0417
.0115	.0252	.0252	.0252	.0515	.0514	.0515	.0514
.0102	.0317	.0316	.0317	.0447	.0445	.0447	.0445
.0090	.0275	.0275	.0275	.0538	.0536	.0538	.0536
.0117	.0335	.0335	.0335	.0465	.0464	.0465	.0464
.0102	.0292	.0291	.0292	.0549	.0547	.0549	.0547
.0099	.0253	.0253	.0253	.0474	.0473	.0474	.0473
.0112	.0302	.0302	.0302	.0551	.0549	.0551	.0549
.0098	.0252	.0261	.0252	.0477	.0475	.0477	.0475

I = 8

J	1%		3%		5%		10%	
	MIN	UPPER LOWER EXACT	MIN	UPPER LOWER EXACT	MIN	UPPER LOWER EXACT	MIN	UPPER LOWER EXACT
3	3	.0156	3	.0156	4	.0625	4	.0625
4	5	.0098	6	.0293 .0292	7	.0684 .0679	7	.0684 .0678
5	8	.0137	9	.0308 .0307	10	.0615 .0513	11	.1128 .1116
6	10	.0064	12	.0282 .0281	13	.0524 .0521	14	.0915 .0904
7	14	.0131	15	.0245	15	.0434 .0433	18	.1182 .1169
8	17	.0115	19	.0354 .0353	20	.0582 .0578	21	.0921 .0911
9	20	.0099	22	.0296 .0285	23	.0460 .0459	25	.1087 .1077
10	23	.0093	26	.0353 .0302	27	.0560 .0557	28	.0840 .0833
11	27	.0114	29	.0285	30	.0437 .0435	32	.0948 .0941
12	30	.0092	33	.0341 .0340	34	.0505 .0503	36	.1042 .1031
13	34	.0115	36	.0265	38	.0567 .0565	40	.1122 .1112
14	37	.0092 .0091	40	.0305	41	.0440 .0439	43	.0868 .0861
15	41	.0110	44	.0342 .0341	45	.0483 .0482	47	.0923 .0916
16	44	.0086	47	.0256 .0265	49	.0522 .0520	51	.0970 .0961
17	48	.0100	51	.0292	53	.0557 .0555	55	.1009 .1002
18	52	.0113	55	.0315	55	.0434 .0433	59	.1042 .1033
19	55	.0089	59	.0339	60	.0459 .0458	63	.1069 .1061
20	59	.0099	52	.0264	64	.0481 .0479	67	.1091 .1081
21	63	.0108	56	.0280	69	.0500 .0499	71	.1108 .1100
22	66	.0085	70	.0295 .0294	72	.0517 .0515	75	.1121 .1111
23	70	.0092	74	.0308 .0307	75	.0532 .0530	78	.0887 .0882
24	74	.0099	78	.0320 .0319	80	.0544 .0542	82	.0896 .0890
25	78	.0106 .0105	82	.0330	84	.0555 .0553	86	.0903 .0898

LEAF 108 OMITTED IN PAGE NUMBERING.

I = 9

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	27	3	27	4	26	5	25
4	5	35	6	34	7	33	8	32
5	8	42	10	40	10	40	12	38
6	11	49	13	47	14	46	15	45
7	15	55	17	53	18	52	19	51
8	18	62	20	60	21	59	23	57
9	22	68	24	66	25	65	27	63
10	25	75	28	72	29	71	31	69
11	29	81	32	78	33	77	35	75
12	33	87	35	84	37	83	39	81
13	37	93	40	90	41	89	43	87
14	41	99	44	96	45	95	48	92
15	45	105	48	102	49	101	52	98
16	49	111	52	108	54	106	56	104
17	53	117	56	114	58	112	60	110
18	57	123	60	120	62	118	65	115
19	61	129	64	126	66	124	69	121
20	65	135	68	131	70	130	73	127
21	69	141	72	137	75	135	78	132
22	73	147	76	143	79	141	82	138
23	77	153	80	149	83	147	86	144
24	81	159	84	155	87	152	91	149
25	86	164	90	160	92	158	95	155

1

I = 10

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	3	30	4	29	4	29	5	28
4	6	38	7	37	7	37	9	35
5	9	46	10	45	11	44	12	43
6	12	54	14	52	15	51	16	50
7	16	61	18	59	19	58	21	56
8	19	69	22	66	23	65	25	63
9	23	75	26	73	27	72	29	70
10	27	83	30	80	32	78	34	76
11	31	90	34	87	35	85	38	83
12	36	96	39	93	40	92	43	89
13	40	103	43	100	45	98	47	96
14	44	110	47	107	49	105	52	102
15	48	117	52	113	54	111	56	109
16	53	123	56	120	58	118	61	115
17	57	130	61	126	63	124	66	121
18	62	135	66	133	67	131	70	128
19	66	143	70	139	72	137	75	134
20	70	150	75	145	77	143	80	140
21	75	155	79	152	81	150	85	146
22	79	163	84	158	85	156	89	153
23	84	169	89	164	91	162	94	159
24	89	175	93	171	95	169	99	165
25	93	182	98	177	100	175	104	171

1%		3%		5%		10%	
UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER
.0100	.0400	.0350	.0349	.0400	.0400	.1000	.0993
.0150	.0350	.0252	.0299	.0350	.0349	.1260	.1233
.0126	.0252	.0300	.0299	.0462	.0460	.0792	.0786
.0092	.0300	.0318	.0317	.0500	.0497	.0800	.0791
.0114	.0318	.0316	.0315	.0501	.0499	.1140	.1126
.0075	.0316	.0302	.0301	.0480	.0477	.1030	.1017
.0081	.0302	.0282	.0281	.0447	.0445	.0918	.0910
.0082	.0282	.0259	.0258	.0580	.0577	.1116	.1101
.0081	.0259	.0329	.0328	.0515	.0514	.0969	.0960
.0114	.0329	.0292	.0292	.0455	.0454	.1117	.1104
.0105	.0292	.0258	.0257	.0542	.0540	.0959	.0952
.0096	.0258	.0306	.0305	.0471	.0469	.1072	.1060
.0087	.0306	.0257	.0260	.0540	.0538	.0916	.0909
.0107	.0257	.0305	.0305	.0465	.0464	.1002	.0993
.0095	.0305	.0265	.0264	.0520	.0518	.1082	.1073
.0113	.0265	.0297	.0297	.0446	.0445	.0972	.0914
.0099	.0297	.0329	.0328	.0490	.0489	.0984	.0977
.0087	.0329	.0282	.0282	.0532	.0530	.1041	.1032
.0100	.0282	.0308	.0307	.0455	.0454	.1094	.1085
.0087	.0308	.0333	.0332	.0488	.0487	.0932	.0925
.0098	.0333	.0285	.0285	.0520	.0519	.0973	.0967
.0109	.0285	.0305	.0305	.0550	.0548	.1011	.1003
.0094	.0305	.0470	.0469	.0470	.0469	.1046	.1039

I = 11

J	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT
3	3	33	.0083	.0331	4	32	.0253	.0262	4	32	.0331	.0331	5	31	.0826	.0821
4	6	42	.0113	.0316	7	41	.0253	.0262	8	40	.0526	.0522	9	39	.0947	.0932
5	9	51	.0086	.0315	11	49	.0316	.0315	12	48	.0541	.0538	13	47	.0879	.0871
6	13	59	.0107	.0310	15	57	.0311	.0310	15	56	.0497	.0494	18	54	.1150	.1131
7	17	67	.0110	.0284	19	65	.0284	.0283	20	64	.0435	.0434	22	62	.0950	.0940
8	21	75	.0104	.0250	23	73	.0250	.0249	25	71	.0543	.0539	27	69	.1086	.1071
9	25	83	.0094	.0312	28	80	.0312	.0311	29	79	.0447	.0445	31	77	.0870	.0863
10	29	91	.0093	.0259	32	88	.0259	.0258	34	85	.0508	.0505	36	84	.0940	.0930
11	34	98	.0105	.0297	37	95	.0297	.0296	39	93	.0555	.0553	41	91	.0988	.0979
12	38	105	.0088	.0329	42	102	.0329	.0328	43	101	.0443	.0441	46	98	.1019	.1008
13	43	113	.0102	.0254	45	110	.0254	.0263	48	108	.0470	.0468	51	105	.1036	.1027
14	48	120	.0115	.0282	51	117	.0282	.0282	53	115	.0489	.0487	56	112	.1041	.1031
15	52	128	.0093	.0297	56	124	.0297	.0296	58	122	.0502	.0500	61	119	.1038	.1029
16	57	135	.0102	.0308	51	131	.0308	.0307	63	129	.0510	.0507	66	126	.1026	.1017
17	62	142	.0109	.0315	56	138	.0315	.0315	68	136	.0513	.0511	71	133	.1010	.1002
18	66	150	.0087	.0320	71	145	.0320	.0319	73	143	.0512	.0510	76	140	.0989	.0980
19	71	157	.0092	.0322	76	152	.0322	.0321	78	150	.0508	.0507	81	147	.0964	.0957
20	76	164	.0095	.0322	81	159	.0322	.0321	83	157	.0502	.0500	86	154	.0937	.0930
21	81	171	.0098	.0320	86	166	.0320	.0319	88	164	.0494	.0492	91	161	.0908	.0902
22	86	178	.0100	.0317	91	173	.0317	.0316	93	171	.0484	.0482	97	167	.1060	.1052
23	91	185	.0102	.0312	96	180	.0312	.0312	98	178	.0472	.0471	102	174	.1020	.1012
24	96	192	.0102	.0307	101	187	.0307	.0306	103	185	.0460	.0458	107	181	.0979	.0972
25	101	199	.0102	.0300	106	194	.0300	.0300	109	191	.0541	.0539	112	188	.0938	.0933

I = 13

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	39	39	38	38	37	37	36	36
4	50	50	48	48	47	47	46	46
5	60	60	58	58	57	57	55	55
6	70	70	67	67	66	66	64	64
7	79	79	76	76	75	75	73	73
8	88	88	85	85	84	84	82	82
9	98	98	94	94	93	93	90	90
10	107	107	103	103	101	101	99	99
11	115	115	112	112	110	110	107	107
12	124	124	120	120	119	119	116	116
13	133	133	129	129	127	127	124	124
14	142	142	137	137	135	135	132	132
15	150	150	146	146	144	144	140	140
16	159	159	154	154	152	152	148	148
17	167	167	163	163	160	160	157	157
18	176	176	171	171	169	169	165	165
19	184	184	179	179	176	176	173	173
20	193	193	187	187	185	185	181	181
21	201	201	195	195	193	193	189	189
22	209	209	204	204	201	201	197	197
23	217	217	212	212	209	209	205	205
24	225	225	220	220	217	217	213	213
25	234	234	228	228	225	225	221	221



I = 14

J	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT
3	3	42	.0051	.0204	4	41	.0255	.0254	5	40	.0510	.0508
4	7	53	.0128	.0127	8	52	.0335	.0334	9	51	.0459	.0456
5	11	64	.0093	.0344	13	62	.0345	.0344	14	61	.0521	.0518
6	15	75	.0103	.0324	18	72	.0325	.0324	19	71	.0504	.0501
7	20	85	.0102	.0292	23	82	.0293	.0292	24	81	.0459	.0456
8	25	95	.0096	.0255	28	92	.0292	.0255	30	90	.0545	.0542
9	30	105	.0087	.0291	33	102	.0319	.0291	35	100	.0461	.0459
10	35	115	.0104	.0316	39	111	.0337	.0316	41	109	.0502	.0499
11	41	124	.0089	.0275	45	120	.0293	.0275	47	119	.0529	.0525
12	46	134	.0100	.0286	51	129	.0286	.0286	53	127	.0544	.0541
13	52	143	.0093	.0285	56	139	.0285	.0285	59	136	.0549	.0547
14	58	152	.0096	.0283	62	148	.0283	.0283	65	145	.0547	.0545
15	63	162	.0098	.0278	68	157	.0278	.0278	71	154	.0540	.0537
16	69	171	.0098	.0328	74	166	.0317	.0328	77	163	.0528	.0525
17	75	180	.0097	.0305	80	175	.0305	.0305	83	172	.0512	.0510
18	81	189	.0096	.0291	86	184	.0291	.0291	89	181	.0494	.0492
19	87	199	.0093	.0278	93	192	.0278	.0278	95	190	.0475	.0473
20	93	207	.0091	.0279	99	201	.0279	.0279	102	198	.0542	.0539
21	99	216	.0091	.0279	105	210	.0279	.0279	108	207	.0515	.0513
22	105	225	.0091	.0279	111	219	.0279	.0279	114	216	.0488	.0486
23	111	234	.0091	.0279	117	228	.0279	.0279	120	225	.0471	.0460
24	117	243	.0091	.0279	124	236	.0279	.0279	127	233	.0511	.0509
25	124	251	.0091	.0279	130	245	.0279	.0279	133	242	.0480	.0479

I = 15

J	1% MIN	1% MAX	1% UPPER EXACT	1% LOWER	3% MIN	3% MAX	3% UPPER EXACT	3% LOWER	5% MIN	5% MAX	5% UPPER EXACT	5% LOWER	10% MIN	10% MAX	10% UPPER EXACT	10% LOWER
3	4	44	.0178		4	44	.0179		5	43	.0444	.0443	6	42	.0889	.0880
4	7	57	.0104		9	55	.0373	.0371	10	54	.0622	.0616	11	53	.0978	.0961
5	11	69	.0091		13	67	.0254		15	65	.0593	.0589	16	64	.0863	.0854
6	16	80	.0105		18	78	.0244		20	76	.0510	.0507	22	74	.0982	.0969
7	21	91	.0102		24	88	.0304	.0303	25	85	.0575	.0572	28	84	.1029	.1017
8	26	102	.0091		30	98	.0340	.0338	31	97	.0456	.0453	34	94	.1027	.1014
9	32	112	.0109	.0108	35	109	.0270	.0269	37	107	.0468	.0466	40	104	.0995	.0985
10	37	123	.0089		41	119	.0278	.0277	43	117	.0465	.0462	46	114	.0945	.0935
11	43	133	.0096		47	129	.0277		49	127	.0451	.0449	53	123	.1094	.1083
12	49	143	.0099		53	139	.0271	.0270	55	136	.0537	.0534	59	133	.1006	.0996
13	55	153	.0100		50	148	.0326	.0325	62	145	.0501	.0499	65	143	.0919	.0912
14	61	163	.0099	.0098	56	158	.0307	.0306	68	155	.0465	.0463	72	152	.1007	.0997
15	67	173	.0096		72	168	.0296	.0285	75	165	.0520	.0518	79	161	.1084	.1074
16	73	183	.0092		79	177	.0323	.0322	81	175	.0474	.0472	85	171	.0971	.0963
17	80	192	.0108		85	187	.0295		88	184	.0516	.0515	92	180	.1027	.1019
18	86	202	.0101		92	196	.0326	.0325	94	194	.0465	.0465	99	189	.1074	.1065
19	92	212	.0094		98	206	.0296	.0295	101	203	.0498	.0497	105	199	.0955	.0949
20	99	221	.0106		105	215	.0318		108	212	.0527	.0525	112	208	.0989	.0981
21	105	231	.0098		111	225	.0287	.0286	114	222	.0472	.0470	119	217	.1017	.1010
22	112	240	.0108		118	234	.0305	.0304	121	231	.0493	.0491	126	226	.1041	.1033
23	118	250	.0098		125	243	.0321	.0320	128	240	.0512	.0511	133	235	.1060	.1052
24	125	259	.0106		131	253	.0287		135	249	.0529	.0527	139	245	.0937	.0931
25	131	269	.0096		138	262	.0300	.0299	141	259	.0470	.0469	146	254	.0949	.0944

I = 16

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3	4	47	5	46	5	45	6	45
4	7	61	9	59	10	58	12	56
5	12	73	14	71	15	70	17	68
6	17	85	19	83	21	81	23	79
7	22	97	25	94	27	92	29	90
8	28	108	31	105	33	103	36	100
9	33	120	37	116	39	114	42	111
10	39	131	43	127	45	124	49	121
11	45	142	50	137	52	135	55	132
12	52	152	56	148	59	145	62	142
13	58	163	63	158	65	156	69	152
14	64	174	70	168	72	165	76	162
15	71	184	76	179	79	175	83	172
16	77	195	83	189	86	185	90	182
17	84	205	90	199	93	195	97	192
18	91	215	97	209	100	205	104	202
19	97	225	104	219	107	215	111	212
20	104	235	111	229	114	225	118	222
21	111	245	118	239	121	235	126	231
22	118	256	125	249	129	245	133	241
23	125	266	132	259	135	256	140	251
24	132	275	139	269	142	266	148	260
25	139	285	146	279	150	275	155	270

EXACT UPPER LOWER
MIN MAX

EXACT UPPER LOWER
MIN MAX

EXACT UPPER LOWER
MIN MAX

EXACT UPPER LOWER
MIN MAX

J

I = 17

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3								
4	4	50	5	49	5	49	7	47
5	7	65	9	63	10	62	12	60
6	12	78	15	75	16	74	18	72
7	17	91	20	88	22	86	24	84
8	23	103	26	100	28	98	31	95
9	29	115	33	111	35	109	37	107
10	35	127	39	123	41	121	44	118
11	41	139	45	134	48	132	51	129
12	48	150	52	146	55	143	58	140
13	54	162	59	157	62	154	65	151
14	61	173	66	168	69	165	73	161
15	68	184	73	179	75	175	80	172
16	75	195	80	190	83	187	87	183
17	81	207	87	201	90	198	95	193
18	88	218	95	211	98	208	102	204
19	96	229	102	222	105	219	110	214
20	103	239	109	233	112	230	117	225
21	110	250	116	244	120	240	125	235
22	117	261	124	254	127	251	132	246
23	124	272	131	265	135	261	140	256
24	132	282	139	275	142	272	148	266
25	139	293	146	286	150	282	155	277
	146	304	154	296	159	292	163	287

J	1%		3%		5%		10%	
	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER
3								
4	.0138	.0345	.0346	.0345	.0346	.0345	.1211	.1191
5	.0071	.0255	.0256	.0255	.0427	.0425	.1008	.0990
6	.0095	.0358	.0350	.0358	.0523	.0520	.1026	.1013
7	.0087	.0272	.0273	.0272	.0526	.0522	.0948	.0935
8	.0102	.0101	.0272	.0335	.0490	.0487	.1079	.1067
9	.0104	.0104	.0336	.0335	.0565	.0561	.0916	.0906
10	.0101	.0297	.0297	.0483	.0485	.0483	.0956	.0947
11	.0093	.0326	.0327	.0326	.0514	.0511	.0967	.0957
12	.0107	.0276	.0277	.0276	.0528	.0526	.0958	.0949
13	.0074	.0287	.0288	.0287	.0530	.0529	.0934	.0926
14	.0101	.0292	.0293	.0292	.0524	.0521	.1070	.1060
15	.0106	.0292	.0293	.0292	.0511	.0509	.1016	.1007
16	.0109	.0288	.0288	.0491	.0493	.0491	.0959	.0952
17	.0091	.0281	.0281	.0473	.0473	.0471	.1051	.1042
18	.0091	.0322	.0323	.0322	.0529	.0527	.0980	.0973
19	.0109	.0308	.0309	.0308	.0499	.0497	.1052	.1043
20	.0106	.0293	.0293	.0469	.0469	.0467	.0973	.0966
21	.0103	.0277	.0278	.0277	.0509	.0507	.1029	.1021
22	.0100	.0304	.0305	.0304	.0474	.0473	.0947	.0941
23	.0096	.0285	.0286	.0285	.0507	.0505	.0991	.0984
24	.0107	.0308	.0308	.0470	.0470	.0468	.1030	.1024
25	.0102	.0286	.0287	.0286	.0497	.0495	.0943	.0937
	.0096	.0305	.0305	.0522	.0522	.0520	.0974	.0968

I = 19

J	1%		3%		5%		10%	
	MIN	LOWER EXACT	MIN	LOWER EXACT	MIN	LOWER EXACT	MIN	LOWER EXACT
3	4	.0111	5	.0277	5	.0554	7	.0970
4	8	.0102	10	.0306	11	.0481	13	.0700
5	13	.0099	16	.0335	17	.0475	19	.0892
6	19	.0110	22	.0301	24	.0544	26	.0930
7	25	.0102	29	.0332	31	.0558	33	.0903
8	31	.0088	35	.0336	38	.0540	41	.0894
9	38	.0095	43	.0325	45	.0505	48	.0926
10	45	.0097	50	.0305	52	.0462	56	.0986
11	52	.0095	57	.0279	60	.0502	64	.0986
12	59	.0090	55	.0306	68	.0531	72	.1022
13	67	.0103	73	.0325	75	.0462	80	.1037
14	74	.0094	80	.0284	83	.0473	88	.1038
15	82	.0102	88	.0294	91	.0478	96	.1026
16	90	.0108	95	.0293	99	.0475	104	.1006
17	97	.0095	104	.0299	107	.0477	112	.0978
18	105	.0099	112	.0299	115	.0471	121	.0947
19	113	.0101	120	.0299	124	.0533	129	.1036
20	121	.0102	128	.0288	132	.0517	137	.0991
21	129	.0102	135	.0281	140	.0500	146	.0944
22	137	.0102	145	.0281	149	.0482	154	.1011
23	145	.0101	153	.0312	157	.0525	163	.0956
24	153	.0079	161	.0301	165	.0501	171	.1012
25	161	.0097	170	.0290	174	.0477	180	.0952
				.0315		.0510		.1005
				.0314		.0508		.0999

I = 20

J	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT
3																
4	4	59	.0100	.0100	5	58	.0250	.0250	5	57	.0500	.0497	7	56	.0875	.0865
5	8	76	.0087	.0087	10	74	.0253	.0261	11	73	.0412	.0410	13	71	.0894	.0881
6	13	92	.0080	.0080	16	89	.0273	.0272	19	87	.0535	.0532	20	85	.0969	.0958
7	19	107	.0085	.0085	23	103	.0315	.0314	25	101	.0553	.0549	27	99	.0925	.0914
8	26	121	.0103	.0103	30	117	.0318	.0317	32	115	.0525	.0522	35	112	.1044	.1032
9	33	135	.0108	.0108	37	131	.0300	.0299	39	129	.0476	.0473	43	125	.1102	.1089
10	40	149	.0106	.0106	44	145	.0272	.0272	47	142	.0516	.0513	50	139	.0928	.0920
11	47	163	.0094	.0094	52	158	.0296	.0296	55	155	.0535	.0533	58	152	.0927	.0918
12	54	177	.0090	.0090	50	171	.0310	.0309	63	168	.0540	.0537	67	164	.1069	.1059
13	62	190	.0099	.0099	58	184	.0315	.0314	71	181	.0533	.0530	75	177	.1021	.1012
14	70	203	.0105	.0105	76	197	.0313	.0313	79	194	.0517	.0515	83	190	.0966	.0959
15	78	216	.0108	.0108	84	210	.0307	.0306	87	207	.0497	.0495	92	202	.1047	.1038
16	86	229	.0109	.0109	92	223	.0297	.0296	95	220	.0473	.0471	100	215	.0973	.0966
17	94	242	.0106	.0106	100	236	.0295	.0284	104	232	.0515	.0514	109	227	.1028	.1020
18	102	255	.0106	.0106	109	248	.0314	.0313	112	245	.0482	.0481	117	240	.0945	.0939
19	110	268	.0103	.0103	117	261	.0295	.0295	121	257	.0514	.0512	126	252	.0982	.0975
20	118	281	.0099	.0099	126	273	.0317	.0317	129	270	.0475	.0474	135	264	.1012	.1005
21	126	294	.0095	.0095	134	286	.0295	.0295	139	282	.0498	.0497	144	276	.1036	.1028
22	135	306	.0104	.0104	143	298	.0312	.0311	147	294	.0518	.0515	153	288	.1053	.1046
23	143	319	.0098	.0098	151	311	.0288	.0287	155	307	.0474	.0472	161	301	.0953	.0947
24	152	331	.0106	.0106	150	323	.0300	.0299	154	319	.0487	.0485	170	313	.0963	.0957
25	160	344	.0098	.0098	159	335	.0311	.0310	173	331	.0498	.0495	179	325	.0968	.0962
	169	356	.0105	.0104	177	348	.0284	.0284	182	343	.0507	.0505	188	337	.0970	.0965

I = 21

J	1%		3%		5%		10%	
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX
3								
4	4	62	5	61	6	60	7	59
5	8	80	11	77	12	75	14	74
6	14	96	17	93	18	92	21	89
7	20	112	24	108	25	106	28	104
8	27	127	31	123	33	121	36	118
9	34	142	38	138	41	135	44	132
10	41	157	46	152	49	149	52	146
11	49	171	54	166	57	163	61	159
12	57	185	62	180	65	177	69	173
13	65	199	71	193	74	190	78	186
14	73	213	79	207	82	204	87	199
15	81	227	87	221	91	217	96	212
16	89	241	96	234	99	231	104	226
17	98	254	105	247	108	244	113	239
18	106	268	113	261	117	257	123	251
19	115	281	122	274	125	270	132	264
20	123	295	131	287	135	283	141	277
21	132	308	140	300	144	295	150	290
22	141	321	149	313	153	309	159	303
23	149	335	158	326	162	322	169	315
24	158	348	167	339	171	335	178	328
25	167	361	176	352	181	347	187	341
	176	374	185	365	190	360	197	353

J	1%		3%		5%		10%	
	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER
3								
4	.0091	.0227	.0227	.0226	.0454	.0451	.0794	.0785
5	.0076	.0356	.0356	.0354	.0535	.0530	.1081	.1062
6	.0103	.0318	.0318	.0317	.0441	.0438	.1046	.1033
7	.0095	.0330	.0330	.0328	.0564	.0560	.0922	.0911
8	.0104	.0306	.0306	.0306	.0497	.0495	.0968	.0958
9	.0101	.0270	.0270	.0270	.0525	.0522	.0962	.0952
10	.0092	.0265	.0265	.0265	.0527	.0524	.0925	.0917
11	.0102	.0289	.0289	.0289	.0512	.0509	.1029	.1018
12	.0107	.0284	.0284	.0283	.0485	.0484	.0945	.0937
13	.0108	.0324	.0324	.0323	.0535	.0533	.1000	.0991
14	.0106	.0303	.0303	.0303	.0492	.0490	.1037	.1029
15	.0102	.0281	.0281	.0280	.0520	.0518	.1060	.1051
16	.0097	.0301	.0301	.0300	.0469	.0467	.0938	.0932
17	.0108	.0316	.0316	.0315	.0484	.0482	.0943	.0936
18	.0100	.0284	.0284	.0284	.0494	.0492	.1063	.1055
19	.0107	.0293	.0293	.0292	.0499	.0497	.1048	.1040
20	.0097	.0298	.0298	.0298	.0500	.0498	.1028	.1021
21	.0102	.0302	.0302	.0301	.0497	.0496	.1004	.0997
22	.0106	.0302	.0302	.0301	.0492	.0491	.0977	.0971
23	.0095	.0302	.0302	.0301	.0485	.0483	.1054	.1047
24	.0097	.0299	.0299	.0299	.0476	.0474	.1017	.1011
25	.0098	.0295	.0295	.0295	.0519	.0518	.0980	.0974
	.0099	.0290	.0290	.0290	.0505	.0503	.1040	.1034

I = 23

J	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	MIN	MAX	UPPER EXACT	LOWER EXACT	
3	4	68	.0076	.0378	.0377	5	.66	.0378	.0377	5	.66	.0378	.0377	8	64	.1059	.1042
4	9	87	.0104	.0271	.0270	13	85	.0589	.0582	13	83	.0589	.0582	15	81	.1122	.1102
5	15	105	.0107	.0306	.0305	19	102	.0554	.0550	19	100	.0554	.0550	22	98	.0941	.0930
6	21	123	.0084	.0275	.0274	25	119	.0460	.0457	27	117	.0460	.0457	30	114	.0922	.0912
7	29	139	.0105	.0299	.0288	33	135	.0454	.0452	35	133	.0454	.0452	39	129	.1034	.1023
8	36	156	.0089	.0290	.0279	41	151	.0515	.0513	44	148	.0515	.0513	48	144	.1083	.1071
9	44	172	.0090	.0315	.0314	50	166	.0458	.0455	52	164	.0458	.0455	56	160	.0923	.0915
10	53	187	.0107	.0290	.0279	58	182	.0474	.0472	61	179	.0474	.0472	66	174	.1065	.1054
11	61	203	.0098	.0290	.0289	67	197	.0476	.0474	70	194	.0476	.0474	75	189	.1022	.1014
12	70	218	.0105	.0292	.0291	76	212	.0467	.0465	79	209	.0467	.0465	84	204	.0968	.0960
13	78	234	.0092	.0289	.0289	85	227	.0522	.0520	89	223	.0522	.0520	94	218	.1037	.1029
14	87	249	.0094	.0240	.0279	94	242	.0495	.0493	98	238	.0495	.0493	103	233	.0960	.0953
15	96	264	.0094	.0309	.0308	104	256	.0531	.0529	108	252	.0531	.0529	113	247	.1000	.0993
16	105	279	.0093	.0292	.0291	113	271	.0494	.0492	117	267	.0494	.0492	123	251	.1029	.1021
17	115	293	.0144	.0312	.0312	123	285	.0517	.0515	127	281	.0517	.0515	133	275	.1048	.1040
18	124	308	.0100	.0290	.0290	132	300	.0475	.0473	135	295	.0475	.0473	143	289	.1058	.1050
19	133	323	.0095	.0304	.0303	142	314	.0489	.0487	145	310	.0489	.0487	152	304	.0954	.0949
20	143	337	.0103	.0315	.0314	152	328	.0499	.0497	155	324	.0499	.0497	162	318	.0955	.0949
21	152	352	.0096	.0298	.0287	151	343	.0505	.0504	165	338	.0505	.0504	173	331	.1051	.1045
22	162	366	.0102	.0294	.0294	171	357	.0509	.0507	175	352	.0509	.0507	183	345	.1039	.1032
23	171	381	.0094	.0299	.0298	181	371	.0509	.0508	185	365	.0509	.0508	193	359	.1024	.1018
24	181	395	.0078	.0302	.0301	191	385	.0508	.0506	195	380	.0508	.0506	203	373	.1005	.0999
25	191	409	.0101	.0303	.0303	201	399	.0504	.0503	205	394	.0504	.0503	213	387	.0984	.0979

I = 24

J	1%		3%		5%		10%					
	MIN	MAX	MIN	MAX	MIN	MAX	MIN	MAX				
	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER	UPPER	LOWER				
	EXACT	EXACT	EXACT	EXACT	EXACT	EXACT	EXACT	EXACT				
3	5	70	.0174	.0173	.0347	.0346	.0608	.0603	8	67	.0972	.0958
4	9	91	.0091	.0091	.0358	.0356	.0517	.0513	15	85	.0987	.0972
5	15	110	.0091	.0090	.0258	.0257	.0467	.0465	23	102	.1014	.1002
6	22	128	.0094	.0094	.0299	.0288	.0473	.0470	31	119	.0925	.0914
7	30	145	.0107	.0106	.0281	.0281	.0538	.0535	40	135	.0971	.0962
8	38	162	.0107	.0106	.0315	.0314	.0466	.0464	49	151	.0964	.0955
9	46	179	.0100	.0100	.0273	.0272	.0471	.0469	59	166	.1084	.1073
10	55	195	.0109	.0109	.0276	.0275	.0541	.0539	68	182	.1007	.0997
11	63	212	.0094	.0094	.0318	.0317	.0511	.0509	78	197	.1061	.1052
12	72	228	.0095	.0095	.0302	.0301	.0474	.0472	87	213	.0957	.0949
13	81	244	.0094	.0094	.0282	.0282	.0501	.0499	97	228	.0975	.0968
14	91	259	.0106	.0106	.0300	.0300	.0518	.0515	107	243	.0979	.0972
15	100	275	.0100	.0100	.0313	.0313	.0528	.0526	117	258	.0972	.0966
16	109	291	.0094	.0094	.0293	.0282	.0531	.0529	127	273	.0957	.0950
17	119	306	.0100	.0100	.0298	.0287	.0529	.0527	138	287	.1044	.1036
18	129	321	.0104	.0104	.0290	.0289	.0523	.0521	148	302	.1011	.1004
19	138	337	.0095	.0095	.0290	.0289	.0513	.0511	158	317	.0974	.0969
20	148	352	.0097	.0097	.0247	.0286	.0500	.0499	169	331	.1034	.1028
21	158	367	.0098	.0098	.0316	.0315	.0485	.0484	179	346	.0988	.0982
22	168	382	.0099	.0099	.0308	.0308	.0521	.0519	190	360	.1035	.1029
23	178	397	.0099	.0099	.0300	.0299	.0501	.0499	200	375	.0982	.0977
24	188	412	.0098	.0098	.0290	.0290	.0480	.0479	211	389	.1019	.1013
25	198	427	.0097	.0097	.0310	.0309	.0505	.0504	221	404	.0963	.0958

I = 25

J	1%		3%		5%		10%	
	MIN	UPPER EXACT	MIN	UPPER EXACT	MIN	UPPER EXACT	MIN	UPPER EXACT
3	5	.0160	6	.0320	7	.0560	8	.0896
4	9	.0081	12	.0317	13	.0458	15	.0874
5	16	.0112	19	.0298	21	.0521	24	.1088
6	23	.0103	27	.0303	29	.0486	32	.0928
7	31	.0108	35	.0275	38	.0516	41	.0918
8	39	.0101	44	.0289	47	.0511	51	.1023
9	48	.0109	53	.0285	55	.0485	61	.1081
10	56	.0092	53	.0323	65	.0524	70	.0956
11	65	.0091	72	.0296	75	.0471	80	.0961
12	75	.0103	82	.0312	85	.0482	90	.0948
13	84	.0096	92	.0320	95	.0483	101	.1034
14	94	.0103	101	.0287	105	.0477	111	.0998
15	104	.0107	112	.0318	115	.0525	121	.0949
16	113	.0095	122	.0310	125	.0506	132	.1000
17	123	.0096	132	.0302	135	.0483	143	.1041
18	133	.0095	142	.0290	147	.0512	153	.0970
19	143	.0094	153	.0310	157	.0482	164	.0995
20	154	.0104	153	.0294	168	.0503	175	.1013
21	164	.0100	174	.0309	179	.0519	186	.1026
22	174	.0097	184	.0290	183	.0483	197	.1033
23	185	.0103	195	.0300	200	.0493	208	.1030
24	195	.0098	206	.0309	211	.0501	219	.1034
25	206	.0104	216	.0287	222	.0507	230	.1029

APPENDIX B

COMPUTER PROGRAM DOCUMENTATION

```

$CONTROL FILE=5,FILE=6
C
C   A FORTRAN PROGRAM FOR CALCULATING ONE-SIDED PERCENTAGE
C   POINTS FOR YU DEN'S TEST BY EXHAUSTIVE ENUMERATION
C   USING BOTH LEHMER'S METHOD AND WELLS'S METHOD OF
C   COMBINATORIC ENUMERATION.
C
C   THE CODING IS SPECIFIC TO THE HP3000 SYSTEM.
C
      CHARACTER ANS*1
      DIMENSION KOUNT(24),NSUMS(6),DIST(80)
      COMMON LIST(760,6),NPERM
5     DISPLAY "ENTER NUMBER OF OBJECTS AND JUDGES"
      ACCEPT NORJ,JUDGES
      IF(NOBJ.EQ.0)STOP
      NPERM=1
      DO 7 I=2,NOBJ
7     NPERM=NPERM*I
      DISPLAY "LEHMER OR WELLS"
      ACCEPT ANS
      IF(ANS.EQ."W") GOTO 10
      CALL LEHMER(NOBJ)
      GOTO 15
10    CALL MRWELLS(NOBJ)
15    DO 20 J=1,JUDGES
20    KOUNT(J)=1
      DO 22 K=1,80
22    DIST(K)=0.0
25    DO 30 I=1,NOBJ
30    NSUMS(I)=I
      DO 40 J=1,JUDGES
      K=KOUNT(J)
      DO 35 I=1,NOBJ
35    NSUMS(I)=NSUMS(I)+LIST(K,I)
40    CONTINUE
      MIN=999
      DO 45 I=1,NOBJ
      IF(NSUMS(I).LT.MIN)MIN=NSUMS(I)
45    CONTINUE
      DIST(MIN)=DIST(MIN)+1.0
      DO 50 J=1,JUDGES
      KOUNT(J)=KOUNT(J)+1
      IF(KOUNT(J).LE.NPERM)GOTO 25
      KOUNT(J)=1
50    CONTINUE
      WRITE(6,100)(DIST(K),K=1,80)
100   FORMAT(" RAW FREQUENCY COUNTS"/,(1X,10F7.0/))
      DO 55 K=2,80
55    DIST(K)=DIST(K)+DIST(K-1)
      WRITE(6,110)(DIST(K),K=1,80)

```

```

110  FORMAT(" ACCUMULATED FREQUENCY COUNTS"/,(1X,10F7.0/))
      P=DISI(80)
      DO 60 K=1,80
60    DISI(K)=DISI(K)/P
      WRITE(6,120)(DISI(K),K=1,80)
120  FORMAT(" STANDARDIZED FREQUENCY COUNTS"/,(1X,10F7.5/))
      GOTO 5
      END
      SUBROUTINE LEHMER(NOBJ)
      DIMENSION NDIGITS(6),NDIGITS2(6)
      COMMON LIST(760,6),NPEPM
      DO 30 K=0,NPEPM-1
      K1=K1+1
      DO 5 I=1,NOBJ
5     NDIGITS(I)=0
      ND=1
10    ND=ND+1
      M=N2/ND
      NDIGITS(ND-1)=N2-M*ND+1
      N2=M
      IF(M.GE.1)GOTO 10
      IF((ND-1).EQ.NOBJ)GOTO 15
      L2=ND-1
      DO 13 I=NOBJ,1,-1
      IF(L2.EQ.0)GOTO 12
      NDIGITS(I)=NDIGITS(L2)
      L2=L2-1
      GOTO 13
12    NDIGITS(I)=0
13    CONTINUE
      LIST(K1,1)=NDIGITS(NOBJ)
15    DO 25 K=1,NOBJ-1
      DO 20 J=I+1,NOBJ
      NDIGITS2(J)=NDIGITS(J-1)
      IF(NDIGITS2(J).GE.NDIGITS(J))NDIGITS2(J)=NDIGITS2(J)+1
20    CONTINUE
      LIST(K1,I+1)=NDIGITS2(NOBJ)
25    CONTINUE
30    CONTINUE
      RETURN
      END
      SUBROUTINE MWELLS(N)
      DIMENSION A(10),XM1(10)
      COMMON D(720,6),NF
      DO 10 I=1,10
      A(I)=0.0
10    XM1(I)=0.0
      DO 12 I=1,N
12    XM1(I)=I
      DO 90 M=1,NF

```

```
      N1=M-1
      J=0
      NH=0
15     J=J+1
      I=J+1
      NQ=N1/I
      NR=N1- $\sqrt{NQ}$ *I
      IF (NH) 30,20,30
20     IF (NR-.J) 25,30,25
25     NH=J
30     A(J)=NR
      IF (NQ) 35,40,35
35     N1=NQ
      GO TO 15
40     IF (NH) 50,45,50
45     NH=J+1
50     L=M
      DO 47 I=1,N
47     D(L,I)=XM1(I)
      NS=NH+1
      IF (NH-(NH/2)*2) 60,55,60
55     IF (A(NH+1)-2.) 60,70,70
60     W=XM1(NS)
      XM1(NS)=XM1(NS-1)
      XM1(NS-1)=W
      GO TO 90
70     NX=NH-A(NH+1)
      IF (NX) 75,75,80
75     NX=1
80     W=XM1(NS)
      XM1(NS)=XM1(NX)
      XM1(NX)=W
90     CONTINUE
      RETURN
      END
```

```
PROGRAM TST(INPUT,OUTPUT,TAPE5=INPUT,TAPE6=0(OUTPUT))
```

```
C
C A FORTRAN PROGRAM FOR CALCULATING ONE-SIDED PERCENTAGE
C POINTS FOR YOUNG'S TEST IN THE CASE OF  $1 = 2$ .
C
C THIS PROGRAM IS BASED ON THE BINOMIAL THEOREM AND IS
C SPECIFIC TO THE CDC6400 SYSTEM.
C
```

```

      DIMENSION NA(200),NB(200)
10  READ(5,100)LO,NHI
100 FORMAT(2I3)
      IF(EOF(5))1000,20
20  *WRITE(5,110)
110 FORMAT(1H1)
      CALL NBINOM(LO-2,NA)
      NS=2
      MS=2*LO+1
      DO 80 J=LO,NHI
      NS=NS*2
      K=J-2
      DO 40 I=1,K
40  NB(I+1)=NA(I)+NA(I+1)
      NB(1)=1
      NB(J)=1
      DO 50 I=1,J
50  NA(I)=NB(I)
      DO 55 I=1,200
      L=J/2
      DO 60 I=1,J
      K=J+I
      IF(I.GT.L)K=K+J-1-I
60  NB(K)=NB(K)+NA(I)
      L=L+J
      *WRITE(5,120)J,NS
120 FORMAT(* C = 2, J = *,I3,* SUM = *,I14)
      C=0.0
      MS2=MS
      DO 70 I=J,L
      P=FLOAT(NB(I))/FLOAT(NS)
      C=C+P
      MS2=MS2-1
70  *WRITE(5,130)I,MS2,NB(I),P,C
130 FORMAT(1X,I2,2X,I2,2X,I14,2X,F8.6,2X,F8.6)
      MS=MS+2
80  CONTINUE
      GOTO 10
1000 STOP
      END
      SUBROUTINE NBINOM(N,LIST)
      DIMENSION LIST(1)
```



```
MID=N/2+1
LIST(1)=1
LIST(N+1)=1
IF(MID.LE.1)RETURN
DO 10 NX=2,MID
LIST(NX)=LIST(NX-1)*(N-NX+1)/NX
10 LIST(N-NX+2)=LIST(NX)
RETURN
END
```

```
$CONTROL FILE=5,FILE=1,FILE=2,MAP,LABEL
```

```
C
```

```
C A FORTRAN PROGRAM FOR CALCULATING ONE-SIDED PERCENTAGE  
C POINTS FOR YOUNG'S TEST USING AN ITERATIVE PROCEDURE  
C BASED ON BONFERRONI'S INEQUALITIES.
```

```
C
```

```
C
```

```
C
```

```
THE CODING IS SPECIFIC TO THE HP3000 SYSTEM.
```

```
COMMON /UPPER/ CINV,VC,PROB(8400),MPTR(25)
```

```
COMMON /LOWER/ CINVL,NCL,MLPTR(25)
```

```
5 DISPLAY "ENTER LOW VALUES FOR C AND J"
```

```
ACCEPT NCL0,JL0
```

```
IF(NCL0.EQ.0)STOP
```

```
DISPLAY "ENTER HIGH VALUES FOR C AND J"
```

```
ACCEPT NCHI,JHI
```

```
DISPLAY "ENTER LOWER & UPPER CUM. PROB. VALUES"
```

```
ACCEPT CUMPL0,CUMPHI
```

```
DO 30 VC=NCL0,NCHI
```

```
NCL=VC-1
```

```
MPTR(1)=1
```

```
MLPTR(1)=1
```

```
DO 7 K=2,25
```

```
MLPTR(K)=MLPTR(K-1)+NCL*K-K
```

```
7 MPTR(K)=MPTR(K-1)+NC*K-K
```

```
LIMITP=MPTR(25)+NC*26-27
```

```
LIMITL=MLPTR(25)+NCL*26-27
```

```
DO 8 K=1,LIMITP
```

```
8 PROB(K)=0.0
```

```
REWIND 2
```

```
DUM=0.0
```

```
DO 9 K=1,LIMITL
```

```
9 WRITE(2)DUM
```

```
C=FLOAT(NC)
```

```
CL=C-1.
```

```
CINV=1./C
```

```
CINVL=1./CL
```

```
DO 20 J=JL0,JHI
```

```
MAXEXACT=(J+1)/2+J
```

```
KMAX=NC*J
```

```
KHI=(KMAX-J)/2+J
```

```
CUMP=0.0
```

```
CUMPL=0.0
```

```
DO 10 K=J,KHI
```

```
CALL T%U(J,K,P)
```

```
PR=C*P
```

```
CUMP=C*JMP+PR
```

```
ALPHAL=CUMP
```

```
IF(K.LT.MAXEXACT)GO TO 15
```

```
CALL T%L(J,K-MAXEXACT+J,P)
```

```
PRL=CL*P
```

```

CUMPL=CUMPL+PRL
ALPHAL=CUMPL-CUMPL*PR
15 IF(CUM>.L1.CUMPL0) GO TO 10
   IF(CJM>.GT.CUMPHI)GO TO 20
   *WRITE(1,100)NC,J,K,KMAX,CUM,ALPHAL
100 FORMAT(12,1X,12,1X,13,1X,13,2F9.6)
10 KMAX=KMAX-1
20 CONTINJE
30 CONTINJE
   GO TO 5
   END
SUBROUTINE TWU(J,K,P)
COMMON /UPPER/ CINV,NC,PROB(8400),MPTR(25)
P=0.0
IF(K.GT.NC*J)RETURN
IF(J.NE.0) GO TO 15
IF(K.EQ.0)P=1.0
RETURN
15 IF(J.NE.1)GO TO 18
   P=CINV
   RETURN
18 IF(J.NE.K)GO TO 20
   P=CINV**J
   RETURN
20 N1=MPTR(J)+K-1
   IF(PROB(N1).EQ.0.0)GO TO 25
   P=PROB(N1)
   RETURN
25 L=K-J+1
   IF(L.GT.NC)L=NC
   DO 30 I=1,L
   CALL TWU(J-1,K-I,PS)
30 P=P+PS
   P=P*CINV
   PROB(N1)=P
   RETURN
   END
SUBROUTINE TWL(J,K,P)
COMMON /LOWER/ CINVL,NCL,MLPTR(25)
P=0.0
IF(K.GT.NCL*J)RETURN
IF(J.NE.0)GO TO 15
IF(K.EQ.0)P=1.0
RETURN
15 IF(J.NE.1)GO TO 18
   P=CINVL
   RETURN
18 IF(J.NE.K)GO TO 20
   P=CINVL**J
   RETURN

```

```
20  N1=MLPTR(J)+K-1
    READ(2@N1)PROBL
    IF(PROBL.EQ.0.0)GO TO 25
    P=PROBL
    RETURN
25  L=K-J+1
    IF(L.GT.NCL)L=NCL
    DO 30 I=1,L
    CALL TNL(J-1,K-I,PS)
30  P=P+PS
    P=P*CI/VL
    WRITE(2@N1)P
    RETURN
    END
```

```

SCONTROL MAP,LABEL,FILE=5,FILE=6
C
C   A FORTRAN PROGRAM FOR GENERATING POWER CURVES FOR
C   NAIR'S TEST AND YOUNG'S UNDER SPECIFIED PARENT
C   POPULATION ASSUMPTIONS.
C
C   THE CODING IS HP3000 SPECIFIC.
C
      DIMENSION A(25,25),B(25,25)
5     WRITE(5,110)
110   FORMAT(" ENTER I,J, # OF SIMULATIONS")
      READ(5,*)I,J,NITER
      IF(I.EQ.0)STOP
      LIM=1-1
      XNITER=FLOAT(NITER)/100.
      WRITE(5,120)
120   FORMAT(" ENTER CRITICAL 1 & 5% VALUES FOR NAIR & YOUNG")
      READ(5,*)C1,C2,NY1,NY2
      WRITE(5,130)
130   FORMAT(" ENTER LOWEST DELTA, STEPSIZE & NUMBER OF STEPS")
      READ(5,*)XLOW,STEP,NUMDELTA
      WRITE(6,140)
140   FORMAT(" ENTER SEED")
      READ(5,*)SEED
      WRITE(6,150)I,J,NITER,C1,C2,NY1,NY2,SEED
150   FORMAT(" I= ",I2," J= ",I2," SIMULATIONS= ",I4/
+ " CRITICAL POINTS= ",F4.2,3X,F4.2,3X,I4,3X,I4/
+ " RANDOM NO. SEED= ",F10.5//)
      IF(SEED.EQ.0.0)GO TO 7
      DUM=RAVD(SEED)
7     DELTA=XLOW-STEP
      DO 90 ND=1,NUMDELTA
      DELTA=DELTA+STEP
      MINN3=0
      MINN4=0
      MINN1=0
      MINN2=0
      MINY1=0
      MINY2=0
      DO 80 ITER=1,NITER
      DO 20 L=1,J
      DO 17 K=1,LIM
C
C   CHANGE THE FOLLOWING TWO LINES
C   TO REFLECT THE PARENT POPULATION DESIRED.
C
17    A(K,L)=UNIFORM(0.0,1.0,DUM)
20    A(I,L)=UNIFORM(0.0,1.0,DUM)-DELTA
      CALL NAIR(I,J,A,STATVR,MINOBJ)
      IF(MINOBJ.NE.I)GO TO 30

```

```

IF (STATNR.GE.C1) MINN3=MINN3+1
IF (STATNR.GE.C2) MINN4=MINN4+1
30 DO 60 L=1,J
XM=0.0
DO 50 M=1,I
XM=XM+1.
XMIN=99.E70
DO 40 K=1,I
IF (A(K,L)-XM) 35,40,40
35 XMIN=A(K,L)
NPTR=K
40 CONTINUE
A(NPTR,L)=99.E70
50 B(NPTR,L)=XM
60 CONTINUE
CALL NAIR(I,J,B,STATNR,MINOBJ)
IF (MINOBJ.NE.I) GO TO 78
IF (STATNR.GE.C1) MINN1=MINN1+1
IF (STATNR.GE.C2) MINN2=MINN2+1
78 CALL YJUDEN(I,J,B,MINOBJ,NOBJSUM)
IF (MINOBJ.NE.I) GO TO 80
IF (NOBJSUM.LE.NY1) MINY1=MINY1+1
IF (NOBJSUM.LE.NY2) MINY2=MINY2+1
80 CONTINUE
YMIN1=FLOAT(MINY1)/XVITER
YMIN2=FLOAT(MINY2)/XVITER
XMIN1=FLOAT(MINN1)/XVITER
XMIN2=FLOAT(MINN2)/XVITER
XMIN3=FLOAT(MINN3)/XVITER
XMIN4=FLOAT(MINN4)/XVITER
WRITE(5,100) DELTA,XMIN1,XMIN2,YMIN1,YMIN2
100 FORMAT(1X,F3.1,4F10.1)
WRITE(5,105) XMIN3,XMIN4
105 FORMAT(4X,2F10.1)
IF (XMIN1.EQ.100..AND.XMIN1.EQ.100.) GO TO 5
90 CONTINUE
GO TO 5
END
SUBROUTINE NAIR(I,J,D,STATNR,NR)
DIMENSION C(25),D(25,25)
QI=I
QJ=J
SSQR5=0.0
SSQCS=0.0
GSSQ=0.0
GS=0.0
DO 5 L=1,J
5 C(L)=0.0
XM=99.E70
DO 20 K=1,I

```

```

RS=0.0
DO 10 L=1,J
P=D(K,L)
RS=RS+P
C(L)=C(L)+P
10 GSSQ=GSSQ+P*P
GS=GS+RS
SSQRS=SSQRS+RS*RS
IF(RS.GE.XMIN)GO TO 20
XMIN=RS
NR=K
20 CONTINUE
DO 30 L=1,J
30 SSQCS=SSQCS+C(L)*C(L)
GM=GS/(QI*QJ)
EV=(GSSQ-SSQCS/QI-SSQRS/QJ+GS*GS/(QI*QJ))/((QI-1.)*(QJ-1.))
STATNR=(GM-XMIN/QJ)/SQRT(EV/QJ)
RETURN
END
SUBROUTINE YOUNEN(I,J,A,MINOBJ,NOBJSUM)
DIMENSION A(25,25),TEMP(25)
C
DO 3 K=1,I
-3 TEMP(K)=0.0
DO 50 K=1,I
DO 40 L=1,J
40 TEMP(K)=TEMP(K)+A(K,L)
50 CONTINUE
XMIN=1000.
DO 60 K=1,I
IF(TEMP(K).GE.XMIN)GO TO 60
XMIN=TEMP(K)
MINOBJ=K
60 CONTINUE
NOBJSUM=IFIX(XMIN)
RETURN
END
FUNCTION UNIFORM(XMEAN,VARIANCE,DUM)
WIDTH=SQRT(12.*VARIANCE)
HALFWIDTH=WIDTH/2.
UNIFORM=RAND(DUM)*WIDTH+XMEAN-HALFWIDTH
RETURN
END
FUNCTION DNORM(DUM)
DNORM=-6.0
DO 10 I=1,12
10 DNORM=DNORM+RAND(DUM)
RETURN
END

```

```

$CONTROL FILE=5,FILE=6
C
C   A FORTRAN PROGRAM FOR ANALYSING DATA ORGANISED
C   IN A TWO WAY CLASSIFICATION BY YOUJEN'S TEST AND
C   NAIR'S TEST.
C
C   THIS PROGRAM WAS WRITTEN FOR THE HP3000 COMPUTER
C   SYSTEM. MODIFICATIONS MAY BE NECESSARY BEFORE
C   SUCCESSFULLY EXECUTING THIS PROGRAM ON ANOTHER
C   COMPUTER.
C
      COMMON A(25,25),XXX,MMM,DUM
      DIMENSION B(25,25)
10  DISPLAY "ENTER ROWS & COLUMNS"
C
C   BOTH NROW AND NCOL HAVE A MAXIMUM VALUE
C   OF 25 AND DEPEND ON THE DIMENSIONS OF
C   ARRAYS A AND B.
C   NROW <= 0 TERMINATES THE PROGRAM.
C
      ACCEPT NROW,NCOL
      IF(NROW.LE.0)STOP
      DUM=RAVD(17.)
      NC=NCOL
      NR=NROW
      DO 15 I=1,NROW
15  ACCEPT(B(I,J),J=1,NCOL)
      GOTO 30
20  DISPLAY "OPTION?"
      ACCEPT NOP
C
C   THE OPTIONS ARE:
C   1  ENTRY OF NEW DATA
C   2  REINITIALIZE WITH EXISTING DATA
C   3  CALCULATE USING NAIR AND YOUJEN'S TESTS.
C   4  INVERT MAIRIX
C   5  TEMPORARILY DELETE A ROW
C   6  DISPLAY PRESENT STATUS.
C   7  SPECIFY EITHER MAX OR MIN EXTREME DEVIATE
C
      GOTO(10,30,50,70,90,110,130),NOP
30  DO 40 I=1,NROW
      DO 40 J=1,NCOL
40  A(I,J)=B(I,J)
      NC=NCOL
      NR=NROW
      GOTO 20
50  CALL YOUJEN(NR,NC,MINROW,MINSUM)
      DISPLAY"YOUJEN.,ROW=",MINROW,"RANK SUM=",MINSUM

```



```

CALL NAIR(NR,NC,STAT,MINROW)
DISPLAY"NAIR..ROW=",MINROW,"STATISTIC=",STAT
GOTO 20
70  LARGE=VR
   IF(NC.GT.NR)LARGE=NC
   DO 80 J=1,LARGE
   DO 80 J=I+1,LARGE
   TEMP=A(I,J)
80  A(I,J)=A(J,I)
   A(J,I)=TEMP
   NTEMP=VR
   NR=NC
   NC=NTEMP
   GOTO 20
90  DISPLAY "ROW?"
   ACCEPT ND
   DO 100 I=ND,NK-1
   DO 100 J=1,NC
100 A(I,J)=A(I+1,J)
   NR=NR-1
   GOTO 20
110 DISPLAY"NR/NC",NR,"/",NC,"MINMAX=",MMM
   DISPLAY ((A(I,J),J=1,NC),I=1,NR)
   GOTO 20
130 DISPLAY "MAX=1,MIN=0"
   ACCEPT NOP
C
C   TO TEST FOR MAX. EXTREME DEVIATE ANSWER 1
C   "   "   "   MIN.   "   "   "   2
C   THIS WILL STAY IN EFFECT UNTIL CHANGED
C   BUT MUST BE INITIALIZED WHEN NEW DATA IS ENTERED
C
   XXX=1.
   MMM=1
   IF(NOP,EQ,0)GOTO 20
   XXX=-1.
   MMM=-1
   GOTO 20.
END
SUBROUTINE NAIR(I,J,STATNR,NR)
COMMON D(25,25),XXX,MMM
DIMENSION C(25)
QI=I
QJ=J
SSQRS=0.0
SSQCS=0.0
GSSQ=0.0
GS=0.0
DO 5 L=1,J
5  C(L)=0.0

```

```

XMIN=99.E70
DO 20 K=1,I
RS=0.0
DO 10 L=1,J
P=D(K,L)
RS=RS+P
C(L)=C(L)+P
10 GSSQ=GSSQ+P*P
GS=GS+RS
SSQRS=SSQRS+RS*RS
IF((RS*XXX).GE.XMIN)GO TO 20
XMIN=RS*XXX
NR=K
20 CONTINUE
XMIN=XMIN*XXX
DO 30 L=1,J
30 SSQCS=SSQCS+C(L)*C(L)
GM=GS/(QI*QJ)
EV=(GSSQ-SSQCS/QI-SSQRS/QJ+GS*GS/(QI*QJ))/((QI-1.)*(QJ-1.))
STATNR=(GM-XMIN/QJ)/SQRT(EV/QJ)
RETURN
END
SUBROUTINE YOUDEN(NR,NC,MINOBJ,NOBJSUM)
COMMON A(25,25),XXX,MMM,DUM
DIMENSION N(25,25),NTEMP(25)
DO 10 I=1,NR
DO 10 J=1,NC
10 N(I,J)=0
DO 30 LC=1,NC
M=0
NSTRT=FLOAT(NR)*RAND(DUM)+1.
DO 30 LR=1,NR
M=M+1
XMIN=99.E70
DO 20 K=1,NR
K1=MOD(NSTRT+K,NR)+1
IF(A(K1,LC).GT.XMIN.OR.N(K1,LC).NE.0)GOTO 20
XMIN=A(K1,LC)
NPTR=K1
20 CONTINUE
30 N(NPTR,LC)=M
DO 40 I=1,NR
NTEMP(I)=0
DO 40 J=1,NC
40 NTEMP(I)=NTEMP(I)+N(I,J)
NOBJSUM=1000
NSTRT=FLOAT(NR)*RAND(DUM)+1.
DO 50 I=1,NR
KI=MOD(NSTRT+I,NR)+1
IF((NTEMP(KI)*MMM).GE.NOBJSUM)GOTO 50

```

```
NOBJSUM=NTEMP(KI)*MMM  
MINOBJ=KI  
50 CONTINUE  
NOBJSUM=NOBJSUM*MMM  
DISPLAY"RANKSUM VALUES"  
DISPLAY (NTEMP(I),I=1,NR)  
RETURN  
END
```