## AN EVALUATION OF THE RELATIONSHIP BETWEEN INCOME AND HEALTH

# A QUANTITATIVE EVALUATION OF THE RELATIONSHIP BETWEEN INCOME AND SELF-REPORTED HEALTH IN CANADA, 1996-1999 

## By

## SIMO GEORGIEV GOSHEV

A Project<br>Submitted to the School of Graduate Studies in Partial Fulfillment of the Requirements for the Degree<br>Master of Science

McMaster University
December 2002

Hamilton, Ontario

TITLE: A Quantitative Evaluation of the Relationship Between Income and Self-reported Health in Canada, 1996-1999

AUTHOR: Simo Georgiev Goshev

SUPERVISOR: Dr. Michael Veall, Department of Economics, McMaster University


#### Abstract

The focus of the research is on the quantitative characterization of a potential relationship between income and health. The study is conducted in both levels and increments contexts. As a background, the research relies on Grossman's standard economic model of health as well as on previous empirical investigations. The source of the data is the Survey of Income and Labour Dynamics, administered by Statistics Canada. The set of instruments implemented includes ordered probit and ordered logit models, gamma GLM models, the OLS method and bootstrapping techniques in linear models. The results of the study suggest that there exists a strong association between the levels of health and income, whereas no evidence of significant association between their increments is present.


## Acknowledgement

I would like to express my sincere gratitude to my supervisor, Professor Michael Veall, for providing me with the opportunity to work in the research area of income and health relationship, for his expert guidance and support at all levels.

I would also like to especially thank Professor P. M. Macdonald for his perfect teaching, his constant support and encouragement throughout my studies.

Thanks are extended to all my instructors for their teaching and guidance in the course work as well as to Cindy Cook and James Chowhan at the Statistics Canada Research Data Centre at McMaster University for their assistance in the data management and results disclosure. Thanks go to my friends for their permanent encouragement and help.

Finally, I would like to thank my parents for their love and constant support.

## Contents

1 Introduction ..... 1
1.1 Problem Foundations ..... 1
1.2 Objectives ..... 1
2 Data Description ..... 3
2.1 SLID Overview ..... 3
2.2 Variables Under Consideration ..... 4
2.3 Data Cleaning and Manipulation ..... 5
3 The Approach ..... 9
4 Methodology ..... 10
4.1 Economic Methodology ..... 10
4.1.1 Levels Analysis ..... 10
4.1.2 Increments Analysis ..... 11
4.1.3 Auxiliary Models ..... 13
4.1.4 Income Deflation Procedure ..... 13
4.2 Statistical Methodology ..... 15
4.2.1 Ordinary Least Squares ..... 15
4.2.2 Latent Variable Approach in Categorical Data Analysis ..... 16
4.2.3 Ordered Probit/Logit Models ..... 19
4.2.4 Regression with Gamma Distribution in the Generalized Linear Models Case ..... 20
4.2.5 Marginal Effects in Ordered Probit/Logit Models ..... 21
4.2.6 Marginal Effects in the Gamma distribution GLM ..... 22
4.2.7 Likelihood Ratio Test ..... 23
4.2.8 Wald Test ..... 24
4.2.9 Bootstrap Confidence Intervals of Parameters Estimates in Lin- ear Regression ..... 25
4.3 Assumptions in the Analysis ..... 27
4.4 Research Hypotheses ..... 28
4.5 Problems ..... 29
5 Levels Models Results ..... 30
5.1 Exploration of the Variable Relationships ..... 30
5.2 The Health Models ..... 39
5.3 The Income Model ..... 47
6 Increments Models Results ..... 53
6.1 Exploration of the Variable Relationships ..... 53
6.2 The Health Model ..... 59
6.3 The Income Model ..... 63
7 Discussion ..... 67
8 Future Directions ..... 70
9 Conclusion ..... 71
10 References ..... 73
11 Appendix ..... 75

## 1 Introduction

### 1.1 Problem Foundations

The two-way relationship between income and health has been a main focus of research attention for the last half century. Although researchers from different fields have provided evidence of the existence of the association, a complete economic theoretical model consistent with the relationship was not available until 1972 when Michael Grossman created the standard economic model of health (Grossman, 1972).

The model includes a health production function in which the health at any given time is a result of factors, including health in previous time period, medical care received, adoption of healthy personal behaviors, etc. (Smith, 1999).

The economic reasoning behind his approach is that health is a stock. The current inputs and chosen behaviors are treated like investments producing increments to that stock. The increments themselves are affected by the current personal choices and thus today's health stock is a function of the entire history of all current and past prices, income, health behaviors and personal health endowment (Smith, 1999). However, health can also affect available economic resources. Thus health enters the model in two ways producing a two-way association (Smith, 1999).

### 1.2 Objectives

The objective of the research is to characterize the relationship between income and health, conditioned on income and health from previous periods, for Canada in the
following aspects (McLeod and Veall, 2002)

- Levels analysis: The relationship between current health, lagged health and lagged income as well as its direction,
- Increments analysis: The relationship between change in health, change in income, lagged change in health and lagged change in income as well as its direction,
- Discovering supplemental predictors and investigating channels that influence the relationships.

The analysis is based on the longitudinal section of the Survey of Labour and Income Dynamics (SLID) administered by Statistics Canada.

The overall temporal framework of the data spans from 1996 to 1999 by years. In particular, for the levels models, the years from 1996 to 1999 were used, unless otherwise indicated. The increments analysis, on the other hand, was performed for the changes between 1996 to 1998.

The number of respondents for the years 1996 to 1998 is 67364 and for the years 1996 to 1999 it is 71278 .

The analysis does not take into account the SLID sampling weights. In addition, there is no adjustment for the clustering in the sample.

## 2 Data Description

### 2.1 SLID Overview

The Survey of Labour and Income Dynamics was initiated in 1993 with the purpose of understanding the economic well-being of Canadians by collecting data on labor market activity, income and related socio-economic and demographic characteristics. The survey consists of a cross-sectional block and a longitudinal block.

As in all recurring surveys, in the cross-sectional SLID survey a new sample of people is being interviewed each time, thus making the data collected more representative of the population and revealing the levels and trends of income or labor for the whole population or sub-groups.

SLID's longitudinal aspect, on the other hand, provides information on the changes experienced by same individuals through time, in that way creating a powerful background for studying transitions, durations, and repeat occurrences in individual financial and work situations.

To keep the longitudinal sample representative, the respondents in each panel stay in the survey for 6 years. A new panel of longitudinal respondents is selected every three years, so there is always an overlap between two panels of respondents. Each panel includes about 15,000 households, including about 30,000 adults.

The SLID is administered by Statistics Canada and the data are collected by computer-assisted telephone interviews. A preliminary interview takes place at the beginning of each panel to collect background information. Each of the six years
has a split-interview format, with labor topics covered in January and income topics in May. In both cases, questions refer to the previous calendar year. The income interview occurs in May to take advantage of income tax time when respondents are more familiar with their records.

### 2.2 Variables Under Consideration

The two basic raw variables under consideration in the analysis are the current state of health variable and the economic family after-tax-income variable.

The health variable (crhlt26) represents a response to the question "Compared to other people his/her age, how would you describe [respondent]'s state of health? Would you say it is... " with possible answers "...excellent?", "...very good?", "...good?", "...fair?", "...poor?", coded as one to five respectively. The health variable is observed for persons aged 16 or older.

The income variable (atinc27) represents the economic family after-tax income for a specific year. The respondents have the option of answering the questions on income in an interview, or giving permission to Statistics Canada to allow SLID to use the information from their income tax return. Over $80 \%$ of the respondents give their consent for accessing their administrative records. The income variable is continuous, in the range -9999999: 99999995, it is measured in dollars (CDN) and is recorded for all persons.

Additional variables considered in the analysis are

- Highest level of education (hlev2g18) recorded for persons aged 15 or older, having four categories corresponding to less than high school education (LHS), graduated high school (GHS), non-university postsecondary certificate (NUPC)
and university degree (UD), coded as one to four respectively,
- Annual labor force status (alfst28) observed for persons aged 16-69, having seven categories: Employed all year; Unemployed all year; Not in the labor force all year (NLF); Employed part-year, unemployed part-year (EPY-UPY); Employed part-year, not in labor force part-year (EPY-NLFPY); Unemployed part-year, not in labor force part-year (UPY-NLFPY); Employed, unemployed and not in labor force during year (AOA), coded as one to seven respectively,
- Person's age as of December 31 of the reference year (age26) is observed for all persons, it is continuous and within the range $0: 150$,
- Sex (sex21) variable is recorded for all person and has two categories: Male; Female, coded as one and two respectively,
- Current stress level in persons life (crstr26) refers to persons aged 16 or older, having four categories: Very stressful; Somewhat stressful; Not very stressful; Not at all stressful, coded as one to four respectively.

In the modelling part, good health, LHS, employed all year, male, very stressful are most frequently used as reference categories.

### 2.3 Data Cleaning and Manipulation

Data management procedures were implemented in the following order:

1. Based on the variables characterizing the respondent status in the SLID, investigation of possible dropouts due to various reasons was implemented. Due to a partial substitution of subject in the panel, which took place in 1999, there
were 32430 dropouts for that year. The resulting useful portion of the data for the period 1996 to 1999 decreased from 67364 to 34934 individuals.
2. The review of the raw variables used in the analysis showed that all of them contained the additional categories

$$
\begin{array}{ll}
\diamond \text { Not in Sample, } & \diamond \text { Refusal, } \\
\diamond \text { Don't Know, } & \diamond \text { Not Applicable, }
\end{array}
$$

and the variable Current stress level also having
$\diamond$ No Opinion.

As these categories could be treated as uninformative, the subjects that had corresponding records were removed. The removal however, was implemented selectively, depending on the model considered and the factors included in it, thus ensuring that no useful data have been lost.
3. The income variables in the SLID are nominal income, but real income is the one of interest. To overcome that problem, a deflation procedure was employed.
4. Differenced health, differenced stress and differenced income variables were created

- As the variable Current state of health (crhlt26) is categorical, the differenced health variables were constructed on the basis of code subtraction. As a result, nine categories where obtained, which were aggregated into five by the rule

$$
\Delta h_{i, t}= \begin{cases}1, & \text { if } h_{i, t}-h_{i, t-1} \in\{4,3\} ; \\ 2, & \text { if } h_{i, t}-h_{i, t-1} \in\{1,2\} ; \\ 3, & \text { if } h_{i, t}-h_{i, t-1} \in\{0\} ; \\ 4, & \text { if } h_{i, t}-h_{i, t-1} \in\{-1,-2\} ; \\ 5, & \text { if } h_{i, t}-h_{i, t-1} \in\{-3,-4\} ;\end{cases}
$$

$$
\text { where } i=1, \ldots, n, \quad t=1997,1998
$$

Thus, according to the new coding the categories defined for the change in health variable (health) are: Improved substantially (1); Improved a bit (2); Unchanged (3); Worsened (4); Worsened badly (5). Unchanged is used as a reference category in the modelling part.

- The differenced stress variable (stress) was defined by grouping the differences from the category codes of crstr26 in the following way

$$
\begin{gathered}
\Delta s_{i, t}= \begin{cases}1, & \text { if } s_{i, t}-s_{i, t-1} \in\{-3\} \\
2, & \text { if } s_{i, t}-s_{i, t-1} \in\{-1,-2\} \\
3, & \text { if } s_{i, t}-s_{i, t-1} \in\{0\} \\
4, & \text { if } s_{i, t}-s_{i, t-1} \in\{1,2\} \\
5, & \text { if } s_{i, t}-s_{i, t-1} \in\{3\}\end{cases} \\
\text { where } i=1, \ldots, n, \quad t=1997,1998
\end{gathered}
$$

The newly formed categories of change in stress level were defined as: Substantially reduced stress (1), Reduced stress (2), Unchanged (3), Increased stress (4), Substantially increased (5). Substantially reduced stress and, less frequently, unchanged stress are used as a reference categories in the modelling part.

- The differenced income variables (Datinc27) were obtained by straightforward subtraction by the rule:

$$
\Delta I_{i, t}=I_{i, t}-I_{i, t-1}, \quad i=1, \ldots, n, \quad t=1997,1998
$$

5. As some extremely small and extremely large observations of income were present in the data, the raw income for the two years of interest and the differenced income were trimmed by removing all cases having records less than the first or greater than the last percentile ( $2 \%$ trimming was performed).
6. For the purposes of the analysis of the income health relationships by age groups, the respondents have been divided into four groups as follows:

- Under 20 years,
- From 20 to 49,
- From 50 to 59,
- 60 and over.

The individuals have been allocated to the different groups based on their age within a three year period. People who in most of the years were within one age group were regarded as members of that group. The border cases were treated as members of the lower age group.

## 3 The Approach

Research on the two-way relationship between health and income has a long history but many questions concerning the degree of direct and indirect causality, incremental contributions and relationships still remain. The present research tries to provide an answer to some of these issues.

In that context, the approach involves a preliminary investigation of the relationships in the data by means of conventional descriptive tools and, at a second step, a study of the causality, directly or indirectly influenced by other predictors, in the levels models and the increments models by regression techniques application.

## 4 Methodology

### 4.1 Economic Methodology

### 4.1.1 Levels Analysis

## Health models

The levels models used in the analysis are characterized by the use of current health status as a response variable and a number of explanatory variables, including income and lagged health, as predictors.

The first model under consideration is a levels model that investigates the relationship between current health, lagged health and lagged income defined as follows

$$
\begin{equation*}
h_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t-2}+\beta_{i, 2} I_{i, t-2}+\mathbf{p}_{i, t-2} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1998, \quad i=1, \ldots, n_{1} \tag{1}
\end{equation*}
$$

where:

- $h_{i, t}$ is current health status,
- $h_{i, t-2}$ is lagged health status,
- $I_{i, t-2}$ is lagged income,
- $\mathbf{p}_{i, t-2}$ is a vector of other lagged predictors like level of education, labor force status, age,
- $\mathbf{z}_{i}$ is the vector of regression parameters corresponding to the elements of $\mathbf{p}_{i, t-2}$,
- $\varepsilon_{i, t}$ are the random errors,
- $i=1, \ldots, n_{1}$ where $n_{1}$ is the number of individuals in the period 1996-1998.

The other model considered is the one investigating the relationship between current health on the left-hand side and health in 1996 and income in 1996 complemented by additional variables on the right-hand side. It is represented by

$$
\begin{equation*}
h_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t-3}+\beta_{i, 2} I_{i, t-3}+\mathbf{p}_{i, t-3} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1999, \quad i=1, \ldots, n_{2} \tag{2}
\end{equation*}
$$

where $i=1, \ldots, n_{2}, n_{2}$ being the number of individuals in the period 1996-1999.

## Income models

The income levels models in the analysis are used for predicting current income conditioned on health, income and other covariates from previous periods. The two specific forms of investigation are

$$
\begin{equation*}
I_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t-2}+\beta_{i, 2} I_{i, t-2}+\mathbf{p} \prime_{i, t-2} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1998, \quad i=1, \ldots, n_{1} \tag{3}
\end{equation*}
$$

where current income $I_{i, t}$ is predicted by the two-year lagged health $h_{i, t-2}$, the twoyear lagged income $I_{i, t-2}$ and other two-year lagged covariates in the vector $\mathbf{p}_{i, t-2}$, and

$$
\begin{equation*}
I_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t-3}+\beta_{i, 2} I_{i, t-3}+\mathbf{p} \prime_{i, t-3} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1999, \quad i=1, \ldots, n_{2} \tag{4}
\end{equation*}
$$

where the time span is increased by one year.

### 4.1.2 Increments Analysis

In the increments analysis two basic models, the health model and the income model, are considered. The first relates change in health as a response to lagged change in
health and lagged change in income along with other predictors (Veall, 2002)

## Health model

$$
\begin{equation*}
\Delta h_{i, t}=\beta_{i, 1} \Delta h_{i, t-1}+\beta_{i, 2} \Delta I_{i, t-1}+\mathbf{p}_{i} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1998, \quad i=1, \ldots, n \tag{5}
\end{equation*}
$$

where:

- $\Delta h_{i, t}=h_{i, t}-h_{i, t-1}$ is change in health status,
- $\Delta h_{i, t-1}=h_{i, t-1}-h_{i, t-2}$ is lagged change in health status,
- $\Delta I_{i, t-1}=I_{i, t-1}-I_{i, t-2}$ is lagged change in income,
- $\mathbf{p}_{i}$ is a vector of other predictors like level of education, labor force status, age,
- $\mathbf{z}_{i}$ is the vector of regression parameters corresponding to the elements of $\mathbf{p}_{i}$,
- $\varepsilon_{i, t}$ are the random errors.
- $i=1, \ldots, n$ where $n$ is the number of individuals in the period 1996-1998.

The second model investigates the relationship between change in income as a response and the lagged change in income, the lagged change in health predictors complemented by additional explanatory variables (Veall, 2002)

## Income model

$$
\begin{equation*}
\Delta I_{i, t}=\beta_{i, 1} \Delta I_{i, t-1}+\beta_{i, 2} \Delta h_{i, t-1}+\mathbf{p}_{i}^{\prime} \mathbf{z}_{i}+\varepsilon_{i, t}, \quad t=1998, \quad i=1, \ldots, n \tag{6}
\end{equation*}
$$

where $\Delta I_{i, t}=I_{i, t}-I_{i, t-1}$ is change in income.

### 4.1.3 Auxiliary Models

In the analysis, some auxiliary models have been used for preliminary assessment of the relationships between health and income and vise versa existing in the data.

The models

$$
\begin{aligned}
& h_{i, t}=\beta_{i, 0}+\beta_{i, 1} I_{i, t}+\varepsilon_{i, t}, \\
& I_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t}+\varepsilon_{i, t},
\end{aligned}
$$

where $t=1996, \ldots, 1999$ and

$$
\begin{aligned}
& h_{i, t}=\beta_{i, 0}+\beta_{i, 1} I_{i, t-j}+\varepsilon_{i, t}, \\
& I_{i, t}=\beta_{i, 0}+\beta_{i, 1} h_{i, t-j}+\varepsilon_{i, t},
\end{aligned}
$$

where $t=1997, \ldots, 1999, \quad j=1, \ldots, 3, \quad 1996 \leq t-j \leq 1999$,
have been estimated separately for each $t$.

### 4.1.4 Income Deflation Procedure

For obtaining the real (deflated) income the following formula has been used

$$
D I_{t}=\left(\frac{I_{t}}{C P I_{t}}\right) \times 100
$$

where:

- $D I_{t}$ is the real (deflated) income in the period $t$,
- $I_{t}$ is the nominal income in the period $t$,
- $C P I_{t}$ is the consumer price index in $t$ calculated with reference to a predefined base.

The base change of the chain CPI's has been done in the following way

$$
C P I_{t_{N}}=100+\left(\frac{C P I_{t}-C P I_{N B}}{C P I_{N B}}\right) \times 100
$$

where:

- $C P I_{t_{N}}$ is the $C P I$ in the period $t$ having a base $C P I_{N B}$,
- $C P I_{N B}$ is the new base of the chain indices.

As a base of the $C P I$ 's in the analysis the $C P I_{1996}$ has been used.

### 4.2 Statistical Methodology

### 4.2.1 Ordinary Least Squares

The Ordinary Least Squares procedure is based on the minimization of the sum (Draper and Smith, 1998)

$$
\begin{equation*}
\varepsilon^{\prime} \varepsilon=(\mathrm{Y}-\mathrm{X} \boldsymbol{\beta})^{\prime}(\mathrm{Y}-\mathrm{X} \boldsymbol{\beta}) \tag{7}
\end{equation*}
$$

where

- $\mathbf{Y}$ is a $(n \times 1)$ vector of observations,
- $\mathbf{X}$ is a $(n \times p)$ known matrix,
- $\boldsymbol{\beta}$ is a $(p \times 1)$ vector of parameters,
- $\varepsilon$ is a $(n \times 1)$ vector of errors, and $E(\varepsilon)=0, V(\varepsilon)=\mathbf{I} \sigma^{2}$

By taking the partial derivatives of the above equation with respect to the elements of $\boldsymbol{\beta}$, the normal equations

$$
\left(\mathbf{X}^{\prime} \mathbf{X}\right) \mathrm{b}=\mathrm{X}^{\prime} \mathrm{Y}
$$

are obtained. The least squares estimates of $\boldsymbol{\beta}$ is $\mathbf{b}$, which in case that there are $p$ independent equations in the normal equations system, are determined by

$$
\begin{equation*}
\mathbf{b}=\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1} \mathbf{X}^{\prime} \mathbf{Y} \tag{8}
\end{equation*}
$$

The properties of the solution $\mathbf{b}$, irrespective of the distribution of the errors are:

- b minimizes the error sum of squares $\varepsilon^{\prime} \varepsilon$,
- $\mathbf{b}$ is an unbiased estimate of $\boldsymbol{\beta}$ and has minimum variance in the class of linear unbiased estimators $\mathbf{W}^{\prime} \mathbf{Y}$, where $\mathbf{W}$ is not a function of $\mathbf{Y}$.


### 4.2.2 Latent Variable Approach in Categorical Data Analysis

Let the model of interest be (Greene, 1990)

$$
\mathrm{Y}^{*}=\mathrm{X} \boldsymbol{\beta}+\boldsymbol{\varepsilon}
$$

where

- $\mathrm{Y}^{*}$ is the vector of observations of the variable $Y^{*}$, having a continuous distribution and
- $E\left(\mathbf{Y}^{*}\right)=\mathbf{X} \boldsymbol{\beta}$.

However, $Y^{*}$ is unobserved and there are no values recorded. Instead, there is a variable $Y$ which is discrete and ordered. The relationship between the latent variable $Y^{*}$ and the realized discrete and ordered outcome is

$$
y_{i}=j \quad \text { if } \quad \delta_{j-1}<y_{i}^{*} \leq \delta_{j}
$$

where $\delta_{j}$ is a threshold value such that

$$
\delta_{0}<\delta_{1}<\ldots<\delta_{J-1}<\delta_{J}, \quad \delta_{0}=-\infty, \delta_{J}=\infty
$$

and $J$ is the number of categories of the discrete and ordered outcome.
The model for $Y^{*}$ could be expressed as a cumulative probability model (Powers and Xie, 2000)

$$
P\left(y_{i} \leq j \mid \mathbf{x}_{i}\right)=P\left(y^{*} \leq \delta_{j} \mid \mathbf{x}_{i}\right)=P\left(\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}+\varepsilon_{i} \leq \delta_{j}\right)=P\left(\varepsilon_{i} \leq \delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)=F\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)
$$

where $F(\cdot)$ is the distribution function of $\varepsilon$ and $\mathbf{x}_{i}$ is the vector of predictors corresponding to $y_{i}$.

Then the cumulative probabilities under the latent variable model (without an intercept) have the form (Powers and Xie, 2000)

$$
P\left(y_{i}=j \mid \mathbf{x}_{i}\right)= \begin{cases}F\left(\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & j=1 ;  \tag{9}\\ F\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-F\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & 1<j \leq J-1 ; \\ 1-F\left(\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & j=J .\end{cases}
$$

The estimation of the parameters $\delta_{j}$ and $\boldsymbol{\beta}$ of the ordered probability model is usually done by the Newton-Raphson method. A set of ordinal variables (Maddala, 1985)

$$
Z_{i j}=\left\{\begin{array}{ll}
1, & \text { if } Y_{i} \text { falls in the } j \text { th category; } \\
0, & \text { otherwise }
\end{array} \quad(i=1, \ldots, n, j=1, \ldots, J)\right.
$$

is defined such that

$$
P\left(Z_{i j}=1\right)=F\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-F\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)
$$

The likelihood function for the model is then

$$
L=\prod_{i=1}^{n} \prod_{j=1}^{J}\left\{F\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-F\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)\right\}^{Z_{i j}}
$$

and the log-likelihood function is

$$
l=\sum_{i=1}^{n} \sum_{j=1}^{J} Z_{i j} \log \left\{F\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-F\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)\right\}
$$

The maximum likelihood estimates of the parameters are the solutions to

$$
\left.\frac{\partial l}{\partial \boldsymbol{\beta}}\right|_{\boldsymbol{\beta}=\widehat{\boldsymbol{\beta}}}=0,\left.\quad \frac{\partial l}{\partial \boldsymbol{\delta}}\right|_{\delta=\widehat{\delta}}=0
$$

where no constraints are posed, the second partial derivatives are less than 0 and the second partial cross derivatives are equal to 0 , where these are defined as:

$$
\frac{\partial^{2} l}{\partial \boldsymbol{\beta}^{2}}, \quad \frac{\partial^{2} l}{\partial \boldsymbol{\beta} \partial \boldsymbol{\beta}^{\prime}}, \quad \frac{\partial^{2} l}{\partial \delta^{2}}, \quad \frac{\partial^{2} l}{\partial \boldsymbol{\delta} \partial \boldsymbol{\delta}^{\prime}}, \quad \frac{\partial^{2} l}{\partial \boldsymbol{\beta} \partial \delta} .
$$

The Newton-Raphson procedure is

$$
\widehat{\boldsymbol{\theta}} \approx \boldsymbol{\theta}^{*}+J^{-1}\left(\boldsymbol{\theta}^{*}\right) U\left(\boldsymbol{\theta}^{*}\right)
$$

where

- $\widehat{\boldsymbol{\theta}}$ is the vector of Newton-Raphson MLE's,
- $\boldsymbol{\theta}^{*}$ is the vector with initial values,
- $J(\cdot)$ is the observed information matrix based on the second derivatives,
- $U(\cdot)$ is the vector of first derivatives,
and is implemented by

1. Choosing a sensible starting value $\boldsymbol{\theta}^{(0)}$,
2. Updating the estimate

$$
\widehat{\boldsymbol{\theta}}^{(k+1)}=\boldsymbol{\theta}^{(k)}+J^{-1}\left(\boldsymbol{\theta}^{(k)}\right) U\left(\boldsymbol{\theta}^{(k)}\right)
$$

until

$$
\left\|\widehat{\boldsymbol{\theta}}^{(k+1)}-\widehat{\boldsymbol{\theta}}^{(k)}\right\|<\varepsilon
$$

for obtaining the MLE's of the parameters. The variances and covariances of the estimates are obtained by

$$
V(\widehat{\boldsymbol{\theta}})=\left(\left.J\right|_{\boldsymbol{\theta}=\widehat{\boldsymbol{\theta}}}\right)^{-1}
$$

### 4.2.3 Ordered Probit/Logit Models

The distribution of the error term $F(\cdot)$, or the distribution of the latent variable $Y^{*}$, could be any of the wide variety of continuous distributions. However, a frequent choice of distribution is either the standard normal, $N(0,1)$ or the logistic $\operatorname{logistic}\left(0, \frac{\pi^{2}}{3}\right)$, mainly due to the fact that they both have a mean of 0 and are symmetric.

In the case the latent variable has a standard normal distribution, the cumulative probabilities are given by (Powers and Xie, 2000)

$$
P\left(y_{i}=j \mid \mathbf{x}_{i}\right)= \begin{cases}\Phi\left(\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & j=1  \tag{10}\\ \Phi\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-\Phi\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & 1<j \leq J-1 \\ 1-\Phi\left(\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right), & j=J\end{cases}
$$

where $\Phi(\cdot)$ is the cumulative standard normal distribution function.
In the case the distribution of the latent variable is logistic, the cumulative probabilities are

$$
P\left(y_{i}=j \mid \mathbf{x}_{i}\right)= \begin{cases}\frac{\exp \left\{\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{1+\exp \left\{\delta_{1} \mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}, & j=1 ;  \tag{11}\\ \frac{\exp \left\{\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{1+\exp \left\{\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}-\frac{\exp \left\{\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{1+\exp \left\{\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}, & 1<j \leq J-1 \\ 1-\frac{\exp \left\{\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{1+\exp \left\{\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}, & j=J .\end{cases}
$$

Unlike the ordinal probit model, the ordered logit model is linear in the logistic scale. This means that (Powers and Xie, 2000)

$$
\log \left(\frac{P\left(y_{i} \leq j \mid \mathbf{x}_{i}\right)}{1-P\left(y_{i} \leq j \mid \mathbf{x}_{i}\right)}\right)=\log \left(\frac{P\left(y_{i} \leq j \mid \mathbf{x}_{i}\right)}{P\left(y_{i}>j \mid \mathbf{x}_{i}\right)}\right)=\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}
$$

In addition, it is also often called proportional odds model, since given two covariate vectors $\mathbf{x}_{i 1}$ and $\mathbf{x}_{i 2}$, the odds of a response $y_{i} \leq j$ versus $y_{i}>j$ are proportionally
higher or lower in the situations $\mathbf{x}_{i}=\mathbf{x}_{1}$ and $\mathbf{x}_{i}=\mathbf{x}_{2}$. This could be seen if the cumulative odds ratio (COR) is defined

$$
\frac{\omega\left(\mathbf{x}_{1}\right)}{\omega\left(\mathbf{x}_{2}\right)}=\frac{P\left(y \leq j \mid \mathbf{x}_{1}\right) / P\left(y>j \mid \mathbf{x}_{1}\right)}{P\left(y \leq j \mid \mathbf{x}_{2}\right) / P\left(y>j \mid \mathbf{x}_{2}\right)}=\frac{\exp \left(\delta_{j}-\mathbf{x}_{1}^{\prime} \boldsymbol{\beta}\right)}{\exp \left(\delta_{j}-\mathbf{x}_{2}^{\prime} \boldsymbol{\beta}\right)}=\exp \left(\mathbf{x}_{1}-\mathbf{x}_{2}\right)^{\prime} \boldsymbol{\beta}
$$

where $\omega(\cdot)$ denotes the cumulative odds associated with either $\mathbf{x}_{1}$ or $\mathbf{x}_{2}$. The COR is proportional to the distances between the values of the explanatory variables. It is important also to notice that the effects of $\mathbf{x}$ are invariant with respect to the outcome categories, namely $\boldsymbol{\beta}$ is not indexed by $j$. The cumulative log-odds ratio is also invariant with respect to $j$

$$
\frac{\omega\left(\mathbf{x}_{1}\right)}{\omega\left(\mathbf{x}_{2}\right)}=\left(\mathbf{x}_{1}-\mathbf{x}_{2}\right)^{\prime} \boldsymbol{\beta} .
$$

The proportional log-odds model defines that the odds of being in each category $j$ or higher compared to all previous categories is constant. Also, the estimated coefficients $\widehat{\boldsymbol{\beta}}$ represent log-odds ratios.

### 4.2.4 Regression with Gamma Distribution in the Generalized Linear Models Case

The gamma distribution in the GLM modelling is employed whenever the errors of a regression model follow a gamma distribution. Thus, $Y \sim \operatorname{Gamma}(\mu, \nu)$ and according to the parametrization of the gamma distribution used in the GLM estimation (McCullagh and Nelder, 1989)

$$
f(y ; \nu, \mu)=\frac{\left(\frac{\nu y}{\mu}\right)^{\nu} \exp \left\{-\frac{\nu y}{\mu}\right\}}{\Gamma(\nu) y}, \quad y>0, \nu>0
$$

where $\nu$ is the shape parameter, the $\log$ likelihood for each $Y_{i}$ has the form

$$
l_{i}\left(\theta_{i} ; y_{i}\right)=\nu\left(-\frac{y_{i}}{\mu_{i}}-\log \mu_{i}\right)+(\nu-1) \log y_{i}+\nu \log \nu-\log \Gamma(\nu)
$$

Using the general form of the log-likelihood for the exponential family (McCullagh and Nelder, 1989)

$$
l(\theta, \varphi ; y)=\frac{(y \theta-\gamma(\theta))}{\varphi(\theta)}+\tau(y, \varphi(\theta))
$$

where $\theta$ is a natural parameter, $\varphi$ is the dispersion or scale parameter, it is possible to see that for the gamma distribution $\theta_{i}=-\frac{1}{\mu_{i}}, \gamma(\theta)=-\log \left(-\theta_{i}\right), \varphi\left(\theta_{i}\right)=\frac{1}{\nu}$ and the $E\left(Y_{i}\right)=\gamma^{\prime}\left(\theta_{i}\right)=\mu_{i}$ and $\operatorname{Var}\left(Y_{i}\right)=\gamma^{\prime \prime}\left(\theta_{i}\right)=\mu_{i}^{2}$.

Commonly used links for the gamma distribution are

$$
\begin{gathered}
\eta\left(\mu_{i}\right)=\mu_{i}^{-1} \\
\eta\left(\mu_{i}\right)=\log \mu_{i} \\
\eta\left(\mu_{i}\right)=\mu_{i}
\end{gathered}
$$

which are called inverse, $\log$ and identity links respectively.
The deviance, which is proportional to twice the difference between the $\log$ likelihood achieved under the model and the maximum attainable value is (McCullagh and Nelder, 1989)

$$
D(\mathbf{y} ; \widehat{\mu})=-2 \sum_{i=1}^{n}\left\{\log \left(\frac{y_{i}}{\widehat{\mu_{i}}}\right)-\frac{y_{i}-\widehat{\mu_{i}}}{\widehat{\mu_{i}}}\right\}
$$

where $y_{i}$ are the observed response values and $\widehat{\mu_{i}}$ are the fitted mean values.
The estimation of the model based on the gamma distribution is realized by applying the Newton-Raphson procedure.

### 4.2.5 Marginal Effects in Ordered Probit/Logit Models

Marginal effects account for the change in the dependent variable for a unit change in the independent variable in the simple linear regression model. In the multiple regression setting, the marginal effects account for the change in the dependent variable
for a unit change in a specific independent variables, where all other predictors are held constant. In the OLS regression models where no transformation has been made on any of the predictors or the response, the marginal effects are the estimates of the regression parameters. However, in the case of the ordered probit/logit models, the situation is slightly different.

For the ordered probit model the marginal effects of $x_{k}$ are given by (Powers and Xie, 2000)

$$
\frac{\partial P\left(y_{i}=j \mid \mathbf{x}_{i}\right)}{\partial x_{i k}}= \begin{cases}-\phi\left(\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right) \beta_{k}, & j=1  \tag{12}\\ \left\{\phi\left(\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)-\phi\left(\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right)\right\} \beta_{k}, & 1<j \leq J-1 \\ \phi\left(\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right) \beta_{k}, & j=J\end{cases}
$$

where $\phi(\cdot)$ is the probability density function of the standard normal distribution.
In the ordered logit situation the marginal effects of $x_{k}$ are given by

$$
\frac{\partial P\left(y_{i}=j \mid \mathbf{x}_{i}\right)}{\partial x_{i k}}= \begin{cases}-\left\{\frac{\exp \left\{\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{\left[1+\exp \left\{\delta_{1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}\right]^{2}}\right\} \beta_{k}, & j=1 ;  \tag{13}\\ \left\{\frac{\exp \left\{\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{\left[1+\exp \left\{\delta_{j-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}\right]^{2}}-\frac{\exp \left\{\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}}{\left[1+\exp \left\{\delta_{j}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}\right]^{2}}\right\} \beta_{k}, & 1<j \leq J-1 ; \\ \left\{\frac{\exp \left\{\delta_{J-1}-\mathbf{x}_{\mathbf{\prime}}^{\prime} \boldsymbol{\beta}\right\}}{\left[1+\exp \left\{\delta_{J-1}-\mathbf{x}_{i}^{\prime} \boldsymbol{\beta}\right\}\right]^{2}}\right\} \beta_{k}, & j=J .\end{cases}
$$

### 4.2.6 Marginal Effects in the Gamma distribution GLM

The way the marginal effects for the fitted gamma GLM are estimated depends on the link function used. Considering the links already presented, the marginal effects could be presented as follows:

- Link $\eta(\mu)=\mu^{-1}$

$$
\frac{\partial \mu}{\partial X_{j}}=-(\mathbf{X} \boldsymbol{\beta})^{-2} \beta_{j}
$$

since the model under consideration has the form $\mu_{i}=(\mathbf{X} \boldsymbol{\beta})^{-1}$,

- Link $\eta(\mu)=\log \mu$

$$
\frac{\partial \mu}{\partial X_{j}}=\beta_{j} \exp \{\mathbf{X} \boldsymbol{\beta}\}
$$

since the model under consideration has the form $\mu_{i}=\exp \{\mathbf{X} \boldsymbol{\beta}\}$,

- Link $\eta(\mu)=\mu$

$$
\frac{\partial \mu}{\partial X_{j}}=\beta_{j}
$$

since the model under consideration has the form $\mu_{i}=\mathbf{X} \boldsymbol{\beta}$.

### 4.2.7 Likelihood Ratio Test

Let us suppose that there are two models considered, $\omega_{1}$ and $\omega_{2}$, such that $\omega_{1} \subset \omega_{2}$. The main idea is to compare the maximized likelihoods of the two models in order to determine whether $\omega_{1}$ is an acceptable model compared to $\omega_{2}$ for the data available.

The maximized likelihood under the model $\omega_{1}$ is

$$
\max _{\theta \in \omega_{1}} L(\theta ; y)=L\left(\widehat{\theta}_{\omega_{1}} ; y\right)
$$

where $\widehat{\theta}_{\omega_{1}}$ denotes the maximum likelihood estimate of $\theta$ under the model $\omega_{1}$.
The maximized likelihood under the model $\omega_{2}$ is

$$
\max _{\theta \in \omega_{2}} L(\theta ; y)=L\left(\widehat{\theta}_{\omega_{2}} ; y\right)
$$

where $\widehat{\theta}_{\omega_{2}}$ denotes the maximum likelihood estimate of $\theta$ under the model $\omega_{2}$.
The likelihood ratio

$$
\lambda=\frac{L\left(\widehat{\theta}_{\omega_{1}} ; y\right)}{L\left(\widehat{\theta}_{\omega_{2}} ; y\right)}
$$

is defined and, under suitable regularity conditions, minus twice the $\log$ of the likelihood ratio has asymptotically a chi-square distribution with degrees of freedom equal to the difference in the number of parameters between the two models. Thus,

$$
-2 \log \lambda=2 l\left(\widehat{\theta}_{\omega_{2}} ; y\right)-2 l\left(\widehat{\theta}_{\omega_{1}} ; y\right) \sim \chi_{\nu}^{2}
$$

where the degrees of freedom $\nu$ are

$$
\nu=\operatorname{dim}\left(\omega_{2}\right)-\operatorname{dim}\left(\omega_{1}\right) .
$$

### 4.2.8 Wald Test

Under certain regularity conditions, the maximum likelihood estimator $\widehat{\theta}$ has approximately in large samples a (multivariate) normal distribution with mean equal to the true parameter value and variance-covariance matrix given by the inverse of the information matrix, so that

$$
\widehat{\theta} \sim A N_{p}\left(\theta, I^{-1}(\theta)\right)
$$

The regularity conditions include the following:

- The true parameter value $\theta$ must be interior to the parameter space,
- The log-likelihood function must be thrice differentiable,
- The third derivatives must be bounded.

This result provides a basis for constructing tests of hypotheses and confidence regions. For example under the hypothesis

$$
H_{0}: \theta=\theta_{0}
$$

for a fixed value $\theta_{0}$, the quadratic form

$$
W=\left(\widehat{\theta}-\theta_{0}\right)^{\prime} \operatorname{Var}^{-1}(\widehat{\theta})\left(\widehat{\theta}-\theta_{0}\right)
$$

has approximately in large samples a chi-squared distribution with $p$ degrees of freedom. These results are extendable to arbitrary linear combinations of $\theta$, including sets of elements of $\theta$.

When the subset has only one element the square root of the Wald statistic is taken and the ratio

$$
z=\frac{\widehat{\theta}-\theta_{0}}{\sqrt{\operatorname{Var}(\widehat{\theta})}}
$$

is treated as a z-statistic.
Frequently, for estimating the variances the observed information matrix is used.

### 4.2.9 Bootstrap Confidence Intervals of Parameters Estimates in Linear Regression

Standard parametric confidence intervals (SPCI's) provide a measure of significance for the estimates of the regression parameters. However, a requirement for the construction of the SPCI's is the acceptance of Gaussian assumptions regarding the estimates. In cases when such an assumption is inappropriate, an alternative to the SPCI's could be CI's built on the basis of nonparametric methods using bootstrap estimates of the variability of the estimated coefficients.

For the purpose of the analysis the nonparametric bootstrap normal, percentile and bias-corrected (BC) confidence intervals have been considered.

The $100(1-2 \alpha) \%$ bootstrap normal confidence interval is defined as

$$
\left[\widehat{\theta}-t_{1-\alpha, R-1} \sqrt{\widehat{v}_{B}}, \quad \widehat{\theta}-t_{\alpha, R-1} \sqrt{\widehat{v}_{B}}\right]
$$

where $\alpha$ is the significance level, $t_{(\cdot), R-1}$ is the $(\cdot)$ quantile of the Student's $t$ distribution with $R-1$ degrees of freedom and

- $R$ is the number of replications,
- $\widehat{\theta}$ is the parameter estimate,
- $\widehat{v}_{B}$ is the bootstrap estimate of the variance of $\widehat{\theta}$, defined as $\operatorname{Var}(\widehat{\theta}) \approx \operatorname{Var}\left(\widehat{\theta}^{*}-\widehat{\theta} \mid X^{*}\right)$.

A crucial requirement for the usage of the bootstrap normal interval is that, in the worst case, the distribution of $\widehat{\theta}$ has to be symmetric. If this is not the case, the inference based on that CI could be misleading.

The two sided ( $1-2 \alpha$ ) bootstrap percentile interval is defined to be (Efron and Tibshirani, 1993)

$$
\left[\begin{array}{ll}
\widehat{\theta}_{(\alpha)}^{*}, & \widehat{\theta}_{(1-\alpha)}^{*}
\end{array}\right]
$$

where $\widehat{\theta}_{(.)}^{*}$ is the $(\cdot)$ percentile of the $\widehat{\theta}^{*}$ distribution.
Now, let

$$
z_{0}=\Phi^{-1}\left\{\frac{\#\left(\hat{\theta}_{i}^{*} \leq \widehat{\theta}\right)}{R}\right\}
$$

where $\#\left(\widehat{\theta}_{i}^{*} \leq \widehat{\theta}\right)$ is the number of elements of the bootstrap distribution that are less than or equal to the estimated statistic and $\Phi$ is the standard cumulative normal. Let

$$
\begin{aligned}
& p_{1}=\Phi\left(2 z_{0}-z_{1-\alpha}\right) \\
& p_{2}=\Phi\left(2 z_{0}+z_{1-\alpha}\right)
\end{aligned}
$$

where $z_{1-\alpha}$ is the $(1-\alpha)$ th quantile of the standard normal distribution. Then the BC interval is defined as (Efron and Tibshirani, 1993)

$$
\left[\begin{array}{ll}
\widehat{\theta}_{p_{1}}^{*}, & \widehat{\theta}_{p_{2}}^{*}
\end{array}\right]
$$

where $\widehat{\theta}_{p}^{*}$ is the $p$ th quantile of the bootstrap distribution.
The multiple linear regression was performed with 1000 bootstrap replications, by fixing the design matrix and resampling from $\left(y_{i}, \mathbf{x}_{i}\right)$, the possible responses conditional on the vector of predictors that referred to them. If the confidence interval for a parameter estimate failed to include 0 , then the p-value was deemed to be less than or equal to $\alpha=0.05$, and the coefficient was regarded as significant.

The sampling weights were not taken into account in the bootstrapping. In addition, no clustering adjustment was performed.

### 4.3 Assumptions in the Analysis

The assumptions on the basis of which the analysis has been conducted could be summarized as follows

- The observations on the different subjects are independent within one time period and between periods,
- Under the ordered probit model, the latent variable $\mathbf{Y}^{*} \sim N(\mathbf{X} \boldsymbol{\beta}, \mathbf{I})$ whereas under the ordered $\operatorname{logit} \mathbf{Y}^{*} \sim \operatorname{logistic}\left(\mathbf{X} \boldsymbol{\beta}, \frac{\pi^{2}}{3} \mathbf{I}\right)$, where $\mathbf{I}$ is the identity matrix,
- The coefficients in the ordered probit model are the same for all groups of the response,
- For the ordinal logit model, the odds across response categories are proportional,
- In the GLM situation the errors are assumed to have gamma distribution. The canonical link is taken to be log.


### 4.4 Research Hypotheses

The economic theory, the numerous relative research results from previous studies and the objective reasoning suggest the following research hypotheses:

1. Health and income are positively related which implies that when income increases health improves, as well as when health improves income increases,
2. Change in health and change in income are positively related in a way that a positive increment in income produces a positive increment in health as well as vise versa,
3. There are additional predictors like level of education, age, labor force status, stress, sex that contribute to the explanation of the health-income relationships so that
(a) A higher educational level produces a higher income as well as better health,
(b) As age increases income increases, health improves and health and income are not closely related up to a certain age, when the relationship between health and income becomes stronger due to the age effect on health and the effect of income as a resource for health,
(c) When labor force status deteriorates, income decreases and health worsens,
(d) For higher stress levels health worsens and as stress declines health improves and income increases,
(e) A female gender suggests a lower income and better health as the male gender implies higher income and poorer health,
4. The additional predictors affect the change in health change in income relationship in same way as they affect the health-income relationship.

### 4.5 Problems

Like almost every applied research dealing with real data collected from real people, the present investigation also faces some problems:

1. Current health status is self-assessed and self-reported in the SLID and hence may not be consistent with an "objective" measure of health. Although it is assumed that the entries are made after comparing the person's own health to the health of the people his/her age, this is not always the case,
2. The Current stress level, used as a channel variable, is also self-reported,
3. In the SLID there is no wealth variable of any kind that could be used as a substitute for income in special cases like the analysis of the health-income relationship for older people, whose wealth is a determining factor of economic means,
4. The income used in the analysis is transitive and may reflect short term fluctuations. Thus any changes in income might not affect health,
5. There could be correlation between members of the same family, thus producing clusters not having iid errors.

## 5 Levels Models Results

### 5.1 Exploration of the Variable Relationships

The study of the relationship between health and income begins by graphical inspection of their association. Figures 1 and 2 present box plots of income 1998 and income 1999 by groups of health 1996.


Figure 1: Box Plot of Income 1998 by Groups of Health Status 1996


Figure 2: Box Plot of Income 1999 by Groups of Health Status 1996

One can see that income tends to be relatively higher for people reporting excellent (1) or very good (2) health, whereas it seems to decrease for persons who characterized their health as fair (4) or poor (5). The overall decreasing pattern of income by categories of health suggests that there exists a positive association between them. In addition, the variance in income appears to be higher for the subjects in the first two groups which might be a result of a third factor influence on the relationship of interest.

Along with the immediate association between income and health, their "partial" relationships with the other variables used in the analysis are also of interest. Figures 3 and 4 present box plots of income 1998 and income 1999 by groups of the variable highest level of education.


Figure 3: Box Plot of Income 1998 by Groups of Highest Level of Education 1996

The group of people who have less than high school education (1) and who are deemed to be the least qualified have the lowest income among all groups. On the contrary, the group of the university graduates (4) beats the income of all other groups rather respectably. The high school graduates (2) and the non-university postsecondary certificate holders (3) seem to be hardly distinguishable in terms of income. The overall increasing pattern of income by groups of education suggests


Figure 4: Box Plot of Income 1999 by Groups of Highest Level of Education 1996
that there exists a positive association between income and education.
In addition, Figure 5 presents box plots of income 1998 and income 1999 by the groups of the variables annual labor force status 1996, sex and current stress level 1996 respectively.

The income shows dependence on the categories of annual labor force status 1996. For the group of people employed for the whole year 1996 (1), both income 1998 and income 1999 are considerably higher than for the rest of the groups. The difference between the unemployed (2) and all other categories is greatest, which is a result of lower income that the unemployed have.

The variable sex, on the other hand, appears to possess an insignificant discriminating power. The family income for both male (1) and female (2) groups is almost identical, being, however, slightly higher for the males.

The box plots of income by categories of current stress level suggest that stress might have an influential power on income or vice versa. People who reported high (1) or moderate (2) stress levels seem to have higher income than the ones that characterized their current 1996 stress levels as somewhat stressful (3) or not at all stressful (4). A possible explanation might be that persons tend to associate stress


Figure 5: Box Plots of Income 1998 and Income 1999 by Groups of Various Variables
with their work responsibilities and environment rather than other factors. Thus higher remuneration corresponds to greater responsibility resulting in a higher level of stress.

It is important to note that the box plots patterns by variable do not change with the lag. This means that the there are steady correspondences between the groups of every variable under consideration within a two and three year lag and no rapid changes in the status of the respondents as a whole have taken place. This could be evidence for a stable medium run association between the events under investigation.

The health "partial" relationships with annual labor force status, education, sex and stress are presented in Tables 1 to 8.

Table 1: Table of Current State of Health 1998 by Labor Force Status 1996

| CSH 1998 |  |  | Annual Labor Force Status 1996 |  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: |
|  | Employed | Unemployed | NLF | EPY-UPY | EPY-NLFPY | UPY-NLFPY | AOA |  |  |
| Excellent | 0.3027 | 0.2073 | 0.1866 | 0.2810 | 0.3671 | 0.3065 | 0.3407 |  |  |
| Very good | 0.4152 | 0.3316 | 0.3030 | 0.4106 | 0.3807 | 0.3417 | 0.3888 |  |  |
| Good | 0.2216 | 0.2694 | 0.2650 | 0.2292 | 0.1767 | 0.2312 | 0.2004 |  |  |
| Fair | 0.0493 | 0.1295 | 0.1558 | 0.0628 | 0.0589 | 0.0804 | 0.0521 |  |  |
| Por | 0.0112 | 0.0570 | 0.0895 | 0.0164 | 0.0151 | 0.0417 | 0.0190 |  |  |
| Fraction | 0.5636 | 0.0193 | 0.2079 | 0.0733 | 0.0662 | 0.0199 | 0.0499 |  |  |
| \# obs. | 41097 |  |  |  |  |  |  |  |  |

Health 1998 and health 1999, in Tables 1 and 2, appear to be related to labor force status 1996. Employed people in general report better health than the one reported

Table 2: Table of Current State of Health 1999 by Labor Force Status 1996

| CSH 1999 |  | Annual Labor Force Status 1996 |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Employed | Unemployed | NLF | EPY-UPY | EPY-NLFPY | UPY-NLFPY | AOA |
| Excellent | 0.2857 | 0.1935 | 0.1882 | 0.2661 | 0.3449 | 0.2766 | 0.3206 |
| Very good | 0.4168 | 0.3441 | 0.2920 | 0.3901 | 0.3540 | 0.3457 | 0.4081 |
| Good | 0.2281 | 0.2581 | 0.2670 | 0.2571 | 0.2103 | 0.2447 | 0.1839 |
| Fair | 0.0557 | 0.1452 | 0.1613 | 0.0658 | 0.0711 | 0.1011 | 0.0717 |
| Poor | 0.0138 | 0.0538 | 0.0920 | 0.0209 | 0.0197 | 0.0266 | 0.0157 |
| Fraction | 0.5730 | 0.0186 | 0.2120 | 0.0669 | 0.0661 | 0.0188 | 0.0446 |
| \# obs. | 19871 |  |  |  |  |  |  |

by all other groups: $71.71 \%$ and $70.25 \%$ of the employed in 1998 and 1999 respectively report excellent or very good health. An exception of the rule turns out to be the group of persons who have a non-constant labor status (AOA) who report even better health. Poorer health, on the other hand, is reported mainly by individuals who are unemployed or not in the labor force in 1996. About $13 \%$ of the unemployed and $16 \%$ of the NLF in 1996 reported fair health in 1998, compared to only around $5 \%$ of the employed. The differences in 1999 health are close to that margin for the same groups. Overall, the poorest health in 1998 and 1999 seems to be reported by the individuals who are not in the labor force in the year 1996.

An additional feature of the health labor force association is the deterioration of health within the labor force groups as the lag changes from two to three years. Most likely, this is a result of the age effect.

The cross tabulations, Tables 3 and 4, of health 1998 and health 1999 by level of education in 1996 suggest the existence of an association.

Table 3: Table of Current State of Health 1998 by Highest Level of Education 1996

| CSH 1998 | Education 1996 |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  | LHS | GHS | NUPC | UD |
| Excellent | 0.1919 | 0.2927 | 0.2668 | 0.3429 |
| Very good | 0.3138 | 0.3925 | 0.4016 | 0.4076 |
| Good | 0.2794 | 0.2260 | 0.2340 | 0.1910 |
| Fair | 0.1499 | 0.0645 | 0.0725 | 0.0444 |
| Poor | 0.0653 | 0.0242 | 0.0255 | 0.0140 |
| Fraction | 0.3168 | 0.2805 | 0.2744 | 0.1283 |
| \# obs. | 45979 |  |  |  |

The individuals having a university degree (UD) in the two years report better health than the individuals in the other three groups: $34.29 \%$ and $40.76 \%$ of the UD report excellent and very good health for $1998,34.36 \%$ and $40.90 \%$ are the respective numbers for health 1999. Furthermore, the percentage of the persons in the UD group

Table 4: Table of Current State of Health 1999 by Highest Level of Education 1996

| CSH 1999 | Education 1996 |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  | LHS | GHS | NUPC | UD |
| Excellent | 0.1792 | 0.2759 | 0.2578 | 0.3436 |
| Very good | 0.3102 | 0.3951 | 0.3964 | 0.4090 |
| Good | 0.2871 | 0.2311 | 0.2411 | 0.1865 |
| Fair | 0.1529 | 0.0726 | 0.0793 | 0.0496 |
| Poor | 0.0710 | 0.0253 | 0.0251 | 0.0120 |
| Fraction | 0.3114 | 0.2921 | 0.2634 | 0.1330 |
| \# obs. | 22031 |  |  |  |

reporting good, fair and poor health is considerably smaller than the one in the rest of the groups. The same situation is repeated for health 1999. Poorest health 1998 and 1999 is reported by the individuals having less than a high school education.

Tables 5 and 6 present health 1998 and health 1999 by sex.

Table 5: Table of Current State of Health 1998 by Sex

| CSH 1998 | Sex |  |
| ---: | ---: | ---: |
|  | Male | Female |
| Excellent | 0.282430 | 0.239779 |
| Very good | 0.369964 | 0.373835 |
| Good | 0.228644 | 0.251379 |
| Fair | 0.083527 | 0.098118 |
| Poor | 0.035436 | 0.036889 |
| Fraction | 0.4741 | 0.5259 |
| \# obs. | 45979 |  |

Table 6: Table of Current State of Health 1998 by Sex

| CSH 1999 | Sex |  |
| ---: | ---: | ---: |
|  | Male | Female |
| Excellent | 0.266204 | 0.235137 |
| Very good | 0.367424 | 0.373857 |
| Good | 0.237795 | 0.251905 |
| Fair | 0.090067 | 0.101753 |
| Poor | 0.038300 | 0.037157 |
| Fraction | 0.4752 | 0.5248 |
| \# obs. | 22031 |  |

The data suggest that male respondents are more inclined to report excellent health compared to the female ones. Along with that, female individuals report more
frequently very good, good and fair health. Overall however, it is difficult to determine whether sex might have a predictive power on health.

The data presented in Tables 7 and 8 suggest that stress in 1996 is probably of importance to the state of health in 1998 and 1999.

Table 7: Table of Current State of Health 1998 by Current Stress Level 1996

| CSH 1998 |  | Current Stress Level 1996 |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  | Very Stressful | Somewhat stressful | Not very stressful | Not at all stressful |
| Excellent | 0.2062 | 0.2630 | 0.2761 | 0.2844 |
| Very good | 0.3368 | 0.3945 | 0.3720 | 0.3435 |
| Good | 0.2596 | 0.2318 | 0.2388 | 0.2489 |
| Fair | 0.1202 | 0.0809 | 0.0867 | 0.0962 |
| Poor | 0.0779 | 0.0297 | 0.0264 | 0.0278 |
| Fraction | 0.1348 | 0.4547 | 0.2919 | 0.1185 |
| \# obs. | 44936 |  |  |  |

Table 8: Table of Current State of Health 1999 by Current Stress Level 1996

| CSH 1999 | Current Stress Level 1996 |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  | Very Stressful | Somewhat stressful | Not very stressful | Not at all stressful |
| Excellent | 0.1940 | 0.2502 | 0.2634 | 0.2832 |
| Very good | 0.3205 | 0.3867 | 0.3814 | 0.3491 |
| Good | 0.2696 | 0.2466 | 0.2362 | 0.2299 |
| Fair | 0.1337 | 0.0865 | 0.0901 | 0.1014 |
| Poor | 0.0821 | 0.0300 | 0.0289 | 0.0363 |
| Fraction | 0.1376 | 0.4533 | 0.2908 | 0.1183 |
| \# obs. | 21875 |  |  |  |

An evidence of that is the considerably poorer health 1998 and 1999 of the individuals who described their stress level as very high in comparison to the rest. In addition, the individuals who characterized their 1996 stress level as somewhat stressful, not very stressful and not at all stressful are relatively similar healthwise suggesting a threshold effect of high stress.

The reported health 1998 and health 1999 seem also associated to age as could be seen on Figure 6.


Figure 6: Box Plot of Age 1996 by the Groups of Current Health Status 1998 and 1999

It appears that younger persons report better health which is a natural predisposition. The overall increasing pattern of age by categories of health suggests that a negative relationship may exist between them.

### 5.2 The Health Models

Based on economic theory and supported by the graphical analysis evidence of existing natural associations in the data, models (1) and (2) have been constructed and estimated as ordered probit models.

The estimates of the regression parameters of model (1) are presented in the following table.

Table 9: Ordered Probit Estimates of Model (1) for Health, 1998

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Health(Excellent)* | -1.105665 | 0.015954 | 4802.49 | 0.000 | 10010.79 | 0.0000 |
| Health(Very good)* | -0.518119 | 0.014431 | 1288.81 | 0.000 |  |  |
| Health(Fair)* | 0.672299 | 0.023102 | 846.81 | 0.000 |  |  |
| Health(Poor)* | 1.538780 | 0.036755 | 1753.10 | 0.000 |  |  |
| Income | -1.95E-06 | $2.05 \mathrm{E}-07$ | 90.06 | 0.000 | 90.38 | 0.0000 |
| LFS (Unemployed)* | 0.211324 | 0.039254 | 28.94 | 0.000 | 112.93 | 0.0000 |
| LFS (NLF)* | 0.146239 | 0.015627 | 87.61 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.060866 | 0.021342 | 8.12 | 0.004 |  |  |
| LFS (EPY-NLFPY)* | -0.010670 | 0.022922 | 0.22 | 0.642 |  |  |
| LFS (UPY-NLFPY)* | 0.098924 | 0.039548 | 6.25 | 0.012 |  |  |
| LFS (AOA)* | 0.041587 | 0.025907 | 2.59 | 0.108 |  |  |
| Sex(Female)* | 0.035976 | 0.011022 | 10.63 | 0.001 | 10.63 | 0.0010 |
| HLE(GHS)* | -0.076183 | 0.014776 | 26.63 | 0.000 | 78.00 | 0.0000 |
| HLE(NUPC)* | -0.068385 | 0.014914 | 21.07 | 0.000 |  |  |
| HLE(UD)* | -0.165838 | 0.019058 | 75.69 | 0.000 |  |  |
| Age | 0.011467 | 0.000428 | 717.70 | 0.000 | 719.57 | 0.0000 |
| cut1 | -0.806429 | 0.026980 |  |  |  |  |
| cut2 | 0.409545 | 0.026793 |  |  |  |  |
| cut3 | 1.481797 | 0.027716 |  |  |  |  |
| cut4 | 2.363663 | 0.030737 |  |  |  |  |
| N | 41097 |  |  |  |  |  |
| Log likelihood | -48376.61 |  |  |  |  |  |
| LR chi2(16) | 15147.57 |  |  |  |  |  |

All predictors included in the model prove to have a significant explanatory power based on the likelihood ratio test. Not all estimates, however, seem to be significant. The coefficient of labor force category ${ }^{1}$ EPY-NLFPY seems to be insignificant based on the Wald test.

[^0]For exploring the direct relationships of the predictors and their categories on health 1998, the marginal effects presented in Table 10 were estimated.

Table 10: Ordered Probit Estimates of Marginal Effects in Model (1)

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.371444 | -0.034250 | -0.235250 | -0.084650 | -0.017300 |
| Health(Very good)* | 0.164853 | 0.011196 | -0.119850 | -0.046750 | -0.009460 |
| Health(Fair)* | -0.159770 | -0.098570 | 0.137249 | 0.092798 | 0.028292 |
| Health(Poor)* | -0.232710 | -0.312460 | 0.124413 | 0.254504 | 0.166258 |
| Income | $5.96 \mathrm{E}-07$ | $9.47 \mathrm{E}-08$ | -4.6E-07 | -1.91E-07 | -3.95E-08 |
| LFS (Unemployed)* | -0.059670 | -0.018490 | 0.048945 | 0.023700 | 0.005510 |
| LFS (NLF)* | -0.043290 | -0.009530 | 0.034324 | 0.015199 | 0.003304 |
| LFS (EPY-UPY)* | -0.018260 | -0.003580 | 0.014334 | 0.006184 | 0.001316 |
| LFS (EPY-NLFPY)* | 0.003274 | 0.000499 | -0.002520 | -0.001040 | -0.000210 |
| LFS (UPY-NLFPY)* | -0.029180 | -0.006640 | 0.023225 | 0.010344 | 0.002255 |
| LFS (AOA)* | -0.012540 | -0.002330 | 0.009802 | 0.004184 | 0.000883 |
| Sex(Female)* | -0.011010 | -0.001740 | 0.008491 | 0.003525 | 0.000729 |
| HLE(GHS)* | 0.023558 | 0.003227 | -0.017980 | -0.007310 | -0.001490 |
| HLE(NUPC)* | 0.021129 | 0.002930 | -0.016140 | -0.006570 | -0.001340 |
| HLE(UD)* | 0.052863 | 0.004057 | -0.039050 | -0.014950 | -0.002920 |
| Age | -0.003510 | -0.000560 | 0.002707 | 0.001124 | 0.000233 |
| $\left.{ }^{*}\right) \mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

The results obtained are quite expected and not at all surprising. The probability of characterizing personal health as excellent in 1998, conditioned on excellent health in 1996, increases by over $37 \%$, whereas conditioned on poor health, decreases with above $23 \%$. The overall direction in the health 1996 health 1999 relationship turns out to be positive as a higher health status in 1996 is associated with an increase in the probability for a higher health status in 1999. This could also be seen from the regression coefficients in Table 9 where the change from good to very good or excellent health in 1996 is related to an increase in health 1998, whereas the change from good to fair or poor health in 1996 produces an opposite change in health 1998.

Income 1996, although having a quite small effect, contributes to improvement of health status 1998. A unit increase in income produces an increase in the probability of reporting excellent health by $5.96 \mathrm{E}-07$ and reporting very good health by $9.47 \mathrm{E}-08$.

At the same time the probabilities for characterizing the health as good, fair and poor decrease.

Labor force status seems to have some predictive power on health 1998. For the unemployed persons in 1996 the probability for reporting poorer health increases. This is common to all categories of the predictor with the exception of the EPYNLFPY. However, as the regression coefficient of that dummy was not significant, any inference could be misleading. Thus, in general, health 1998 is affected by labor status 1996 and this relationship is positive. This could also be seen in Table 9 where the change in the status from employed to any other category produces a decrease in health 1998.

The regression coefficient of the predictor sex from Table 9 suggests that males report better health than females. In addition, the marginal effect of sex shows that being a female increases the probability of reporting good, fair or poor health by $0.85 \%, 0.35 \%$ and $0.07 \%$ respectively.

The highest level of education 1996 is a significant predictor of health 1998. Table 9 shows that the change from the lowest level of education, less than high school, to any other level produces an increase in health 1998. The marginal effects also suggest that higher education is associated with an increase in the probability of reporting better health.

Age 1996, as already discovered graphically, is negatively associated with health 1998. As seen in both Table 9 and 10, a unit increase in age 1996 produces a decrease in health 1998 as well as a decrease in the probability of reporting excellent and very good health.

The estimates of the regression parameters of model (2) are presented in Table
11. They generally follow the pattern from model (1), with income 1996 having a predictive power on health 1999. Along with that, health 1996 is a significant predictor of health 1999. In this setting, sex seems to be insignificant at the $10 \%$ level. In addition, some of the labor force dummies turn out to have no significant predictive power on health 1999, although the predictor labor force 1996 is significant.

Table 11: Ordered Probit Estimates of Model (2) for Health, 1999

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Health(Excellent) | -1.024221 | 0.022671 | 2041.232 | 0.000 | 4339.68 | 0.0000 |
| Health(Very good)* | -0.466134 | 0.020647 | 509.856 | 0.000 |  |  |
| Health(Fair)* | 0.673368 | 0.032617 | 426.010 | 0.000 |  |  |
| Health(Poor)* | 1.396575 | 0.052572 | 705.434 | 0.000 |  |  |
| Income | $-2.17 \mathrm{E}-06$ | $2.86 \mathrm{E}-07$ | 57.305 | 0.000 | 57.58 | 0.0000 |
| Sex(Female)* | 0.024652 | 0.015737 | 2.465 | 0.117 | 2.45 | 0.1172 |
| LFS (Unemployed)* | 0.191487 | 0.057234 | 11.223 | 0.001 | 41.44 | 0.0000 |
| LFS (NLF)* | 0.122294 | 0.022052 | 30.803 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.077847 | 0.031666 | 6.052 | 0.014 |  |  |
| LFS (EPY-NLFPY)* | 0.012623 | 0.032579 | 0.152 | 0.698 |  |  |
| LFS (UPY-NLFPY)* | 0.049609 | 0.057947 | 0.740 | 0.392 |  |  |
| LFS (AOA)* | -0.003407 | 0.038880 | 0.008 | 0.930 |  |  |
| HLE(GHS)* | -0.087197 | 0.021004 | 17.223 | 0.000 | 76.38 | 0.0000 |
| HLE(NUPC)* | -0.108645 | 0.021585 | 25.301 | 0.000 |  |  |
| HLE(UD)* | -0.233892 | 0.027077 | 74.650 | 0.000 |  | 334.54 |
| Age | 0.011144 | 0.000610 | 333.793 | 0.000 |  |  |
| cut1 | -0.854501 | 0.038400 |  |  |  |  |
| cut2 | 0.334427 | 0.038062 |  |  |  |  |
| cut3 | 1.371913 | 0.039215 |  |  |  |  |
| cut4 | 2.242905 | 0.043093 |  |  |  |  |
| N | 19871 |  |  |  |  |  |
| Log likelihood | -24025.114 |  |  |  |  |  |
| LR chi2(16) | 6727.43 |  |  |  |  |  |

Education and age predictors retain their significance as higher education suggests better health, whereas the increase of age leads to a decrease in health.

The marginal effects of model (2), presented in Table 12, support the hypothesis that health 1999 is associated with income 1996. The probabilities of reporting excellent or very good health increase by $6.46 \mathrm{E}-07$ and $1.44 \mathrm{E}-07$ respectively for a dollar increase in income 1996. At the same time, the probabilities for reporting worse health decrease.

Table 12: Ordered Probit Estimates of Marginal Effects in Model (2)

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.337657 | -0.009570 | -0.217225 | -0.090204 | -0.020659 |
| Health(Very good)* | 0.144570 | 0.019554 | -0.105153 | -0.047936 | -0.011037 |
| Health(Fair)* | -0.154909 | -0.106331 | 0.125693 | 0.100476 | 0.035071 |
| Health(Poor)* | -0.216764 | -0.286186 | 0.110792 | 0.235976 | 0.156183 |
| Income | $6.46 \mathrm{E}-07$ | $1.44 \mathrm{E}-07$ | -4.93E-07 | -2.40E-07 | -5.67E-08 |
| Sex(Female)* | -0.007351 | -0.001631 | 0.005611 | 0.002726 | 0.000645 |
| LFS (Unemployed)* | -0.052956 | -0.019209 | 0.042273 | 0.023665 | 0.006228 |
| LFS (NLF)* | -0.035459 | -0.009727 | 0.027570 | 0.014133 | 0.003483 |
| LFS (EPY-UPY)* | -0.022602 | -0.006154 | 0.017570 | 0.008979 | 0.002206 |
| LFS (EPY-NLFPY)* | -0.003747 | -0.000863 | 0.002870 | 0.001406 | 0.000335 |
| LFS (UPY-NLFPY)* | -0.014517 | -0.003739 | 0.011229 | 0.005654 | 0.001373 |
| LFS (AOA)* | 0.001017 | 0.000224 | -0.000776 | -0.000376 | -0.000089 |
| HLE(GHS)* | 0.026323 | 0.005213 | -0.019894 | -0.009444 | -0.002198 |
| HLE(NUPC)* | 0.032977 | 0.006180 | -0.024808 | -0.011656 | -0.002693 |
| HLE(UD)* | 0.073991 | 0.007819 | -0.053524 | -0.023216 | -0.005070 |
| Age | -0.003322 | -0.000739 | 0.002536 | 0.001233 | 0.000292 |
| (*) $\mathrm{dy} / \mathrm{dx}$ i | or discrete | ange of du | my variab | from 0 to |  |

The health state in 1996 is a significant predictor of the future health. The probability of reporting excellent health in 1999, conditioned on having reported excellent health in 1996 increases by $33.77 \%$ whereas the probability drops conditioned on having reported worse health in 1996. The pattern in the probability for reporting better health 1999 than the one reported in 1996 suggests an overall improvement of health.

Being unemployed or out of the labor force in 1996 also seems to decrease the probability of reporting better health in 1999. Having a higher educational level, on the other hand, increases the probability of characterizing health 1999 as very good or excellent. The healthiest people, keeping all other predictors and conditions fixed, seem to be those with a university degree.

The predictor age appears to have the same negative effect on health 1999 as the one on health 1998. With the increase of age, the probability of reporting good, fair or poor health increases.

These results suggest that income has a significant predictive power on health 1998 and health 1999. In addition to income 1996, lagged health, labor force status, education, sex and age also contribute significantly to the model and could be used as predictors of future health.

## Age Groups and Stress in the Health Model

The age groups analysis, the results from which are presented in Tables A. 1 to A.16, was undertaken in order to investigate the health income relationship within the different age ranges. The justification for a treatment of that kind lies in the different economic and health behaviors of the subjects of different ages. Controlling for these aspects makes the results of the investigation much more reliable and concrete.

The two effects that deserve special attention in model (1) age groups analysis are the significant (dropping with age) income effect on future health and the positive effect of current health on future health. Model (2) shares the same features, but the three year lag seems to affect this pattern for the group of the 60 and over, where the suggested tendency does not hold.

The effect of income on health seems to be highest for the people in the group range under 20 . For models (1) and (2) it is $2.57 \mathrm{E}-06$ and $5.45 \mathrm{E}-06$ respectively, which are the highest values of income for all age groups. Its effect tends to drop until it reaches $1.20 \mathrm{E}-06$ and $1.02 \mathrm{E}-06$ for model (1) and model (2) respectively in the 60 and over and the 50-59 groups correspondingly. In model (2), the people over 60 years of age seem to experience an income effect on health of $2.00 \mathrm{E}-06$, which is higher only than the one in the third group.

Current health appears to have an increasing influence on future health by age
groups. The probability of being in the upper health group from the currently reported one is also increasing, which reconfirms the positive, strengthening with age, association between the current and future health.

The overall pattern of current health, current income and future health association suggests, that as people grow older, their future health is increasingly affected by their current health, whereas the income effect is diminishing.

Another special investigation of the health-income relationship has been done by introducing the current level of stress variable into the models (1) and (2) as an alternative to the current health variable. Tables A. 17 to A. 36 show the regression parameters' estimates and marginal effects of the predictors on future health for the whole panel.

Stress proves to have a significant predictive power on health. Moreover, with the decrease in the stress level, future health improves. The probabilities of reporting excellent health 1998 increase by $9.34 \%, 15.97 \%$ and $19.96 \%$ if the stress level has been been reduced to somewhat stressful, not very stressful or not at all stressful in 1996 respectively. At the same time these probabilities are $9.41 \%, 17.07 \%$ and $25.63 \%$ for health 1999.

Current income continues to predict health significantly with the stress variable in the model. A unit increase in income produces a future health improvement of $3.54 \mathrm{E}-06$ in 1998 and $3.63 \mathrm{E}-06$ in 1999. In addition, the marginal effect of income suggests that the probabilities for reporting better health increase when income rises.

Within each age group, the association between stress effect on future health and age appears to be positive. Decreasing stress has a bigger effect on future health for older people than those in the below 20 group. Persons in the range 60 and over seem
to be an exception, but still a decrease in stress improves their future health. The current income effect in these models, on the other hand, has the same, decreasing with age, effect on future health.

### 5.3 The Income Model

So far, the research was directed toward the predictive power of income on future health. The results achieved suggested that current income 1996 has a significant effect on health 1998 and 1999 and that effect changes with age.

An equally important question is whether current health is a significant predictor of future income. For the investigation of that aspect of the relationship, models (3) and (4) have been constructed and estimated. The marginal effects of the predictors have been calculated. Tables 13 and 14 present the results obtained for model (3).

Table 13: Gamma GLM estimates of Model (3) for Income, 1998

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000012 | $1.63 \mathrm{E}-07$ | 5470.082 | 0.000 | 4257.50 | 0.0000 |
| Age | -0.001669 | 0.000419 | 15.840 | 0.000 | 17.68 | 0.0000 |
| Health(Excellent)* | 0.041932 | 0.015204 | 7.618 | 0.006 | 70.98 | 0.0000 |
| Health(Very good)* | 0.026893 | 0.014337 | 3.534 | 0.061 |  |  |
| Health(Fair)* | -0.088463 | 0.023048 | 14.746 | 0.000 |  |  |
| Health(Poor)* | -0.177645 | 0.034528 | 26.420 | 0.000 |  |  |
| Sex(Female)* | -0.029812 | 0.010747 | 7.673 | 0.006 | 8.59 | 0.0034 |
| LFS (Unemployed)* | -0.276161 | 0.038848 | 50.552 | 0.000 | 111.72 | 0.0000 |
| LFS (NLF)* | -0.100318 | 0.015420 | 42.380 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.092233 | 0.020786 | 19.714 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.061560 | 0.022226 | 7.673 | 0.006 |  |  |
| LFS (UPY-NLFPY)* | -0.126895 | 0.038734 | 10.758 | 0.001 |  |  |
| LFS (AOA)* | -0.098939 | 0.025078 | 15.603 | 0.000 |  |  |
| HLE(GHS)* | 0.071919 | 0.014357 | 25.100 | 0.000 | 123.47 | 0.0000 |
| HLE(NUPC)* | 0.113585 | 0.014544 | 60.996 | 0.000 |  |  |
| HLE(UD)* | 0.177987 | 0.018300 | 94.673 | 0.000 |  |  |
| Constant | 10.145930 | 0.025790 | 154763.560 | 0.000 |  |  |
| N | 40275 |  |  |  |  |  |
| Log likelihood | -468572.2 |  | AIC | 23.26948 |  |  |
| Deviance | 7488.29138 |  |  |  |  |  |

All predictors have a significant explanatory power on the variation in income 1998 at $10 \%$ level. In particular, the results for health 1996 suggest that when health improves, income increases. The actual amount of increase in income when health increases from good to very good or excellent is $\$ 1120.85$ and $\$ 1755.81$ respectively.

A deterioration of health to fair and poor is related to a drop of income by $\$ 3538.40$ and $\$ 6792.89$.

Table 14: Marginal Effects of Model (3)

|  | Income 1998 |
| ---: | ---: |
| Variables 1996 | Marginal Effect |
| Income | 0.49970 |
| Age | -69.32924 |
| Health(Excellent)* | 1755.81200 |
| Health(Very good)* | 1120.85200 |
| Health(Fair)* | -3538.39500 |
| Health(Poor)* | -6792.88700 |
| Sex(Female)* $^{*}$ | -1238.94200 |
| LFS (Unemployed)* | -10076.59000 |
| LFS (NLF)* | -4048.00200 |
| LFS (EPY-UPY)* | -3684.60100 |
| LFS (EPY-NLFPY)* | -2489.87400 |
| LFS (UPY-NLFPY)* | -4962.38700 |
| LFS (AOA)* | -3932.12700 |
| HLE(GHS)* | 3032.37300 |
| HLE(NUPC)* | 4833.91700 |
| HLE(UD)* | 7900.44200 |
| (*) dy/dx is for discrete change of DV from 0 to 1 |  |

Unemployment and level of education seem to have a considerable effect on income. The unemployed persons in 1996 receive $\$ 10076.59$ less than the employed for 1998 , whereas the individuals having any other than a university degree have less income in comparison to university graduates.

Sex and age have a negative effect on income 1998. An increase in age 1996 by one year is associated with a decrease in income by about $\$ 70$ and being a female is associated with a drop of $\$ 1238.94$ compared to being male.

Similar patterns can be seen in the results of model (4) estimation, presented in Tables 15 and 16. The predictors are still significant at the $10 \%$ level and their directions of influence are alike but there are some changes in the effects.

With good health as the reference category, income differentials associated with

Table 15: Gamma GLM estimates of Model (4) for Income, 1999

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | $9.94 \mathrm{E}-06$ | $1.84 \mathrm{E}-07$ | 2926.810 | 0.000 | 1666.42 | 0.0000 |
| Age | -0.001820 | 0.000495 | 13.542 | 0.000 | 10.76 | 0.0010 |
| Health(Excellent)* | 0.037771 | 0.017892 | 4.452 | 0.035 | 39.45 | 0.0000 |
| Health(Very good)* | 0.032219 | 0.016896 | 3.648 | 0.057 |  |  |
| Health(Fair)* | -0.084920 | 0.026875 | 9.986 | 0.002 |  |  |
| Health(Poor)* | -0.197840 | 0.040965 | 23.329 | 0.000 |  |  |
| Sex(Female)* | -0.034319 | 0.012661 | 7.344 | 0.007 | 5.84 | 0.0157 |
| LFS (Unemployed)* | -0.269671 | 0.046098 | 34.223 | 0.000 | 63.71 | 0.0000 |
| LFS (NLF)* | -0.105562 | 0.017948 | 34.574 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.107222 | 0.025354 | 17.893 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.063615 | 0.026067 | 5.954 | 0.015 |  |  |
| LFS (UPY-NLFPY)* | -0.113439 | 0.046580 | 5.954 | 0.015 |  |  |
| LFS (AOA)* | -0.131231 | 0.030741 | 18.233 | 0.000 |  |  |
| HLE(GHS)* | 0.088290 | 0.016813 | 27.563 | 0.000 | 95.47 | 0.0000 |
| HLE(NUPC)* | 0.137684 | 0.017359 | 62.885 | 0.000 |  |  |
| HLE(UD)* | 0.219431 | 0.021483 | 104.244 | 0.000 |  |  |
| Constant | 10.265120 | 0.030128 | 116090.118 | 0.000 |  |  |
| N | 20635 |  |  |  |  |  |
| Log likelihood | -240831.4826 |  | AIC | 23.34369 |  |  |
| Deviance | 4569.56371 |  |  |  |  |  |

Table 16: Marginal Effects of Model (4)

|  | Income 1999 |
| :---: | :---: |
| Variables 1996 | Marginal Effect |
| Income | 0.428317 |
| Age | -78.411810 |
| Health(Excellent)* | 1639.435000 |
| Health(Very good)* | 1394.120000 |
| Health(Fair)* | -3529.531000 |
| Health(Poor)* | -7775.131000 |
| Sex(Female)* | -1479.660000 |
| LFS (Unemployed)* | -10237.170000 |
| LFS (NLF)* | -4414.281000 |
| LFS (EPY-UPY)* | -4413.012000 |
| LFS (EPY-NLFPY)* | -2666.920000 |
| LFS (UPY-NLFPY)* | -4630.779000 |
| LFS (AOA)* | -5330.986000 |
| HLE(GHS)* | 3870.627000 |
| HLE(NUPC)* | 6122.916000 |
| HLE(UD)* | 10260.070000 |
| (*) dy/dx is for discrete change of DV from 0 to 1 |  |

health are $\$ 1639.44$ (excellent) and $\$ 1394.12$ (very good). Again, a low health is associated with reductions in income of $\$ 3529.53$ (fair) and $\$ 7775.13$ (poor).

Unemployment and level of education seem to have a bigger effect on income 1999. Unemployment in 1996 produces a drop of $\$ 10237.17$ compared to employment. On
the other hand, the university degree premium increases to $\$ 10260.07$.
Within the context of the results of the analysis, it is possible to conclude that health 1996 has a significant predictive power on income 1998 and income 1999. The association is positive which suggests that an increase in the current health produces an increase in future income.

The additional predictors included in the models also prove to significantly effect income 1998 and income 1999. The ones that have a considerable impact on future income are labor force status and level of education.

## Age Groups and Stress in the Income Model

The analysis of the association between future income and current health has been performed within age groups. The results presented in Tables A. 37 to A. 41 reveal that health 1996 affects income 1998. Better health is associated with a higher income. In addition, it appears that the improved health carries higher income for the people in the range 50-59 years than ones in the 20-49 range. For the persons 60 and over, a drop in health brings about a significant decrease in income.

The highest level of education and labor force status, whenever significant predictors, provide evidence that employed individuals and persons with higher education have higher future income.

Age has a significant effect on future income for the 20-49 group only, where a year increase in age produces about $\$ 123$ increase in income 1998. Sex, on the other hand, is significant only for the 50-59 age group. Being a female in that group is associated with a $\$ 1311$ decrease in income.

With an increase of the lag by one year, health becomes insignificant as a predictor
of income for the groups below 20, 50-59 and 60 and over ages. Nonetheless, some of the categories of the current 1996 health are significant and there is still the pattern, although a bit vague, that associates better health with higher future income.

The age group 20-49 is distinguishable from the rest by the significance of the health 1996. Reporting excellent or very good health suggests an increase in future income by around $\$ 1489$ and $\$ 1831$ respectively. Poor health, on the other hand, produces a negative effect of $\$ 2110$ (fair) and $\$ 6523$ (poor).

Employment and education are both significant for 20-49 and 50-59 age groups and produce the expected effect on health. However, the people in the range 50-59 reporting a UPY-NLFPY labor status experience a drop of $\$ 12033$ in future income, which, quite unusually, is greater than the drop in the income of the permanently unemployed in 1996. The level of education, being also significant for the persons in the group 60 and over, has a positive relationship with future income where having a university degree produces an increase of $\$ 4182$.

The age predictor again has a positive effect on income 1999, of $\$ 140$ dollars, for the group 20-49. Sex once more is significant for the 50-59 age range and shows that being a female produces a decrease in future income by $\$ 1447$.

Overall, the age groups analysis of the effect of current health on future income provides evidence for differences in the significance and the magnitude of the association within the studied age ranges. The general pattern of relationship: better health resulting in higher income, is present. However, in model (4) it seems to be affected by the increase of the time lag.

Stress, as another dimension of health, was introduced in models (3) and (4) as a substitute of current state of health and its effect on future health was examined.

The results for the whole panel presented in Tables A. 47 to A. 50 suggest that stress 1996 is a significant predictor of income 1998 and income 1999.

The observed pattern of association between the different categories of stress and income relates a highest stress level and a lack of stress to low income. For model (3), the somewhat stressful category is associated with $\$ 2066$ increase in income 1998 compared to the stressful. An even greater increase in income, $\$ 2272$, is predicted if stress decreases to not very stressful level. On the other hand, the not at all stressful category corresponds to a lower income, $\$ 1706$, than to either of the intermediate stress levels. The same pattern is observed for model (4) but the effects differ in magnitude. The highest income 1999 increase, $\$ 2628$, is predicted for the people in the not very stressful category.

In the age groups context, stress is not a significant predictor of income 1998 only for the group 60 and over. Everywhere else, the decrease of stress is generally associated with an increase in income. In contrast with the health predictor, stress seems to become more significant in some of the age ranges as the time lag increases.

The results from the introduction of the stress variable provide another perspective of the health income relationship and generally suggest that when stress decreases, income increases. In addition, the age groups analysis provides evidence that age should be controlled for in the stress income relationship and that there might be a time effect of stress on future income.

## 6 Increments Models Results

The purpose of the health and income increment analysis undertaken in this part of the project is to discover and quantitatively characterize any relationships that exist between change in health and change in income.

### 6.1 Exploration of the Variable Relationships

The investigation of the relationship between income and health increments begins by graphical exploration of associations between the variables included in the study.

Figure 7 presents a box plot of the increments of income 1998 by the increments of health 1997.


Figure 7: Box Plot of Change in Income 1998 by Groups of Change in Health 1997

There seems to be no considerable difference in change in income 1998 by groups of change in health 1997. The persons whose health has improved substantially (1) seem to have an equal median of the change in income with the individuals reported badly
worsened health (5). The groups improved a bit (2), unchanged (3) and worsened (4) appear to have similar medians.

Figure 8 presents box plots of change in income by groups of highest level of education 1998, labor force status 1998, sex and change in stress 1997.


Figure 8: Box Plots of Change in Income 1998 by Groups of Various Variables

Change in income seems to be different across the groups of some of the variables considered. Slight differences exist between the medians of the highest level of education categories. The individuals holding a university degree (4) have a higher change
in income 1997 than the rest of the groups. On the other hand, the persons having less than high school education (1) tend to have a lower median of change in income.

Small differences are also observed between the groups of labor force status 1998. It appears that the persons who are unemployed (2), not in the labor force all year (3) or unemployed part-year, not in the labor force part-year (6) have smaller changes in income than the rest of the people.

The investigation of change in health starts with a box plot of the age 1998 by the groups of change in health 1998, Figure 9.


Figure 9: Box Plot of Age 1998 by Groups of Change in Health 1998

Persons, who are experiencing a drastic change either positive (1) or negative (5) in their health seem to have a higher age. By comparison, younger people tend to have an unchanged (3) health status or undergo moderate positive (2) or negative (4) health transitions.

Almost half of the male and female respondents have reported unchanged health in 1998 in comparison with 1997, according to Table 17. Overall, there appears to be no difference in the change of health by sex.

Persons holding a university degree seem to be more likely to report unchanged

Table 17: Table of Change in Health 1998 by Sex

| Change in Health 1998 | Sex |  |
| ---: | ---: | ---: |
|  | Male | Female |
| Significantly improved | 0.0080 | 0.0080 |
| Improved a bit | 0.2575 | 0.2558 |
| Unchanged | 0.4925 | 0.4913 |
| Worsened | 0.2364 | 0.2398 |
| Worsened badly | 0.0057 | 0.0053 |
| Fraction | 0.4741 | 0.5259 |
| \# obs. | 45979 |  |

health than the rest of the groups, based on Table 18. They also do not report drastic changes in their health in comparison with the individuals in the other groups. Most rapid changes in health appears to be experienced by the respondents having less than high school education. These observations suggest that change in health might be influenced by the level of education.

Table 18: Table of Change in Health 1998 by Groups of Education 1998

| Change in Health 1998 | Education 1998 |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  | LHS | GHS | NUPC | UD |
| Significantly improved | 0.0107 | 0.0074 | 0.0071 | 0.0058 |
| Improved a bit | 0.2677 | 0.2513 | 0.2578 | 0.2417 |
| Unchanged | 0.4543 | 0.5069 | 0.4931 | 0.5346 |
| Worsened | 0.2591 | 0.2299 | 0.2373 | 0.2150 |
| Worsened badly | 0.0086 | 0.0042 | 0.0051 | 0.0024 |
| Fraction | 0.2802 | 0.2845 | 0.2967 | 0.1386 |
| \# obs. | 45979 |  |  |  |

Change in health appears to be different by categories of labor status. A bigger fraction of the unemployed persons experience considerable changes in health by comparison with the rest of the groups. In contrast, the employed respondents tend to have unchanged health or undergo moderate changes. In general, the largest fraction report unchanged health, but the other changes within some groups suggest that labor force status 1998 could affect change in health 1998.

Table 19: Table of Change in Health 1998 by Groups of Labor Status 1998

| Change in Health 1998 | Annual Labor Force Status 1998 |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  |  | Unemployed | NLF | EPY-UPY | EPY-NLFPY | UPY-NLFPY |
|  |  |  |  |  |  | $\&$ AOA |
| Significantly improved |  | 0.0149 | 0.0099 | 0.0066 | 0.0150 | 0.0099 |
| Improved a bit |  | 0.2762 | 0.2495 | 0.2608 | 0.2430 | 0.2448 |
| Unchanged |  | 0.4420 | 0.4850 | 0.4822 | 0.5077 | 0.5103 |
| Worsened | 0.2305 | 0.2597 | 0.2474 | 0.2430 | 0.2307 | 0.2273 |
| Worsened badly | 0.0033 | 0.0122 | 0.0083 | 0.0073 | 0.0039 | 0.0075 |
| Fraction | 0.5831 | 0.0181 | 0.1928 | 0.0786 | 0.0646 | 0.0629 |
| \# obs. | 40197 |  |  |  |  |  |

The same conclusion could be made for change in stress 1997 variable, based on Table 20.

Table 20: Table of Change in Health 1998 by Groups of Change in Stress 1997

| Change in Health 1998 | Change in Stress Level 1997 |  |  |
| ---: | ---: | ---: | ---: |
|  | Subst. reduced |  |  |
| \& Reduced | Unchanged |  |  |
| Significantly improved | 0.0079 | 0.0068 | 0.0109 |
| Improved a bit | 0.2460 | 0.2491 | 0.2833 |
| Unchanged | 0.4809 | 0.5046 | 0.4794 |
| Worsened | 0.2562 | 0.2358 | 0.2215 |
| Worsened badly | 0.0091 | 0.0037 | 0.0048 |
| Fraction | 0.2533 | 0.5182 | 0.2284 |
| \# obs. | 44936 |  |  |

The respondents in the different stress categories appear to experience different changes in health 1998. The persons having the largest positive change in health 1998 are the ones who have undergone an increase in stress 1997. On the other hand, the largest negative change in health is reported by the individuals whose stress has reduced. The people in the unchanged stress category appear to report more stable health.

The results of the change in income 1998 exploratory analysis suggest that even though there are some variable categories over which change in income varies, the differences seem quite small, which might affect the investigation of significant predictors and their effects. The same issue seems to exist with the change in health
variable though it appears that a higher variation of change in health is present in some cases. In addition, the exploratory analysis suggests that it is unlikely a strong relationship between change in income and change in health to be present.

### 6.2 The Health Model

The health model represented by equation (5) was estimated using order probit model and selection of the significant variable including change in income 1997 was performed. The results of the estimation and selection are presented in Table 21 and the marginal effects are shown in Table 22.

Table 21: Ordered Probit Estimates of Model (5) for Change in Health

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHealth 97 (Imp. a bit)* | -0.794476 | 0.065725 | 146.168 | 0.000 | 9230.74 | 0.0000 |
| CHealth 97 (Unchanged)* | -1.599494 | 0.065643 | 593.897 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -2.190215 | 0.066258 | 1092.964 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* | -3.010705 | 0.095519 | 993.510 | 0.000 |  |  |
| Change in Income 97 | $2.05 \mathrm{E}-08$ | $3.08 \mathrm{E}-07$ | 0.005 | 0.947 | 0.00 | 0.9468 |
| HLE 98(GHS)* | -0.028999 | 0.014791 | 3.842 | 0.050 | 7.79 | 0.0506 |
| HLE 98(NUPC)* | -0.024141 | 0.014302 | 2.856 | 0.091 |  |  |
| HLE 98(UD)* | -0.046558 | 0.017558 | 7.023 | 0.008 |  |  |
| Age 98 | -0.000668 | 0.000321 | 4.326 | 0.037 | 4.34 | 0.0372 |
| cut1 | -4.262409 | 0.071744 |  |  |  |  |
| cut2 | -2.294413 | 0.068752 |  |  |  |  |
| cut3 | -0.797390 | 0.068418 |  |  |  |  |
| cut4 | 1.319707 | 0.069534 |  |  |  |  |
| N | 45979 |  |  |  |  |  |
| Log likelihood | -46257.636 |  |  |  |  |  |
| LR chi2(9) | 9236.13 |  |  |  |  |  |

The predictor of primary interest, change in income 1997, appears to be insignificant. On the other hand, change in health from a previous period, age and level of education turn out to be statistically significant predictors of change in health 1998.

The pattern of the effect of change in health from a previous period over change in health 1998 suggests that in general irrespective to the direction of the health change in 1997 with respect to 1996 health, 1998 health has improved with respect to 1997 health. In addition, a worsened badly 1997 health change is associated with a $61.81 \%$ increase in the probability that in 1998 the health would be significantly improved with respect to 1997.

The effects of level of education provide enough evidence to conclude that a higher

Table 22: Marginal Effects of Model (5)

|  | Change in Health 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.016148 | 0.254893 | -0.073882 | -0.193857 | -0.003302 |
| CHealth 97 (Unchanged)* | 0.030357 | 0.441576 | -0.024982 | -0.429661 | -0.017291 |
| CHealth 97 (Worsened)* | 0.156147 | 0.566047 | -0.334560 | -0.379082 | -0.008552 |
| CHealth 97 (Wors. badly)* | 0.618145 | 0.138709 | -0.534502 | -0.220371 | -0.001981 |
| Change in Income 97 | $-2.24 \mathrm{E}-10$ | -6.12E-09 | $3.07 \mathrm{E}-10$ | 5.91E-09 | $1.23 \mathrm{E}-10$ |
| HLE 98(GHS)* | 0.000322 | 0.008676 | -0.000516 | -0.008311 | -0.000170 |
| HLE 98(NUPC)* | 0.000267 | 0.007216 | -0.000415 | -0.006926 | -0.000142 |
| HLE 98(UD)* | 0.000531 | 0.014021 | -0.001049 | -0.013238 | -0.000265 |
| Age 98 | 7.28E-06 | 0.000199 | -1.00E-05 | -0.000192 | $-3.99 \mathrm{E}-06$ |
| ${ }^{(*)} \mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

level of education is more probable to produce a higher positive change in health. In addition, a unit change in age increases the probability of health improvement.

As a possible reason for the insignificant prediction power of change in income 1997 could be that some of the people in certain age groups in the panel have more or less unchanged income. To deal with that problem an age analysis has been implemented, the results of which are presented in Tables A. 61 to A.70.

Change in income appears to have close to $10 \%$ level of significance only for the persons in the group 50-59 years of age. Along with that, the sign of the change in income predictor is negative which suggests a positive association with change in health 1998. Based on these results, it is possible to conclude that change in income is probably important for the health of the people in the age 50-59. A unit change in income 1997 is associated with a change in the probability of a substantial improvement, an improvement and unchanged health of $8.37 \mathrm{E}-09,2.13 \mathrm{E}-07$ and $9.62 \mathrm{E}-10$ respectively for the ordered probit model. For the ordered logit, these probabilities are 9.76E-09, $2.78 \mathrm{E}-07$ and $3.09 \mathrm{E}-09$ respectively.

In all age groups, change in health from a previous period is of significance for change in health 1998. The trend is consistent with the trend already discovered for
the whole panel. The additional predictors included in the age models like highest level of education, labor force status, age and sex have the expected effect on the change in health, whenever significant.

To investigate another dimension of the relationship between present and future change in health and the possible effects on the significance of change in income as a predictor of future change in health, the variable change in stress 1997 was introduced in the increments models as a substitute of change in health 1998. The estimates of the regression parameters for the significant predictors are presented in Table 23.

Table 23: Ordered Probit Estimates of Model (5) with Stress

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| CStress 97 (Reduced) | -0.266627 | 0.084970 | 9.860 | 0.002 | 89.14 | 0.0000 |
| CStress 97 (Unchanged) | -0.312336 | 0.084673 | 13.616 | 0.000 |  |  |
| CStress 97 (Increased)* | -0.393311 | 0.085046 | 21.344 | 0.000 |  |  |
| CStress 97 (Subst. increased) | -0.421267 | 0.122023 | 11.903 | 0.001 |  |  |
| Change in Income 97 | $-1.90 \mathrm{E}-07$ | $3.05 \mathrm{E}-07$ | 0.384 | 0.533 | 0.39 | 0.5328 |
| Age 98 | 0.005107 | 0.001584 | 10.368 | 0.001 | 10.4 | 0.0013 |
| Age 98sq | -0.000054 | 0.000016 | 11.628 | 0.001 | 11.62 | 0.0007 |
| HLE 98(GHS)* | -0.029458 | 0.014689 | 4.040 | 0.045 | 7.74 | 0.0518 |
| HLE 98(NUPC) | -0.030677 | 0.014384 | 4.537 | 0.033 |  |  |
| HLE 98(UD)* | -0.044051 | 0.017593 | 6.250 | 0.012 |  |  |
| cut1 | -2.648842 | 0.093803 |  |  |  |  |
| cut2 | -0.868928 | 0.092064 |  |  |  |  |
| cut3 | 0.460147 | 0.092027 |  |  |  |  |
| cut4 | 2.323233 | 0.094281 |  |  |  |  |
| N | 44936 |  |  |  |  |  |
| Log likelihood | -49624.963 |  |  |  |  |  |
| LR chi2(10) | 105.63 |  |  |  |  |  |

Though change in income 1997 is still an insignificant predictor of change in health 1998 for the whole panel, change in stress appears to be of importance to future change in health. In addition, the improvement in future health is bigger for the persons with high negative changes in stress 1997. The effects of these stress changes are presented in Table 24.

The increase in stress 1997 rises the probability for a health improvement in the future. Although insignificant, a unit change in the change in income is also associ-

Table 24: Marginal Effects of Model (5) with Stress

|  | Change in Health 1998 |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Variables | Improved <br> substantially | Improved a bit | Unchanged | Worsened | Worsened <br> badly |
| CStress 97 (Reduced)* | 0.006901 | 0.083611 | -0.011190 | -0.075892 | -0.003430 |
| CStress 97 (Unchanged)* | 0.006846 | 0.094625 | -0.003603 | -0.092997 | -0.004871 |
| CStress 97 (Increased)* | 0.011367 | 0.124827 | -0.023042 | -0.108520 | -0.004632 |
| CStress 97 (Subst. increased)* | 0.015297 | 0.137503 | -0.041547 | -0.107447 | -0.003806 |
| Change in Income 97 | $4.14 \mathrm{E}-09$ | $5.79 \mathrm{E}-08$ | $-2.67 \mathrm{E}-09$ | $-5.66 \mathrm{E}-08$ | $-2.86 \mathrm{E}-09$ |
| Age 98 | -0.000111 | -0.001558 | 0.000072 | 0.001520 | 0.000077 |
| Age 98sq | $1.17 \mathrm{E}-06$ | 0.000016 | $-7.58 \mathrm{E}-07$ | -0.000016 | $-8.10 \mathrm{E}-07$ |
| HLE 98(GHS)* | 0.000652 | 0.009015 | -0.000494 | -0.008737 | -0.000436 |
| HLE 98(NUPC)* | 0.000679 | 0.009386 | -0.000512 | -0.009099 | -0.000454 |
| HLE 98(UD)* | 0.000998 | 0.013542 | -0.000916 | -0.012988 | -0.000636 |
| (*) dy/dx is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |
|  |  |  |  |  |  |

ated with health improvement. As might be expected, a higher level of education is associated with a higher probability for health improvement, whereas an increase in age by one unit decreases that probability.

The age groups analysis in this setting shows that change in income 1997 is still an insignificant predictor of change in health 1998 at around $10 \%$ level for all groups with the exception of the $50-60$ years one. The change by a unit in the change of income increases the probabilities of a substantial increase or increase of future health by $2.16 \mathrm{E}-08$ and $2.83 \mathrm{E}-07$ respectively for the ordered probit model.

Change in stress becomes a significant predictor of future health for the people above 20 years of age. Its effect on change in health is coherent with the one for the whole panel. The additional predictors, whenever significant follow the already discovered pattern.

The results from the health model estimation under the increments analysis suggest that change in income 1997 is a significant predictor of future change in health at around $10 \%$ level only for the age group 50-59 years.

### 6.3 The Income Model

The income model (6) in the increments aspect, was estimated using ordinary least squares. The confidence intervals of the estimates were obtained by applying the bootstrap methods in linear regression. Table 25 briefly $^{2}$ presents the results obtained.

Table 25: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) for Change in Income

| Variables | Reps | Observed | Std. Err. | $95 \%$ Normal CI |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Change in Income 97 | 1000 | -0.1519452 | 0.0075865 | -0.1668325 | -0.1370579 |
| CHealth 97 (Subst. imp.) | 1000 | -1129.833 | 830.164 | -2758.898 | 499.232 |
| ${\text { CHealth } 97 \text { (Imp. a bit) }^{*}}^{2}$ | 1000 | -189.3181 | 137.7626 | -459.655 | 81.019 |
| CHealth 97 (Worsened) | 1000 | 2.848337 | 145.0803 | -281.849 | 287.545 |
| CHealth 97 (Wors. badly)* | 1000 | -836.9861 | 874.7398 | -2553.524 | 879.552 |
| LFS 98(Unemployed)* | 1000 | -3613.345 | 369.1722 | -4337.787 | -2888.904 |
| LFS 98 (NLF)* | 1000 | -1859.29 | 156.5067 | -2166.410 | -1552.170 |
| LFS 98(EPY-UPY)* | 1000 | -1171.935 | 226.8242 | -1617.041 | -726.828 |
| LFS 98 (EPY-NLFPY)* | 1000 | -1073.788 | 300.9145 | -1664.285 | -483.291 |
| LFS 98 (UPY-NLFPY)* | 1000 | -3360.684 | 416.608 | -4178.211 | -2543.157 |
| LFS 98 (AOA)* | 1000 | -1562.81 | 334.5782 | -2219.366 | -906.253 |
| HLE 98(GHS)* | 1000 | 258.7125 | 161.2039 | -57.625 | 575.050 |
| HLE 98(NUPC)* | 1000 | 202.953 | 154.0629 | -99.371 | 505.277 |
| HLE 98(UD)* | 1000 | 631.1646 | 215.1229 | 209.020 | 1053.309 |
| Age 98 | 1000 | -28.98312 | 5.107968 | -39.007 | -18.960 |
| Sex(Female)* | 1000 | 212.238 | 119.0517 | -21.382 | 445.858 |
| Constant | 1000 | 2698.509 | 275.2028 | 2158.467 | 3238.551 |

The regression coefficients of the change in health categories suggest that change in income 1998 is only positive if the health has worsened in 1997 in comparison with 1996. The expected change in income in 1998 is highly negative if health has

[^1]changed rapidly in 1997 in either positive or negative direction. The significance test for these coefficients however, does not provide enough evidence to conclude that they are significant. Thus change in health 1997 appears to be an insignificant predictor of change in income 1998.

Labor status and education turn out to be significant. Unemployed persons are expected to experience a negative change in income whereas the persons holding a university degree are more likely to have a higher income in comparison with the reference category. The increase in age is associated with a negative change in income. On the other hand, sex does not appear to be a significant predictor.

The age groups analysis provides evidence, presented in Tables A. 81 to A.84, that the effect of change in health is not identical in direction, though still insignificant for all age groups. An increase in health for the individuals below 20 years is associated with a considerable positive change in income. Badly worsened health is also related to a positive change in income, which is also true for for the group of the over 60 years of age. The change in health effects pattern for the persons within the groups 20-49 and 50-59 follows the one for the whole panel. The 60 and over years old individuals experience a positive change in income when their 1997 health improves a bit with respect to their 1996 health.

The effect of change in stress, as an alternative aspect of health, on change in income was studied. The brief ${ }^{3}$ results are presented in Table 26.

An increase in stress is generally associated with a negative change in future income. Though insignificant, a substantially reduced stress provides a positive change in income. Labor force status and the highest level of education affect change in

[^2]Table 26: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress

| Variables | Reps | Observed | Std. Err. | 95\% Normal CI |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1539399 | 0.0079097 | -0.169461 | -0.138419 |
| CStress 97 (Subst. reduced)* | 1000 | 860.317 | 1025.756 | -1152.566 | 2873.200 |
| CStress 97 (Reduced)* | 1000 | -52.20552 | 142.1739 | -331.199 | 226.788 |
| CStress 97 (Increased)* | 1000 | -357.7228 | 155.1973 | -662.273 | -53.173 |
| CStress 97 (Subst. increased)* | 1000 | -17.1635 | 999.3138 | -1978.158 | 1943.831 |
| LFS 98(Unemployed)* | 1000 | -3658.192 | 350.4684 | -4345.930 | -2970.453 |
| LFS 98 (NLF)* | 1000 | -1875.532 | 166.3463 | -2201.960 | -1549.104 |
| LFS 98(EPY-UPY)* | 1000 | -1154.012 | 234.7163 | -1614.605 | -693.419 |
| LFS 98 (EPY-NLFPY)* | 1000 | -1147.382 | 285.3843 | -1707.403 | -587.360 |
| LFS 98 (UPY-NLFPY)* | 1000 | -3359.781 | 430.0322 | -4203.651 | -2515.911 |
| LFS 98 (AOA)* | 1000 | -1692.293 | 320.7599 | -2321.733 | -1062.852 |
| HLE 98(GHS)* | 1000 | 267.0355 | 166.8652 | -60.411 | 594.482 |
| HLE 98(NUPC)* | 1000 | 191.8385 | 157.7822 | -117.784 | 501.461 |
| HLE 98(UD)* | 1000 | 613.8696 | 201.8337 | 217.803 | 1009.936 |
| Age 98 | 1000 | -29.90941 | 5.36597 | -40.439 | -19.380 |
| Sex(Female)* | 1000 | 204.0447 | 124.72 | -40.699 | 448.788 |
| Constant | 1000 | 2789.919 | 291.0311 | 2218.816 | 3361.021 |

income according to the already discovered pattern, where higher education and employment are associated with higher positive change in income. A unit change in age provides a negative change in income, whereas sex is again not a significant predictor.

Change in stress continues to be insignificant by age groups. Irrespectively, the marginal effects suggest that for the below 20 years group, there is a positive change in income only when substantial increase in stress 1997 takes place. In the range 2049 years only when stress changes most rapidly in a positive or a negative direction an increase in income 1998 with respect to income 1997 is predicted. While change in stress 1997 and change in income 1998 exhibit a negative association for the group
of the people between 50 and 59 years, for the individuals over 60 the association is positive.

Considering the results of the income model estimation in the increments context, there is no significant prediction power of either change in health or change in stress on future change in income. However, the marginal effects provide information that suggests a possible direction of association and magnitude.

## 7 Discussion

The research results provide evidence that the levels of health and income are related within the time framework implemented. The overall pattern of the relationship presumed that an increase in income is associated with an increase in health and vice versa. Along with that, a number of additional variables also turn out to have a significant predictive power on both health and income, for the whole panel as well as by age groups.

The increments analysis, on the other hand, suggested that the increments of income are a significant predictor of the increments of health for the persons in 50-59 age range only, whereas the increments of health appeared to have no predictive power on the increments of income. The additional variables exhibit unsteady association with both increments by age groups.

These results were obtained by the implementation of a number of statistical techniques for data analysis. In the health models, the analysis involved the latent variable technique as well as ordered probit and ordered logit statistical models, which are deemed to be the most suitable whenever the response has more than two categories. A possible OLS approach, for instance, which applied to the present case would

- Only provide relatively good results if the "cut points" for the categories are about the same distance apart, a condition which is not likely to be true,
- Yield heteroscedastic, non-normal errors, and pose a results reliability problem,
the techniques implemented provide much more reliable results, though there are well-known drawbacks of likelihood estimation.

In the income levels models, the study included a log link gamma GLM. This type of modelling was mainly determined by the shape of the trimmed income 1998 and income 1999 distributions, for the whole panel, presented on Figure 10. The modelling of the income increments, on the other hand, was implemented by means of OLS for the regression parameters estimation and bootstrap techniques for construction of confidence intervals of these estimates. The unconventional distribution of the change in income 1998, seen again on Figure 10, suggested that type of treatment.


Figure 10: Histograms of Income 1998, Income 1999 and Change in Income 1998

In both cases, the residuals analysis supported the choice of the modelling techniques
and the correctness of the results obtained.
Probably, a better estimation of the income increments models could be achieved, if the distribution ambiguity is overcome. However, this and some other issues are left for further research.

## 8 Future Directions

The future directions in the study could be defined in a common economic and statistical context.

A possible direction is a more comprehensive research of the increments relationships between income and health based on data for subsequent years. To enable that however, it would be necessary to investigate ways of achieving relative conformity between the subjects within different waves of the SLID.

Another research direction (Veall, 2002) is the investigation of possible positive and negative outcomes of an application of different underlying distributions of the latent variable.

Identifying the distribution, modelling and estimation of the change in income regression could be another possible future research, which would enable parametric treatment instead of the non-parametric one used in the project.

## 9 Conclusion

The objective of the research was to characterize the relationship between income and health, conditioned on income and health from previous periods, for Canada. The study included levels analysis and increments analysis of that relationship as well as investigation of supplementary predictors and channels of influence.

The set of instruments implemented included ordered probit and ordered logit models, gamma GLM models, OLS method and bootstrapping techniques in linear models.

The results of the study provided support to the hypotheses that

- The levels of income and health are positively related,
- A higher educational level produces a higher income as well as better health,
- Labor force status is positively related to both income and health in the levels as well as in the increments context,
- For higher stress levels health worsens and as stress declines health improves and income increases.

The study could not provide definite evidence that change in health and change in income are, in general, positively related, though this turned out to be true for people in the age range 50-59 years.

The hypothesis that females have a lower income and better health whereas males have higher income and poorer health was overruled by the results, which in the
general case suggested that sex is not an important predictor of either health or family income levels or their increments.

As part of a comprehensive study of the health income relationship, the value of the present research is that it provides additional quantitative details of their levels and increments relationship as well as possible directions for future investigation.

## 10 References

1. Draper, N., Smith, H. (1998). Applied Regression Analysis. John Wiley \& Sons, Inc., Toronto.
2. Efron, B., Tibshirani, R.J. (1993). An introduction to the bootstrap. Chapman \& Hall, New York, pp. 183
3. Ettner, S. (1996). New Evidence on the Relationship Between Income and Health. Journal of Health Economics, Vol. 15, pp. 67-85.
4. Greene, W. (1990). Econometric Analysis. Macmillan Publishing Company.
5. Grossman, M. (1972). On the Concept of Health Capital and the Demand for Health. Journal of Political Economy, Vol. 80, Issue 2, pp. 223-255.
6. Hurd, M., McFadden, D., Merrill A. (1999). Predictors of Mortality Among the Elderly. NBER Working Paper No. w7440
7. Maddala, G. (1983). Limited-dependent and Qualitative Variables in Econometrics. Cambridge University Press.
8. McCullagh, P., Nelder., J. (1989). Generalized Linear Models. London: Chapman and Hall.
9. McLeod, L., Veall, M. (2002). The Dynamics of Food Deprivation and Overall Health: Evidence from the Canadian National Population Health Survey. SEDAP: Research Papers Released in 2002, No. 86
10. Myers, R., Montgomery, D. (1997). A Tutorial on Generalized Linear Models. Journal of Quality Technology, Vol. 29, No. 3, pp. 274-291.
11. Powers, D., Xie, Y. (2000). Statistical Methods for Categorical Data Analysis. Academic Press.
12. Smith, J. (1999). Healthy Bodies and Thick Wallets: The Dual Relation Between Health and Economic Status. Journal of Economic Perspectives, Vol. 13, No. 2, pp. 145-166.
13. Stata Corporation (2001). Stata Reference Manual. Stata Press, Vol. 1-4.
14. Veall, M. (2002). Income and Health: Leads and Lags. McMaster University.

## 11 Appendix

Follows the complete set of reference tables containing the models estimation results.

Table 1: Ordered Probit Estimates of Model (1) for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Health(Excellent) | -0.906525 | 0.062779 | 208.51 | 0.000 | 379.06 | 0.0000 |
| Health(Very good)* | -0.365351 | 0.064494 | 32.04 | 0.000 |  |  |
| Health(Fair)* | 0.537415 | 0.142259 | 14.29 | 0.000 |  |  |
| Health(Poor)* $^{*}$ | 1.283492 | 0.278074 | 21.34 | 0.000 |  |  |
| Income | $-3.57 \mathrm{E}-06$ | $7.03 \mathrm{E}-07$ | 25.81 | 0.000 | 26.02 | 0.0000 |
| Sex(Female)* | 0.188397 | 0.039867 | 22.37 | 0.000 | 22.34 | 0.0000 |
| HLE(GHS)* | -0.164971 | 0.055051 | 9.00 | 0.003 | 11.98 | 0.0075 |
| HLE(NUPC)* | 0.090751 | 0.144687 | 0.40 | 0.531 |  |  |
| HLE(UD)* | -0.726283 | 0.771913 | 0.88 | 0.347 | 8.13 | 0.0044 |
| Age | 0.067499 | 0.023673 | 8.12 | 0.004 |  |  |
| cut1 | 0.387415 | 0.410384 |  |  |  |  |
| cut2 | 1.537852 | 0.410924 |  |  |  |  |
| cut3 | 2.430943 | 0.412569 |  |  |  |  |
| cut4 | 3.347068 | 0.422404 |  |  |  |  |
| N | 3340 |  |  |  |  |  |
| Log likelihood | -3452.3205 |  |  |  |  |  |
| LR chi2(10) | 497.58 |  |  |  |  |  |

Table 2: Ordered Probit Estimates of Model (1) for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Health(Excellent)* | -1.082787 | 0.019606 | 3050.35 | 0.000 | 5565.30 | 0.0000 |
| Health(Very good)* | -0.507955 | 0.018040 | 792.99 | 0.000 |  |  |
| Health(Fair)* | 0.639230 | 0.032791 | 379.86 | 0.000 |  |  |
| Health(Poor)* | 1.536151 | 0.056559 | 737.67 | 0.000 |  |  |
| Income | -1.92E-06 | $2.66 \mathrm{E}-07$ | 52.42 | 0.000 | 52.63 | 0.0000 |
| LFS (Unemployed)* | 0.165999 | 0.047167 | 12.39 | 0.000 | 135.68 | 0.0000 |
| LFS (NLF)* | 0.271558 | 0.024123 | 126.79 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.047104 | 0.023809 | 3.92 | 0.048 |  |  |
| LFS (EPY-NLFPY)* | 0.031960 | 0.030658 | 1.08 | 0.297 |  |  |
| LFS (UPY-NLFPY)* | 0.144517 | 0.053175 | 7.40 | 0.007 |  |  |
| LFS (AOA)* | 0.068872 | 0.030882 | 4.97 | 0.026 |  |  |
| HLE(GHS)* | -0.107198 | 0.020568 | 27.14 | 0.000 | 72.82 | 0.0000 |
| HLE(NUPC)* | -0.109007 | 0.020412 | 28.52 | 0.000 |  |  |
| HLE(UD)* | -0.209490 | 0.024588 | 72.59 | 0.000 |  |  |
| Age | 0.014527 | 0.000875 | 275.56 | 0.000 | 276.12 | 0.0000 |
| cut1 | -0.767355 | 0.041032 |  |  |  |  |
| cut2 | 0.472117 | 0.040868 |  |  |  |  |
| cut3 | 1.582963 | 0.042065 |  |  |  |  |
| cut4 | 2.397218 | 0.046048 |  |  |  |  |
| N | 26588 |  |  |  |  |  |
| Log likelihood | -30693.69 |  |  |  |  |  |
| LR chi2(15) | 7459.02 |  |  |  |  |  |

Table 3: Ordered Probit Estimates of Model (1) for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Health(Excellent)* | -1.190674 | 0.041967 | 804.86 | 0.000 | 2008.90 | 0.0000 |
| Health(Very good)* | -0.562095 | 0.035350 | 252.81 | 0.000 |  |  |
| Health(Fair)* | 0.662142 | 0.048487 | 186.60 | 0.000 |  |  |
| Health(Poor)* | 1.522299 | 0.071538 | 452.84 | 0.000 |  |  |
| Income | -1.48E-06 | $4.49 \mathrm{E}-07$ | 10.82 | 0.001 | 10.89 | 0.0010 |
| LFS (Unemployed)* | 0.396676 | 0.085428 | 21.53 | 0.000 | 82.85 | 0.0000 |
| LFS (NLF)* | 0.294712 | 0.036898 | 63.84 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.069590 | 0.060464 | 1.32 | 0.250 |  |  |
| LFS (EPY-NLFPY)* | 0.021420 | 0.069433 | 0.10 | 0.758 |  |  |
| LFS (UPY-NLFPY)* | 0.446102 | 0.126809 | 12.39 | 0.000 |  |  |
| LFS (AOA)* | 0.134229 | 0.091697 | 2.13 | 0.143 |  |  |
| Sex(Female)* | -0.082328 | 0.028354 | 8.41 | 0.004 | 8.43 | 0.0037 |
| HLE(GHS)* | -0.175277 | 0.037703 | 21.62 | 0.000 | 39.62 | 0.0000 |
| HLE(NUPC)* | -0.136799 | 0.034973 | 15.29 | 0.000 |  |  |
| HLE(UD)* | -0.251983 | 0.046052 | 29.92 | 0.000 |  |  |
| cut1 | -1.471056 | 0.045847 |  |  |  |  |
| cut2 | -0.286027 | 0.042923 |  |  |  |  |
| cut3 | 0.737639 | 0.043979 |  |  |  |  |
| cut4 | 1.671037 | 0.051070 |  |  |  |  |
| N | 6289 |  |  |  |  |  |
| Log likelihood | -7812.1216 |  |  |  |  |  |
| LR chi2(15) | 2808.1 |  |  |  |  |  |

Table 4: Ordered Probit Estimates of Model (1) for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Health(Excellent)* | -1.109348 | 0.037489 | 875.57 | 0.000 | 2823.45 | 0.0000 |
| Health(Very good)* | -0.500238 | 0.027552 | 329.79 | 0.000 |  |  |
| Health(Fair)* | 0.612961 | 0.031693 | 374.04 | 0.000 |  |  |
| Health(Poor)* $_{\text {Income }}$ | 1.281860 | 0.049026 | 683.82 | 0.000 | 4.60 | 0.0320 |
| HLE(GHS)* | $-1.20 \mathrm{E}-06$ | $5.58 \mathrm{E}-07$ | 4.58 | 0.032 | 0.0000 |  |
| HLE(NUPC)* | -0.123667 | 0.030150 | 6.30 | 0.012 | 38.56 | 0.000 |
| HLE(UD)* | -0.241780 | 0.030224 | 0.045613 | 16.73 | 0.000 |  |
| Age | 0.013718 | 0.001500 | 88.09 | 0.000 | 83.75 | 0.0000 |
| cut1 | -0.710210 | 0.112016 |  | 0.000 |  |  |
| cut2 | 0.459887 | 0.111650 |  |  |  |  |
| cut3 | 1.511110 | 0.112332 |  |  |  |  |
| cut4 | 2.501410 | 0.113879 |  |  |  |  |
| N | 9762 |  |  |  |  |  |
| Log likelihood | -12808.25 | 3314.38 |  |  |  |  |
| LR chi2(9) |  |  |  |  |  |  |

Table 5: Marginal Effects of Model (1) for the Persons of Age Under 20

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.348332 | -0.141590 | -0.148055 | -0.051302 | -0.007385 |
| Health(Very good)* | 0.144930 | -0.070933 | -0.056189 | -0.015989 | -0.001820 |
| Health(Fair)* | -0.202178 | 0.052191 | 0.101187 | 0.041938 | 0.006863 |
| Health(Poor)* | -0.389969 | -0.046264 | 0.221686 | 0.164867 | 0.049681 |
| Income | $1.42 \mathrm{E}-06$ | -6.45E-07 | -5.83E-07 | -1.75E-07 | -2.07E-08 |
| Sex (Female)* | -0.074938 | 0.033698 | 0.030797 | 0.009326 | 0.001117 |
| HLE(GHS)* | 0.065717 | -0.030727 | -0.026360 | -0.007732 | -0.000898 |
| HLE(NUPC)* | -0.036031 | 0.015295 | 0.015311 | 0.004826 | 0.000599 |
| HLE(UD)* | 0.270649 | -0.168700 | -0.082035 | -0.018234 | -0.001680 |
| Age | -0.026893 | 0.012178 | 0.011016 | 0.003308 | 0.000392 |
| (*) $\mathrm{dy} / \mathrm{dx}$ | for discre | change of | mmy variab | from 0 to |  |

Table 6: Marginal Effects of Model (1) for the People between 20 and 49 Years of Age

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.376198 | -0.066556 | -0.230348 | -0.065100 | -0.014194 |
| Health(Very good)* | 0.170498 | -0.008924 | -0.118669 | -0.035371 | -0.007534 |
| Health(Fair)* | -0.166996 | -0.071463 | 0.145550 | 0.071101 | 0.021808 |
| Health(Poor)* | -0.256799 | -0.296771 | 0.177229 | 0.228784 | 0.147557 |
| Income | $6.30 \mathrm{E}-07$ | 7.52E-09 | -4.63E-07 | -1.44E-07 | -3.09E-08 |
| LFS (Unemployed)* | -0.051482 | -0.005998 | 0.040231 | 0.013984 | 0.003266 |
| LFS (NLF)* | -0.082332 | -0.013033 | 0.065640 | 0.023921 | 0.005804 |
| LFS (EPY-UPY)* | -0.015227 | -0.000552 | 0.011363 | 0.003624 | 0.000793 |
| LFS (EPY-NLFPY)* | -0.010365 | -0.000310 | 0.007704 | 0.002440 | 0.000532 |
| LFS (UPY-NLFPY)* | -0.045131 | -0.004664 | 0.035020 | 0.012001 | 0.002773 |
| LFS (AOA)* | -0.022095 | -0.001132 | 0.016640 | 0.005393 | 0.001195 |
| HLE(GHS)* | 0.035474 | -0.000384 | -0.025659 | -0.007786 | -0.001645 |
| HLE(NUPC)* | 0.036046 | -0.000337 | -0.026099 | -0.007932 | -0.001678 |
| HLE(UD)* | 0.071409 | -0.005187 | -0.049365 | -0.014025 | -0.002832 |
| Age | -0.004755 | -0.000057 | 0.003493 | 0.001085 | 0.000233 |
| (*) $\mathrm{dy} / \mathrm{dx}$ | for discret | hange of d | my variabl | from 0 to 1 |  |

Table 7: Marginal Effects of Model(1) for the People between 50 and 59 Years of Age

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.359067 | 0.046131 | -0.240744 | -0.132749 | -0.031705 |
| Health(Very good)* | 0.141814 | 0.073932 | -0.112675 | -0.081492 | -0.021579 |
| Health(Fair)* | -0.115787 | -0.142155 | 0.084703 | 0.123310 | 0.049928 |
| Health(Poor)* | -0.162589 | -0.329709 | -0.009112 | 0.261982 | 0.239427 |
| Income | $3.44 \mathrm{E}-07$ | $2.40 \mathrm{E}-07$ | -2.86E-07 | -2.33E-07 | -6.46E-08 |
| LFS (Unemployed)* | -0.074840 | -0.082246 | 0.059424 | 0.071960 | 0.025703 |
| LFS (NLF)* | -0.063258 | -0.053600 | 0.052221 | 0.049426 | 0.015211 |
| LFS (EPY-UPY)* | -0.015671 | -0.011902 | 0.013062 | 0.011267 | 0.003243 |
| LFS (EPY-NLFPY)* | -0.004933 | -0.003533 | 0.004107 | 0.003404 | 0.000955 |
| LFS (UPY-NLFPY)* | -0.081220 | -0.094844 | 0.063076 | 0.082285 | 0.030703 |
| LFS (AOA)* | -0.029161 | -0.024168 | 0.024275 | 0.022345 | 0.006709 |
| Sex(Female)* | 0.019142 | 0.013345 | -0.015912 | -0.012974 | -0.003602 |
| HLE(GHS)* | 0.042899 | 0.025604 | -0.035269 | -0.026324 | -0.006911 |
| HLE(NUPC)* | 0.032805 | 0.020891 | -0.027113 | -0.020958 | -0.005625 |
| HLE(UD)* | 0.064125 | 0.033372 | -0.052019 | -0.036333 | -0.009146 |
| (*) $\mathrm{dy} / \mathrm{dx}$ | for discret | ange of | my variabl | from 0 to 1 |  |

Table 8: Marginal Effects of Model (1) for the People 60 and Over Years of Age

| Health status 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.249309 | 0.169938 | -0.186154 | -0.179596 | -0.053496 |
| Health(Very good)* | 0.077088 | 0.115957 | -0.052693 | -0.101735 | -0.038617 |
| Health(Fair)* | -0.061277 | -0.150955 | 0.003309 | 0.131653 | 0.077270 |
| Health(Poor)* | -0.076443 | -0.266110 | -0.129496 | 0.207661 | 0.264388 |
| Income | $1.58 \mathrm{E}-07$ | $2.96 \mathrm{E}-07$ | -8.97E-08 | -2.56E-07 | -1.08E-07 |
| HLE(GHS)* | 0.010368 | 0.018521 | -0.006309 | -0.016063 | -0.006518 |
| HLE(NUPC)* | 0.017366 | 0.030095 | -0.010968 | -0.026112 | -0.010381 |
| HLE(UD)* | 0.037183 | 0.056906 | -0.026297 | -0.049587 | -0.018205 |
| Age | -0.001814 | -0.003392 | 0.001029 | 0.002942 | 0.001235 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 9: Ordered Probit Estimates of Model (2) for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Health(Excellent) | -0.735941 | 0.090396 | 66.260 | 0.000 | 113.36 | 0.0000 |
| Health(Very good)* | -0.342843 | 0.094555 | 13.177 | 0.000 |  |  |
| Health(Fair)* | 0.332061 | 0.209605 | 2.496 | 0.113 |  |  |
| Health(Poor) $^{*}$ | 1.265176 | 0.438272 | 8.352 | 0.004 |  |  |
| Income | $-5.45 \mathrm{E}-06$ | $1.07 \mathrm{E}-06$ | 25.806 | 0.000 | 9.19 | 0.0000 |
| Sex(Female)* | 0.175557 | 0.058151 | 9.120 | 0.003 |  |  |
| cut1 | -0.831011 | 0.103128 |  |  |  |  |
| cut2 | 0.299273 | 0.101775 |  |  |  |  |
| cut3 | 1.260868 | 0.111031 |  |  |  |  |
| cut4 | 1.945212 | 0.145822 |  |  |  |  |
| N | 1518 |  |  |  |  |  |
| Log likelihood | -1650.5089 |  |  |  |  |  |
| LR chi2(6) | 164.82 |  |  |  |  |  |

Table 10: Ordered Probit Estimates of Model (2) for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Health(Excellent)* | -1.004311 | 0.027875 | 1298.161 | 0.000 | 2491.17 | 0.0000 |
| Health(Very good)* | -0.453385 | 0.025672 | 311.876 | 0.000 |  |  |
| Health(Fair)* | 0.676079 | 0.046207 | 214.037 | 0.000 |  |  |
| Health(Poor)* | 1.457789 | 0.079307 | 337.824 | 0.000 |  |  |
| Income | -2.21E-06 | $3.75 \mathrm{E}-07$ | 34.928 | 0.000 | 35.01 | 0.0000 |
| LFS (Unemployed)* | 0.168270 | 0.067222 | 6.250 | 0.012 | 70.88 | 0.0000 |
| LFS (NLF)* | 0.270368 | 0.033752 | 64.160 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.063802 | 0.035189 | 3.276 | 0.070 |  |  |
| LFS (EPY-NLFPY)* | 0.045998 | 0.043381 | 1.124 | 0.289 |  |  |
| LFS (UPY-NLFPY)* | 0.079518 | 0.077555 | 1.061 | 0.305 |  |  |
| LFS (AOA)* | -0.037866 | 0.046540 | 0.656 | 0.416 |  |  |
| HLE(GHS)* | -0.131641 | 0.029159 | 20.340 | 0.000 | 61.73 | 0.0000 |
| HLE(NUPC)* | -0.146865 | 0.029456 | 24.900 | 0.000 |  |  |
| HLE(UD)* | -0.272922 | 0.034859 | 61.309 | 0.000 |  |  |
| Age | 0.012355 | 0.001250 | 97.812 | 0.000 | 97.87 | 0.0000 |
| cut1 | -0.866374 | 0.058535 |  |  |  |  |
| cut2 | 0.343329 | 0.058177 |  |  |  |  |
| cut3 | 1.421231 | 0.059629 |  |  |  |  |
| cut4 | 2.250021 | 0.064942 |  |  |  |  |
| N | 12956 |  |  |  |  |  |
| Log likelihood | -15300.991 |  |  |  |  |  |
| LR chi2(15) | 3459.83 |  |  |  |  |  |

Table 11: Ordered Probit Estimates of Model (2) for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Health(Excellent)* | -1.156261 | 0.060221 | 368.640 | 0.000 | 877.47 | 0.0000 |
| Health(Very good)* | -0.521807 | 0.051337 | 103.226 | 0.000 |  |  |
| Health(Fair)* | 0.658870 | 0.068914 | 91.394 | 0.000 |  |  |
| Health(Poor)* | 1.295018 | 0.102309 | 160.276 | 0.000 |  |  |
| Income | $-1.02 \mathrm{E}-06$ | $5.88 \mathrm{E}-07$ | 3.028 | 0.082 | 3.04 | 0.0815 |
| LFS (Unemployed)* | 0.271819 | 0.134235 | 4.080 | 0.043 | 19.77 | 0.003 |
| LFS (NLF)* | 0.211217 | 0.050958 | 17.140 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.120436 | 0.093736 | 1.638 | 0.199 |  |  |
| LFS (EPY-NLFPY)* | 0.074192 | 0.094603 | 0.608 | 0.433 |  |  |
| LFS (UPY-NLFPY)* | 0.052626 | 0.186688 | 0.078 | 0.778 |  |  |
| LFS (AOA)* | 0.057462 | 0.137763 | 0.176 | 0.677 |  |  |
| HLE(GHS)* | -0.135462 | 0.053120 | 6.503 | 0.011 | 25.75 | 0.0000 |
| HLE(NUPC)* | -0.175569 | 0.050787 | 11.972 | 0.001 |  |  |
| HLE(UD)* | -0.312010 | 0.066044 | 22.278 | 0.000 |  |  |
| cut1 | -1.484897 | 0.062976 |  |  |  |  |
| cut2 | -0.328522 | 0.058316 |  |  |  |  |
| cut3 | 0.645909 | 0.059455 |  |  |  |  |
| cut4 | 1.576161 | 0.068579 |  |  |  |  |
| N | 2994 |  |  |  |  |  |
| Log likelihood | -3841.8602 |  |  |  |  |  |
| LR chi2(14) | 1184.38 |  |  |  |  |  |

Table 12: Ordered Probit Estimates of Model (2) for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Health(Excellent)* | -1.036182 | 0.052793 | 385.337 | 0.000 | 1184.14 | 0.0000 |
| Health(Very good)* | -0.470980 | 0.040331 | 136.422 | 0.000 |  |  |
| Health(Fair)* | 0.559971 | 0.046506 | 144.962 | 0.000 |  |  |
| Health(Poor)* | 1.257874 | 0.076004 | 273.903 | 0.000 |  |  |
| Income | -2.00E-06 | $8.15 \mathrm{E}-07$ | 6.003 | 0.014 | 6.02 | 0.0142 |
| HLE(GHS)* | -0.079089 | 0.044082 | 3.204 | 0.073 | 27.99 | 0.0000 |
| HLE(NUPC)* | -0.143843 | 0.043971 | 10.693 | 0.001 |  |  |
| HLE(UD)* | -0.313898 | 0.066833 | 22.090 | 0.000 |  |  |
| cut1 | -1.676028 | 0.045331 |  |  |  |  |
| cut2 | -0.521913 | 0.039439 |  |  |  |  |
| cut3 | 0.474129 | 0.039330 |  |  |  |  |
| cut4 | 1.439583 | 0.045692 |  |  |  |  |
| N | 4563 |  |  |  |  |  |
| Log likelihood | -6090.2129 |  |  |  |  |  |
| LR chi2(8) | 1392.48 |  |  |  |  |  |

Table 13: Marginal Effects of Model (2) for the Persons of Age Under 20

|  | Health status 1999 |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.282567 | -0.095656 | -0.136674 | -0.038747 | -0.011489 |
| Health(Very good) | 0.135645 | -0.056578 | -0.060617 | -0.014701 | -0.003750 |
| Health(Fair) | -0.125826 | 0.031522 | 0.067067 | 0.020797 | 0.006441 |
| Health(Poor)* | -0.363560 | -0.084167 | 0.226252 | 0.140028 | 0.081447 |
| Income | $2.15 \mathrm{E}-06$ | $-8.08 \mathrm{E}-07$ | $-1.01 \mathrm{E}-06$ | $-2.60 \mathrm{E}-07$ | $-6.88 \mathrm{E}-08$ |
| Sex(Female)* | -0.069207 | 0.026057 | 0.032564 | 0.008366 | 0.002221 |
| (*) dy/dx is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 14: Marginal Effects of Model (2) for the People between 20 and 49 Years of Age

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.345202 | -0.043380 | -0.215819 | -0.070284 | -0.015719 |
| Health(Very good)* | 0.149397 | 0.000851 | -0.105309 | -0.036771 | -0.008168 |
| Health(Fair)* | -0.169661 | -0.087255 | 0.143443 | 0.085207 | 0.028267 |
| Health(Poor)* | -0.245164 | -0.283647 | 0.155900 | 0.226934 | 0.145977 |
| Income | $7.12 \mathrm{E}-07$ | $4.72 \mathrm{E}-08$ | -5.25E-07 | -1.91E-07 | -4.30E-08 |
| LFS (Unemployed)* | -0.051078 | -0.008945 | 0.039766 | 0.016262 | 0.003995 |
| LFS (NLF)* | -0.080354 | -0.017144 | 0.063486 | 0.027112 | 0.006901 |
| LFS (EPY-UPY)* | -0.020136 | -0.002034 | 0.015128 | 0.005719 | 0.001324 |
| LFS (EPY-NLFPY)* | -0.014578 | -0.001356 | 0.010906 | 0.004088 | 0.000940 |
| LFS (UPY-NLFPY)* | -0.024888 | -0.002915 | 0.018854 | 0.007250 | 0.001699 |
| LFS (AOA)* | 0.012302 | 0.000547 | -0.008956 | -0.003188 | -0.000706 |
| HLE(GHS)* | 0.042851 | 0.001706 | -0.031087 | -0.011030 | -0.002439 |
| HLE(NUPC)* | 0.047957 | 0.001601 | -0.034652 | -0.012216 | -0.002689 |
| HLE(UD)* | 0.092495 | -0.004068 | -0.063478 | -0.020648 | -0.004300 |
| Age | -0.003970 | -0.000263 | 0.002927 | 0.001066 | 0.000240 |
| (*) $\mathrm{d} y / \mathrm{dx}$ | for discret | hange of d | my variab | from 0 to 1 |  |

Table 15: Marginal Effects of Model (2) for the People between 50 and 59 Years of Age

|  | Health status 1999 |  |  |  | Fair |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Variables 1996 | Excellent | Very good | Good | Poor |  |
| Health(Excellent)* | 0.336070 | 0.071979 | -0.218359 | -0.148738 | -0.040952 |
| Health(Very good)* | 0.125703 | 0.077821 | -0.092810 | -0.084681 | -0.026033 |
| Health(Fair)* | -0.109680 | -0.144669 | 0.064511 | 0.128765 | 0.061073 |
| Health(Poor)* | -0.144143 | -0.287325 | -0.007789 | 0.232759 | 0.206498 |
| Income | $2.27 \mathrm{E}-07$ | $1.80 \mathrm{E}-07$ | $-1.71 \mathrm{E}-07$ | $-1.78 \mathrm{E}-07$ | $-5.76 \mathrm{E}-08$ |
| LFS (Unemployed)* | -0.052072 | -0.055867 | 0.036843 | 0.051311 | 0.019785 |
| LFS (NLF)* | -0.044198 | -0.039893 | 0.032692 | 0.038108 | 0.013290 |
| LFS (EPY-UPY)* | -0.025176 | -0.022833 | 0.018685 | 0.021764 | 0.007561 |
| LFS (EPY-NLFPY)* | -0.015877 | -0.013691 | 0.011878 | 0.013221 | 0.004469 |
| LFS (UPY-NLFPY)* | -0.011361 | -0.009607 | 0.008521 | 0.009327 | 0.003120 |
| LFS (AOA)* | -0.012380 | -0.010516 | 0.009280 | 0.010197 | 0.003419 |
| HLE(GHS)* | 0.031279 | 0.022393 | -0.023669 | -0.022893 | -0.007110 |
| HLE(NUPC)* | 0.040694 | 0.028806 | -0.030760 | -0.029576 | -0.009164 |
| HLE(UD)* | 0.077673 | 0.044201 | -0.058261 | -0.049412 | -0.014201 |
| (*)dy/dx is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 16: Marginal Effects of Model (2) for the People 60 and Over Years of Age

|  | Health status 1999 |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Health(Excellent)* | 0.235427 | 0.158759 | -0.163197 | -0.173351 | -0.057637 |
| Health(Very good)* | 0.077320 | 0.105788 | -0.048430 | -0.095211 | -0.039468 |
| Health(Fair)* | -0.062474 | -0.136896 | 0.008171 | 0.118217 | 0.072981 |
| Health(Poor)* | -0.082410 | -0.264259 | -0.118655 | 0.195420 | 0.269906 |
| Income | $2.86 \mathrm{E}-07$ | $4.81 \mathrm{E}-07$ | $-1.52 \mathrm{E}-07$ | $-4.23 \mathrm{E}-07$ | $-1.93 \mathrm{E}-07$ |
| HLE(GHS) $^{*}$ | 0.011764 | 0.018804 | -0.006669 | -0.016591 | -0.007307 |
| HLE(NUPC)* $^{*}$ | 0.022049 | 0.033798 | -0.013083 | -0.029935 | -0.012829 |
| HLE(UD)* | 0.054190 | 0.069278 | -0.036580 | -0.062723 | -0.024165 |
| (*) dy/dx is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 17: Ordered Probit Estimates of Model (1) with Stress

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.281827 | 0.016050 | 308.354 | 0.000 | 1248.91 | 0.0000 |
| Stress (Not very stressful)* | -0.507550 | 0.017583 | 833.477 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.695707 | 0.022861 | 925.985 | 0.000 |  |  |
| Income | -3.54E-06 | $2.05 \mathrm{E}-07$ | 297.563 | 0.000 | 300.32 | 0.0000 |
| LFS (Unemployed)* | 0.323375 | 0.039375 | 67.404 | 0.000 | 701.38 | 0.0000 |
| LFS (NLF)* | 0.401283 | 0.015634 | 658.949 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.103072 | 0.021293 | 23.426 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | 0.044348 | 0.022896 | 3.764 | 0.053 |  |  |
| LFS (UPY-NLFPY)* | 0.253405 | 0.039414 | 41.345 | 0.000 |  |  |
| LFS (AOA)* | 0.114179 | 0.025818 | 19.536 | 0.000 |  |  |
| Sex(Female)* | 0.031146 | 0.010973 | 8.066 | 0.005 | 8.06 | 0.0045 |
| HLE(GHS)* | -0.171572 | 0.014734 | 135.490 | 0.000 | 401.66 | 0.0000 |
| HLE(NUPC)* | -0.170414 | 0.014896 | 130.874 | 0.000 |  |  |
| HLE(UD)* | -0.374377 | 0.019002 | 388.090 | 0.000 |  |  |
| Age | 0.019438 | 0.000421 | 2137.213 | 0.000 | 2153.84 | 0.0000 |
| cut1 | -0.378357 | 0.027390 |  |  |  |  |
| cut2 | 0.716455 | 0.027572 |  |  |  |  |
| cut3 | 1.635724 | 0.028338 |  |  |  |  |
| cut4 | 2.358909 | 0.030192 |  |  |  |  |
| N | 40296 |  |  |  |  |  |
| Log likelihood | -51634.161 |  |  |  |  |  |
| LR chi2(15) | 6207.69 |  |  |  |  |  |

Table 18: Marginal Effect of Model (1) with Stress

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.092173 | 0.008355 | -0.057422 | -0.030832 | -0.012274 |
| Stress (Not very stressful)* | 0.175475 | -0.005024 | -0.103235 | -0.049145 | -0.018070 |
| Stress (Not at all stressful)* | 0.256241 | -0.047351 | -0.136598 | -0.054580 | -0.017712 |
| Income | $1.16 \mathrm{E}-06$ | $1.13 \mathrm{E}-07$ | -7.25E-07 | -3.90E-07 | -1.54E-07 |
| LFS (Unemployed)* | -0.094397 | -0.028207 | 0.061610 | 0.041449 | 0.019545 |
| LFS (NLF)* | -0.120412 | -0.029593 | 0.077636 | 0.049670 | 0.022698 |
| LFS (EPY-UPY)* | -0.032696 | -0.004969 | 0.020845 | 0.011892 | 0.004929 |
| LFS (EPY-NLFPY)* | -0.014304 | -0.001736 | 0.009041 | 0.004986 | 0.002012 |
| LFS (UPY-NLFPY)* | -0.075966 | -0.019214 | 0.049380 | 0.031531 | 0.014269 |
| LFS (AOA)* | -0.036036 | -0.005813 | 0.023029 | 0.013274 | 0.005545 |
| Sex(Female)* | -0.010174 | -0.000990 | 0.006382 | 0.003427 | 0.001356 |
| HLE(GHS)* | 0.057208 | 0.003226 | -0.035302 | -0.018163 | -0.006969 |
| HLE(NUPC)* | 0.056838 | 0.003175 | -0.035067 | -0.018031 | -0.006914 |
| HLE(UD)* | 0.131413 | -0.006952 | -0.076837 | -0.035189 | -0.012435 |
| Age | -0.006348 | -0.000622 | 0.003983 | 0.002140 | 0.000847 |
| (*) $\mathrm{d} y / \mathrm{dx}$ is | discrete ch | ge of dumm | variable fr | 0 to 1 |  |

Table 19: Ordered Probit Estimates of Model (2) with Stress

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Stress (Somewhat stressful)* | -0.296532 | 0.022742 | 170.042 | 0.000 | 625.34 | 0.0000 |
| Stress (Not very stressful)* | -0.506579 | 0.024844 | 415.752 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.707155 | 0.032336 | 478.297 | 0.000 |  |  |
| Income | $-3.63 E-06$ | $2.84 \mathrm{E}-07$ | 163.328 | 0.000 | 165.10 | 0.0000 |
| LFS (Unemployed)* | 0.301997 | 0.056783 | 28.302 | 0.000 | 307.48 |  |
| LFS (NLF)* | 0.367382 | 0.021556 | 290.362 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.108650 | 0.031482 | 11.903 | 0.001 |  |  |
| LFS (EPY-NLFPY)* | 0.068334 | 0.032312 | 4.452 | 0.034 |  |  |
| LFS (UPY-NLFPY)* | 0.240805 | 0.057391 | 17.640 | 0.000 |  |  |
| LFS (AOA)* | 0.080625 | 0.038485 | 4.368 | 0.036 | 0.000 | 263.98 |
| HLE(GHS)* | -0.190123 | 0.020766 | 83.906 | 0.000 |  |  |
| HLE(NUPC)* | -0.200819 | 0.021359 | 88.360 | 0.000 |  |  |
| HLE(UD)* | -0.429119 | 0.026827 | 256.000 | 0.000 | 0.000 | 971.84 |
| Age | 0.018514 | 0.000596 | 965.345 |  | 0.0000 |  |
| cut1 | -0.517117 | 0.038252 |  |  |  |  |
| cut2 | 0.567524 | 0.038349 |  |  |  |  |
| cut3 | 1.474977 | 0.039328 |  |  |  |  |
| cut4 | 2.210695 | 0.041849 |  |  |  |  |
| N | 19750 |  |  |  |  |  |
| Log likelihood | -25723.478 |  |  |  |  |  |
| LR chi2(14) | 2981.75 |  |  |  |  |  |

Table 20: Marginal Effect of Model (2) with Stress

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.094182 | 0.014145 | -0.058679 | -0.035174 | -0.014474 |
| Stress (Not very stressful)* | 0.170705 | 0.004806 | -0.101357 | -0.053751 | -0.020403 |
| Stress (Not at all stressful)* | 0.256339 | -0.036223 | -0.138715 | -0.061077 | -0.020323 |
| Income | $1.15 \mathrm{E}-06$ | $1.85 \mathrm{E}-07$ | -7.21E-07 | -4.34E-07 | -1.78E-07 |
| LFS (Unemployed)* | -0.085613 | -0.030466 | 0.055013 | 0.041154 | 0.019911 |
| LFS (NLF)* | -0.107409 | -0.032035 | 0.068473 | 0.048394 | 0.022578 |
| LFS (EPY-UPY)* | -0.033276 | -0.007364 | 0.021167 | 0.013610 | 0.005863 |
| LFS (EPY-NLFPY)* | -0.021195 | -0.004209 | 0.013428 | 0.008419 | 0.003557 |
| LFS (UPY-NLFPY)* | -0.069940 | -0.021995 | 0.044926 | 0.032055 | 0.014954 |
| LFS (AOA)* | -0.024881 | -0.005169 | 0.015791 | 0.010001 | 0.004258 |
| HLE(GHS)* | 0.061630 | 0.007086 | -0.038150 | -0.021891 | -0.008675 |
| HLE(NUPC)* | 0.065425 | 0.006887 | -0.040371 | -0.022925 | -0.009016 |
| HLE(UD)* | 0.148273 | -0.002604 | -0.086737 | -0.043332 | -0.015600 |
| Age | -0.005862 | -0.000942 | 0.003683 | 0.002215 | 0.000907 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is f | discrete ch | e of dumm | variable fr | 0 to 1 |  |

Table 21: Ordered Probit Estimates of Model (1) with Stress for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.213219 | 0.083770 | 6.503 | 0.011 | 31.2 | 0.0000 |
| Stress (Not very stressful)* | -0.354510 | 0.084967 | 17.389 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.439202 | 0.093350 | 22.090 | 0.000 |  |  |
| Income | -4.58E-06 | $7.14 \mathrm{E}-07$ | 41.088 | 0.000 | 41.49 | 0.0000 |
| LFS (Unemployed)* | 0.502055 | 0.211628 | 5.617 | 0.018 | 15.09 | 0.0196 |
| LFS (NLF)* | 0.109921 | 0.060388 | 3.312 | 0.069 |  |  |
| LFS (EPY-UPY)* | 0.316312 | 0.114523 | 7.618 | 0.006 |  |  |
| LFS (EPY-NLFPY)* | 0.031588 | 0.060760 | 0.270 | 0.603 |  |  |
| LFS (UPY-NLFPY)* | 0.117047 | 0.082879 | 1.988 | 0.158 |  |  |
| LFS (AOA)* | 0.126186 | 0.070049 | 3.240 | 0.072 |  |  |
| Sex(Female)* | 0.225913 | 0.040385 | 31.248 | 0.000 | 31.32 | $0.0000$ |
| HLE(GHS)* | -0.233789 | 0.055903 | 17.472 | 0.000 | 19.26 | $0.0002$ |
| HLE(NUPC)* | -0.028068 | 0.145453 | 0.036 | 0.847 |  |  |
| HLE(UD)* | -0.724128 | 0.740895 | 0.960 | 0.328 |  |  |
| Age | 0.091670 | 0.025353 | 13.104 | 0.000 | 13.08 | 0.0003 |
| cut1 | 1.136849 | 0.448269 |  |  |  |  |
| cut2 | 2.220766 | 0.449245 |  |  |  |  |
| cut3 | 3.036654 | 0.450831 |  |  |  |  |
| cut4 | 3.863475 | 0.458579 |  |  |  |  |
| N | 3237 |  |  |  |  |  |
| Log likelihood | -3508.4482 |  |  |  |  |  |
| LR chi2(15) | 167.49 |  |  |  |  |  |

Table 22: Ordered Probit Estimates of Model (1) with Stress for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.274780 | 0.018866 | 212.285 | 0.000 | 507.37 | 0.0000 |
| Stress (Not very stressful)* | -0.443952 | 0.021544 | 424.772 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.532240 | 0.032076 | 275.228 | 0.000 |  |  |
| Income | -3.25E-06 | $2.65 \mathrm{E}-07$ | 150.553 | 0.000 | 151.77 | 0.0000 |
| LFS (Unemployed)* | 0.281193 | 0.047138 | 35.641 | 0.000 | 449.59 | 0.0000 |
| LFS (NLF)* | 0.496404 | 0.024190 | 421.070 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.076013 | 0.023737 | 10.240 | 0.001 |  |  |
| LFS (EPY-NLFPY)* | 0.054702 | 0.030512 | 3.204 | 0.073 |  |  |
| LFS (UPY-NLFPY)* | 0.291953 | 0.052954 | 30.360 | 0.000 |  |  |
| LFS (AOA)* | 0.107216 | 0.030744 | 12.180 | 0.000 |  |  |
| Sex(Female)* | 0.035050 | 0.013751 | 6.503 | 0.011 | 6.50 | 0.0108 |
| HLE(GHS)* | -0.208500 | 0.020571 | 102.820 | 0.000 | 285.28 | 0.0000 |
| HLE(NUPC)* | -0.220491 | 0.020457 | 116.208 | 0.000 |  |  |
| HLE(UD)* | -0.414677 | 0.024606 | 283.923 | 0.000 |  |  |
| Age | 0.020747 | 0.000867 | 573.124 | 0.000 | 575.06 | 0.0000 |
| cut1 | -0.372401 | 0.042267 |  |  |  |  |
| cut2 | 0.752048 | 0.042430 |  |  |  |  |
| cut3 | 1.719348 | 0.043406 |  |  |  |  |
| cut4 | 2.390986 | 0.045947 |  |  |  |  |
| N | 26171 |  |  |  |  |  |
| Log likelihood | -32643.15 |  |  |  |  |  |
| LR chi2(15) | 2311.69 |  |  |  |  |  |

Table 23: Ordered Probit Estimates of Model (1) with Stress for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.269727 | 0.039893 | 45.698 | 0.000 | 272.74 | 0.0000 |
| Stress (Not very stressful)* | -0.561774 | 0.043146 | 169.520 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.763141 | 0.055021 | 192.377 | 0.000 |  |  |
| Income | -3.02E-06 | $4.51 \mathrm{E}-07$ | 45.024 | 0.000 | 45.48 | 0.0000 |
| LFS (Unemployed)* | 0.521116 | 0.086247 | 36.482 | 0.000 | 479.54 | 0.0000 |
| LFS (NLF)* | 0.750896 | 0.035909 | 437.228 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.110715 | 0.060585 | 3.349 | 0.068 |  |  |
| LFS (EPY-NLFPY)* | 0.202507 | 0.069449 | 8.526 | 0.004 |  |  |
| LFS (UPY-NLFPY)* | 0.957657 | 0.125573 | 58.217 | 0.000 |  |  |
| LFS (AOA)* | 0.266106 | 0.091784 | 8.410 | 0.004 |  |  |
| Sex(Female)* | -0.125424 | 0.028132 | 19.892 | 0.000 | $19.89$ |  |
| HLE(GHS)* | -0.304413 | 0.037446 | 66.097 | 0.000 | $133.48$ | $0.0000$ |
| HLE(NUPC)* | -0.239174 | 0.034863 | 47.060 | 0.000 |  |  |
| HLE(UD)* | -0.474154 | 0.045747 | 107.330 | 0.000 |  |  |
| cut1 | -1.421490 | 0.049713 |  |  |  |  |
| cut2 | -0.383565 | 0.048087 |  |  |  |  |
| cut3 | 0.474076 | 0.048315 |  |  |  |  |
| cut4 | 1.240774 | 0.051560 |  |  |  |  |
| N | 6143 |  |  |  |  |  |
| Log likelihood | -8454.6983 |  |  |  |  |  |
| LR chi2(14) | 1052.72 |  |  |  |  |  |

Table 24: Ordered Probit Estimates of Model (1) with Stress for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Stress (Somewhat stressful)* | -0.235316 | 0.061614 | 14.592 | 0.000 | 305.49 | 0.0000 |
| Stress (Not very stressful)* | -0.619335 | 0.060476 | 104.858 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.859238 | 0.064190 | 179.292 | 0.000 |  |  |
| Income | $-3.01 \mathrm{E}-06$ | $7.29 \mathrm{E}-07$ | 17.057 | 0.000 | 17.08 | 0.0000 |
| LFS (Unemployed)* | 0.043088 | 0.170983 | 0.063 | 0.801 | 166.07 | 0.0000 |
| LFS (NLF)* | 0.496433 | 0.040759 | 148.352 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.195031 | 0.125044 | 2.434 | 0.119 |  |  |
| LFS (EPY-NLFPY)* | 0.216650 | 0.078999 | 7.508 | 0.006 |  |  |
| LFS (UPY-NLFPY)* | 0.036621 | 0.187859 | 0.036 | 0.845 |  |  |
| LFS (AOA)* | 0.073797 | 0.164705 | 0.203 | 0.654 | 0.000 | 87.93 |
| HLE(GHS)* | -0.205701 | 0.041875 | 24.108 | 0.000 |  |  |
| HLE(NUPC)* | -0.251948 | 0.041151 | 37.454 | 0.0000 |  |  |
| HLE(UD)* | -0.472899 | 0.060349 | 61.466 | 0.000 |  |  |
| cut1 | -1.540250 | 0.071739 |  |  |  |  |
| cut2 | -0.492900 | 0.070002 |  |  |  |  |
| cut3 | 0.420803 | 0.069955 |  |  |  |  |
| cut4 | 1.259002 | 0.072104 |  |  |  |  |
| N | 4745 |  |  |  |  |  |

Table 25: Marginal Effects of Model (1) with Stress for the Persons of Age Under 20

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.084880 | -0.036433 | -0.033770 | -0.012453 | -0.002225 |
| Stress (Not very stressful)* | 0.140692 | -0.062027 | -0.055080 | -0.020032 | -0.003553 |
| Stress (Not at all stressful)* | 0.172966 | -0.085750 | -0.062923 | -0.020885 | -0.003408 |
| Income | $1.82 \mathrm{E}-06$ | -7.60E-07 | -7.38E-07 | -2.77E-07 | -5.01E-08 |
| LFS (Unemployed)* | -0.190111 | 0.044324 | 0.089901 | 0.045015 | 0.010873 |
| LFS (NLF)* | -0.043731 | 0.017575 | 0.017966 | 0.006909 | 0.001281 |
| LFS (EPY-UPY)* | -0.123380 | 0.038644 | 0.055126 | 0.024408 | 0.005202 |
| LFS (EPY-NLFPY)* | -0.012583 | 0.005173 | 0.005119 | 0.001938 | 0.000354 |
| LFS (UPY-NLFPY)* | -0.046463 | 0.017857 | 0.019450 | 0.007691 | 0.001465 |
| LFS (AOA)* | -0.050098 | 0.019335 | 0.020935 | 0.008259 | 0.001569 |
| Sex(Female)* | -0.089827 | 0.037062 | 0.036425 | 0.013807 | 0.002533 |
| HLE(GHS)* | 0.093053 | -0.040702 | -0.036653 | -0.013342 | -0.002356 |
| HLE(NUPC)* | 0.011192 | -0.004755 | -0.004481 | -0.001659 | -0.000297 |
| HLE(UD)* | 0.269474 | -0.158897 | -0.083681 | -0.023556 | -0.003340 |
| Age | -0.036534 | 0.015215 | 0.014771 | 0.005544 | 0.001004 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is f | discrete ch | ge of dumm | variable fr | 0 to 1 |  |

Table 26: Marginal Effects of Model (1) with Stress for the People between 20 and 49 Years of Age

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.093444 | -0.000098 | -0.059219 | -0.024306 | -0.009820 |
| Stress (Not very stressful)* | 0.159747 | -0.020279 | -0.093599 | -0.033539 | -0.012330 |
| Stress (Not at all stressful)* | 0.199613 | -0.046610 | -0.107697 | -0.033943 | -0.011363 |
| Income | $1.11 \mathrm{E}-06$ | -1.17E-08 | -7.05E-07 | -2.84E-07 | -1.13E-07 |
| LFS (Unemployed)* | -0.088405 | -0.013296 | 0.059707 | 0.028842 | 0.013152 |
| LFS (NLF)* | -0.148384 | -0.034891 | 0.101700 | 0.054460 | 0.027114 |
| LFS (EPY-UPY)* | -0.025569 | -0.000628 | 0.016466 | 0.006909 | 0.002823 |
| LFS (EPY-NLFPY)* | -0.018473 | -0.000314 | 0.011854 | 0.004931 | 0.002001 |
| LFS (UPY-NLFPY)* | -0.091371 | -0.014493 | 0.061860 | 0.030155 | 0.013849 |
| LFS (AOA)* | -0.035701 | -0.001578 | 0.023196 | 0.009949 | 0.004133 |
| Sex(Female)* | -0.012006 | 0.000136 | 0.007596 | 0.003061 | 0.001213 |
| HLE(GHS)* | 0.072674 | -0.003584 | -0.044950 | -0.017419 | -0.006722 |
| HLE(NUPC)* | 0.076788 | -0.003673 | -0.047523 | -0.018455 | -0.007136 |
| HLE(UD)* | 0.150980 | -0.022948 | -0.087051 | -0.030183 | -0.010799 |
| Age | -0.007104 | 0.0000744 | 0.004496 | 0.001814 | 0.000719 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is f | discrete ch | ge of dumm | variable fr | 0 to 1 |  |

Table 27: Marginal Effects of Model (1) with Stress for the People between 50 and 59 Years of Age

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.073939 | 0.031766 | -0.042395 | -0.041750 | -0.021560 |
| Stress (Not very stressful)* | 0.167197 | 0.045731 | -0.095130 | -0.079974 | -0.037824 |
| Stress (Not at all stressful)* | 0.253475 | 0.015286 | -0.138738 | -0.092530 | -0.037492 |
| Income | $8.20 \mathrm{E}-07$ | $3.71 \mathrm{E}-07$ | -4.72E-07 | -4.73E-07 | -2.46E-07 |
| LFS (Unemployed)* | -0.110650 | -0.094145 | 0.050935 | 0.089044 | 0.064816 |
| LFS (NLF)* | -0.169037 | -0.123467 | 0.080617 | 0.124244 | 0.087644 |
| LFS (EPY-UPY)* | -0.028719 | -0.015161 | 0.016228 | 0.017828 | 0.009824 |
| LFS (EPY-NLFPY)* | -0.050397 | -0.030076 | 0.027777 | 0.033292 | 0.019404 |
| LFS (UPY-NLFPY)* | -0.158918 | -0.193208 | 0.032210 | 0.155068 | 0.164847 |
| LFS (AOA)* | -0.064029 | -0.041765 | 0.034390 | 0.044349 | 0.027055 |
| Sex(Female)* | 0.033985 | 0.015397 | -0.019521 | -0.019634 | -0.010226 |
| HLE(GHS)* | 0.088634 | 0.028761 | -0.051296 | -0.044777 | -0.021322 |
| HLE(NUPC)* | 0.067729 | 0.025377 | -0.039176 | -0.036120 | -0.017811 |
| HLE(UD)* | 0.146470 | 0.031224 | -0.083819 | -0.065079 | -0.028796 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is | discrete ch | e of dumm | variable fr | 0 to 1 |  |

Table 28: Marginal Effects of Model (1) with Stress for the People 60 and Over Years of Age

|  | Health status 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.050159 | 0.043255 | -0.027540 | -0.042852 | -0.023023 |
| Stress (Not very stressful)* | 0.135326 | 0.107533 | -0.073007 | -0.109781 | -0.060070 |
| Stress (Not at all stressful)* | 0.219782 | 0.110898 | -0.127297 | -0.138189 | -0.065195 |
| Income | $6.08 \mathrm{E}-07$ | $5.86 \mathrm{E}-07$ | -3.16E-07 | -5.60E-07 | -3.18E-07 |
| LFS (Unemployed)* | -0.008478 | -0.008539 | 0.004258 | 0.008057 | 0.004702 |
| LFS (NLF)* | -0.112552 | -0.083453 | 0.063797 | 0.087340 | 0.044869 |
| LFS (EPY-UPY)* | -0.035114 | -0.040802 | 0.015019 | 0.036969 | 0.023929 |
| LFS (EPY-NLFPY)* | -0.038822 | -0.045408 | 0.016410 | 0.041065 | 0.026755 |
| LFS (UPY-NLFPY)* | -0.007232 | -0.007237 | 0.003651 | 0.006842 | 0.003976 |
| LFS (AOA)* | -0.014262 | -0.014813 | 0.006971 | 0.013853 | 0.008252 |
| HLE(GHS)* | 0.044713 | 0.037082 | -0.025063 | -0.037219 | -0.019512 |
| HLE(NUPC)** | 0.055499 | 0.044657 | -0.031399 | -0.045300 | -0.023456 |
| HLE(UD)* | 0.117761 | 0.068611 | -0.070917 | -0.079265 | -0.036190 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is f | discrete ch | e of dumm | variable fr | 0 to 1 |  |

Table 29: Ordered Probit Estimates of Model (2) with Stress for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Stress (Somewhat stressful)* | -0.159261 | 0.127333 | 1.563 | 0.211 | 13.83 | 0.0031 |
| Stress (Not very stressful)* | -0.340717 | 0.127953 | 7.076 | 0.008 |  |  |
| Stress (Not at all stressful)* | -0.353967 | 0.139793 | 6.401 | 0.011 |  |  |
| Income | $-6.55 \mathrm{E}-06$ | $1.07 \mathrm{E}-06$ | 37.454 | 0.000 | 38.22 | 0.0000 |
| Sex(Female)* | 0.181083 | 0.058738 | 9.486 | 0.002 | 9.51 | 0.0020 |
| cut1 | -0.630899 | 0.133717 |  |  |  |  |
| cut2 | 0.461960 | 0.133559 |  |  |  |  |
| cut3 | 1.369789 | 0.140923 |  |  |  |  |
| cut4 | 2.002284 | 0.165677 |  |  |  |  |
| N | 1501 |  |  |  |  |  |
| Log likelihood | -1678.6349 |  |  |  |  |  |
| LR chi2 $(5)$ | 65.09 |  |  |  |  |  |

Table 30: Ordered Probit Estimates of Model (2) with Stress for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.291619 | 0.026751 | 118.810 | 0.000 | 295.81 | 0.0000 |
| Stress (Not very stressful)* | -0.466209 | 0.030394 | 235.316 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.600891 | 0.045309 | 175.828 | 0.000 |  |  |
| Income | -3.47E-06 | $3.72 \mathrm{E}-07$ | 87.236 | 0.000 | 87.91 | 0.0000 |
| LFS (Unemployed)* | 0.287379 | 0.066631 | 18.576 | 0.000 | 266.46 | 0.0000 |
| LFS (NLF)* | 0.522358 | 0.032858 | 252.810 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.090932 | 0.034980 | 6.760 | 0.009 |  |  |
| LFS (EPY-NLFPY)* | 0.076068 | 0.042903 | 3.133 | 0.076 |  |  |
| LFS (UPY-NLFPY)* | 0.262056 | 0.077096 | 11.560 | 0.001 |  |  |
| LFS (AOA)* | 0.026388 | 0.046011 | 0.325 | 0.566 |  |  |
| HLE(GHS)* | -0.245737 | 0.028896 | 72.250 | 0.000 | 191.53 | 0.0000 |
| HLE(NUPC)* | -0.258819 | 0.029231 | 78.323 | 0.000 |  |  |
| HLE(UD)* | -0.477519 | 0.034605 | 190.440 | 0.000 |  |  |
| Age | 0.018696 | 0.001231 | 230.432 | 0.000 | 231.06 | 0.0000 |
| cut1 | -0.564175 | 0.059644 |  |  |  |  |
| cut2 | 0.546907 | 0.059678 |  |  |  |  |
| cut3 | 1.500223 | 0.060853 |  |  |  |  |
| cut4 | 2.201674 | 0.064338 |  |  |  |  |
| N | 12884 |  |  |  |  |  |
| Log likelihood | -16304.075 |  |  |  |  |  |
| LR chi2(14) | 1252.77 |  |  |  |  |  |

Table 31: Ordered Probit Estimates of Model (2) with Stress for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.315747 | 0.056852 | 30.803 | 0.000 | 117.63 | 0.0000 |
| Stress (Not very stressful)* | -0.525669 | 0.061460 | 73.103 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.760225 | 0.078219 | 94.478 | 0.000 |  |  |
| Income | -2.49E-06 | $5.82 \mathrm{E}-07$ | 18.233 | 0.000 | 18.48 | 0.0000 |
| LFS (Unemployed)* | 0.422680 | 0.132128 | 10.240 | 0.001 | 163.15 | 0.0000 |
| LFS (NLF)* | 0.620099 | 0.050451 | 151.044 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.129746 | 0.092970 | 1.960 | 0.163 |  |  |
| LFS (EPY-NLFPY)* | 0.258833 | 0.094174 | 7.563 | 0.006 |  |  |
| LFS (UPY-NLFPY)* | 0.743108 | 0.183301 | 16.403 | 0.000 |  |  |
| LFS (AOA)* | 0.305703 | 0.136989 | 4.973 | 0.026 |  |  |
| Sex(Female)* | -0.075103 | 0.039990 | 3.534 | 0.060 | 3.53 | 0.0604 |
| HLE(GHS)* | -0.264952 | 0.052448 | 25.503 | 0.000 | 74.57 | 0.0000 |
| HLE(NUPC)* | -0.271334 | 0.050352 | 29.052 | 0.000 |  |  |
| HLE(UD)* | -0.528092 | 0.065243 | 65.448 | 0.000 |  |  |
| cut1 | -1.482674 | 0.070433 |  |  |  |  |
| cut2 | -0.469017 | 0.067847 |  |  |  |  |
| cut3 | 0.362615 | 0.067641 |  |  |  |  |
| cut4 | 1.154406 | 0.071694 |  |  |  |  |
| N | 2976 |  |  |  |  |  |
| Log likelihood | -4195.3202 |  |  |  |  |  |
| LR chi2(14) | 427.78 |  |  |  |  |  |

Table 32: Ordered Probit Estimates of Model (2) with Stress for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stress (Somewhat stressful)* | -0.210559 | 0.083082 | 6.401 | 0.011 | 119.29 | 0.0000 |
| Stress (Not very stressful)* | -0.514940 | 0.081814 | 39.564 | 0.000 |  |  |
| Stress (Not at all stressful)* | -0.757817 | 0.087066 | 75.690 | 0.000 |  |  |
| Income | -3.63E-06 | $1.02 \mathrm{E}-06$ | 12.603 | 0.000 | 12.61 | 0.0004 |
| LFS (Unemployed)* | 0.361434 | 0.225460 | 2.560 | 0.109 | 54.83 | 0.0000 |
| LFS (NLF)* | 0.410013 | 0.056593 | 52.418 | 0.000 |  |  |
| LFS (EPY-UPY)* | 0.253523 | 0.167486 | 2.280 | 0.130 |  |  |
| LFS (EPY-NLFPY)* | 0.163143 | 0.110256 | 2.190 | 0.139 |  |  |
| LFS (UPY-NLFPY)* | 0.407122 | 0.292743 | 1.932 | 0.164 |  |  |
| LFS (AOA)* | 0.269788 | 0.224816 | 1.440 | 0.230 |  |  |
| HLE(GHS)* | -0.254296 | 0.059270 | 18.404 | 0.000 | 50.93 | 0.0000 |
| HLE(NUPC)* | -0.229274 | 0.057839 | 15.682 | 0.000 |  |  |
| HLE(UD)* | -0.522319 | 0.086602 | 36.361 | 0.000 |  |  |
| cut1 | -1.559055 | 0.096555 |  |  |  |  |
| cut2 | -0.502618 | 0.093438 |  |  |  |  |
| cut3 | 0.388354 | 0.093592 |  |  |  |  |
| cut4 | 1.202305 | 0.096792 |  |  |  |  |
| N | 2389 |  |  |  |  |  |
| Log likelihood | -3400.9562 |  |  |  |  |  |
| LR chi2(13) | 268.6 |  |  |  |  |  |

Table 33: Marginal Effects of Model (2) with Stress for the Persons of Age Under 20

|  |  | Health status 1999 |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |  |  |
| Stress (Somewhat stressful)* | 0.063003 | -0.023363 | -0.028740 | -0.008173 | -0.002727 |  |  |
| Stress (Not very stressful)* | 0.134572 | -0.051356 | -0.060508 | -0.017037 | -0.005671 |  |  |
| Stress (Not at all stressful)* | 0.140451 | -0.060968 | -0.059320 | -0.015375 | -0.004787 |  |  |
| Income | $2.59 \mathrm{E}-06$ | $-9.29 \mathrm{E}-07$ | $-1.20 \mathrm{E}-06$ | $-3.45 \mathrm{E}-07$ | $-1.17 \mathrm{E}-07$ |  |  |
| Sex(Female)* | -0.071469 | 0.025714 | 0.032984 | 0.009542 | 0.003229 |  |  |
| ( $^{*}$ ) dy/dx is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |  |  |

Table 34: Marginal Effects of Model (2) with Stress for the People between 20 and 49 Years of Age

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.097064 | 0.004667 | -0.061864 | -0.028622 | -0.011246 |
| Stress (Not very stressful)* | 0.165038 | -0.013987 | -0.097789 | -0.039276 | -0.013986 |
| Stress (Not at all stressful)* | 0.224055 | -0.048483 | -0.120771 | -0.041564 | -0.013238 |
| Income | $1.16 \mathrm{E}-06$ | $4.74 \mathrm{E}-08$ | -7.42E-07 | -3.38E-07 | -1.30E-07 |
| LFS (Unemployed)* | -0.087743 | -0.018529 | 0.059132 | 0.032575 | 0.014564 |
| LFS (NLF)* | -0.150712 | -0.045810 | 0.102212 | 0.063160 | 0.031150 |
| LFS (EPY-UPY)* | -0.029755 | -0.002537 | 0.019334 | 0.009252 | 0.003706 |
| LFS (EPY-NLFPY)* | -0.024955 | -0.002010 | 0.016187 | 0.007705 | 0.003073 |
| LFS (UPY-NLFPY)* | -0.080653 | -0.015902 | 0.054199 | 0.029385 | 0.012972 |
| LFS (AOA)* | -0.008775 | -0.000479 | 0.005633 | 0.002605 | 0.001016 |
| HLE(GHS)* | 0.084009 | -0.000278 | -0.052366 | -0.022817 | -0.008548 |
| HLE(NUPC)* | 0.088783 | -0.000935 | -0.055128 | -0.023837 | -0.008883 |
| HLE(UD)* | 0.172104 | -0.021474 | -0.099473 | -0.038084 | -0.013073 |
| Age | -0.006262 | -0.000255 | 0.003996 | 0.001819 | 0.000702 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is | discrete ch | ge of dumm | variable fr | 0 to 1 |  |

Table 35: Marginal Effects of Model (2) with Stress for the People between 50 and 59 Years of Age

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.083204 | 0.042804 | -0.043166 | -0.053400 | -0.029443 |
| Stress (Not very stressful)* | 0.149968 | 0.054789 | -0.079715 | -0.083232 | -0.041811 |
| Stress (Not at all stressful)* | 0.243489 | 0.033183 | -0.128649 | -0.103750 | -0.044273 |
| Income | $6.20 \mathrm{E}-07$ | $3.38 \mathrm{E}-07$ | -3.22E-07 | -4.09E-07 | -2.27E-07 |
| LFS (Unemployed)* | -0.088230 | -0.077393 | 0.035160 | 0.075172 | 0.055292 |
| LFS (NLF)* | -0.131704 | -0.101648 | 0.056094 | 0.104022 | 0.073236 |
| LFS (EPY-UPY)* | -0.031820 | -0.020352 | 0.015761 | 0.022775 | 0.013636 |
| LFS (EPY-NLFPY)* | -0.057271 | -0.041669 | 0.026662 | 0.043939 | 0.028339 |
| LFS (UPY-NLFPY)* | -0.127146 | -0.144438 | 0.028674 | 0.124665 | 0.118245 |
| LFS (AOA)* | -0.066094 | -0.051331 | 0.029509 | 0.052582 | 0.035335 |
| HLE(GHS)* | 0.073079 | 0.031678 | -0.039046 | -0.043414 | -0.022297 |
| HLE(NUPC)* | 0.074819 | 0.033382 | -0.039818 | -0.045007 | -0.023376 |
| HLE(UD)* | 0.157930 | 0.042623 | -0.085160 | -0.078941 | -0.036453 |
| Age | -0.000957 | -0.000521 | 0.000497 | 0.000631 | 0.000350 |
| (*) $d y / d x$ is | discrete ch | ge of dumn | variable fr | 0 to 1 |  |

Table 36: Marginal Effects of Model (2) with Stress for the People 60 and Over Years of Age

|  | Health status 1999 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | Excellent | Very good | Good | Fair | Poor |
| Stress (Somewhat stressful)* | 0.043575 | 0.039341 | -0.022819 | -0.037711 | -0.022386 |
| Stress (Not very stressful)* | 0.109320 | 0.091800 | -0.057228 | -0.090270 | -0.053623 |
| Stress (Not at all stressful)* | 0.185576 | 0.106883 | -0.104877 | -0.123272 | -0.064310 |
| Income | $7.21 \mathrm{E}-07$ | $7.15 \mathrm{E}-07$ | -3.58E-07 | -6.65E-07 | -4.14E-07 |
| LFS (Unemployed)* | -0.055799 | -0.076091 | 0.016460 | 0.064116 | 0.051315 |
| LFS (NLF)* | -0.093721 | -0.075152 | 0.050814 | 0.075173 | 0.042885 |
| LFS (EPY-UPY)* | -0.040153 | -0.049527 | 0.015214 | 0.043038 | 0.031429 |
| LFS (EPY-NLFPY)* | -0.030121 | -0.034728 | 0.012743 | 0.030798 | 0.021308 |
| LFS (UPY-NLFPY)* | -0.064324 | -0.093144 | 0.014941 | 0.076980 | 0.065547 |
| LFS (AOA)* | -0.044611 | -0.056753 | 0.015870 | 0.048859 | 0.036635 |
| HLE(GHS)* | 0.056030 | 0.045941 | -0.030677 | -0.045565 | -0.025729 |
| HLE(NUPC)* | 0.049899 | 0.042056 | -0.027039 | -0.041297 | -0.023618 |
| HLE(UD)* | 0.130932 | 0.073807 | -0.076940 | -0.085736 | -0.042062 |
| Age | 0.002109 | 0.002093 | -0.001046 | -0.001944 | -0.001211 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is | discrete c | e of dumm | variable fr | 0 to 1 |  |

Table 37: Gamma GLM Estimates of Model (3) for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.0000149 | $3.38 \mathrm{E}-07$ | 1933.361 | 0.000 | 450.98 | 0.0000 |
| Age | -0.028402 | 0.007333 | 14.977 | 0.000 | 3.25 | 0.0716 |
| Health(Very good)* | -0.035486 | 0.018214 | 3.803 | 0.051 | 5.45 | 0.2440 |
| Health(Good)* | -0.093076 | 0.026475 | 12.390 | 0.000 |  |  |
| Health(Fair)* | -0.164214 | 0.059380 | 7.673 | 0.006 |  |  |
| Health(Poor)* | -0.347124 | 0.120761 | 8.237 | 0.004 |  |  |
| Constant | 10.533660 | 0.129115 | 6655.296 | 0.000 |  |  |
| N | 3271 |  |  |  |  |  |
| Log likelihood | -38423.743 |  | AIC | 23.49786 |  |  |
| Deviance | 1087.77012 |  |  |  |  |  |

Table 38: Gamma GLM Estimates of Model (3) for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.0000148 | $1.15 \mathrm{E}-07$ | 16635.840 | 0.000 | 2573.66 | 0.0000 |
| Age | 0.002836 | 0.000310 | 83.723 | 0.000 | 12.03 | 0.0005 |
| Health(Excellent)* | 0.036963 | 0.006664 | 30.803 | 0.000 | 19.9 | 0.0005 |
| Health(Very good)* | 0.035908 | 0.006385 | 31.584 | 0.000 |  |  |
| Health(Fair)* | -0.040031 | 0.011829 | 11.424 | 0.001 |  |  |
| Health(Poor)* | -0.133950 | 0.019216 | 48.581 | 0.000 |  |  |
| LFS (Unemployed)* | -0.225176 | 0.016740 | 180.903 | 0.000 | 51.46 | 0.0000 |
| LFS (NLF)* | -0.109690 | 0.008520 | 165.894 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.051691 | 0.008331 | 38.440 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.044463 | 0.010693 | 17.306 | 0.000 |  |  |
| LFS (UPY-NLFPY)* | -0.122400 | 0.018935 | 41.732 | 0.000 |  |  |
| LFS (AOA)* | -0.070494 | 0.010766 | 42.903 | 0.000 |  |  |
| HLE(GHS)* | 0.072761 | 0.007238 | 101.003 | 0.000 | 45.89 | 0.0000 |
| HLE(NUPC)* | 0.092814 | 0.007176 | 167.185 | 0.000 |  |  |
| HLE(UD)* | 0.150022 | 0.008572 | 306.250 | 0.000 |  |  |
| Constant | 9.842089 | 0.014387 | 467992.810 | 0.000 |  |  |
| N | 26055 |  |  |  |  |  |
| Log likelihood | -304219.35 |  | AIC | 23.35332 |  |  |
| Deviance | 4023.77085 |  |  |  |  |  |

Table 39: Gamma GLM Estimates of Model (3) for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.0000163 | $2.46 \mathrm{E}-07$ | 4387.738 | 0.000 | 869.84 |  |
| Age | -0.009391 | 0.001912 | 24.108 | 0.000 | 4.25 | 0.0392 |
| Health(Excellent)* | 0.055337 | 0.015766 | 12.320 | 0.000 | 10.16 | 0.0377 |
| Health(Very good)* | 0.047113 | 0.013957 | 11.424 | 0.001 |  |  |
| Health(Fair)* | -0.034787 | 0.019176 | 3.276 | 0.070 |  |  |
| Health(Poor)* | -0.115625 | 0.026519 | 19.010 | 0.000 |  |  |
| Sex(Female)* | -0.033199 | 0.011178 | 8.821 | 0.003 | 1.56 | 0.2122 |
| LFS (Unemployed)* | -0.280142 | 0.033994 | 67.898 | 0.000 | 23.19 | 0.0007 |
| LFS (NLF)** | -0.113580 | 0.014798 | 58.982 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.072878 | 0.023855 | 9.364 | 0.002 |  |  |
| LFS (EPY-NLFPY)* | -0.158281 | 0.027219 | 33.872 | 0.000 |  |  |
| LFS (UPY-NLFPY)* | -0.187240 | 0.051037 | 13.469 | 0.000 |  |  |
| LFS (AOA)* | -0.041035 | 0.036422 | 1.277 | 0.260 |  |  |
| HLE(GHS)* | 0.096345 | 0.014853 | 42.120 | 0.000 | 12.02 | 0.0073 |
| HLE(NUPC)* | 0.096869 | 0.013808 | 49.280 | 0.000 |  |  |
| HLE(UD)* | 0.097920 | 0.018063 | 29.376 | 0.000 |  |  |
| Constant | 10.333920 | 0.106409 | 9430.352 | 0.000 |  |  |
| N | 6163 |  |  |  |  |  |
| Log likelihood | -71387.548 |  | AIC | 23.17201 |  |  |
| Deviance | 1170.70661 |  |  |  |  |  |

Table 40: Gamma GLM Estimates of Model (3) for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.0000103 | $2.99 \mathrm{E}-07$ | 1198.544 | 0.000 | 1070.65 | 0.0000 |
| Health(Excellent) $^{*}$ | 0.014647 | 0.041347 | 0.123 | 0.723 | 18.12 | 0.0012 |
| Health(Very good)* $^{*}$ | -0.004685 | 0.031687 | 0.023 | 0.882 |  |  |
| Health(Fair)* | -0.104917 | 0.036236 | 8.410 | 0.004 |  |  |
| Health(Poor)* | -0.085835 | 0.053718 | 2.560 | 0.110 | 116.18 | 0.0000 |
| HLE(GHS)* | 0.144466 | 0.034391 | 17.640 | 0.000 |  |  |
| HLE(NUPC)* | 0.204850 | 0.034626 | 35.046 | 0.000 |  |  |
| HLE(UD)* | 0.346910 | 0.050477 | 47.197 | 0.000 |  |  |
| Constant | 9.880105 | 0.024794 | 158794.280 | 0.000 |  |  |
| N | 9762 |  |  |  |  |  |
| Log likelihood | -109948.24 |  |  | 22.52761 |  |  |
| Deviance | 1841.2377 |  |  |  |  |  |

Table 41: Marginal Effects of Model (3)

|  |  | Income 1998 |  |  |
| ---: | ---: | ---: | ---: | ---: |
| Variables 1996 | ME below 20 | ME 20-49 | ME 50-59 | ME 60 and over |
| Income | 0.69075 | 0.64219 | 0.64416 | 0.29615 |
| Age | -1320.12700 | 122.81470 | -370.64270 |  |
| Health(Excellent)* |  | 1610.88700 | 2218.43500 | 421.88630 |
| Health(Very good)* | -1638.93600 | 1561.01200 | 1872.41800 | -134.08590 |
| Health(Good)* | -4175.87700 |  |  |  |
| Health(Fair)* | -7061.62100 | -1702.67800 | -1354.82500 | -2908.94100 |
| Health(Poor)* | -13653.62000 | -5441.24800 | -4336.22500 | -2369.40900 |
| Sex(Female)* |  | -1310.70600 |  |  |
| LFS (Unemployed)* | -8772.59900 | -9715.57100 |  |  |
| LFS (NLF)* | -4546.00200 | -4356.58300 |  |  |
| LFS (EPY-UPY)* | -2192.19300 | -2785.39700 |  |  |
| LFS (EPY-NLFPY)* | -1887.84800 | -5816.13800 |  |  |
| LFS (UPY-NLFPY)* | -4998.89300 | -6753.92200 |  |  |
| LFS (AOA)* | -2958.83100 | -1588.27100 |  |  |
| HLE(GHS)* | 3191.22300 | 3910.45000 | 4345.51100 |  |
| HLE(NUPC)* | 4079.50400 | 3905.99700 | 6290.82300 |  |
| HLE(UD)* | 6842.95300 | 4005.97200 | 11594.39000 |  |
| (*) dy/dx is for discrete change of DV from 0 to 1 |  |  |  |  |

Table 42: Gamma GLM Estimates of Model (4) for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.0000131 | $5.40 \mathrm{E}-07$ | 586.608 | 0.000 | 181.84 | 0.0000 |
| Age | -0.041076 | 0.012068 | 11.560 | 0.001 | 3.32 | 0.0686 |
| Health(Excellent)* | 0.059600 | 0.043086 | 1.904 | 0.167 | 1.41 | 0.8424 |
| Health(Very good)* | 0.057580 | 0.045571 | 1.588 | 0.206 |  |  |
| Health(Fair)* | 0.072830 | 0.101173 | 0.518 | 0.472 |  |  |
| Health(Poor)* | -0.344257 | 0.222355 | 2.403 | 0.122 |  |  |
| Constant | 10.771180 | 0.217958 | 2442.336 | 0.000 |  |  |
| N | 1621 |  |  |  |  |  |
| Log likelihood | -19063.06682 |  |  |  |  |  |
| Deviance | 710.5784858 |  |  |  |  |  |

Table 43: Gamma GLM Estimates of Model (4) for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.0000131 | $1.65 \mathrm{E}-07$ | 6272.640 | 0.000 | 1095.45 | 0.0000 |
| Age | 0.003110 | 0.000459 | 45.833 | 0.000 | 7.49 | 0.0062 |
| Health(Excellent)* | 0.032772 | 0.009846 | 11.089 | 0.001 | 13.03 | 0.0112 |
| Health(Very good)* | 0.040338 | 0.009404 | 18.404 | 0.000 |  |  |
| Health(Fair)* | -0.047715 | 0.017332 | 7.563 | 0.006 |  |  |
| Health(Poor)* | -0.155422 | 0.027673 | 31.584 | 0.000 |  |  |
| LFS (Unemployed)* | -0.221138 | 0.024474 | 81.722 | 0.000 | 31.69 | 0.0000 |
| LFS (NLF)* | -0.127213 | 0.012284 | 107.330 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.065039 | 0.012755 | 26.010 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.021775 | 0.015709 | 1.932 | 0.166 |  |  |
| LFS (UPY-NLFPY)* | -0.094671 | 0.028197 | 11.290 | 0.001 |  |  |
| LFS (AOA)* | -0.090804 | 0.016670 | 29.703 | 0.000 |  |  |
| HLE(GHS)* | 0.086296 | 0.010626 | 65.934 | 0.000 | 32.74 | 0.0000 |
| HLE(NUPC)* | 0.106590 | 0.010714 | 99.003 | 0.000 |  |  |
| HLE(UD)* | 0.176315 | 0.012638 | 194.603 | 0.000 |  |  |
| Constant | 9.939054 | 0.021081 | 222283.961 | 0.000 |  |  |
| N | 13513 |  |  |  |  |  |
| Log likelihood | -158352.9351 |  | AIC | 23.43949 |  |  |
| Deviance | 2405.309195 |  |  |  |  |  |

Table 44: Gamma GLM Estimates of Model (4) for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.0000139 | $3.72 \mathrm{E}-07$ | 1404.001 | 0.000 | 329.29 | 0.0000 |
| Age | -0.011573 | 0.002952 | 15.366 | 0.000 | 3.24 | 0.0721 |
| Health(Excellent)* | 0.058766 | 0.024498 | 5.760 | 0.016 | 4.18 | 0.3826 |
| Health(Very good)* | 0.081318 | 0.021816 | 13.913 | 0.000 |  |  |
| Health(Fair)* | 0.008423 | 0.029143 | 0.084 | 0.773 |  |  |
| Health(Poor)* | -0.034410 | 0.041204 | 0.706 | 0.404 |  |  |
| Sex(Female)* | -0.036186 | 0.017229 | 4.410 | 0.036 | 0.93 | 0.3347 |
| LFS (Unemployed)* | -0.284412 | 0.055628 | 26.112 | 0.000 | 11.85 | 0.0654 |
| LFS (NLF)** | -0.094346 | 0.022801 | 17.140 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.083574 | 0.039041 | 4.580 | 0.032 |  |  |
| LFS (EPY-NLFPY)* | -0.127179 | 0.040124 | 10.049 | 0.002 |  |  |
| LFS (UPY-NLFPY)* | -0.356398 | 0.080537 | 19.625 | 0.000 |  |  |
| LFS (AOA)* | -0.038938 | 0.059709 | 0.423 | 0.514 |  |  |
| HLE(GHS)* | 0.097989 | 0.022593 | 18.836 | 0.000 | 8.18 | 0.0424 |
| HLE(NUPC)* | 0.107504 | 0.021714 | 24.503 | 0.000 |  |  |
| HLE(UD)* | 0.141741 | 0.028084 | 25.503 | 0.000 |  |  |
| Constant | 10.532300 | 0.163895 | 4129.348 | 0.000 |  |  |
| N | 3068 |  |  |  |  |  |
| Log likelihood | -35576.84909 |  | AIC | 23.20329 |  |  |
| Deviance | 724.6516749 |  |  |  |  |  |

Table 45: Gamma GLM Estimates of Model (4) for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.0000209 | $3.46 \mathrm{E}-07$ | 3634.884 | 0.000 | 592.46 | 0.0000 |
| Age | -0.007200 | 0.000793 | 82.446 | 0.000 | 11.62 | 0.0007 |
| Health(Excellent)* | 0.013106 | 0.018120 | 0.518 | 0.470 | 2.28 | 0.6840 |
| Health(Very good)* | 0.007913 | 0.014222 | 0.314 | 0.578 |  |  |
| Health(Fair)* | -0.029480 | 0.016425 | 3.204 | 0.073 |  |  |
| Health(Poor)* | -0.079313 | 0.025665 | 9.548 | 0.002 |  |  |
| Sex(Female)* | -0.058507 | 0.011291 | 26.832 | 0.000 | 3.83 | 0.0502 |
| HLE(GHS)* | 0.074600 | 0.015609 | 22.848 | 0.000 | 8.31 | 0.0399 |
| HLE(NUPC)* | 0.075425 | 0.015593 | 23.426 | 0.000 |  |  |
| HLE(UD)* | 0.138112 | 0.023540 | 34.457 | 0.000 |  |  |
| Constant | 10.120650 | 0.058737 | 29687.290 | 0.000 |  |  |
| N | 4617 |  | AIC | 22.52001 |  |  |
| Log likelihood | -51976.4471 |  |  |  |  |  |
| Deviance | 562.8120285 |  |  |  |  |  |

Table 46: Marginal Effects of Model (4)

|  | Income 1999 |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | ME below 20 | ME 20-49 | ME 50-59 | ME 60 and over |
| Income | 0.61623 | 0.59160 | 0.55722 | 0.59436 |
| Age | -1934.75600 | 140.52250 | -462.68970 | -205.18930 |
| Health(Excellent)* | 2800.84300 | 1489.41400 | 2388.34300 | 375.25380 |
| Health(Very good)* | 2741.04000 | 1830.83100 | 3291.76800 | 225.85400 |
| Health(Fair)* | 3553.15100 | -2110.49600 | 337.81190 | -832.37740 |
| Health(Poor)* | -13736.03000 | -6522.54200 | -1354.71000 | -2182.64600 |
| Sex(Female)* |  |  | -1447.03200 | -1672.82500 |
| LFS (Unemployed)* |  | -9007.30200 | -9963.68900 |  |
| LFS (NLF)* |  | -5467.97300 | -3683.57300 |  |
| LFS (EPY-UPY)* |  | -2861.28600 | -3218.69400 |  |
| LFS (EPY-NLFPY)* |  | -974.48240 | -4802.82500 |  |
| LFS (UPY-NLFPY)* |  | -4087.79900 | -12033.33000 |  |
| LFS (AOA)* |  | -3939.61300 | -1527.96600 |  |
| HLE(GHS)** |  | 3953.77600 | 4024.32400 | 2179.75500 |
| HLE(NUPC)* |  | 4909.95400 | 4408.88400 | 2204.27400 |
| HLE(UD)* |  | 8464.76400 | 5969.79700 | 4181.84200 |
| (*) | /dx is for dis | te change of | V from 0 to 1 |  |

Table 47: Gamma GLM Estimates of Model (3) with Stress

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000012 | $1.66 \mathrm{E}-07$ | 5286.744 | 0.000 | 4252.43 | 0.0000 |
| Age | -0.002267 | 0.000421 | 28.944 | 0.000 | 33.78 | 0.0000 |
| Stress (Somewhat stressful)* | 0.049543 | 0.016375 | 9.181 | 0.002 | 12.35 | 0.0063 |
| Stress (Not very stressful)* | 0.053862 | 0.017859 | 9.120 | 0.003 |  |  |
| Stress (Not at all stressful)* | 0.040298 | 0.022984 | 3.063 | 0.080 |  |  |
| Sex(Female)* | -0.031192 | 0.011087 | 7.896 | 0.005 | 9.21 | 0.0024 |
| LFS (Unemployed)* | -0.285977 | 0.040278 | 50.410 | 0.000 | 131.33 | 0.0000 |
| LFS (NLF)* | -0.121304 | 0.015954 | 57.760 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.094444 | 0.021431 | 19.448 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.064885 | 0.023014 | 7.952 | 0.005 |  |  |
| LFS (UPY-NLFPY)* | -0.139102 | 0.040079 | 12.041 | 0.001 |  |  |
| LFS (AOA)* | -0.106538 | 0.025873 | 16.974 | 0.000 |  |  |
| HLE(GHS)* | 0.079024 | 0.014810 | 28.516 | 0.000 | 141.20 | 0.0000 |
| HLE(NUPC)* | 0.121700 | 0.015027 | 65.610 | 0.000 |  |  |
| HLE(UD)* | 0.193189 | 0.018824 | 105.268 | 0.000 |  |  |
| Constant | 10.139490 | 0.027476 | 136183.141 | 0.000 |  |  |
| N | 39489 |  |  |  |  |  |
| Log likelihood | -459552.1809 |  | AIC | 23.27576 |  |  |
| Deviance | 7357.206914 |  |  |  |  |  |

Table 48: Marginal Effects of Model (3) with Stress

|  | Income 1998 |
| :---: | :---: |
| Variables 1996 | Marginal Effect |
| Income | 0.50220 |
| Age | -94.44846 |
| Stress (Somewhat stressful)* | 2066.46700 |
| Stress (Not very stressful)* | 2271.55100 |
| Stress (Not at all stressful)* | 1706.63600 |
| Sex(Female)* | -1300.46800 |
| LFS (Unemployed)* | -10420.10000 |
| LFS (NLF)* | -4879.99800 |
| LFS (EPY-UPY)* | -3781.36500 |
| LFS (EPY-NLFPY)* | -2628.81700 |
| LFS (UPY-NLFPY)* | -5425.65200 |
| LFS (AOA)* | -4233.32000 |
| HLE(GHS)* | 3347.46000 |
| HLE(NUPC)* | 5203.86100 |
| HLE(UD)* | 8649.88400 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete chan | DV from 0 to 1 |

Table 49: Gamma GLM Estimates of Model (4) with Stress

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | $9.99 \mathrm{E}-06$ | $1.85 \mathrm{E}-07$ | 2909.524 | 0.000 | 1697.80 | 0.0000 |
| Age | -0.002369 | 0.000490 | 23.329 | 0.000 | 19.13 | 0.0000 |
| Stress (Somewhat stressful)* | 0.043085 | 0.018916 | 5.198 | 0.023 | 8.07 | 0.0445 |
| Stress (Not very stressful)* | 0.060102 | 0.020609 | 8.526 | 0.004 |  |  |
| Stress (Not at all stressful)* | 0.018050 | 0.026429 | 0.462 | 0.495 |  |  |
| Sex(Female)* | -0.034293 | 0.012894 | 7.076 | 0.008 | 5.79 | 0.0161 |
| LFS (Unemployed)* | -0.278972 | 0.046932 | 35.284 | 0.000 | 75.24 | 0.0000 |
| LFS (NLF)* | -0.126759 | 0.018289 | 48.025 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.106815 | 0.025843 | 17.057 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.068153 | 0.026581 | 6.554 | 0.010 |  |  |
| LFS (UPY-NLFPY)* | -0.125459 | 0.047497 | 6.970 | 0.008 |  |  |
| LFS (AOA)* | -0.139477 | 0.031316 | 19.803 | 0.000 |  |  |
| HLE(GHS)* | 0.097292 | 0.017079 | 32.490 | 0.000 | 107.97 | 0.0000 |
| HLE(NUPC)* | 0.145129 | 0.017680 | 67.404 | 0.000 |  |  |
| HLE(UD)* | 0.235342 | 0.021815 | 116.424 | 0.000 |  |  |
| Constant | 10.257700 | 0.031774 | 104225.666 | 0.000 |  |  |
| N | 20507 |  |  |  |  |  |
| Log likelihood | -239361.6572 |  | AIC | 23.34595 |  |  |
| Deviance | 4573.606662 |  |  |  |  |  |

Table 50: Marginal Effects of Model (4) with Stress

|  | Income 1999 |
| :---: | :---: |
| Variables 1996 | Marginal Effects |
| Income | 0.43102 |
| Age | -102.21280 |
| Stress (Somewhat stressful)* | 1860.82300 |
| Stress (Not very stressful)* | 2627.94100 |
| Stress (Not at all stressful)* | 784.31960 |
| Sex(Female)* | -1480.30600 |
| LFS (Unemployed)* | -10557.52000 |
| LFS (NLF)* | -5275.31500 |
| LFS (EPY-UPY)* | -4402.10600 |
| LFS (EPY-NLFPY)* | -2854.97000 |
| LFS (UPY-NLFPY)* | -5098.46600 |
| LFS (AOA)* | -5651.98500 |
| HLE(GHS)* | 4278.07200 |
| HLE(NUPC)* | 6472.49600 |
| HLE(UD)* | 11082.78000 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change | DV from 0 to 1 |

Table 51: Gamma GLM Estimates of Model (3) with Stress for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.000015 | $3.44 \mathrm{E}-07$ | 1935.120 | 0.000 | 455.76 | 0.0000 |
| Age | -0.025338 | 0.007561 | 11.223 | 0.001 | 2.45 | 0.1175 |
| Stress (Somewhat stressful)* | 0.072270 | 0.035913 | 4.040 | 0.044 | 1.46 |  |
| Stress (Not very stressful)* | 0.071998 | 0.036276 | 3.920 | 0.047 |  |  |
| Stress (Not at all stressful)* | 0.102436 | 0.039421 | 6.760 | 0.009 |  |  |
| Sex(Female)* | -0.029629 | 0.016791 | 3.098 | 0.078 | 0.68 | 0.4097 |
| Constant | 10.382520 | 0.139675 | 5524.949 | 0.000 |  |  |
| N | 3171 |  | AIC | 23.50128 |  |  |
| Log likelihood | -37254.27609 |  |  |  |  |  |
| Deviance | 1067.230569 |  |  |  |  |  |

Table 52: Gamma GLM Estimates of Model (3) with Stress for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000015 | $1.15 \mathrm{E}-07$ | 16731.423 | 0.000 | 2579.68 | 0.0000 |
| Age | 0.002373 | 0.000310 | 58.676 | 0.000 | 8.41 | 0.0037 |
| Stress (Somewhat stressful)* | 0.038116 | 0.006742 | 31.923 | 0.000 | 4.94 | 0.1759 |
| Stress (Not very stressful)* | 0.026624 | 0.007641 | 12.110 | 0.000 |  |  |
| Stress (Not at all stressful)* | 0.013128 | 0.011269 | 1.346 | 0.244 |  |  |
| LFS (Unemployed)* | -0.232180 | 0.016918 | 188.238 | 0.000 | 58.36 | 0.0000 |
| LFS (NLF)* | -0.124237 | 0.008462 | 215.502 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.051300 | 0.008397 | 37.332 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.044583 | 0.010776 | 17.140 | 0.000 |  |  |
| LFS (UPY-NLFPY)* | -0.126836 | 0.019078 | 44.223 | 0.000 |  |  |
| LFS (AOA)* | -0.071875 | 0.010840 | 43.957 | 0.000 |  |  |
| HLE(GHS)* | 0.075295 | 0.007302 | 106.296 | 0.000 | 48.53 | 0.0000 |
| HLE(NUPC)* | 0.095753 | 0.007246 | 174.504 | 0.000 |  |  |
| HLE(UD)* | 0.155936 | 0.008631 | 326.525 | 0.000 |  |  |
| Constant | 9.848422 | 0.014944 | 434281.000 | 0.000 |  |  |
| N | 25647 |  |  |  |  |  |
| Log likelihood | -299508.3799 |  | AIC | 23.35738 |  |  |
| Deviance | 3942.415405 |  |  |  |  |  |

Table 53: Gamma GLM Estimates of Model (3) with Stress for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000016 | $2.47 \mathrm{E}-07$ | 4385.088 | 0.000 | 862.05 | 0.0000 |
| Age | -0.009711 | 0.001927 | 25.402 | 0.000 | 4.43 | 0.0353 |
| Stress (Somewhat stressful)* | 0.055352 | 0.015916 | 12.110 | 0.001 | 2.77 | 0.4288 |
| Stress (Not very stressful)* | 0.046968 | 0.017084 | 7.563 | 0.006 |  |  |
| Stress (Not at all stressful)* | 0.076515 | 0.021586 | 12.532 | 0.000 |  |  |
| Sex(Female)* | -0.031804 | 0.011219 | 8.009 | 0.005 | 1.40 | 0.2359 |
| LFS (Unemployed)* | -0.289862 | 0.034598 | 70.224 | 0.000 | 32.32 | 0.0000 |
| LFS (NLF)* | -0.149191 | 0.014312 | 108.576 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.073035 | 0.024113 | 9.181 | 0.002 |  |  |
| LFS (EPY-NLFPY)* | -0.176672 | 0.027554 | 41.088 | 0.000 |  |  |
| LFS (UPY-NLFPY)* | -0.244464 | 0.051140 | 22.848 | 0.000 |  |  |
| LFS (AOA)* | -0.054707 | 0.036792 | 2.220 | 0.137 |  |  |
| HLE(GHS)* | 0.107832 | 0.014904 | 52.418 | 0.000 | 15.09 | 0.0017 |
| HLE(NUPC)* | 0.107467 | 0.013888 | 59.908 | 0.000 |  |  |
| HLE(UD)* | 0.116776 | 0.018117 | 41.603 | 0.000 |  |  |
| Constant | 10.324250 | 0.106809 | 9343.156 | 0.000 |  |  |
| N | 6019 |  |  |  |  |  |
| Log likelihood | -69750.21706 |  | AIC | 23.18200 |  |  |
| Deviance | 1140.039735 |  |  |  |  |  |

Table 54: Gamma GLM Estimates of Model (3) with Stress for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000010 | $2.79 \mathrm{E}-07$ | 1303.210 | 0.000 | 808.60 | 0.0000 |
| Stress (Somewhat stressful)* | -0.069394 | 0.044652 | 2.403 | 0.120 | 3.39 | 0.3349 |
| Stress (Not very stressful)* | -0.019769 | 0.027310 | 0.518 | 0.469 |  |  |
| Stress (Not at all stressful)* | -0.033108 | 0.027256 | 1.464 | 0.224 |  |  |
| Sex(Female)* | -0.138116 | 0.021993 | 39.438 | 0.000 | 42.91 | 0.0000 |
| HLE(GHS)* | 0.142435 | 0.029958 | 22.563 | 0.000 | 105.99 | 0.0000 |
| HLE(NUPC)* | 0.205671 | 0.030081 | 46.786 | 0.000 |  |  |
| HLE(UD)* | 0.324038 | 0.044031 | 54.170 | 0.000 |  |  |
| Constant | 9.964811 | 0.024031 | 171942.916 | 0.000 |  |  |
| N | 9197 |  |  |  |  |  |
| Log likelihood | -103553.9805 |  | AIC | 22.52104 |  |  |
| Deviance | 1479.383695 |  |  |  |  |  |

Table 55: Marginal Effects of Model (3) with Stress

|  | Income 1998 |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | ME below 20 | ME 20-49 | ME 50-59 | ME 60 and over |
| Income | 0.70439 | 0.64753 | 0.64850 | 0.28766 |
| Age | -1179.66800 | 102.98060 | -385.21020 |  |
| Stress (Somewhat stressful)* | 3390.29700 | 1652.14800 | 2202.26300 | -1923.53700 |
| Stress (Not very stressful)* | 3384.19800 | 1163.25500 | 1882.15400 | -561.85640 |
| Stress (Not at all stressful)* | 4935.38400 | 572.98560 | 3128.54200 | -938.07700 |
| Sex(Female)* | -1378.72700 |  | -1262.05100 | -3983.44700 |
| LFS (Unemployed)* |  | -9034.51500 | -10057.94000 |  |
| LFS (NLF)* |  | -5129.28200 | -5700.07800 |  |
| LFS (EPY-UPY)* |  | -2180.41100 | -2805.32200 |  |
| LFS (EPY-NLFPY)* |  | -1896.75100 | -6471.43600 |  |
| LFS (UPY-NLFPY)* |  | -5179.82100 | -8627.50000 |  |
| LFS (AOA)* |  | -3021.18900 | -2114.50800 |  |
| HLE(GHS)* |  | 3310.71300 | 4413.54900 | 4266.53700 |
| HLE(NUPC)* |  | 4218.86100 | 4365.40000 | 6294.96000 |
| HLE(UD)* |  | 7140.82100 | 4834.71100 | 10685.33000 |
| (*) dy | x is for discret | hange of DV | om 0 to 1 |  |

Table 56: Gamma GLM Estimates of Model (4) with Stress for the Persons of Age Under 20

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.000013 | $5.40 \mathrm{E}-07$ | 590.976 | 0.000 | 183.25 | 0.0000 |
| Age | -0.036722 | 0.012288 | 8.940 | 0.003 | 0.56 | 0.1099 |
| Stress (Somewhat stressful)* | 0.158384 | 0.059804 | 7.023 | 0.008 | 1.99 | 0.5751 |
| Stress (Not very stressful)* | 0.137199 | 0.060118 | 5.198 | 0.022 |  |  |
| Stress (Not at all stressful)* | 0.153621 | 0.065032 | 5.570 | 0.018 |  |  |
| Sex(Female)* $^{*}$ | -0.043978 | 0.027042 | 2.657 | 0.104 | 0.76 | 0.3839 |
| Constant | 10.627520 | 0.225725 | 2216.526 | 0.000 |  |  |
| N | 1602 |  |  |  |  |  |
| Log likelihood | -18841.15457 |  | 23.53078 |  |  |  |
| Deviance | 699.8370331 |  |  |  |  |  |

Table 57: Gamma GLM Estimates of Model (4) with Stress for the People between 20 and 49 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000013 | $1.66 \mathrm{E}-07$ | 6296.423 | 0.000 | 1108.07 | 0.0000 |
| Age | 0.002661 | 0.000458 | 33.756 | 0.000 | 5.54 | 0.0186 |
| Stress (Somewhat stressful)* | 0.024350 | 0.009916 | 6.052 | 0.014 | 1.64 | 0.6515 |
| Stress (Not very stressful)* | 0.020386 | 0.011209 | 3.312 | 0.069 |  |  |
| Stress (Not at all stressful)* | -0.009388 | 0.016457 | 0.325 | 0.568 |  |  |
| LFS (Unemployed)* | -0.228557 | 0.024620 | 86.118 | 0.000 | 37.60 | 0.0000 |
| LFS (NLF)* | -0.144442 | 0.012141 | 141.610 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.063930 | 0.012859 | 24.701 | 0.000 |  |  |
| LFS (EPY-NLFPY)* | -0.022938 | 0.015772 | 2.103 | 0.146 |  |  |
| LFS (UPY-NLFPY)* | -0.101338 | 0.028444 | 12.674 | 0.000 |  |  |
| LFS (AOA)* | -0.095977 | 0.016775 | 32.718 | 0.000 |  |  |
| HLE(GHS)* | 0.092214 | 0.010660 | 74.823 | 0.000 | 35.61 | 0.0000 |
| HLE(NUPC)* | 0.110730 | 0.010776 | 105.678 | 0.000 |  |  |
| HLE(UD)* | 0.184816 | 0.012688 | 212.285 | 0.000 |  |  |
| Constant | 9.954358 | 0.021904 | 206533.892 | 0.000 |  |  |
| N | 13434 |  |  |  |  |  |
| Log likelihood | -157438.099 |  | AIC | 23.44099 |  |  |
| Deviance | 2403.060425 |  |  |  |  |  |

Table 58: Gamma GLM Estimates of Model (4) with Stress for the People between 50 and 59 Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Income | 0.000014 | $3.74 \mathrm{E}-07$ | 1416.017 | 0.000 | 336.17 | 0.0000 |
| Age | -0.011562 | 0.002976 | 15.132 | 0.000 | 3.21 | 0.0730 |
| Stress (Somewhat stressful)* | 0.047155 | 0.024650 | 3.648 | 0.056 | 1.35 | 0.7172 |
| Stress (Not very stressful)* | 0.062604 | 0.026434 | 5.617 | 0.018 |  |  |
| Stress (Not at all stressful)* | 0.065194 | 0.032881 | 3.920 | 0.047 |  |  |
| Sex(Female)* | -0.036459 | 0.017333 | 4.410 | 0.035 | 0.94 | 0.3312 |
| LFS (Unemployed)* | -0.300014 | 0.055878 | 28.837 | 0.000 | 15.09 | 0.0195 |
| LFS (NLF)* | -0.117350 | 0.022225 | 27.878 | 0.000 |  |  |
| LFS (EPY-UPY)* | -0.082863 | 0.039380 | 4.410 | 0.035 |  |  |
| LFS (EPY-NLFPY)* | -0.141859 | 0.040693 | 12.180 | 0.000 |  |  |
| LFS (UPY-NLFPY)* | -0.396000 | 0.080563 | 24.206 | 0.000 |  |  |
| LFS (AOA)* | -0.054986 | 0.060478 | 0.828 | 0.363 |  |  |
| HLE(GHS)* | 0.105838 | 0.022729 | 21.716 | 0.000 | 9.79 | 0.0204 |
| HLE(NUPC)* | 0.116634 | 0.021874 | 28.409 | 0.000 |  |  |
| HLE(UD)* | 0.156556 | 0.028234 | 30.692 | 0.000 |  |  |
| Constant | 10.521320 | 0.164694 | 4080.654 | 0.000 |  |  |
| N | 3050 |  |  |  |  |  |
| Log likelihood | -35368.61346 |  | AIC | 23.20303 |  |  |
| Deviance | 725.2332615 |  |  |  |  |  |

Table 59: Gamma GLM Estimates of Model (4) with Stress for the People 60 and Over Years of Age

| Variables 1996 | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Income | 0.000018 | $4.49 \mathrm{E}-07$ | $1,631.352$ | 0.000 | 278.92 | 0.0000 |
| Stress (Somewhat stressful)* | 0.054629 | 0.030742 | 3.168 | 0.076 | 1.07 | 0.7837 |
| Stress (Not very stressful)* | 0.078337 | 0.030002 | 6.812 | 0.009 |  |  |
| Stress (Not at all stressful)* | 0.062699 | 0.031664 | 3.920 | 0.048 |  |  |
| HLE(GHS)* | 0.067637 | 0.021882 | 9.548 | 0.002 | 6.26 | 0.0997 |
| HLE(NUPC)* | 0.090400 | 0.021238 | 18.148 | 0.000 |  |  |
| HLE(UD)* | 0.163446 | 0.031414 | 27.040 | 0.000 |  |  |
| Constant | 9.647321 | 0.030453 | $100,362.240$ | 0.000 |  |  |
| N | 2416 |  |  |  |  |  |
| Log likelihood | -27425.56212 |  |  | 22.70990 |  |  |
| Deviance | 344.2145216 |  |  |  |  |  |

Table 60: Marginal Effects of Model (4) with Stress

|  | Income 1999 |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Variables 1996 | ME below 20 | ME 20-49 | ME 50-59 | ME 60 and over |
| Income | 0.61897 | 0.59487 | 0.56225 | 0.56776 |
| Age | -1731.31200 | 120.32890 | -462.30470 |  |
| Stress (Somewhat stressful)* | 7602.55800 | 1100.58900 | 1890.59400 | 1729.54700 |
| Stress (Not very stressful)* | 6580.34200 | 926.72970 | 2538.02100 | 2476.73000 |
| Stress (Not at all stressful)* | 7628.55400 | -422.83410 | 2673.61500 | 1996.05500 |
| Sex(Female)* | -2074.02100 |  | -1458.18500 |  |
| LFS (Unemployed)* |  | -9285.28300 | -10438.56000 |  |
| LFS (NLF)* |  | -6171.69100 | -4556.68800 |  |
| LFS (EPY-UPY)* |  | -2816.00800 | -3192.80000 |  |
| LFS (EPY-NLFPY)* |  | -1026.84800 | -5323.09900 |  |
| LFS (UPY-NLFPY)* |  | -4365.29100 | -13132.78000 |  |
| LFS (AOA)* |  | -4157.86600 | -2141.64100 |  |
| HLE(GHS)* |  | 4232.62800 | 4356.43400 | 2164.68900 |
| HLE(NUPC)** |  | 5108.44000 | 4794.54700 | 2911.16400 |
| HLE(UD)* |  | 8906.11100 | 6630.98300 | 5490.38800 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of DV from 0 to 1 |  |  |  |  |

Table 61: Ordered Probit Estimates of Model (5) for the Persons of Age Under 20

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHealth 97 (Imp. a bit)* | -1.225048 | 0.308791 | 15.761 | 0.000 | 711.7 | 0.0000 |
| CHealth 97 (Unchanged)* | -2.314680 | 0.308982 | 56.100 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -2.717033 | 0.311211 | 76.213 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* | -3.738613 | 0.441729 | 71.572 | 0.000 |  |  |
| Change in Income 97 | $6.05 \mathrm{E}-07$ | $9.75 \mathrm{E}-07$ | 0.384 | 0.535 | 0.38 | 0.5350 |
| Sex(Female)* | -0.067384 | 0.039462 | 2.924 | 0.088 | 2.92 | 0.0877 |
| HLE 98(GHS)* | 0.131597 | 0.049486 | 7.076 | 0.008 | 7.75 | 0.0515 |
| HLE 98(NUPC)* | 0.069398 | 0.070459 | 0.960 | 0.325 |  |  |
| HLE 98(UD)* | -0.053622 | 0.243659 | 0.048 | 0.826 |  |  |
| cut1 | -4.792913 | 0.320553 |  |  |  |  |
| cut2 | -2.837408 | 0.311131 |  |  |  |  |
| cut3 | -1.162037 | 0.309790 |  |  |  |  |
| cut4 | 1.273019 | 0.309904 |  |  |  |  |
| N | 3340 |  |  |  |  |  |
| Log likelihood | -3163.8097 |  |  |  |  |  |
| LR chi2(9) | 716.8 |  |  |  |  |  |

Table 62: Ordered Probit Estimates of Model (5) for the People between 20 and 49 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHealth 97 (Imp. a bit)* | -0.864221 | 0.097660 | 78.323 | 0.000 | 5433.04 | 0.0000 |
| CHealth 97 (Unchanged)* | -1.712595 | 0.097584 | 308.003 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -2.299559 | 0.098337 | 546.624 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* | -3.026579 | 0.139089 | 473.498 | 0.000 |  |  |
| Change in Income 97 | $2.14 \mathrm{E}-07$ | $4.02 \mathrm{E}-07$ | 0.281 | 0.594 | 0.28 | $0.5945$ |
| HLE 98(GHS)* | -0.041031 | 0.021587 | 3.610 | 0.057 | 7.49 | $0.0579$ |
| HLE 98(NUPC)* | -0.045589 | 0.020789 | 4.796 | 0.028 |  |  |
| HLE 98(UD)* | -0.062824 | 0.023819 | 6.970 | 0.008 |  |  |
| Age 98 | -0.001689 | 0.000844 | 4.000 | 0.045 | 4.00 | 0.0455 |
| cut1 | -4.479364 | 0.107942 |  |  |  |  |
| cut2 | -2.474969 | 0.104108 |  |  |  |  |
| cut3 | -0.926070 | 0.103695 |  |  |  |  |
| cut4 | 1.256549 | 0.104805 |  |  |  |  |
| N | 26588 |  |  |  |  |  |
| Log likelihood | -26239.949 |  |  |  |  |  |
| LR chi2(9) | 5442.28 |  |  |  |  |  |

Table 63: Ordered Probit Estimates of Model (5) for the People between 50 and 59 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHealth 97 (Imp. a bit)* | -0.734144 | 0.165848 | 19.625 | 0.000 | 1212.32 | 0.0000 |
| CHealth 97 (Unchanged)* | -1.473385 | 0.165437 | 79.388 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -2.097838 | 0.167297 | 157.252 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* | -2.819251 | 0.231273 | 148.596 | 0.000 |  |  |
| Change in Income 97 | -7.30E-07 | $6.95 \mathrm{E}-07$ | 1.103 | 0.293 | 1.11 | 0.2927 |
| LFS 98(Unemployed)* | -0.147096 | 0.094568 | 2.434 | 0.120 | 21.24 | 0.0017 |
| LFS 98 (NLF)* | 0.063663 | 0.031828 | 4.000 | 0.045 |  |  |
| LFS 98(EPY-UPY)* | 0.095842 | 0.065876 | 2.103 | 0.146 |  |  |
| LFS 98 (EPY-NLFPY)* | -0.157424 | 0.063337 | 6.200 | 0.013 |  |  |
| LFS 98 (UPY-NLFPY)* | -0.245613 | 0.130522 | 3.534 | 0.060 |  |  |
| LFS 98 (AOA)* | 0.041809 | 0.090582 | 0.212 | 0.644 |  |  |
| cut1 | -4.079504 | 0.174761 |  |  |  |  |
| cut2 | -2.154853 | 0.166086 |  |  |  |  |
| cut3 | -0.680905 | 0.165294 |  |  |  |  |
| cut4 | 1.349389 | 0.167578 |  |  |  |  |
| N | 6289 |  |  |  |  |  |
| Log likelihood | -6416.6854 |  |  |  |  |  |
| LR chi2(11) | 1226.06 |  |  |  |  |  |

Table 64: Ordered Logit Estimates of Model (5) for the People between 50 and 59 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| CHealth 97 (Imp. a bit)* | -1.370453 | 0.370010 | 13.690 | 0.000 | 1229.13 | 0.0000 |
| CHealth 97 (Unchanged)* | -2.697463 | 0.369303 | 53.290 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -3.799320 | 0.372178 | 104.244 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* $^{*}$ | -5.172960 | 0.498175 | 107.744 | 0.000 |  |  |
| Change in Income 97 | $-1.67 \mathrm{E}-06$ | $1.23 \mathrm{E}-06$ | 1.823 | 0.176 | 1.83 | 0.1758 |
| LFS 98(Unemployed)* | -0.288037 | 0.171104 | 2.822 | 0.092 | 18.54 | 0.0050 |
| LFS 98 (NLF)* | 0.103645 | 0.055180 | 3.534 | 0.060 |  |  |
| LFS 98(EPY-UPY)* | 0.152280 | 0.114242 | 1.769 | 0.183 |  |  |
| LFS 98 (EPY-NLFPY)* | -0.265048 | 0.111698 | 5.617 | 0.018 |  |  |
| LFS 98 (UPY-NLFPY)* | -0.330034 | 0.229649 | 2.074 | 0.151 |  |  |
| LFS 98 (AOA)* | 0.094439 | 0.162017 | 0.336 | 0.560 |  |  |
| cut1 | -7.718634 | 0.394753 |  |  |  |  |
| cut2 | -3.845195 | 0.370461 |  |  |  |  |
| cut3 | -1.354162 | 0.368569 |  |  |  |  |
| cut4 | 2.748791 | 0.383672 |  |  |  |  |
| N | 6289 |  |  |  |  |  |

Table 65: Ordered Probit Estimates of Model (5) for the People 60 and Over Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHealth 97 (Imp. a bit)* | -0.676553 | 0.112454 | 36.240 | 0.000 | 1964.62 | 0.0000 |
| CHealth 97 (Unchanged)* | -1.343047 | 0.112239 | 143.281 | 0.000 |  |  |
| CHealth 97 (Worsened)* | -1.984540 | 0.113660 | 304.852 | 0.000 |  |  |
| CHealth 97 (Wors. badly)* | -2.981246 | 0.173472 | 295.496 | 0.000 |  |  |
| Change in Income 97 | $2.04 \mathrm{E}-07$ | $9.30 \mathrm{E}-07$ | 0.048 | 0.826 | 0.05 | 0.8261 |
| HLE 98(GHS)* | -0.077508 | 0.030887 | 6.300 | 0.012 | 9.98 | 0.0188 |
| HLE 98(NUPC)* | -0.029781 | 0.030921 | 0.922 | 0.335 |  |  |
| HLE 98(UD)* | -0.102262 | 0.044750 | 5.244 | 0.022 |  |  |
| Age 98 | -0.004618 | 0.001542 | 9.000 | 0.003 | 8.97 | 0.0027 |
| cut1 | -4.236611 | 0.165360 |  |  |  |  |
| cut2 | -2.295938 | 0.160020 |  |  |  |  |
| cut3 | -0.964397 | 0.159257 |  |  |  |  |
| cut4 | 1.043342 | 0.161455 |  |  |  |  |
| N | 9762 |  |  |  |  |  |
| Log likelihood | -10288.995 |  |  |  |  |  |
| LR chi2(9) | 1972.46 |  |  |  |  |  |

Table 66: Marginal Effects of Model (5) for the Persons of Age Under 20

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.038543 | 0.388799 | -0.201209 | -0.225141 | -0.000992 |
| CHealth 97 (Unchanged)* | 0.045677 | 0.539472 | 0.033807 | -0.599419 | -0.019537 |
| CHealth 97 (Worsened)* | 0.288775 | 0.536553 | -0.466370 | -0.356350 | -0.002607 |
| CHealth 97 (Wors. badly)* | 0.842729 | -0.066725 | -0.595042 | -0.180554 | -0.000409 |
| Change in Income 97 | -6.22E-09 | -1.76E-07 | $2.51 \mathrm{E}-08$ | $1.56 \mathrm{E}-07$ | $8.49 \mathrm{E}-10$ |
| Sex(Female)* | 0.000697 | 0.019626 | -0.002837 | -0.017390 | -0.000094 |
| HLE 98(GHS)* | -0.001440 | -0.038827 | 0.006755 | 0.033339 | 0.000173 |
| HLE 98(NUPC)* | -0.000666 | -0.019833 | 0.002037 | 0.018355 | 0.000107 |
| HLE 98(UD)* | 0.000593 | 0.015882 | -0.002887 | -0.013519 | -0.000069 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 67: Marginal Effects of Model (5) for the People between 20 and 49 Years of Age

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.016164 | 0.276493 | -0.086894 | -0.203075 | -0.002688 |
| CHealth 97 (Unchanged)* | 0.028987 | 0.460539 | -0.016936 | -0.456056 | -0.016534 |
| CHealth 97 (Worsened)* | 0.162695 | 0.583275 | -0.358368 | -0.380560 | -0.007043 |
| CHealth 97 (Wors. badly)* | 0.601211 | 0.162989 | -0.549688 | -0.213039 | -0.001473 |
| Change in Income 97 | -1.96E-09 | -6.28E-08 | $3.06 \mathrm{E}-09$ | $6.07 \mathrm{E}-08$ | $9.81 \mathrm{E}-10$ |
| HLE 98(GHS)* | 0.000385 | 0.012111 | -0.000744 | -0.011568 | -0.000184 |
| HLE 98(NUPC)* | 0.000425 | 0.013436 | -0.000780 | -0.012876 | -0.000206 |
| HLE 98(UD)* | 0.000608 | 0.018689 | -0.001478 | -0.017548 | -0.000272 |
| Age 98 | 0.000016 | 0.000496 | -0.000024 | -0.000479 | -7.75E-06 |
| ${ }^{*}$ ) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 68: Marginal Effects of Ordered Probit Model (5) for the People between 50 and 59 Years of Age

|  | Change in Health 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.014870 | 0.231466 | -0.053909 | -0.187889 | -0.004538 |
| CHealth 97 (Unchanged)* | 0.026851 | 0.403058 | -0.000614 | -0.408433 | -0.020862 |
| CHealth 97 (Worsened)* | 0.146172 | 0.552366 | -0.307176 | -0.380113 | -0.011249 |
| CHealth 97 (Wors. badly)* | 0.549184 | 0.207925 | -0.519105 | -0.234954 | -0.003050 |
| Change in Income 97 | $8.37 \mathrm{E}-09$ | $2.13 \mathrm{E}-07$ | $9.62 \mathrm{E}-10$ | -2.16E-07 | -6.38E-09 |
| LFS 98(Unemployed)* | 0.002038 | 0.044833 | -0.004417 | -0.041391 | -0.001063 |
| LFS 98 (NLF)* | -0.000707 | -0.018432 | -0.000434 | 0.018996 | 0.000576 |
| LFS 98(EPY-UPY)* | -0.000981 | -0.027155 | -0.001977 | 0.029166 | 0.000945 |
| LFS 98 (EPY-NLFPY)* | 0.002183 | 0.047977 | -0.004729 | -0.044293 | -0.001139 |
| LFS 98 (UPY-NLFPY)* | 0.003900 | 0.076827 | -0.012726 | -0.066442 | -0.001558 |
| LFS 98 (AOA)* | -0.000455 | -0.012043 | -0.000427 | 0.012538 | 0.000386 |
| (*) $^{*} \mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 69: Marginal Effects of Ordered Logit Model (5) for the People between 50 and 59 Years of Age

|  | Change in Health 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.011984 | 0.262066 | -0.076579 | -0.192419 | -0.005053 |
| CHealth 97 (Unchanged)* | 0.020933 | 0.434940 | 0.000796 | -0.439876 | -0.016793 |
| CHealth 97 (Worsened)* | 0.096975 | 0.638321 | -0.340098 | -0.384077 | -0.011120 |
| CHealth 97 (Wors. badly)* | 0.495352 | 0.269439 | -0.534235 | -0.225610 | -0.004946 |
| Change in Income 97 | $9.76 \mathrm{E}-09$ | $2.78 \mathrm{E}-07$ | $3.09 \mathrm{E}-09$ | -2.83E-07 | -7.93E-09 |
| LFS 98(Unemployed)* | 0.001938 | 0.051531 | -0.007019 | -0.045250 | -0.001199 |
| LFS 98 (NLF)* | -0.000595 | -0.017095 | -0.000590 | 0.017777 | 0.000503 |
| LFS 98(EPY-UPY)* | -0.000834 | -0.024447 | -0.002282 | 0.026787 | 0.000776 |
| LFS 98 (EPY-NLFPY)* | 0.001749 | 0.046977 | -0.005497 | -0.042105 | -0.001124 |
| LFS 98 (UPY-NLFPY)* | 0.002276 | 0.059711 | -0.009509 | -0.051134 | -0.001344 |
| LFS 98 (AOA)* | -0.000529 | -0.015364 | -0.000987 | 0.016411 | 0.000470 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 70: Marginal Effects of Model (5) for the People 60 and Over Years of Age

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CHealth 97 (Imp. a bit)* | 0.015941 | 0.221661 | -0.051555 | -0.180961 | -0.005086 |
| CHealth 97 (Unchanged)* | 0.034107 | 0.404551 | -0.057116 | -0.364960 | -0.016581 |
| CHealth 97 (Worsened)* | 0.138380 | 0.536200 | -0.270047 | -0.391086 | -0.013447 |
| CHealth 97 (Wors. badly)* | 0.649310 | 0.081169 | -0.484751 | -0.242212 | -0.003516 |
| Change in Income 97 | -3.06E-09 | -6.42E-08 | $3.89 \mathrm{E}-09$ | $6.13 \mathrm{E}-08$ | $2.03 \mathrm{E}-09$ |
| HLE 98(GHS)* | 0.001242 | 0.024652 | -0.002316 | -0.022861 | -0.000718 |
| HLE 98(NUPC)* | 0.000458 | 0.009399 | -0.000691 | -0.008878 | -0.000288 |
| HLE 98(UD)* | 0.001718 | 0.032811 | -0.003843 | -0.029784 | -0.000902 |
| Age 98 | 0.000069 | 0.001450 | -0.000088 | -0.001386 | -0.000046 |
| $\left({ }^{*}\right) \mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 71: Ordered Probit Estimates of Model (5) with Stress for the Persons of Age Under 20

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p -value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CStress 97 (Reduced)* | -0.292340 | 0.269134 | 1.188 | 0.277 | 5.59 | 0.2316 |
| CStress 97 (Unchanged)* | -0.309213 | 0.268409 | 1.323 | 0.249 |  |  |
| CStress 97 (Increased)* | -0.396058 | 0.270050 | 2.161 | 0.142 |  |  |
| CStress 97 (Subst. increased)* | -0.389151 | 0.558271 | 0.490 | 0.486 |  |  |
| Change in Income 97 | $2.01 \mathrm{E}-07$ | $9.72 \mathrm{E}-07$ | 0.044 | 0.836 | 0.04 | 0.8364 |
| Age 98 | 0.033837 | 0.019175 | 3.098 | 0.078 | 3.11 | 0.0776 |
| Sex(Female)* | -0.042290 | 0.039148 | 1.166 | 0.280 | 1.17 | 0.2800 |
| HLE 98(GHS)* | 0.083599 | 0.051318 | 2.657 | 0.103 | 4.67 | 0.1972 |
| HLE 98(NUPC)* | 0.027390 | 0.075354 | 0.130 | 0.716 |  |  |
| HLE 98(UD)* | -0.223091 | 0.241623 | 0.846 | 0.356 |  |  |
| LFS 98(Unemployed)* | 0.105344 | 0.158358 | 0.449 | 0.506 | 3.53 | 0.7400 |
| LFS 98 (NLF)* | 0.086332 | 0.068007 | 1.613 | 0.204 |  |  |
| LFS 98(EPY-UPY)* | 0.021858 | 0.074073 | 0.090 | 0.768 |  |  |
| LFS 98 (EPY-NLFPY)* | 0.071579 | 0.053974 | 1.769 | 0.185 |  |  |
| LFS 98 (UPY-NLFPY)* | 0.100150 | 0.088779 | 1.277 | 0.259 |  |  |
| LFS 98 (AOA)* | 0.016078 | 0.057585 | 0.078 | 0.780 |  |  |
| cut1 | -2.048583 | 0.453665 |  |  |  |  |
| cut2 | -0.237495 | 0.448659 |  |  |  |  |
| cut3 | 1.237968 | 0.449096 |  |  |  |  |
| cut4 | 3.268794 | 0.462933 |  |  |  |  |
| N | 3237 |  |  |  |  |  |
| Log likelihood | -3401.7296 |  |  |  |  |  |
| LR chi2(16) | 17.18 |  | p-value | 0.3738 |  |  |

Table 72: Ordered Probit Estimates of Model (5) with Stress for the People between 20 and 49 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| CStress 97 (Reduced)* | -0.128029 | 0.136205 | 0.884 | 0.347 | 45.84 | 0.0000 |
| CStress 97 (Unchanged)* | -0.167019 | 0.135799 | 1.513 | 0.219 |  |  |
| CStress 97 (Increased)* | -0.258118 | 0.136288 | 3.572 | 0.058 |  |  |
| CStress 97 (Subst. increased)* | -0.248009 | 0.186352 | 1.769 | 0.183 |  |  |
| Change in Income 97 | $1.38 \mathrm{E}-08$ | $3.97 \mathrm{E}-07$ | 0.001 | 0.972 | 0.00 | 0.9723 |
| HLE 98(GHS)* | -0.046998 | 0.021381 | 4.840 | 0.028 | 7.92 | 0.0476 |
| HLE 98(NUPC)* | -0.051499 | 0.020589 | 6.250 | 0.012 |  |  |
| HLE 98(UD)* | -0.059675 | 0.023540 | 6.452 | 0.011 | 0.093 | 2.82 |
| Age 98 | -0.001399 | 0.000834 | 2.822 |  | 0.0933 |  |
| cut1 | -2.753752 | 0.142577 |  |  |  |  |
| cut2 | -0.926003 | 0.140221 |  |  |  |  |
| cut3 | 0.443481 | 0.140166 |  |  |  |  |
| cut4 | 2.348575 | 0.143166 |  |  |  |  |
| N | 26171 |  |  |  |  |  |

Table 73: Ordered Probit Estimates of Model (5) with Stress for the People between 50 and 59 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CStress 97 (Reduced)* | -0.254274 | 0.198596 | 1.638 | 0.200 | 20.65 | 0.0004 |
| CStress 97 (Unchanged)* | -0.318983 | 0.197553 | 2.592 | 0.106 |  |  |
| CStress 97 (Increased)* | -0.405796 | 0.198585 | 4.162 | 0.041 |  |  |
| CStress 97 (Subst. increased)* | -0.790206 | 0.313629 | 6.350 | 0.012 |  |  |
| Change in Income 97 | -9.47E-07 | $6.84 \mathrm{E}-07$ | 1.904 | 0.166 | 1.92 | 0.1663 |
| LFS 98(Unemployed)* | -0.130979 | 0.094412 | 1.932 | 0.165 | 20.90 | 0.0019 |
| LFS 98 (NLF)* | 0.072528 | 0.032762 | 4.884 | 0.027 |  |  |
| LFS 98(EPY-UPY)* | 0.115078 | 0.065143 | 3.133 | 0.077 |  |  |
| LFS 98 (EPY-NLFPY)* | -0.136931 | 0.062881 | 4.752 | 0.029 |  |  |
| LFS 98 (UPY-NLFPY)* | -0.233164 | 0.129883 | 3.240 | 0.073 |  |  |
| LFS 98 (AOA)* | 0.018758 | 0.091029 | 0.044 | 0.837 |  |  |
| Age 98 | 0.007954 | 0.004985 | 2.560 | 0.111 | 2.55 | 0.1106 |
| cut1 | -2.254811 | 0.344841 |  |  |  |  |
| cut2 | -0.518238 | 0.341725 |  |  |  |  |
| cut3 | 0.795413 | 0.341790 |  |  |  |  |
| cut4 | 2.607242 | 0.345716 |  |  |  |  |
| N | 6143 |  |  |  |  |  |
| Log likelihood | -6846.4901 |  |  |  |  |  |
| LR chi2(12) | 46.75 |  |  |  |  |  |

Table 74: Ordered Logit Estimates of Model (5) with Stress for the People between 50 and 59 Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| CStress 97 (Reduced)* | -0.437718 | 0.357102 | 1.513 | 0.220 | 18.92 | 0.0008 |
| CStress 97 (Unchanged)* | -0.539171 | 0.355324 | 2.310 | 0.129 |  |  |
| CStress 97 (Increased)* | -0.684271 | 0.357096 | 3.686 | 0.055 |  |  |
| CStress 97 (Subst. increased)* | -1.348684 | 0.541194 | 6.200 | 0.013 | 0.924 | 0.088 |
| Change in Income 97 | $-2.06 \mathrm{E}-06$ | $1.20 \mathrm{E}-06$ | 2.924 | 2.91 | 0.0883 |  |
| LFS 98(Unemployed)* | -0.231624 | 0.167020 | 1.932 | 0.166 | 20.95 | 0.0019 |
| LFS 98 (NLF)* | 0.148582 | 0.054662 | 7.398 | 0.007 |  |  |
| LFS 98(EPY-UPY)* | 0.182333 | 0.112531 | 2.624 | 0.105 |  |  |
| LFS 98 (EPY-NLFPY)* | -0.217941 | 0.108548 | 4.040 | 0.045 |  |  |
| LFS 98 (UPY-NLFPY)* | -0.327532 | 0.227641 | 2.074 | 0.150 | 0.830 |  |
| LFS 98 (AOA)* | 0.034316 | 0.160134 | 0.044 |  |  |  |
| cut1 | -5.270136 | 0.380389 |  |  |  |  |
| cut2 | -1.581884 | 0.355591 |  |  |  |  |
| cut3 | 0.557429 | 0.355060 |  |  |  |  |
| cut4 | 4.452762 | 0.384892 |  |  |  |  |
| N | 6143 |  |  |  |  |  |

Table 75: Ordered Probit Estimates of Model (5) with Stress for the People 60 and Over Years of Age

| Variables | Coefficient | Std. Err. | $\chi^{2}$-statistic | p-value | LRT Chi2 | p-value |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| CStress 97 (Reduced) | -0.413009 | 0.148865 | 7.673 | 0.006 | 25.22 | 0.0000 |
| CStress 97 (Unchanged)* | -0.472005 | 0.148322 | 10.112 | 0.001 |  |  |
| CStress 97 (Increased)* | -0.531615 | 0.149040 | 12.745 | 0.000 |  |  |
| CStress 97 (Subst. increased)* | -0.498585 | 0.207733 | 5.760 | 0.016 |  |  |
| Change in Income 97 | $2.78 \mathrm{E}-07$ | $9.27 \mathrm{E}-07$ | 0.090 | 0.764 | 0.09 | 0.7640 |
| HLE 98(GHS)* | -0.072549 | 0.030814 | 5.523 | 0.019 | 6.20 | 0.1021 |
| HLE 98(NUPC)* | -0.011238 | 0.030791 | 0.130 | 0.715 |  |  |
| HLE 98(UD)* | -0.050387 | 0.044561 | 1.277 | 0.258 | 0.021 | 5.33 |
| Age 98 | -0.003587 | 0.001553 | 5.336 |  |  |  |
| cut1 | -3.012700 | 0.192084 |  |  |  |  |
| cut2 | -1.299258 | 0.188669 |  |  |  |  |
| cut3 | -0.112596 | 0.188395 |  |  |  |  |
| cut4 | 1.685971 | 0.191919 |  |  |  |  |
| N | 9385 |  |  |  |  |  |

Table 76: Marginal Effects of Model (5) with Stress for the People between 20 and 49 Years of Age

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CStress 97 (Reduced)* | 0.002569 | 0.039585 | -0.003625 | -0.037020 | -0.001509 |
| CStress 97 (Unchanged)* | 0.003028 | 0.050488 | -0.001700 | -0.049611 | -0.002205 |
| CStress 97 (Increased)* | 0.005770 | 0.081178 | -0.011715 | -0.072468 | -0.002766 |
| CStress 97 (Subst. increased)* | 0.006218 | 0.079616 | -0.016349 | -0.067160 | -0.002325 |
| Change in Income 97 | -2.54E-10 | -4.19E-09 | $1.89 \mathrm{E}-10$ | $4.08 \mathrm{E}-09$ | $1.77 \mathrm{E}-10$ |
| HLE 98(GHS)* | 0.000886 | 0.014358 | -0.000839 | -0.013815 | -0.000589 |
| HLE 98(NUPC)* | 0.000964 | 0.015709 | -0.000857 | -0.015166 | -0.000650 |
| HLE 98(UD)* | 0.001151 | 0.018321 | -0.001309 | -0.017434 | -0.000729 |
| Age 98 | 0.000026 | 0.000425 | -0.000019 | -0.000414 | -0.000018 |
| (*) | $\mathrm{dy} / \mathrm{dx}$ is for discrete cha | of dummy var | ble from 0 to |  |  |

Table 77: Marginal Effects of Ordered Probit Model (5) with Stress for the People between 50 and 59 Years of Age

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CStress 97 (Reduced)* | 0.006864 | 0.078254 | -0.007180 | -0.073808 | -0.004129 |
| CStress 97 (Unchanged)* | 0.007345 | 0.094752 | 0.000439 | -0.096312 | -0.006225 |
| CStress 97 (Increased)* | 0.012193 | 0.126295 | -0.017875 | -0.114511 | -0.006103 |
| CStress 97 (Subst. increased)* | 0.045961 | 0.251329 | -0.115301 | -0.175741 | -0.006249 |
| Change in Income 97 | $2.16 \mathrm{E}-08$ | $2.83 \mathrm{E}-07$ | -3.96E-10 | -2.86E-07 | -1.79E-08 |
| LFS 98(Unemployed)* | 0.003473 | 0.040327 | -0.003502 | -0.038173 | -0.002126 |
| LFS 98 (NLF)* | -0.001600 | -0.021514 | -0.000404 | 0.022097 | 0.001421 |
| LFS 98(EPY-UPY)* | -0.002322 | -0.033410 | -0.002463 | 0.035719 | 0.002475 |
| LFS 98 (EPY-NLFPY)* | 0.003620 | 0.042128 | -0.003575 | -0.039942 | -0.002231 |
| LFS 98 (UPY-NLFPY)* | 0.007003 | 0.073181 | -0.011181 | -0.065655 | -0.003348 |
| LFS 98 (AOA)* | -0.000419 | -0.005581 | -0.000063 | 0.005700 | 0.000363 |
| Age 98 | -0.000181 | -0.002378 | $3.33 \mathrm{E}-06$ | 0.002406 | 0.000150 |
| (*) | $\mathrm{dy} / \mathrm{dx}$ is for discrete cha | of dummy var | le from 0 to |  |  |

Table 78: Marginal Effects of Ordered Logit Model (5) with Stress for the People between 50 and 59 Years of Age

| Change in Health 1998 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CStress 97 (Reduced)* | 0.004170 | 0.083570 | -0.009319 | -0.075714 | -0.002707 |
| CStress 97 (Unchanged)* | 0.004576 | 0.097517 | 0.000563 | -0.098882 | -0.003775 |
| CStress 97 (Increased)* | 0.007014 | 0.133541 | -0.022161 | -0.114376 | -0.004018 |
| CStress 97 (Subst. increased)* | 0.023434 | 0.290248 | -0.139670 | -0.168876 | -0.005136 |
| Change in Income 97 | $1.74 \mathrm{E}-08$ | $3.74 \mathrm{E}-07$ | -5.88E-10 | $-3.77 \mathrm{E}-07$ | -1.41E-08 |
| LFS 98(Unemployed)* | 0.002187 | 0.044218 | -0.004804 | -0.040170 | -0.001430 |
| LFS 98 (NLF)* | -0.001221 | -0.026655 | -0.000759 | 0.027584 | 0.001052 |
| LFS 98(EPY-UPY)* | -0.001423 | -0.031873 | -0.002723 | 0.034659 | 0.001360 |
| LFS 98 (EPY-NLFPY)* | 0.002030 | 0.041375 | -0.003983 | -0.038058 | -0.001363 |
| LFS 98 (UPY-NLFPY)* | 0.003252 | 0.063754 | -0.009729 | -0.055348 | -0.001930 |
| LFS 98 (AOA)* | -0.000285 | -0.006194 | -0.000094 | 0.006335 | 0.000240 |
| (*) $\mathrm{dy} / \mathrm{dx}$ is for discrete change of dummy variable from 0 to 1 |  |  |  |  |  |

Table 79: Marginal Effects of Model (5) with Stress for the People 60 and Over Years of Age

|  | Change in Health 1998 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Improved substantially | Improved a bit | Unchanged | Worsened | Worsened badly |
| CStress 97 (Reduced)* | 0.015832 | 0.131358 | -0.020800 | -0.119281 | -0.007110 |
| CStress 97 (Unchanged)* | 0.015716 | 0.146690 | -0.011265 | -0.141427 | -0.009714 |
| CStress 97 (Increased)* | 0.022516 | 0.169500 | -0.034267 | -0.149250 | -0.008500 |
| CStress 97 (Subst. increased)* | 0.026630 | 0.161327 | -0.053131 | -0.128955 | -0.005871 |
| Change in Income 97 | -8.49E-09 | -8.67E-08 | $4.36 \mathrm{E}-09$ | $8.51 \mathrm{E}-08$ | $5.78 \mathrm{E}-09$ |
| HLE 98(GHS)* | 0.002339 | 0.022807 | -0.001820 | -0.021904 | -0.001422 |
| HLE 98(NUPC)* | 0.000346 | 0.003508 | -0.000192 | -0.003430 | -0.000231 |
| HLE 98(UD)* | 0.001616 | 0.015830 | -0.001219 | -0.015233 | -0.000993 |
| Age 98 | 0.000110 | 0.001118 | -0.000056 | -0.001097 | -0.000075 |
|  | $\mathrm{dy} / \mathrm{dx}$ is for discrete cha | of dummy var | e from 0 to |  |  |

Table 80: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6)

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1519452 | -0.0003977 | 0.0075865 | -0.1668325 | -0.1370579 | (N) |
|  |  |  |  |  | -0.1664828 | -0.1371194 | (P) |
|  |  |  |  |  | -0.1653944 | -0.1358931 | (BC) |
| CHealth 97 (Subst. imp.)* | 1000 | $-1129.833$ | -4.140325 | 830.164 | -2758.898 | 499.232 | (N) |
|  |  |  |  |  | -2768.247 | 473.740 | (P) |
|  |  |  |  |  | -2830.368 | 419.850 | (BC) |
| CHealth 97 (Imp. a bit)* | 1000 | -189.3181 | -3.693215 | 137.7626 | -459.655 | 81.019 | (N) |
|  |  |  |  |  | -459.351 | 66.520 | (P) |
|  |  |  |  |  | -452.255 | 69.445 | (BC) |
| CHealth 97 (Worsened)* | 1000 | 2.848337 | $-2.10237$ | 145.0803 | -281.849 | 287.545 | (N) |
|  |  |  |  |  | -270.359 | 291.192 | (P) |
|  |  |  |  |  | -265.037 | 298.406 | (BC) |
| CHealth 97 (Wors. badly)* | 1000 | -836.9861 | -33.46674 | 874.7398 | -2553.524 | 879.552 | (N) |
|  |  |  |  |  | -2657.995 | 868.725 | (P) |
|  |  |  |  |  | -2582.838 | 982.235 | (BC) |
| LFS 98(Unemployed)*LFS 98 (NLF)* | 1000 | -3613.345 | -9.076668 | 369.1722 | -4337.787 | -2888.904 | (N) |
|  |  |  |  |  | -4319.908 | -2922.365 | (P) |
|  |  |  |  |  | -4317.603 | -2921.755 | (BC) |
|  | 1000 | -1859.29 | 5.010948 | 156.5067 | -2166.410 | -1552.170 | (N) |
|  |  |  |  |  | -2175.257 | -1542.119 | (P) |
|  |  |  |  |  | -2191.781 | -1546.730 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1171.935 | -4.055265 | 226.8242 | -1617.041 | -726.828 | (N) |
|  |  |  |  |  | -1594.462 | -737.018 | (P) |
|  |  |  |  |  | -1579.700 | -734.085 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1073.788 | 11.69886 | 300.9145 | -1664.285 | -483.291 | (N) |
|  |  |  |  |  | -1677.919 | -480.708 | (P) |
|  |  |  |  |  | -1704.996 | -488.702 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -3360.684 | 19.50144 | 416.608 | -4178.211 | -2543.157 | (N) |
|  |  |  |  |  | $\begin{aligned} & -4193.728 \\ & -4222.388 \end{aligned}$ | $\begin{aligned} & -2569.349 \\ & -2593.589 \end{aligned}$ | (P) |
| LFS 98 (AOA)* | 1000 | -1562.81 | -8.920867 | 334.5782 | -2219.366 | -906.253 | (N) |
|  |  |  |  |  | -2213.629 | -926.035 | (P) |
|  |  |  |  |  | -2201.063 | -912.274 | (BC) |
| HLE 98(GHS)* | 1000 | 258.7125 | -8.160569 | 161.2039 | -57.625 | 575.050 | (N) |
|  |  |  |  |  | -57.897 | 553.117 | (P) |
|  |  |  |  |  | -50.972 | 577.841 | (BC) |
| HLE 98(NUPC)* | 1000 | 202.953 | -1.541444 | 154.0629 | -99.371 | 505.277 | (N) |
|  |  |  |  |  | -118.694 | 485.198 | (P) |
|  |  |  |  |  | -125.036 | 473.988 | (BC) |
| HLE 98(UD)* | 1000 | 631.1646 | 0.3782268 | 215.1229 | 209.020 | 1053.309 | (N) |
|  |  |  |  |  | 207.087 | 1041.042 | (P) |
|  |  |  |  |  | 204.118 | 1027.789 | (BC) |
| Age 98 | 1000 | $-28.98312$ | -0.1339627 | 5.107968 | -39.007 | -18.960 | (N) |
|  |  |  |  |  | -39.199 | -19.212 | (P) |
|  |  |  |  |  | -38.661 | -18.348 | (BC) |
| Sex(Female)* | 1000 | 212.238 | -4.927051 | 119.0517 | -21.382 | 445.858 | (N) |
|  |  |  |  |  | -32.774 | 437.560 | (P) |
|  |  |  |  |  | -0.701 | 463.256 | (BC) |
| Constant | 1000 | 2698.509 | 10.12159 | 275.2028 | 2158.467 | 3238.551 | (N) |
|  |  |  |  |  | 2161.973 | 3228.057 | (P) |
|  |  |  |  |  | 2150.981 | 3222.018 | (BC) |
|  |  | normal, P | percentile, | $=$ bias-co | ected |  |  |

Table 81: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) for the Persons of Age Under 20

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 2000 | -0.2192924 | -0.0013764 | 0.0280884 | -0.2743781 | -0.1642068 | (N) |
|  |  |  |  |  | -0.2758427 | -0.1669236 | (P) |
|  |  |  |  |  | -0.2712429 | -0.1624007 | (BC) |
| CHealth 97 (Subst. imp.)* | 2000 | 6348.795 | -72.36743 | 6220.759 | -5851.056 | 18548.650 | (N) |
|  |  |  |  |  | -6123.463 | 18119.950 | (P) |
|  |  |  |  |  | -6092.301 | 18153.050 | (BC) |
| CHealth 97 (Imp. a bit)* | 2000 | 1552.683 | -6.215028 | 800.3973 | -17.017 | 3122.384 | (N) |
|  |  |  |  |  | -52.068 | 3068.047 | (P) |
|  |  |  |  |  | -50.780 | 3071.035 | (BC) |
| CHealth 97 (Worsened)* | 2000 | -221.9899 | $-14.77181$ | 792.2187 | -1775.651 | 1331.671 | (N) |
|  |  |  |  |  | -1802.912 | 1314.422 | (P) |
|  |  |  |  |  | -1797.215 | 1317.634 | (BC) |
| CHealth 97 (Wors. badly)* | 2000 | 962.1774 | -166.2725 | 4012.857 | -6907.643 | 8831.998 | (N) |
|  |  |  |  |  | -6605.624 | 8764.262 | (P) |
|  |  |  |  |  | -5925.383 | 9846.988 | (BC) |
| LFS 98(Unemployed)* | 2000 | -6405.653 | 16.506 | 2029.158 | -10385.140 | -2426.167 | (N) |
|  |  |  |  |  | -10493.780 | -2551.152 | (P) |
|  |  |  |  |  | -10640.370 | -2648.594 | (BC) |
| LFS 98 (NLF)* | 2000 | -1987.945 | -41.31618 | 1031.111 | -4010.110 | 34.220 | (N) |
|  |  |  |  |  | -3996.416 | 20.114 | (P) |
|  |  |  |  |  | -3871.011 | 236.784 | (BC) |
| LFS 98(EPY-UPY)* | 2000 | -1351.728 | 7.720622 | 1148.765 | -3604.630 | 901.174 | (N) |
|  |  |  |  |  | -3707.213 | 812.428 | (P) |
|  |  |  |  |  | -3796.320 | 741.670 | (BC) |
| LFS 98 (EPY-NLFPY)* | 2000 | -1383.372 | -6.174331 | 897.6642 | -3143.828 | 377.083 | (N) |
|  |  |  |  |  | -3193.500 | 379.225 | (P) |
|  |  |  |  |  | -3193.976 | 363.443 | (BC) |
| LFS 98 (UPY-NLFPY)*LFS 98 (AOA)* | 2000 | -3198.783 | 11.09657 | 1057.886 | -5273.458 | -1124.108 | (N) |
|  |  |  |  |  | -5254.465 | -1124.665 | (P) |
|  |  |  |  |  | -5254.896 | -1131.495 | (BC) |
|  | 2000 | -2899.933 | -30.48591 | 961.8877 | -4786.341 | -1013.526 | (N) |
|  |  |  |  |  | -4851.862 | -1028.706 | (P) |
|  |  |  |  |  | -4812.689 | -988.323 | (BC) |
| HLE 98(GHS)* | 2000 | 515.8502 | 31.91381 | 813.8676 | -1080.267 | 2111.968 | (N) |
|  |  |  |  |  | -986.321 | 2162.166 | (P) |
|  |  |  |  |  | -1072.619 | 2087.976 2191.068 | (BC) $(\mathrm{N})$ |
| HLE 98(NUPC)* | 2000 | -234.0744 | 50.45603 | 1236.591 | -2658.496 | 2184.710 | (P) |
|  |  |  |  |  | -2748.307 | 1991.166 | (BC) |
| HLE 98(UD)* | 2000 | 188.9626 | 126.4686 | 4160.12 | -7969.662 | 8347.587 | (N) |
|  |  |  |  |  | -8250.583 | 8153.374 | (P) |
|  |  |  |  |  | -8924.624 | 7667.586 | (BC) |
| Age 98 | 2000 | -339.3291 | -6.544902 | 322.597 | -971.991 | 293.333 | (N) |
|  |  |  |  |  | -986.245 | 288.675 | (P) |
|  |  |  |  |  | -985.614 | 295.675 | (BC) |
| Sex(Female)* | 2000 | -481.5898 | -18.22894 | 625.8452 | -1708.967 | 745.787 | (N) |
|  |  |  |  |  | -1771.957 | 778.066 | (P) |
|  |  |  |  |  | -1704.546 | 816.670 | (BC) |
| Constant | 2000 | 8253.964 | 125.5568 | 6120.76 | -3749.773 | 20257.700 | (N) |
|  |  |  |  |  | -3801.057 | 20572.700 | (P) |
|  |  |  |  |  | -4013.385 | 20308.050 | (BC) |
|  |  | = normal, P | percentile, | C bias-co | ected |  |  |

Table 82: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) for the People between 20 and 49 Years of Age

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1609016 | -0.0009949 | 0.0085645 | -0.1777081 | -0.1440952 | (N) |
|  |  |  |  |  | -0.1792785 | -0.1460818 | (P) |
|  |  |  |  |  | -0.1772059 | -0.1442442 | (BC) |
| CHealth 97 (Subst. imp.)* | 1000 | -928.228 | 36.50397 | 1022.757 | -2935.226 | 1078.770 | (N) |
|  |  |  |  |  | -2942.458 | 1121.741 | (P) |
|  |  |  |  |  | -3074.150 | 1012.676 | (BC) |
| CHealth 97 (Imp. a bit)* | 1000 | -293.6041 | 3.808419 | 161.1103 | -609.757 | 22.549 | (N) |
|  |  |  |  |  | -623.482 | 15.283 | (P) |
|  |  |  |  |  | -642.052 | -0.899 | (BC) |
| CHealth 97 (Worsened)* | 1000 | 118.7314 | -7.292181 | 172.5949 | -219.959 | 457.422 | (N) |
|  |  |  |  |  | -220.669 | 448.932 | (P) |
|  |  |  |  |  | -211.719 | 475.673 | (BC) |
| CHealth 97 (Wors. badly)* | 1000 | -536.4304 | -43.11863 | 1152.726 | -2798.472 | 1725.611 | (N) |
|  |  |  |  |  | -2844.780 | 1651.095 | (P) |
|  |  |  |  |  | -2839.028 | 1640.580 | (BC) |
| LFS 98(Unemployed)*LFS 98 (NLF)* | 1000 | -3613.064 | -11.55389 | 403.764 | -4405.387 | -2820.741 | (N) |
|  |  |  |  |  | -4416.415 | -2826.095 | (P) |
|  |  |  |  |  | -4372.941 | -2797.403 | (BC) |
|  | 1000 | -1774.257 | -2.332749 | 236.61 | -2238.567 | -1309.947 | (N) |
|  |  |  |  |  | -2270.431 | -1331.037 | (P) |
|  |  |  |  |  | -2285.448 | -1341.733 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1074.655 | 5.231742 | 256.3255 | -1577.653 | -571.657 | (N) |
|  |  |  |  |  | -1572.995 | -567.699 | (P) |
|  |  |  |  |  | -1602.694 | -587.205 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1364.775 | 5.495233 | 373.6967 | -2098.095 | -631.454 | (N) |
|  |  |  |  |  | -2116.034 | -599.209 | (P) |
|  |  |  |  |  | -2087.681 | -591.538 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -3368.938 | -6.453768 | 593.3056 | -4533.207 | -2204.670 | (N) |
|  |  |  |  |  | -4518.461 | -2266.940 | (P) |
|  |  |  |  |  | -4524.652 | -2291.107 | (BC) |
| LFS 98 (AOA)* | 1000 | -1582.175 | -5.630452 | 400.272 | -2367.645 | -796.704 | (N) |
|  |  |  |  |  | -2349.447 | -790.975 | (P) |
|  |  |  |  |  | -2337.959 | -780.822 | (BC) |
| HLE 98(GHS)* | 1000 | 329.8902 | -3.093048 | 206.6137 | -75.556 | 735.337 | (N) |
|  |  |  |  |  | -61.873 | 759.026 | (P) |
|  |  |  |  |  | -60.553 | 761.535 | (BC) |
| HLE 98(NUPC)* | 1000 | 411.8238 | -0.6744381 | 205.2756 | 9.003 | 814.645 | (N) |
|  |  |  |  |  | 7.417 | 835.667 | (P) |
|  |  |  |  |  | 6.715 | 835.522 | (BC) |
| HLE 98(UD)* | 1000 | 1216.656 | -11.99339 | 254.4289 | 717.380 | 1715.933 | (N) |
|  |  |  |  |  | 704.424 | 1672.013 | (P) |
|  |  |  |  |  | 743.285 | 1714.031 | (BC) |
| Age 98 | 1000 | -7.026226 | 0.3077807 | 9.979608 | -26.610 | 12.557 | (N) |
|  |  |  |  |  | -27.132 | 12.625 | (P) |
|  |  |  |  |  | -28.063 | 11.829 | (BC) |
| Sex(Female)* | 1000 | 260.2632 | 4.39063 | 145.9076 | -26.057 | 546.584 | (N) |
|  |  |  |  |  | -12.197 | 557.265 | (P) |
|  |  |  |  |  | -10.962 | 558.480 | (BC) |
| Constant | 1000 | 1794.868 | -9.619103 | 446.7699 | 918.153 | 2671.583 | (N) |
|  |  |  |  |  | 871.001 | 2685.631 | (P) |
|  |  |  |  |  | 902.307 | 2707.256 | (BC) |
|  |  | = normal, P | percentile, | $=$ bias-c | cted |  |  |

Table 83: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) for the People between 50 and 59 Years of Age

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1351552 | -0.0023938 | 0.0219249 | -0.1781793 | -0.0921311 | (N) |
|  |  |  |  |  | -0.1793945 | -0.0973532 | (P) |
|  |  |  |  |  | -0.1783892 | -0.0963628 |  |
| CHealth 97 (Subst. imp.)* | 1000 | -779.7293 | -46.2007 | 1836.429 | -4383.429 | 2823.971 | (N) |
|  |  |  |  |  | -4362.534 | 2659.574 | (P) |
|  |  |  |  |  | -4076.641 | 2862.711 | (BC) |
| CHealth 97 (Imp. a bit)* | 1000 | -226.5892 | -16.52547 | 384.8245 | -981.746 | 528.568 | (N) |
|  |  |  |  |  | -1013.532 | 547.682 | (P) |
|  |  |  |  |  | -987.354 | 586.111 | (BC) |
| CHealth 97 (Worsened)* | 1000 | 220.1647 | -12.42602 | 403.9685 | -572.559 | 1012.889 | (N) |
|  |  |  |  |  | -613.260 | 989.175 | (P) |
|  |  |  |  |  | -599.664 | 1017.257 | (BC) |
| CHealth 97 (Wors. badly)* | 1000 | -2955.074 | 21.75063 | 1744.391 | -6378.164 | 468.015 | (N) |
|  |  |  |  |  | -6703.537 | 225.182 | (P) |
|  |  |  |  |  | -7103.743 | -15.586 | (BC) |
| LFS 98 (Unemployed)*LFS 98 (NLF)* | 1000 | -2197.656 | -7.110655 | 866.6102 | -3898.242 | -497.071 | (N) |
|  |  |  |  |  | -3850.656 | -545.228 | (P) |
|  |  |  |  |  | -3848.261 | -493.704 | (BC) |
|  | 1000 | -2627.67 | -9.516379 | 385.2841 | -3383.729 | -1871.611 | (N) |
|  |  |  |  |  | -3416.982 | -1920.952 | (P) |
|  |  |  |  |  | -3405.224 | -1915.806 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1515.095 | -29.63698 | 710.3771 | -2909.097 | -121.092 | (N) |
|  |  |  |  |  | -2859.844 | -141.503 | (P) |
|  |  |  |  |  | -2775.216 | -34.746 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -80.06652 | -35.33712 | 948.8243 | -1941.984 | 1781.851 | (N) |
|  |  |  |  |  | -1914.790 | 1811.197 | (P) |
|  |  |  |  |  | -1887.653 | 1857.501 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -4004.912 | 33.50533 | 1390.679 | -6733.898 | -1275.926 | (N) |
|  |  |  |  |  | -6690.287 | -1427.974 | (P) |
|  |  |  |  |  | -6753.167 | -1465.738 | (BC) |
| LFS 98 (AOA)* | 1000 | 1483.339 | -54.23738 | 1172.767 | -818.030 | 3784.708 | (N) |
|  |  |  |  |  | -896.653 | 3724.815 | (P) |
|  |  |  |  |  | -761.017 | 3830.455 | (BC) |
| HLE 98(GHS)* | 1000 | 564.3912 | -31.05215 | 423.4105 | -266.485 | 1395.267 | (N) |
|  |  |  |  |  | -301.079 | 1358.570 | (P) |
|  |  |  |  |  | -260.250 | 1383.165 | (BC) |
| HLE 98(NUPC)* | 1000 | 395.0593 | -0.6806683 | 392.866 | -375.878 | 1165.997 | (N) |
|  |  |  |  |  | -383.460 | 1200.749 | (P) |
|  |  |  |  |  | -383.868 | 1199.504 | (BC) |
| HLE 98(UD)* | 1000 | -332.4826 | -10.82701 | 614.7724 | -1538.876 | 873.911 | (N) |
|  |  |  |  |  | -1524.327 | 889.137 | (P) |
|  |  |  |  |  | -1504.631 | 914.149 | (BC) |
| Age 98 | 1000 | 65.09822 | -0.645678 | 56.19627 | -45.178 | 175.375 | (N) |
|  |  |  |  |  | -46.602 | 173.892 | (P) |
|  |  |  |  |  | -49.130 | 171.494 | (BC) |
| Sex(Female)* | 1000 | 276.8173 | 13.7065 | 325.4463 | -361.819 | 915.454 | (N) |
|  |  |  |  |  | -332.219 | 936.688 | (P) |
|  |  |  |  |  | -344.270 | 916.323 | (BC) |
| Constant | 1000 | -3412.052 | 49.80656 | 3177.389 | -9647.175 | 2823.071 | (N) |
|  |  |  |  |  | -9279.429 | 3008.231 | (P) |
|  |  |  |  |  | -9202.806 | 3037.691 | (BC) |
|  |  | = normal, P | percentile, | $\mathrm{C}=$ bias-co | ected |  |  |

Table 84: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) for the People 60 and Over Years of Age

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1183763 | -0.0000515 | 0.0186019 | -0.1548795 | -0.0818731 | (N) |
|  |  |  |  |  | -0.1555529 | -0.0817578 | (P) |
|  |  |  |  |  | -0.1563663 | -0.0824588 | (BC) |
| CHealth 97 (Subst. imp.)* | 1000 | -404.8744 | 35.68482 | 915.9692 | -2202.319 | 1392.570 | (N) |
|  |  |  |  |  | -2169.428 | 1374.064 | (P) |
|  |  |  |  |  | -2287.542 | 1348.555 | (BC) |
| CHealth 97 (Imp. a bit)* | 1000 | 117.8372 | -16.81126 | 285.0172 | -441.464 | 677.138 | (N) |
|  |  |  |  |  | -497.913 | 661.662 | (P) |
|  |  |  |  |  | -430.647 | 688.806 | (BC) |
| CHealth 97 (Worsened)* | 1000 | -178.6558 | -15.39815 | 291.8954 | -751.454 | 394.143 | (N) |
|  |  |  |  |  | -766.417 | 384.911 | (P) |
|  |  |  |  |  | -691.326 | 473.154 | (BC) |
| CHealth 97 (Wors. badly)* | 1000 | 1541.582 | -75.41444 | 1391.659 | -1189.329 | 4272.493 | (N) |
|  |  |  |  |  | -1170.869 | 4242.958 | (P) |
|  |  |  |  |  | -939.739 | 4639.166 | (BC) |
| LFS 98(Unemployed)* | 1000 | -1376.322 | 30.92808 | 1132.315 | -3598.310 | 845.666 | (N) |
|  |  |  |  |  | -3494.228 | 727.335 | (P) |
|  |  |  |  |  | -3920.593 | 560.544 | (BC) |
| LFS 98 (NLF)* | 1000 | -552.327 | -10.60472 | 414.8962 | -1366.495 | 261.841 | (N) |
|  |  |  |  |  | -1388.447 | 217.066 | (P) |
|  |  |  |  |  | -1387.171 | 222.874 | ( BC ) |
| LFS 98(EPY-UPY)* | 1000 | 484.6785 | 42.44723 | 1088.385 | -1563.135 | 2704.105 2672.796 | (B) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1037.016 | -23.09869 | 985.9809 | -2971.847 | 897.815 | (N) |
|  |  |  |  |  | -2973.703 | 921.279 | (P) |
|  |  |  |  |  | -2856.869 | 995.002 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -1910.18 | 51.6742 | 1973.594 | $-5783.045$ | 1962.684 | (N) |
|  |  |  |  |  | -5567.413 | 2265.450 | (P) |
|  |  |  |  |  | -5420.628 | 2357.709 | (BC) |
| LFS 98 (AOA)* | 1000 | -501.537 | -19.891 | 1377.168 | -3204.011 | 2200.937 | (N) |
|  |  |  |  |  | -3315.679 | 2078.813 | (P) |
|  |  |  |  |  | -3381.241 | 2061.170 | (BC) |
| HLE 98(GHS)* | 1000 | 290.8663 | -3.658929 | 342.5284 | -381.291 | 963.024 | (N) |
|  |  |  |  |  | -379.874 | 973.862 | (P) |
|  |  |  |  |  | -347.600 | 983.431 | (BC) |
| HLE 98(NUPC)* | 1000 | -342.6597 | -0.0439277 | 321.9435 | -974.423 | 289.103 | (N) |
|  |  |  |  |  | -952.834 | 294.325 | (P) |
|  |  |  |  |  | -942.693 | 297.172 | (BC) |
| HLE 98(UD)* | 1000 | -141.2812 | -11.06124 | 535.7368 | -1192.580 | 910.017 | (N) |
|  |  |  |  |  | -1257.510 | 872.656 | (P) |
|  |  |  |  |  | -1171.445 | 879.554 | (BC) |
| Age 98 | 1000 | 6.618227 | 2.057892 | 54.45128 | -100.234 | 113.470 | (N) |
|  |  |  |  |  | -100.125 | 117.370 | (P) |
| Sex(Female)* |  |  |  |  | -99.768 -329.190 | 116.882 610.799 | $(\mathrm{BC})$ |
|  | 1000 | 140.8042 | -0.9514696 | 239.5069 | -358.721 | 613.509 | (P) |
|  |  |  |  |  | -360.710 | 609.607 | (BC) |
| Constant | 1000 | 199.9536 | -113.2733 | 3499.796 | -6667.842 | 7067.749 | (N) |
|  |  |  |  |  | -6807.844 | 7045.968 | (P) |
|  |  |  |  |  | -6838.305 | 6969.150 | (BC) |
|  |  | = normal, P | percentile, | $=\mathrm{bias}-\mathrm{co}$ | ected |  |  |

Table 85: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1539399 | -0.0001022 | 0.0079097 | -0.169461 | -0.138419 | (N) |
|  |  |  |  |  | -0.169797 | -0.138935 | (P) |
|  |  |  |  |  | -0.169914 | -0.139130 | (BC) |
| CStress 97 (Subst. reduced)* | 1000 | 860.317 | 0.2011496 | 1025.756 | -1152.566 | 2873.200 | (N) |
|  |  |  |  |  | -1087.618 | 2819.471 | (P) |
|  |  |  |  |  | -1082.280 | 2823.311 | (BC) |
| CStress 97 (Reduced)* | 1000 | -52.20552 | -8.133353 | 142.1739 | -331.199 | 226.788 | (N) |
|  |  |  |  |  | -330.067 | 227.706 | (P) |
|  |  |  |  |  | -326.306 | 246.536 | (BC) |
| CStress 97 (Increased)* | 1000 | -357.7228 | -2.181628 | 155.1973 | -662.273 | -53.173 | (N) |
|  |  |  |  |  | -676.294 | -41.473 | (P) |
|  |  |  |  |  | -635.557 | -29.088 | (BC) |
| CStress 97 (Subst. increased)* | 1000 | -17.1635 | 2.86061 | 999.3138 | -1978.158 | 1943.831 | (N) |
|  |  |  |  |  | -2068.131 | 2038.198 | (P) |
|  |  |  |  |  | -2073.185 | 2025.237 | (BC) |
| LFS 98(Unemployed)* | 1000 | -3658.192 | 18.90737 | 350.4684 | -4345.930 | -2970.453 | (N) |
|  |  |  |  |  | -4309.775 | -2950.868 | (P) |
|  |  |  |  |  | -4386.264 | -3015.542 | (BC) |
| LFS 98 (NLF)* | 1000 | -1875.532 | 5.323632 | 166.3463 | -2201.960 | -1549.104 | (N) |
|  |  |  |  |  | -2195.876 | -1549.546 | (P) |
|  |  |  |  |  | -2207.144 | -1551.516 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1154.012 | 1.232206 | 234.7163 | -1614.605 | -693.419 | (N) |
|  |  |  |  |  | -1633.691 | -723.746 | (P) |
|  |  |  |  |  | -1663.229 | -734.856 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1147.382 | 5.547276 | 285.3843 | -1707.403 | -587.360 | (N) |
|  |  |  |  |  | -1685.586 | -564.817 | (P) |
| LFS 98 (UPY-NLFPY)* | 1000 | -3359.781 | -7.548004 | 430.0322 | -4203.651 | -2515.911 | (N) |
|  |  |  |  |  | -4202.245 | -2524.024 | (P) |
|  |  |  |  |  | -4197.563 | -2526.094 | (BC) |
| LFS 98 (AOA)* | 1000 | -1692.293 | 3.520657 | 320.7599 | -2321.733 | -1062.852 | (N) |
|  |  |  |  |  | -2296.930 | -1063.663 | (P) |
|  |  |  |  |  | -2313.421 | -1084.607 | (BC) |
| HLE 98(GHS)* | 1000 | 267.0355 | -0.3975645 | 166.8652 | -60.411 | 594.482 | (N) |
|  |  |  |  |  | -67.193 | 592.613 | (P) |
|  |  |  |  |  | -68.705 | 584.257 | (BC) |
| HLE 98(NUPC)* | 1000 | 191.8385 | 4.243788 | 157.7822 | -117.784 | 501.461 | (N) |
|  |  |  |  |  | -117.775 | 502.183 | (P) |
| HLE 98(UD)* | 1000 | 613.8696 | 2.133933 | 201.8337 | - 217.803 | 1009.936 | (N) |
|  |  |  |  |  | 216.871 | 1010.830 | (P) |
|  |  |  |  |  | 215.525 | 1009.373 | (BC) |
| Age 98 | 1000 | -29.90941 | 0.0665765 | 5.36597 | -40.439 | -19.380 | (N) |
|  |  |  |  |  | -40.474 | -19.551 | (P) |
|  |  |  |  |  | -40.434 | -19.529 | (BC) |
| Sex(Female)* | 1000 | 204.0447 | -3.544012 | 124.72 | -40.699 | 448.788 451.810 | (1) |
|  |  |  |  |  | -41.597 | 455.775 | (BC) |
| Constant | 1000 | 2789.919 | -5.622323 | 291.0311 | 2218.816 | 3361.021 | (N) |
|  |  |  |  |  | 2182.780 | 3335.692 | (P) |
|  |  |  |  |  | 2176.969 | 3330.329 | (BC) |
| $\mathrm{N}=$ normal, $\mathrm{P}=$ percentile, $\mathrm{BC}=$ bias-corrected |  |  |  |  |  |  |  |

Table 86: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress for the Persons of Age Under 20

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.2120466 | -0.0009043 | 0.0285709 | -0.268113 | -0.155981 | (N) |
|  |  |  |  |  | -0.268033 | -0.157898 | (P) |
|  |  |  |  |  | -0.265754 | -0.155179 | (BC) |
| CStress 97 (Subst. reduced)* | 1000 | -3850 | 75.28764 | 4368.896 | -12423.270 | 4723.267 | (N) |
|  |  |  |  |  | -12240.880 | 4850.812 | (P) |
|  |  |  |  |  | -12259.620 | 4845.814 |  |
| CStress 97 (Reduced)* | 1000 | -378.1679 | -1.372062 | 710.9883 | -1773.370 | 1017.034 | (N) |
|  |  |  |  |  | -1735.462 | 1094.100 | (P) |
|  |  |  |  |  | -1696.711 | 1129.495 | (BC) |
| CStress 97 (Increased)* | 1000 | -1434.396 | -6.970977 | 815.1875 | -3034.073 | 165.280 | (N) |
|  |  |  |  |  | -3069.561 | 166.439 | (P) |
|  |  |  |  |  | -3117.547 | 138.934 | (BC) |
| CStress 97 (Subst. increased)* | 991 | 1488.73 | 124.8736 | 5470.156 | -9245.702 | 12223.160 | (N) |
|  |  |  |  |  | -7174.034 | 13425.890 | (P) |
|  |  |  |  |  | -6509.133 | 15897.120 | (BC) |
| LFS 98(Unemployed)* | 1000 | -6579.279 | 89.7245 | 2174.228 | -10845.860 | -2312.702 | (N) |
|  |  |  |  |  | -10935.130 | -2286.003 | (P) |
|  |  |  |  |  | -11395.560 | -2562.557 | (BC) |
| LFS 98 (NLF)* | 1000 | -2160.43 | -13.5301 | 984.068 | -4091.508 | -229.353 | (N) |
|  |  |  |  |  | -4165.497 | -246.714 | (P) |
|  |  |  |  |  | -4218.351 | -251.995 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1291.202 | 18.17684 | 1208.822 | -3663.323 | 1080.919 | (N) |
|  |  |  |  |  | -3624.758 | 1023.418 | (P) |
|  |  |  |  |  | -3661.714 | 964.499 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1732.788 | 26.14142 | 889.7913 | -3478.863 | 13.286 | (N) |
|  |  |  |  |  | -3413.078 | 22.444 | (P) |
|  |  |  |  |  | -3411.043 | 23.429 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -3035.302 | 8.31137 | 1105.945 | -5205.545 | -865.060 | (N) |
|  |  |  |  |  | -5292.752 | -808.385 | (P) |
|  |  |  |  |  | -5444.582 | -993.946 | (BC) |
| LFS 98 (AOA)* | 1000 | -2985.701 | -46.18552 | 959.5631 | -4868.691 | -1102.710 | (N) |
|  |  |  |  |  | -4805.359 | -1066.534 | (P) |
|  |  |  |  |  | -4775.124 | -1041.238 | (BC) |
| HLE 98(GHS)* | 1000 | 397.7511 | 42.45571 | 803.6719 | -1179.328 | 1974.830 | (N) |
|  |  |  |  |  | -1079.741 | 2084.469 | (P) |
|  |  |  |  |  | -1107.293 | 1992.434 | (BC) |
| HLE 98(NUPC)* | 1000 | -481.7366 | 54.15994 | 1260.174 | -2954.628 | 1991.155 | (N) |
|  |  |  |  |  | -2889.008 | 2165.066 | (P) |
|  |  |  |  |  | -3087.116 | 1887.421 | (BC) |
| HLE 98(UD)* | 1000 | -191.4202 | 35.53489 | 4252.422 | -8536.125 | 8153.285 | (N) |
|  |  |  |  |  | -8719.174 | 7261.719 | (P) |
|  |  |  |  |  | -9272.689 | 7110.528 | (BC) |
| Age 98 | 1000 | -304.8104 | -14.73451 | 331.9968 | -956.302 | 346.681 | (N) |
|  |  |  |  |  | -992.069 | 326.345 | (P) |
|  |  |  |  |  | -953.751 | 373.394 | (BC) |
| Sex(Female)* | 1000 | -517.2297 | -21.61107 | 618.9335 | -1731.789 | 697.329 | (N) |
|  |  |  |  |  | -1807.349 | 675.159 727.387 | (P) |
| Constant | 1000 | 8621.119 | 248.3601 | 6291.191 | -3724.345 | 20966.580 | (N) |
|  |  |  |  |  | -3641.752 | 21347.590 | (P) |
|  |  |  |  |  | -5086.867 | 20634.000 | (BC) |
| $\mathrm{N}=$ normal, $\mathrm{P}=$ percentile, $\mathrm{BC}=$ bias-corrected |  |  |  |  |  |  |  |

Table 87: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress for the People between 20 and 49 Years of Age

| Variables | Reps | Observed | Bias | Std. Err. |  |  | Interval type(N)(P) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1632071 | 0.0000757 |  | 95\% Conf. Interval$-0.180122 \quad-0.146292$ |  |  |
|  |  |  |  |  | -0.179588 | -0.146422 |  |
|  |  |  |  |  | -0.179603 | -0.146424 | (BC) |
| CStress 97 (Subst. reduced)* | 1000 | 2358.325 | 105.1909 | 1828.462 | -1229.743000 | 5946.392000 | (N) |
|  |  |  |  |  | -1007.442000 | 6069.043000 | (P) |
|  | 1000 |  | 3.200598 | 176.4975 | -1164.596000 | 5836.748000 | (BC) |
| CStress 97 (Reduced)* |  | -129.373 |  |  | -475.721300 | 216.975300 | (N) |
|  |  |  |  |  | -462.744200 | 215.994200 | (P) |
|  |  |  |  |  | -459.479300 | 220.387800 | (BC) |
| CStress 97 (Increased)* | 1000 | -222.6865 | 2.515913 | 175.268 | -566.622200 | 121.249100 | (N) |
|  |  |  |  |  | -552.283900 | 119.473100 | (P) |
|  |  |  |  |  | -565.729300 | 110.548200 | (BC) |
| CStress 97 (Subst. increased)* | 1000 | 563.1731 | -49.33578 | 1268.129 | -1925.330000 | 3051.676000 | (N) |
|  |  |  |  |  | -1958.882000 | 2999.497000 | (P) |
|  |  |  |  |  | -1916.558000 | 3081.800000 | (BC) |
| LFS 98(Unemployed)* | 1000 | -3681.477 | -17.71834 | 417.7821 | -4501.308000 | -2861.646000 | (N) |
|  |  |  |  |  | -4538.404000 | -2857.804000 | (P) |
|  |  |  |  |  | -4517.993000 | -2845.262000 | (BC) |
| LFS 98 (NLF)* | 1000 | -1726.026 | -0.7583057 | 251.779 | -2220.102000 | -1231.949000 | (N) |
|  |  |  |  |  | -2231.785000 | -1248.744000 | (P) |
|  |  |  |  |  | -2230.892000 | -1252.626000 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1046.934 | -17.37209 | 254.056 | -1545.479000 | -548.389900 | (N) |
|  |  |  |  |  | -1569.038000 | -577.407900 | (P) |
|  |  |  |  |  | -1518.394000 | -546.749600 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -1374.405 | -25.6943 | 358.6365 | -2078.172000 | -670.637700 | (N) |
|  |  |  |  |  | -2108.613000 | -682.363800 | (P) |
|  |  |  |  |  | -2090.945000 | -657.865200 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -3463.809 | -26.20347 | 627.477 | -4695.133000 | -2232.485000 | (N) |
|  |  |  |  |  | -4732.062000 | -2265.447000 | (P) |
|  |  |  |  |  | -4642.657000 | -2236.127000 | (BC) |
| LFS 98 (AOA)* | 1000 | -1733.585 | 17.29434 | 411.7586 | -2541.596000 | -925.573700 | (N) |
|  |  |  |  |  | -2533.491000 | -889.772000 | (P) |
|  |  |  |  |  | -2562.259000 | -927.946500 | (BC) |
| HLE 98(GHS)* | 1000 | 371.7068 | -0.1940119 | 214.3167 | -48.855850 | 792.269400 | (N) |
|  |  |  |  |  | -36.437660 | 779.976400 | (P) |
|  |  |  |  |  | -39.312360 5.200628 | 777.632500 825.044000 | $(\mathrm{BC})$ |
| HLE 98(NUPC)* | 1000 | 415.1223 | 8.522133 | 208.8942 | -27.930550 | 828.005600 | (P) |
|  |  |  |  |  | -52.128240 | 802.207700 | (BC) |
| HLE 98(UD)* | 1000 | 1212.365 | -4.029741 | 259.6047 | 702.932300 | 1721.798000 | (N) |
|  |  |  |  |  | 669.710900 | 1711.380000 | (P) |
|  |  |  |  |  | 673.877700 | 1712.862000 | (BC) |
| Age 98 | 1000 | -7.63113 | 0.3582275 | 10.09858 | -27.447990 | 12.185730 | (N) |
|  |  |  |  |  | -26.548230 | 13.510120 | (P) |
| Sex(Female)* |  |  |  |  | -26.720410 | 12.979910 538.151100 | $(\mathrm{BC})$ |
|  | 1000 | 259.4637 | -1.386675 | 142.0178 | -18.456460 | 561.604000 | (P) |
|  |  |  |  |  | -14.154060 | 564.963800 | (BC) |
| Constant | 1000 | 1826.916 | -16.87668 | 459.899 | 924.437400 | 2729.395000 | (N) |
|  |  |  |  |  | 878.390900 | 2654.143000 | (P) |
|  |  |  |  |  | 891.694900 | 2663.057000 | (BC) |
|  |  | = normal, | = percentile, | C $=$ bias -c | rected |  |  |

Table 88: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress for the People between 50 and 59 Years of Age

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1390227 | -0.0025006 | 0.0228866 | -0.183934 | -0.094111 | (N) |
|  |  |  |  |  | -0.186931 | -0.098408 | (P) |
|  |  |  |  |  | -0.184913 | -0.097690 | (BC) |
| CStress 97 (Subst. reduced)* | 1000 | 1322.976 | -31.64569 | 1685.691 | -1984.925 | 4630.877 | (N) |
|  |  |  |  |  | -1928.546 | 4842.902 | (P) |
|  |  |  |  |  | -1711.268 | 5350.644 | (BC) |
| CStress 97 (Reduced)* | 1000 | 364.3316 | -4.597221 | 372.6166 | -366.869 | 1095.533 | (N) |
|  |  |  |  |  | -404.676 | 1085.003 | (P) |
|  |  |  |  |  | -379.850 | 1103.578 | (BC) |
| CStress 97 (Increased)* | 1000 | -785.4662 | 3.043113 | 433.2496 | -1635.650 | 64.717 | (N) |
|  |  |  |  |  | -1619.027 | 63.795 | (P) |
|  |  |  |  |  | -1611.202 | 66.164 | $(\mathrm{BC})$ |
| CStress 97 (Subst. increased)* | 1000 | -1615.363 | 197.9491 | 3385.817 | -8259.491 | 5028.765 | (N) |
|  |  |  |  |  | -8251.458 | 5094.129 | (P) |
|  |  |  |  |  | -8471.712 | 4748.320 | (BC) |
| LFS 98(Unemployed)* | 1000 | -2248.685 | 0.7035189 | 872.8495 | -3961.513 | -535.856 | (N) |
|  |  |  |  |  | -3878.763 | -524.777 | (P) |
|  |  |  |  |  | -3925.137 | -553.496 | (BC) |
| LFS 98 (NLF)* | 1000 | -2841.18 | -16.66068 | 405.3969 | -3636.707 | -2045.653 | (N) |
|  |  |  |  |  | -3641.385 | -2116.230 | (P) |
|  |  |  |  |  | -3620.568 | -2080.755 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | -1582.972 | 9.873335 | 726.5854 | -3008.780 | -157.163 | (N) |
|  |  |  |  |  | -2967.957 | -181.512 | (P) |
|  |  |  |  |  | -2964.711 | -180.434 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -310.3412 | -0.7303831 | 1031.111 | -2333.733 | 1713.050 | (N) |
|  |  |  |  |  | -2347.432 | 1679.165 | (P) |
|  |  |  |  |  | -2409.481 | 1625.553 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -4154.709 | 7.630219 | 1485.521 | -7069.809 | -1239.610 | (N) |
|  |  |  |  |  | -7095.902 | -1366.254 | (P) |
|  |  |  |  |  | -7248.389 | -1477.676 | (BC) |
| LFS 98 (AOA)* | 1000 | 1309.803 | -45.90249 | 1205.814 | -1056.415 | 3676.022 | (N) |
|  |  |  |  |  | -1105.044 | 3620.505 | (P) |
|  |  |  |  |  | -1080.258 | 3685.520 | (BC) |
| HLE 98(GHS)* | 1000 | 537.9454 | -11.06905 | 410.0689 | -266.750 | 1342.641 | (N) |
|  |  |  |  |  | -328.658 | 1326.724 | (P) |
|  |  |  |  |  | -259.956 | 1372.887 | (BC) |
| HLE 98(NUPC)* | 1000 | 359.6082 | -7.424488 | 383.8747 | -393.685 | 1112.901 | (N) |
|  |  |  |  |  | -389.048 | 1145.795 1148.597 | (P) |
| HLE 98(UD)* | 1000 | -393.1782 | -16.70166 | 614.8957 | -1599.814 | 1148.597 813.457 | (N) |
|  |  |  |  |  | -1674.935 | 761.971 | (P) |
|  |  |  |  |  | -1665.129 | 760.995 | (BC) |
| Age 98 | 1000 | 79.22575 | 1.38527 | 56.33283 | -31.319 | 189.770 | (N) |
|  |  |  |  |  | -28.187 | 192.181 | (P) |
|  |  |  |  |  | -29.144 | 189.005 | (BC) |
| Sex(Female)* | 1000 | 257.6027 | 1.70818 | 332.3066 | -394.496 | 909.702 | (N) |
|  |  |  |  |  | -370.375 | 930.407 | (P) |
|  |  |  |  |  | -374.241 | 925.181 | (BC) |
| Constant | 1000 | -4038.837 | -64.01479 | 3154.975 | -10229.970 | 2152.300 | (N) |
|  |  |  |  |  | -10323.430 | 1917.141 | (P) |
|  |  |  |  |  | -10205.740 | 1945.724 | (BC) |
| $\mathrm{N}=$ normal, $\mathrm{P}=$ percentile, $\mathrm{BC}=$ bias-corrected |  |  |  |  |  |  |  |

Table 89: OLS Parameter Estimates and Bootstrap Parameter CI's of Model (6) with Stress for the People 60 and Over Years of Age

| Variables | Reps | Observed | Bias | Std. Err. | 95\% Conf. | Interval | Interval type |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Change in Income 97 | 1000 | -0.1199579 | -0.0006842 | 0.0194359 | -0.158098 | -0.081818 | (N) |
|  |  |  |  |  | -0.158971 | -0.081425 | (P) |
|  |  |  |  |  | -0.157153 | -0.079670 | (BC) |
| CStress 97 (Subst. reduced)* | 1000 | 974.0878 | -20.34289 | 1713.077 | -2387.554 | 4335.729 | (N) |
|  |  |  |  |  | -2127.412 | 4640.552 | (P) |
|  |  |  |  |  | -2066.298 | 4709.903 | (BC) |
| CStress 97 (Reduced)* | 1000 | -525.7607 | -10.36224 | 288.5704 | -1092.034 | 40.513 | (N) |
|  |  |  |  |  | -1104.543 | 23.980 | (P) |
|  |  |  |  |  | -1096.761 | 36.613 | (BC) |
| CStress 97 (Increased)* | 1000 | -216.7766 | -7.923616 | 317.0309 | -838.899 | 405.346 | (N) |
|  |  |  |  |  | -821.527 | 390.349 | (P) |
|  |  |  |  |  | -822.874 | 381.297 | (BC) |
| CStress 97 (Subst. increased)* | 1000 | -480.3437 | 32.17059 | 1478.685 | -3382.029 | 2421.341 | (N) |
|  |  |  |  |  | -3517.613 | 2440.660 | (P) |
|  |  |  |  |  | -3634.980 | 2244.369 | (BC) |
| LFS 98(Unemployed)* | 1000 | -1368.784 | -45.04868 | 1127.183 | -3580.703 | 843.135 | (N) |
|  |  |  |  |  | -3771.023 | 774.175 | (P) |
|  |  |  |  |  | -3690.599 | 854.812 | (BC) |
| LFS 98 (NLF)* | 1000 | -573.6223 | -29.17606 | 422.9973 | -1403.687 | 256.443 | (N) |
|  |  |  |  |  | -1435.614 | 239.614 | (P) |
|  |  |  |  |  | -1403.135 | 275.799 | (BC) |
| LFS 98(EPY-UPY)* | 1000 | 509.4504 | -51.66911 | 1104.804 | -1658.552 | 2677.453 | (N) |
|  |  |  |  |  | -1571.842 | 2839.665 | (P) |
|  |  |  |  |  | -1467.351 | 2921.167 | (BC) |
| LFS 98 (EPY-NLFPY)* | 1000 | -982.1015 | -46.96082 | 985.0833 | -2915.171 | 950.968 | (N) |
|  |  |  |  |  | -2965.717 | 833.971 | (P) |
|  |  |  |  |  | -2897.001 | 871.998 | (BC) |
| LFS 98 (UPY-NLFPY)* | 1000 | -1934.215 | 28.81277 | 2066.753 | -5989.889 | 2121.459 | (N) |
|  |  |  |  |  | -5712.913 | 2208.934 | (P) |
|  |  |  |  |  | -5583.095 | 2352.439 | (BC) |
| LFS 98 (AOA)* | 1000 | -561.0274 | -24.06745 | 1383.281 | -3275.496 | 2153.441 | (N) |
|  |  |  |  |  | -3172.589 | 2339.869 | (P) |
|  |  |  |  |  | -3083.187 | 2437.088 | (BC) |
| HLE 98(GHS)* | 1000 | 298.0523 | 3.215104 | 343.3356 | $-375.689$ | 971.794 | (N) |
|  |  |  |  |  | -392.470 | 983.487 | (P) |
|  |  |  |  |  | -418.924 | 960.145 | (BC) |
| HLE 98(NUPC)* | 1000 | -332.167 | -18.18585 | 328.0956 | -976.003 | 311.669 | (N) |
|  |  |  |  |  | -958.265 | 310.811 | (P) |
|  |  |  |  |  | -908.892 | 365.455 | (BC) |
| HLE 98(UD)* | 1000 | -220.4368 | 0.3705085 | 559.6251 | -1318.612 | 877.739 | (N) |
|  |  |  |  |  | -1334.268 | 841.078 | (P) |
|  |  |  |  |  | -1358.089 | 813.956 | (BC) |
| Age 98 | 1000 | 14.44918 | 0.0149341 | 57.1992 | -97.795 | 126.694 | (N) |
|  |  |  |  |  | -100.438 | 125.040 123.796 | (B) |
| Sex(Female)* | 1000 | 82.24911 | 2.028571 | 261.7612 | -431.416 | 595.914 | (N) |
|  |  |  |  |  | -463.983 | 596.280 | (P) |
|  |  |  |  |  | -475.600 | 587.228 | (BC) |
| Constant | 1000 | -66.15809 | 36.99401 | 3706.894 | -7340.350 | 7208.034 | (N) |
|  |  |  |  |  | -7255.831 | 7165.172 | (P) |
|  |  |  |  |  | -7295.601 | 7129.949 | (BC) |
| $\mathrm{N}=$ normal, $\mathrm{P}=$ percentile, $\mathrm{BC}=$ bias-corrected |  |  |  |  |  |  |  |


[^0]:    ${ }^{1}$ Throughout the analysis, the star symbol next to a category of a predictor represents a dummy variable

[^1]:    ${ }^{2}$ Comprehensive results could be found in Table 80 in the Appendix

[^2]:    ${ }^{3}$ Comprehensive results could be found in Table 85 in the Appendix

